

# Supplementary Materials for “Heteroscedastic Nested Error Regression Models with Variance Functions”

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In this supplementary file, we provide the additional figure in the data application and the proofs of the results in the paper.

## S1 Proof of Theorem 1

Since  $\mathbf{y}_1, \dots, \mathbf{y}_m$  are mutually independent, the consistency of  $\hat{\gamma}$  follows from the standard argument, so that  $\hat{\tau}^2$  and  $\hat{\beta}$  are also consistent. In what follows, we derive the asymptotic expressions of the estimators.

First we consider the asymptotic approximation of  $\hat{\tau}^2 - \tau^2$ . From (6), we obtain

$$\begin{aligned}
 \hat{\tau}^2 - \tau^2 &= \frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \left\{ (y_{ij} - \mathbf{x}'_{ij} \hat{\beta}_{\text{OLS}})^2 - \hat{\sigma}_{ij}^2 \right\} - \tau^2 \\
 &= \frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \left\{ (y_{ij} - \mathbf{x}'_{ij} \beta)^2 - \sigma_{ij}^2 \right\} - \tau^2 - \frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \sigma_{ij(1)}^2 \mathbf{z}'_{ij} (\hat{\gamma} - \gamma) \\
 &\quad - \frac{2}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} (y_{ij} - \mathbf{x}'_{ij} \beta) \mathbf{x}'_{ij} (\hat{\beta}_{\text{OLS}} - \beta) + o_p(\hat{\gamma} - \gamma) + o_p(\hat{\beta}_{\text{OLS}} - \beta) \\
 &= \frac{1}{m} \sum_{i=1}^m u_{1i} - \frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \sigma_{ij(1)}^2 \mathbf{z}'_{ij} (\hat{\gamma} - \gamma) + o_p(m^{-1/2}) + o_p(\hat{\gamma} - \gamma), \tag{S1}
 \end{aligned}$$

where  $u_{1i} = mN^{-1} \sum_{j=1}^{n_i} \left\{ (y_{ij} - \mathbf{x}'_{ij} \beta)^2 - \sigma_{ij}^2 \right\} - \tau^2$  and we used the fact that  $\hat{\beta}_{\text{OLS}} - \beta = O_p(m^{-1/2})$  and  $N^{-1} \sum_{i=1}^m \sum_{j=1}^{n_i} (y_{ij} - \mathbf{x}'_{ij} \beta) \mathbf{x}_{ij} = O_p(m^{-1/2})$  from the central limit theorem.

For the asymptotic expansion of  $\hat{\gamma}$ , remember that the estimator  $\hat{\gamma}$  is given as the solution of the estimating equation

$$\frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \left[ \left\{ y_{ij} - \bar{y}_i - (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' \hat{\beta}_{\text{OLS}} \right\}^2 \mathbf{z}_{ij} - \sigma_{ij}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \right] = \mathbf{0}$$

Using Taylor expansions, we have

$$\begin{aligned}
 \mathbf{0} &= \frac{1}{m} \sum_{i=1}^m \mathbf{u}_{2i} - \frac{2}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \left\{ y_{ij} - \bar{y}_i - (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' \beta \right\} \mathbf{z}_{ij} (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' (\hat{\beta}_{\text{OLS}} - \beta) \\
 &\quad - \frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \sigma_{ij(1)}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \mathbf{z}'_{ij} (\hat{\gamma} - \gamma) + o_p(\hat{\gamma} - \gamma) + o_p(m^{-1/2}),
 \end{aligned}$$

where

$$\mathbf{u}_{2i} = mN^{-1} \sum_{j=1}^{n_i} \left[ \{y_{ij} - \bar{y}_i - (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' \boldsymbol{\beta}\}^2 \mathbf{z}_{ij} - \sigma_{ij}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \right].$$

From the central limit theorem, it follows that

$$\frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \{y_{ij} - \bar{y}_i - (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' \boldsymbol{\beta}\} \mathbf{z}_{ij} (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' = O_p(m^{-1/2}),$$

so that the second terms in the expansion formula is  $o_p(m^{-1/2})$ . Then we get

$$\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma} = \frac{N}{m} \left( \sum_{i=1}^m \sum_{j=1}^{n_i} \sigma_{ij(1)}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \mathbf{z}'_{ij} \right)^{-1} \sum_{i=1}^m \mathbf{u}_{2i} + o_p(\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma}) + o_p(m^{-1/2}).$$

Under (A1)-(A5), we have

$$\sum_{i=1}^m \sum_{j=1}^{n_i} \sigma_{ij(1)}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \mathbf{z}'_{ij} = O(m).$$

From the independence of  $\mathbf{y}_1, \dots, \mathbf{y}_m$  and the fact  $E(\mathbf{u}_{2i}) = \mathbf{0}$ , we can use the central limit theorem to show that the leading term in the expansion of  $\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma}$  is  $O_p(m^{-1/2})$ . Thus,

$$\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma} = \frac{N}{m} \left( \sum_{i=1}^m \sum_{j=1}^{n_i} \sigma_{ij(1)}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \mathbf{z}'_{ij} \right)^{-1} \sum_{i=1}^m \mathbf{u}_{2i} + o_p(m^{-1/2}).$$

Using the approximation of  $\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma}$  and  $\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma} = O_p(m^{-1/2})$ , we get the asymptotic expression of  $\hat{\tau}^2 - \tau^2$  from (S1), which establishes the result for  $\hat{\tau}^2$  and  $\hat{\boldsymbol{\gamma}}$ .

Finally we consider the asymptotic expansion of  $\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}$ . From the expression in (4), it follows that

$$\hat{\boldsymbol{\beta}} - \boldsymbol{\beta} = \tilde{\boldsymbol{\beta}} - \boldsymbol{\beta} + \sum_{s=1}^q \left( \frac{\partial}{\partial \gamma_s} \tilde{\boldsymbol{\beta}} \right)' (\hat{\gamma}_s - \gamma) + \left( \frac{\partial}{\partial \tau^2} \tilde{\boldsymbol{\beta}} \right)' (\hat{\tau}^2 - \tau^2) + o_p(\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma}) + o_p(\hat{\tau}^2 - \tau^2).$$

Since

$$\frac{\partial}{\partial \tau^2} \boldsymbol{\Sigma}_i = \mathbf{J}_{n_i}, \quad \frac{\partial}{\partial \gamma_s} \boldsymbol{\Sigma}_i = \mathbf{W}_{i(s)}, \quad s = 1, \dots, q,$$

for  $\mathbf{W}_{i(s)} = \text{diag}(\sigma_{i1(1)}^2 z_{i1s}, \dots, \sigma_{in_i(1)}^2 z_{in_i s})$ , we have

$$\begin{aligned} \frac{\partial}{\partial \tau^2} \tilde{\boldsymbol{\beta}} &= (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \left( \sum_{i=1}^m \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} \mathbf{J}_{n_i} \boldsymbol{\Sigma}_i^{-1} \mathbf{X}_i \right) (\tilde{\boldsymbol{\beta}}_{\tau}^* - \tilde{\boldsymbol{\beta}}), \\ \frac{\partial}{\partial \gamma_s} \tilde{\boldsymbol{\beta}} &= (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \left( \sum_{i=1}^m \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} \mathbf{W}_{i(s)} \boldsymbol{\Sigma}_i^{-1} \mathbf{X}_i \right) (\tilde{\boldsymbol{\beta}}_{\gamma_s}^* - \tilde{\boldsymbol{\beta}}), \quad s = 1, \dots, q, \end{aligned} \tag{S2}$$

where

$$\begin{aligned}\tilde{\boldsymbol{\beta}}_{\tau}^* &= \left( \sum_{i=1}^m \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} \mathbf{J}_{n_i} \boldsymbol{\Sigma}_i^{-1} \mathbf{X}_i \right)^{-1} \sum_{i=1}^m \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} \mathbf{J}_{n_i} \boldsymbol{\Sigma}_i^{-1} \mathbf{y}_i, \\ \tilde{\boldsymbol{\beta}}_{\gamma_s}^* &= \left( \sum_{i=1}^m \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} \mathbf{W}_{i(s)} \boldsymbol{\Sigma}_i^{-1} \mathbf{X}_i \right)^{-1} \sum_{i=1}^m \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} \mathbf{W}_{i(s)} \boldsymbol{\Sigma}_i^{-1} \mathbf{y}_i, \quad s = 1, \dots, q.\end{aligned}$$

Under (A1)-(A5), we have  $\tilde{\boldsymbol{\beta}}_a^* - \boldsymbol{\beta} = O_p(m^{-1/2})$  for  $a \in \{\tau, \gamma_1, \dots, \gamma_q\}$ , whereby  $\tilde{\boldsymbol{\beta}}^* - \tilde{\boldsymbol{\beta}} = O_p(m^{-1/2})$ . Since  $\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma} = O_p(m^{-1/2})$  and  $\hat{\tau}^2 - \tau^2 = O_p(m^{-1/2})$  as shown above, we get

$$\hat{\boldsymbol{\beta}} - \boldsymbol{\beta} = (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \sum_{i=1}^m \mathbf{X}_i \boldsymbol{\Sigma}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) + o_p(m^{-1/2}),$$

which completes the proof.

## S2 Proof of Corollary 1

Let  $\boldsymbol{\theta} = (\theta_1, \dots, \theta_{p+q+1})' = (\boldsymbol{\beta}', \boldsymbol{\gamma}', \tau^2)'$ . Note that  $\psi_i^{\theta_k}, k = 1, \dots, p+q+1$  does not depend on  $\mathbf{y}_1, \dots, \mathbf{y}_{i-1}, \mathbf{y}_{i+1}, \dots, \mathbf{y}_m$  and that  $\mathbf{y}_1, \dots, \mathbf{y}_m$  are mutually independent. Then,

$$\begin{aligned}\frac{1}{m^2} E \left[ \left( \sum_{j=1}^m \psi_j^{\theta_k} \right) \left( \sum_{j=1}^m \psi_j^{\theta_l} \right) \middle| \mathbf{y}_i \right] &= \frac{1}{m^2} \sum_{j=1, j \neq i}^m E \left[ \psi_j^{\theta_k} \psi_j^{\theta_l} \right] + \frac{1}{m^2} \psi_i^{\theta_k} \psi_i^{\theta_l} \\ &= \boldsymbol{\Omega}_{kl} + \frac{1}{m^2} \left\{ \psi_i^{\theta_k} \psi_i^{\theta_l} - E \left[ \psi_i^{\theta_k} \psi_i^{\theta_l} \right] \right\},\end{aligned}$$

where  $\boldsymbol{\Omega}_{kl}$  is the  $(k, l)$ -element of  $\boldsymbol{\Omega}$  and we used the fact that  $E[\psi_j^{\theta_k} | \mathbf{y}_i] = E[\psi_j^{\theta_k}] = 0$  for  $j \neq i$ . Hence, we get the result from the asymptotic approximation of  $\hat{\boldsymbol{\theta}}$  given in Theorem 1.

## S3 Proof of Theorem 2

We begin by deriving the conditional asymptotic bias of  $\hat{\boldsymbol{\gamma}}$ . Let  $\tilde{\boldsymbol{\gamma}}$  be the solution of the equation

$$\mathbf{F}(\boldsymbol{\gamma}; \boldsymbol{\beta}) \equiv \frac{1}{N} \sum_{i=1}^m \sum_{j=1}^{n_i} \left[ \{y_{ij} - \bar{y}_i - (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)' \boldsymbol{\beta}\}^2 \mathbf{z}_{ij} - \sigma_{ij}^2 (\mathbf{z}_{ij} - 2n_i^{-1} \mathbf{z}_{ij} + n_i^{-1} \bar{\mathbf{z}}_i) \right] = \mathbf{0}$$

with  $\sigma_{ij}^2 = \sigma^2(\mathbf{z}'_{ij} \boldsymbol{\gamma})$ . For notational simplicity, we use  $\mathbf{F}$  instead of  $\mathbf{F}(\boldsymbol{\gamma}; \boldsymbol{\beta})$  without any confusion and  $F_r, r = 1, \dots, q$  denotes the  $r$ -th component of  $\mathbf{F}$ , namely  $\mathbf{F} = (F_1, \dots, F_q)'$ . Define the derivatives  $\mathbf{F}_{(a)}$  and  $F_{h(ab)}$  by

$$\mathbf{F}_{(a)} = \frac{\partial \mathbf{F}}{\partial \mathbf{a}'}, \quad F_{r(ab)} = \frac{\partial^2 F_r}{\partial a \partial b'}.$$

It is noted that  $F_{h(\boldsymbol{\beta}\boldsymbol{\gamma})} = 0$ . Expanding  $\mathbf{F}(\hat{\boldsymbol{\gamma}}; \hat{\boldsymbol{\beta}}_{\text{OLS}}) = \mathbf{0}$ , we obtain

$$\mathbf{0} = \mathbf{F} + \mathbf{F}_{(\boldsymbol{\gamma})}(\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma}) + \mathbf{F}_{(\boldsymbol{\beta})}(\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}) + \frac{1}{2} \mathbf{t}_1 + \frac{1}{2} \mathbf{t}_2 + o_p(m^{-1}),$$

where  $\mathbf{t}_s = (t_{s1}, \dots, t_{sq})$ ,  $s = 1, 2$  for

$$t_{1r} = (\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma})' F_{r(\boldsymbol{\gamma}\boldsymbol{\gamma})} (\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma}), \quad t_{2r} = (\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta})' F_{r(\boldsymbol{\beta}\boldsymbol{\beta})} (\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}).$$

It is also noted that

$$\begin{aligned} \mathbf{F}_{(\boldsymbol{\gamma})} &= -\frac{1}{m} \sum_{k=1}^m \sum_{j=1}^{n_k} \sigma_{kj(1)}^2 (\mathbf{z}_{kj} - 2n_k^{-1} \mathbf{z}_{kj} + n_k^{-1} \bar{\mathbf{z}}_k) \mathbf{z}'_{kj} \\ \mathbf{F}_{(\boldsymbol{\beta})} &= -\frac{2}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} \{y_{kj} - \bar{y}_k - (\mathbf{x}_{kj} - \bar{\mathbf{x}}_k)' \boldsymbol{\beta}\} \mathbf{z}_{ij} (\mathbf{x}_{kj} - \bar{\mathbf{x}}_k)', \end{aligned}$$

so that  $\mathbf{F}_{(\boldsymbol{\gamma})}$  is non-stochastic. Thus we have

$$E[\hat{\boldsymbol{\gamma}} - \boldsymbol{\gamma} | \mathbf{y}_i] = -(\mathbf{F}_{(\boldsymbol{\gamma})})^{-1} \left\{ E[\mathbf{F}(\boldsymbol{\gamma}; \boldsymbol{\beta}) | \mathbf{y}_i] + E[\mathbf{F}_{(\boldsymbol{\beta})}(\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}) | \mathbf{y}_i] + \frac{1}{2} E[\mathbf{t}_1 | \mathbf{y}_i] + \frac{1}{2} E[\mathbf{t}_2 | \mathbf{y}_i] \right\} + o_p(m^{-1}).$$

In what follows, we shall evaluate the each term in the parenthesis in the above expression. For the first term, since  $\mathbf{y}_1, \dots, \mathbf{y}_m$  are mutually independent and  $E(\mathbf{u}_{2i}) = \mathbf{0}$ , we have

$$E[\mathbf{F}(\boldsymbol{\gamma}; \boldsymbol{\beta}) | \mathbf{y}_i] = \frac{1}{m} \mathbf{u}_{2i}.$$

For evaluation of the second term, we define  $\mathbf{Z}_{kr} = \text{diag}(z_{k1r}, \dots, z_{kn_k r})$ , where  $z_{kjr}$  denotes the  $r$ -th element of  $\mathbf{z}_{kj}$ . Then it follows that

$$\begin{aligned} E[\mathbf{F}_{r(\boldsymbol{\beta})}(\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}) | \mathbf{y}_i] &= -\frac{2}{N} \sum_{k=1}^m E \left[ (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\beta})' \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k (\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}) \middle| \mathbf{y}_i \right] \\ &= -\frac{2}{N} \sum_{k=1, k \neq i}^m E \left[ (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\beta})' \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k (\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}) \middle| \mathbf{y}_i \right] - \frac{2}{N} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta})' \mathbf{E}_i \mathbf{Z}_{ir} \mathbf{E}_i \mathbf{X}_i E \left[ \hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta} \middle| \mathbf{y}_i \right]. \end{aligned}$$

Noting that it holds for  $\ell = 1, \dots, m$  and  $k \neq i$

$$E \left[ (\mathbf{y}_\ell - \mathbf{X}_\ell \boldsymbol{\beta}) (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\beta})' \middle| \mathbf{y}_i \right] = 1_{\{\ell=k\}} \boldsymbol{\Sigma}_k, \quad E[\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta} | \mathbf{y}_i] = (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}'_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}),$$

we have

$$\begin{aligned} &E \left[ (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\beta})' \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k (\hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta}) \middle| \mathbf{y}_i \right] \\ &= \sum_{\ell=1}^m \text{tr} \left\{ \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}'_k E \left[ (\mathbf{y}_\ell - \mathbf{X}_\ell \boldsymbol{\beta}) (\mathbf{y}_k - \mathbf{X}_k \boldsymbol{\beta})' \middle| \mathbf{y}_i \right] \right\} \\ &= \text{tr} \left\{ (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}'_k \boldsymbol{\Sigma}_k \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k \right\}, \end{aligned}$$

which is  $O(m^{-1})$  and

$$\frac{1}{N} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta})' \mathbf{E}_i \mathbf{Z}_{ir} \mathbf{E}_i \mathbf{X}_i E \left[ \hat{\boldsymbol{\beta}}_{\text{OLS}} - \boldsymbol{\beta} \middle| \mathbf{y}_i \right] = o_p(m^{-1}).$$

Thus, we get

$$E \left[ \mathbf{F}_{r(\beta)}(\widehat{\beta}_{\text{OLS}} - \beta) \mid \mathbf{y}_i \right] = -\frac{2}{m} \sum_{k=1}^m \sum_{j=1}^{n_k} \text{tr} \left\{ (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'_k \boldsymbol{\Sigma}_k \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k \right\} + o_p(m^{-1}), \quad (\text{S3})$$

where the leading term is  $O(m^{-1})$ . For the third and fourth terms, note that

$$F_{r(\gamma\gamma)} = -\frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} \sigma_{kj(2)}^2 (z_{kj} - 2n_k^{-1} z_{kj} + n_k^{-1} \bar{z}_k) z'_{kj} z_{kjr} \quad F_{r(\beta\beta)} = \frac{2}{N} \sum_{k=1}^m \mathbf{X}'_k \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k,$$

which are non-stochastic. Then for  $h = 1, \dots, q$ ,

$$\begin{aligned} E[t_{1r} \mid \mathbf{y}_i] &= -\frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} z_{kjr} \sigma_{kj(2)}^2 (z_{kj} - 2n_k^{-1} z_{kj} + n_k^{-1} \bar{z}_k)' \boldsymbol{\Omega}_{\gamma\gamma} z_{kj} + o_p(m^{-1}), \\ E[t_{2r} \mid \mathbf{y}_i] &= \frac{2}{N} \sum_{k=1}^m \text{tr} (\mathbf{X}'_k \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k \mathbf{V}_{\text{OLS}}) + o_p(m^{-1}), \end{aligned}$$

for  $\mathbf{V}_{\text{OLS}} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\boldsymbol{\Sigma}\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}$ , where we used Corollary 1 and

$$E \left[ (\widehat{\beta}_{\text{OLS}} - \beta)(\widehat{\beta}_{\text{OLS}} - \beta)' \mid \mathbf{y}_i \right] = \mathbf{V}_{\text{OLS}} + o_p(m^{-1}), \quad (\text{S4})$$

which follows from the similar argument to the proof of Corollary 1. Thus we obtain

$$\begin{aligned} E[\mathbf{t}_1 \mid \mathbf{y}_i] &= -\frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} z_{kj} \sigma_{kj(2)}^2 (z_{kj} - 2n_k^{-1} z_{kj} + n_k^{-1} \bar{z}_k)' \boldsymbol{\Omega}_{\gamma\gamma} z_{kj} + o_p(m^{-1}), \\ E[\mathbf{t}_2 \mid \mathbf{y}_i] &= \frac{2}{N} \sum_{k=1}^m \left\{ \text{tr} (\mathbf{X}'_k \mathbf{E}_k \mathbf{Z}_{kr} \mathbf{E}_k \mathbf{X}_k \mathbf{V}_{\text{OLS}}) \right\}_r + o_p(m^{-1}), \end{aligned}$$

where  $\{\mathbf{a}_r\}_r$  denotes the  $q$ -dimensional vector  $(a_1, \dots, a_q)$ . Therefore, we have established the result for  $\widehat{\gamma}$  in (13).

We next derive the result for  $\widehat{\tau}^2$ . Let

$$\widetilde{\tau}^2 = \frac{1}{N} \sum_{k=1}^m \left\{ (\mathbf{y}_k - \mathbf{X}_k \beta)' (\mathbf{y}_k - \mathbf{X}_k \beta) - \sum_{j=1}^{n_k} \sigma_{kj}^2 \right\}.$$

Using the Taylor series expansion, we have

$$\begin{aligned} \widehat{\tau}^2 &= \widetilde{\tau}^2 + \frac{\partial \widetilde{\tau}^2}{\partial \gamma} (\widehat{\gamma} - \gamma) + \frac{1}{2} (\widehat{\gamma} - \gamma)' \left( \frac{\partial^2 \widetilde{\tau}^2}{\partial \gamma \partial \gamma'} \right) (\widehat{\gamma} - \gamma) \\ &\quad + \frac{\partial \widetilde{\tau}^2}{\partial \beta} (\widehat{\beta}_{\text{OLS}} - \beta) + \frac{1}{2} (\widehat{\beta}_{\text{OLS}} - \beta)' \left( \frac{\partial^2 \widetilde{\tau}^2}{\partial \beta \partial \beta'} \right) (\widehat{\beta}_{\text{OLS}} - \beta) + o_p(m^{-1}), \end{aligned}$$

where we used the fact that  $\partial^2 \tilde{\tau}^2 / \partial \gamma \partial \beta' = 0$ . The straight calculation shows that

$$\frac{\partial \tilde{\tau}^2}{\partial \gamma} = -\frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} \sigma_{kj(1)}^2 \mathbf{z}_{kj}, \quad \frac{\partial^2 \tilde{\tau}^2}{\partial \gamma \partial \gamma'} = -\frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} \sigma_{kj(2)}^2 \mathbf{z}_{kj} \mathbf{z}'_{kj}, \quad \frac{\partial^2 \tilde{\tau}^2}{\partial \beta \partial \beta'} = \frac{2}{N} \sum_{k=1}^m \mathbf{X}'_k \mathbf{X}_k,$$

which are non-stochastic. Thus we obtain

$$\begin{aligned} E[\tilde{\tau}^2 - \tau^2 | \mathbf{y}_i] &= E[\tilde{\tau}^2 - \tau^2 | \mathbf{y}_i] + \left( \frac{\partial \tilde{\tau}^2}{\partial \gamma} \right)' E[\hat{\gamma} - \gamma | \mathbf{y}_i] + \frac{1}{2} \text{tr} \left\{ \left( \frac{\partial^2 \tilde{\tau}^2}{\partial \gamma \partial \gamma'} \right) E[(\hat{\gamma} - \gamma)(\hat{\gamma} - \gamma)' | \mathbf{y}_i] \right\} \\ &+ E \left[ \left( \frac{\partial \tilde{\tau}^2}{\partial \beta} \right)' (\hat{\beta}_{\text{OLS}} - \beta) \middle| \mathbf{y}_i \right] + \frac{1}{2} \text{tr} \left\{ \left( \frac{\partial^2 \tilde{\tau}^2}{\partial \beta \partial \beta'} \right) E[(\hat{\beta}_{\text{OLS}} - \beta)(\hat{\beta}_{\text{OLS}} - \beta)' | \mathbf{y}_i] \right\} + o_p(m^{-1}) \\ &\equiv B_{\tau 1}(\mathbf{y}_i) + B_{\tau 2}(\mathbf{y}_i) + B_{\tau 3}(\mathbf{y}_i) + B_{\tau 4}(\mathbf{y}_i) + B_{\tau 5}(\mathbf{y}_i) + o_p(m^{-1}). \end{aligned}$$

From the expression of  $\tilde{\tau}^2$ , it holds that

$$\begin{aligned} B_{\tau 1}(\mathbf{y}_i) &= \frac{1}{N} \sum_{k=1, k \neq i}^m n_k \tau^2 + \frac{1}{N} \left\{ (\mathbf{y}_i - \mathbf{X}_i \beta)' (\mathbf{y}_i - \mathbf{X}_i \beta) - \sum_{j=1}^{n_i} \sigma_{ij}^2 \right\} - \tau^2 \\ &= \left( 1 - \frac{n_i}{N} \right) \tau^2 + \frac{1}{m} u_{1i} + \frac{n_i}{N} \tau^2 - \tau^2 = \frac{1}{m} u_{1i}, \end{aligned}$$

for  $u_{1i}$  defined in (8). Also, we immediately have

$$B_{\tau 2}(\mathbf{y}_i) = -\frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} \sigma_{kj(1)}^2 \mathbf{z}'_{kj} \mathbf{b}_{\gamma}^{(i)}(\mathbf{y}_i)$$

For evaluation of  $B_{\tau 4}(\mathbf{y}_i)$ , note that

$$\frac{\partial \tilde{\tau}^2}{\partial \beta} = -\frac{2}{N} \sum_{k=1}^m \mathbf{X}'_k (\mathbf{y}_k - \mathbf{X}_k \beta).$$

Similarly to (S3), we get

$$\begin{aligned} B_{\tau 4}(\mathbf{y}_i) &= -\frac{2}{N} \sum_{k=1}^m E \left[ (\mathbf{y}_k - \mathbf{X}_k \beta)' \mathbf{X}_k (\hat{\beta}_{\text{OLS}} - \beta) \middle| \mathbf{y}_i \right] \\ &= -\frac{2}{N} \sum_{k=1}^m \text{tr} \left\{ (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}'_k \Sigma_k \mathbf{X}_k \right\} + o_p(m^{-1}). \end{aligned}$$

Moreover, Corollary 1 and (S4) enable us to obtain the expression of  $B_{\tau 3}(\mathbf{y}_i)$  and  $B_{\tau 5}(\mathbf{y}_i)$ , whereby we get

$$b_{\tau}^{(i)}(\mathbf{y}_i) = m^{-1} u_{1i} - \frac{1}{N} \sum_{k=1}^m \sum_{j=1}^{n_k} \sigma_{kj(1)}^2 \mathbf{z}'_{kj} \left\{ \mathbf{b}_{\gamma}^{(i)}(\mathbf{y}_i) - \mathbf{b}_{\gamma} \right\} + b_{\tau},$$

which completes the proof for  $\hat{\tau}^2$  in (13).

We finally derive the result for  $\widehat{\boldsymbol{\beta}}$ . By the Taylor series expansion,

$$\widehat{\boldsymbol{\beta}} - \boldsymbol{\beta} = \widetilde{\boldsymbol{\beta}} - \boldsymbol{\beta} + \sum_{s=1}^q \left( \frac{\partial}{\partial \gamma_s} \widetilde{\boldsymbol{\beta}} \right) (\widehat{\gamma}_s - \gamma_s) + \left( \frac{\partial}{\partial \tau^2} \widetilde{\boldsymbol{\beta}} \right) (\widehat{\tau}^2 - \tau^2) + o_p(m^{-1}),$$

since

$$\left( \frac{\partial \widetilde{\boldsymbol{\beta}}}{\partial \boldsymbol{\phi}} \right)' (\widehat{\boldsymbol{\phi}} - \boldsymbol{\phi})(\widehat{\boldsymbol{\phi}} - \boldsymbol{\phi})' \left( \frac{\partial \widetilde{\boldsymbol{\beta}}}{\partial \boldsymbol{\phi}} \right) = o_p(m^{-1}),$$

from  $\partial \widetilde{\boldsymbol{\beta}} / \partial \boldsymbol{\phi} = O_p(m^{-1/2})$  as shown in the proof of Theorem 1. From (S2), we have

$$\begin{aligned} & \sum_{s=1}^q \left( \frac{\partial}{\partial \gamma_s} \widetilde{\boldsymbol{\beta}} \right) (\widehat{\gamma}_s - \gamma_s) \\ &= (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \sum_{s=1}^q \left( \sum_{k=1}^m \mathbf{X}'_k \boldsymbol{\Sigma}_k^{-1} \mathbf{W}_{i(s)} \boldsymbol{\Sigma}_k^{-1} \mathbf{X}_k \right) \left\{ (\widetilde{\boldsymbol{\beta}}_{\gamma_s}^* - \boldsymbol{\beta})(\widehat{\gamma}_s - \gamma_s) - (\widetilde{\boldsymbol{\beta}} - \boldsymbol{\beta})(\widehat{\gamma}_s - \gamma_s) \right\}, \end{aligned}$$

and

$$\left( \frac{\partial}{\partial \tau^2} \widetilde{\boldsymbol{\beta}} \right) (\widehat{\tau}^2 - \tau^2) = (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \left( \sum_{k=1}^m \mathbf{X}'_k \boldsymbol{\Sigma}_k^{-1} \mathbf{J}_{n_k} \boldsymbol{\Sigma}_k^{-1} \mathbf{X}_k \right) \left\{ (\widetilde{\boldsymbol{\beta}}_{\tau}^* - \boldsymbol{\beta})(\widehat{\tau}^2 - \tau^2) - (\widetilde{\boldsymbol{\beta}} - \boldsymbol{\beta})(\widehat{\tau}^2 - \tau^2) \right\}.$$

Let  $\boldsymbol{\Omega}_{\beta^* \gamma_s} = E[(\widetilde{\boldsymbol{\beta}}_{\gamma_s}^* - \boldsymbol{\beta})(\widehat{\gamma}_s - \gamma_s)]$  and  $\boldsymbol{\Omega}_{\beta^* \tau} = E[(\widetilde{\boldsymbol{\beta}}_{\tau}^* - \boldsymbol{\beta})(\widehat{\tau} - \tau)]$ . Then it can be shown that

$$E[(\widetilde{\boldsymbol{\beta}}_{\tau}^* - \boldsymbol{\beta})(\widehat{\tau} - \tau) | \mathbf{y}_i] = \boldsymbol{\Omega}_{\beta^* \tau} + o_p(m^{-1}), \quad E[(\widetilde{\boldsymbol{\beta}}_{\gamma_s}^* - \boldsymbol{\beta})(\widehat{\gamma}_s - \gamma_s) | \mathbf{y}_i] = \boldsymbol{\Omega}_{\beta^* \tau} + o_p(m^{-1}),$$

which can be proved by the same arguments as in Corollary 1. Thus from Corollary 1 and the fact that

$$E[\widetilde{\boldsymbol{\beta}} - \boldsymbol{\beta} | \mathbf{y}_i] = (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \mathbf{X}'_i \boldsymbol{\Sigma}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}),$$

we obtain the result for  $\widehat{\boldsymbol{\beta}}$  in (13).

## S4 Proof of (18)

From the expansion of  $\widehat{\mu}_i$ , we have

$$E[(\widehat{\mu}_i - \widetilde{\mu}_i)^2] = E \left[ \left\{ \left( \frac{\partial \widetilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' (\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}) \right\}^2 \right] + \frac{1}{2} U_1 + \frac{1}{4} U_2,$$

where

$$\begin{aligned} U_1 &= E \left[ \left( \frac{\partial \widetilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' (\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta})(\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta})' \left( \frac{\partial^2 \widetilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) (\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}) \right] \\ U_2 &= E \left[ \left\{ (\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta})' \left( \frac{\partial^2 \widetilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) (\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}) \right\}^2 \right]. \end{aligned}$$

It is noted that

$$U_1 = \sum_{j=1}^{p+q+1} \sum_{k=1}^{p+q+1} \sum_{\ell=1}^{p+q+1} E \left[ \left( \frac{\partial \tilde{\mu}_i}{\partial \theta_j} \right) \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \theta_k \partial \theta_\ell} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) (\hat{\theta}_j - \theta_j)(\hat{\theta}_k - \theta_k)(\hat{\theta}_\ell - \theta_\ell) \right] \equiv \sum_{j=1}^{p+q+1} \sum_{k=1}^{p+q+1} \sum_{\ell=1}^{p+q+1} U_{1jkl},$$

and

$$\begin{aligned} |U_{1jkl}| &\leq E \left[ \left| \left( \frac{\partial \tilde{\mu}_i}{\partial \theta_j} \right) \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \theta_k \partial \theta_\ell} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) \right| |(\hat{\theta}_j - \theta_j)(\hat{\theta}_k - \theta_k)(\hat{\theta}_\ell - \theta_\ell)| \right] \\ &\leq E \left[ \left| \left( \frac{\partial \tilde{\mu}_i}{\partial \theta_j} \right) \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \theta_k \partial \theta_\ell} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) \right|^4 \right]^{1/4} E \left[ |(\hat{\theta}_j - \theta_j)(\hat{\theta}_k - \theta_k)(\hat{\theta}_\ell - \theta_\ell)|^{4/3} \right]^{3/4} \end{aligned} \quad (\text{S5})$$

using Holder's inequality. Since both  $\partial \tilde{\mu}_i / \partial \theta_j$  and  $\partial^2 \tilde{\mu}_i / \partial \theta_k \partial \theta_\ell$  are linear functions of  $\mathbf{y}_i$ , the first term of (S5) is finite under (A4). Moreover, from Theorem 1, it follows  $\sqrt{m}|\hat{\theta}_j - \theta_j| \leq C(\mathbf{y})$  for some quadratic function of  $\mathbf{y}$ , so that the second term in (S5) is also finite. Hence, we have  $U_1 = o(m^{-1})$ . Similarly, we also obtain  $U_2 = o(m^{-1})$ . Therefore, using Corollary 1, we have

$$\begin{aligned} E[(\hat{\mu}_i - \tilde{\mu}_i)^2] &= E \left[ \left\{ \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}) \right\}^2 \right] + o(m^{-1}) \\ &= \text{tr} \left\{ E \left[ \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right) \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' E \left( (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})' \Big| \mathbf{y}_i \right) \right] \right\} + o(m^{-1}) \\ &= \text{tr} \left\{ E \left[ \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right) \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' \boldsymbol{\Omega} + \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right) \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' c(\mathbf{y}_i) o(m^{-1}) \right] \right\} + o(m^{-1}) \\ &= \text{tr} \left\{ E \left[ \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right) \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' \right] \boldsymbol{\Omega} \right\} + o(m^{-1}) \end{aligned}$$

since  $c(\mathbf{y}_i)$  is fourth-order function of  $\mathbf{y}_i$  and  $\partial \tilde{\mu}_i / \partial \boldsymbol{\theta}$  is a linear function of  $\mathbf{y}_i$ , which completes the proof.

## S5 Derivation of $R_{31i}(\boldsymbol{\phi}, \boldsymbol{\kappa})$

Since  $\mathbf{y}_i$  given  $v_i, \boldsymbol{\epsilon}_i$  is non-stochastic, we have

$$\begin{aligned} &E \left[ \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}) w_i \right] \\ &= E \left[ E \left[ \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right)' (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}) w_i \Big| v_i, \boldsymbol{\epsilon}_i \right] \right] = E \left[ E(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta} | \mathbf{y}_i)' \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\theta}} \right) w_i \right] \\ &= E \left[ \mathbf{b}_{\boldsymbol{\beta}}^{(i)}(\mathbf{y}_i)' \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\beta}} \right) w_i \right] + E \left[ \mathbf{b}_{\boldsymbol{\gamma}}^{(i)}(\mathbf{y}_i)' \left( \frac{\partial \tilde{\mu}_i}{\partial \boldsymbol{\gamma}} \right) w_i \right] + E \left[ b_{\tau}^{(i)}(\mathbf{y}_i) \left( \frac{\partial \tilde{\mu}_i}{\partial \tau} \right) w_i \right] + o(m^{-1}) \\ &\equiv R_{31i}(\boldsymbol{\phi}) + o(m^{-1}). \end{aligned}$$

It is noted that  $E(w_i) = 0$  and

$$E[(y_{ij} - \mathbf{x}'_{ij} \boldsymbol{\beta}) w_i] = E[(v_i + \varepsilon_{ij}) w_i] = \left( \sum_{j=1}^{n_i} \lambda_{ij} - 1 \right) \tau^2 + \sum_{j=1}^{n_i} \lambda_{ij} \sigma_{ij}^2 = 0. \quad (\text{S6})$$



Using the expression (13) and (17), it follows that

$$\begin{aligned}
E \left[ \mathbf{b}_{\boldsymbol{\beta}}^{(i)}(\mathbf{y}_i)' \left( \frac{\partial \tilde{\boldsymbol{\mu}}_i}{\partial \boldsymbol{\beta}} \right) w_i \right] &= \left( \mathbf{c}_i - \sum_{j=1}^{n_i} \lambda_{ij} \mathbf{x}_{ij} \right)' (\mathbf{X}' \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \mathbf{X}' \boldsymbol{\Sigma}_i^{-1} E[(\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}) w_i] = 0 \\
E \left[ \mathbf{b}_{\boldsymbol{\gamma}}^{(i)}(\mathbf{y}_i)' \left( \frac{\partial \tilde{\boldsymbol{\mu}}_i}{\partial \boldsymbol{\gamma}} \right) w_i \right] &= \eta_i^{-2} \sum_{j=1}^{n_i} \sigma_{ij}^{-2} \boldsymbol{\delta}'_{ij} \left( \sum_{k=1}^m \sum_{h=1}^{n_k} \sigma_{kh(1)}^2 \mathbf{z}_{kh} \mathbf{z}'_{kh} \right)^{-1} \mathbf{M}_{2ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) \\
E \left[ \mathbf{b}_{\boldsymbol{\tau}}^{(i)}(\mathbf{y}_i)' \left( \frac{\partial \tilde{\boldsymbol{\mu}}_i}{\partial \boldsymbol{\tau}} \right) w_i \right] &= m^{-1} \eta_i^{-2} \sum_{j=1}^{n_i} \sigma_{ij}^{-2} \left\{ M_{1ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) - \mathbf{T}_1(\boldsymbol{\gamma})' \mathbf{T}_2(\boldsymbol{\gamma}) \mathbf{M}_{2ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) \right\},
\end{aligned}$$

where

$$\mathbf{M}_{2ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) = E[\mathbf{u}_{2i}(y_{ij} - \mathbf{x}'_{ij} \boldsymbol{\beta}) w_i], \quad M_{1ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) = E[u_{1i}(y_{ij} - \mathbf{x}'_{ij} \boldsymbol{\beta}) w_i].$$

To evaluate  $M_{1ij}$  and  $\mathbf{M}_{2ij}$ , we first prove the following result for fixed  $j, k, \ell \in \{1, \dots, n_i\}$ .

$$\begin{aligned}
E[(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{i\ell}) w_i] &= \tau^2 \eta_i^{-1} \left[ \tau^2 (3 - \kappa_v) + \kappa_\varepsilon \sigma_{ij}^2 1_{\{j=k=\ell\}} + \sigma_{ij}^2 (1_{\{j=k \neq \ell\}} - 1_{\{j=k\}}) \right. \\
&\quad \left. + \sigma_{ij}^2 (1_{\{j=\ell \neq k\}} - 1_{\{j=\ell\}}) + \sigma_{ik}^2 (1_{\{k=\ell \neq j\}} - 1_{\{k=\ell\}}) \right]. \tag{S7}
\end{aligned}$$

To show (S7), we note that the left side can be rewritten as

$$-\eta_i^{-1} E[(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{i\ell}) v_i] + \sum_{h=1}^{n_i} \lambda_{ih} E[(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{i\ell}) \varepsilon_{ih}] \tag{S8}$$

from the definition of  $w_i$ . Using the fact that  $\varepsilon_{i1}, \dots, \varepsilon_{in_i}$  and  $v_i$  are independent, the first term in (S8) is calculated as

$$E[v_i^4 + (\varepsilon_{ij} \varepsilon_{ik} + \varepsilon_{ij} \varepsilon_{i\ell} + \varepsilon_{ik} \varepsilon_{i\ell}) v_i^2] = \kappa_v \tau^4 + \tau^2 (\sigma_{ij}^2 1_{\{j=k\}} + \sigma_{ij}^2 1_{\{j=\ell\}} + \sigma_{ik}^2 1_{\{k=\ell\}}).$$

Moreover, we have

$$\begin{aligned}
E[(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{i\ell}) \varepsilon_{ih}] &= E[\varepsilon_{ih} (\varepsilon_{ij} + \varepsilon_{i\ell} + \varepsilon_{ik}) v_i^2 + \varepsilon_{ij} \varepsilon_{ik} \varepsilon_{i\ell} \varepsilon_{ih}] \\
&= \tau^2 \sigma_{ih}^2 (1_{\{h=j\}} + 1_{\{h=k\}} + 1_{\{h=\ell\}}) + \kappa_\varepsilon \sigma_{ih}^4 1_{\{j=k=\ell=h\}} \\
&\quad + \sigma_{ih}^2 (\sigma_{ij}^2 1_{\{j=k \neq \ell=h\}} + \sigma_{ij}^2 1_{\{j=\ell \neq k=h\}} + \sigma_{ik}^2 1_{\{j=h \neq k=\ell\}}),
\end{aligned}$$

whereby the second term in (S8) can be calculated as

$$\tau^2 \eta_i^{-1} [3\tau^2 + \kappa_\varepsilon \sigma_{ij}^2 1_{\{j=k=\ell\}} + \sigma_{ij}^2 1_{\{j=k \neq \ell\}} + \sigma_{ij}^2 1_{\{j=\ell \neq k\}} + \sigma_{ik}^2 1_{\{k=\ell \neq j\}}],$$

where we used the expression  $\lambda_{ih} = \tau^2 \eta_i^{-1} \sigma_{ih}^{-2}$ . Then we established the result (S7). From (S7), we immediately have

$$\begin{aligned}
\sum_{\ell=1}^{n_i} E[(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{i\ell}) w_i] &= \tau^2 \eta_i^{-1} [n_i \tau^2 (3 - \kappa_v) + \sigma_{ij}^2 (\kappa_\varepsilon - 3) 1_{\{j=k\}}] \\
&= E[(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})^2 w_i].
\end{aligned}$$

Now, we return to the evaluation of  $M_{1ij}$  and  $M_{2ij}$ . It follows that

$$\begin{aligned} M_{1ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) &= \frac{m}{N} \sum_{h=1}^{n_i} E [(y_{ih} - \mathbf{x}'_{ih}\boldsymbol{\beta})^2 (y_{ij} - \mathbf{x}'_{ij}\boldsymbol{\beta}) w_i] \\ &= mN^{-1} \eta_i^{-1} \tau^2 \left\{ n_i \tau^2 (3 - \kappa_v) + \sigma_{ij}^2 (\kappa_\varepsilon - 3) \right\} \end{aligned}$$

and

$$\begin{aligned} M_{2ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) &= \frac{m}{N} \sum_{h=1}^{n_i} \mathbf{z}_{ih} E [\{v_i + \varepsilon_{ih} - (v_i + \bar{\varepsilon}_i)\}^2 (v_i + \varepsilon_{ij}) w_i] \\ &= \frac{m}{N} \sum_{h=1}^{n_i} \mathbf{z}_{ih} \left\{ E [(v_i + \varepsilon_{ih})^2 (v_i + \varepsilon_{ij}) w_i] - 2n_i^{-1} \sum_{k=1}^{n_i} E [(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{ih}) w_i] \right. \\ &\quad \left. + n_i^{-2} \sum_{k=1}^{n_i} \sum_{\ell=1}^{n_i} E [(v_i + \varepsilon_{ij})(v_i + \varepsilon_{ik})(v_i + \varepsilon_{i\ell}) w_i] \right\}. \end{aligned}$$

Using the identity given in (S7), we have

$$\begin{aligned} M_{2ij}(\boldsymbol{\phi}, \boldsymbol{\kappa}) &= mN^{-1} \tau^2 \eta_i^{-1} \sum_{h=1}^{n_i} \mathbf{z}_{ih} \left\{ \sigma_{ij}^2 (\kappa_\varepsilon - 3) (1_{\{j=h\}} - 2n_i^{-1} 1_{\{j=h\}} + n_i^{-2}) \right\} \\ &= mN^{-1} \tau^2 \eta_i^{-1} n_i^{-2} (n_i - 1)^2 (\kappa_\varepsilon - 3) \sigma_{ij}^2 \mathbf{z}_{ij}, \end{aligned}$$

which completes the result in (20).

## S6 Evaluation of $R_{32i}(\boldsymbol{\phi})$

Since  $\mathbf{y}_i$  given  $v_i$  and  $\boldsymbol{\varepsilon}_i$  is non-stochastic, we have

$$\begin{aligned} R_{32i}(\boldsymbol{\phi}) &= \frac{1}{2} E \left[ (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})' \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}) w_i \right] = \frac{1}{2} E \left[ E \left[ (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta})' \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) (\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}) w_i \Big| v_i, \boldsymbol{\varepsilon}_i \right] \right] \\ &= \frac{1}{2} \text{tr} \left\{ \boldsymbol{\Omega} E \left[ \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) w_i \right] \right\} + o(m^{-1}) E \left[ \text{tr} \left\{ c(\mathbf{y}_i) \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) \right\} w_i \right], \end{aligned}$$

where we used Corollary 1 in the last equation. Note that

$$\frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} = \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} + \sum_{k=1}^{p+q+1} (\theta_k^* - \theta_k) \left( \frac{\partial^3 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}' \partial \theta_k} \Big|_{\theta_k=\theta_k^{**}} \right), \quad (\text{S9})$$

where  $\theta_k^{**}$  is an intermediate value between  $\theta_k^*$  and  $\theta_k$ . Further note that the third order partial derivatives of  $\tilde{\mu}_i$  is a linear function of  $\mathbf{y}_i$ , so that the second term of  $R_{32i}$  is  $o(m^{-1})$ . Similarly, it follows that

$$E \left[ \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \Big|_{\boldsymbol{\theta}=\boldsymbol{\theta}^*} \right) w_i \right] = E \left[ \left( \frac{\partial^2 \tilde{\mu}_i}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \right) w_i \right] + o(1) = o(1),$$

since the second order partial derivatives of  $\tilde{\mu}_i$  is a linear function of  $y_{ij} - \mathbf{x}'_{ij}\boldsymbol{\beta}$  and the identity (S6). Therefore, we finally get  $R_{32i}(\boldsymbol{\phi}) = o(m^{-1})$ .

## S7 Predicted values of $\mu_i$ in data analysis

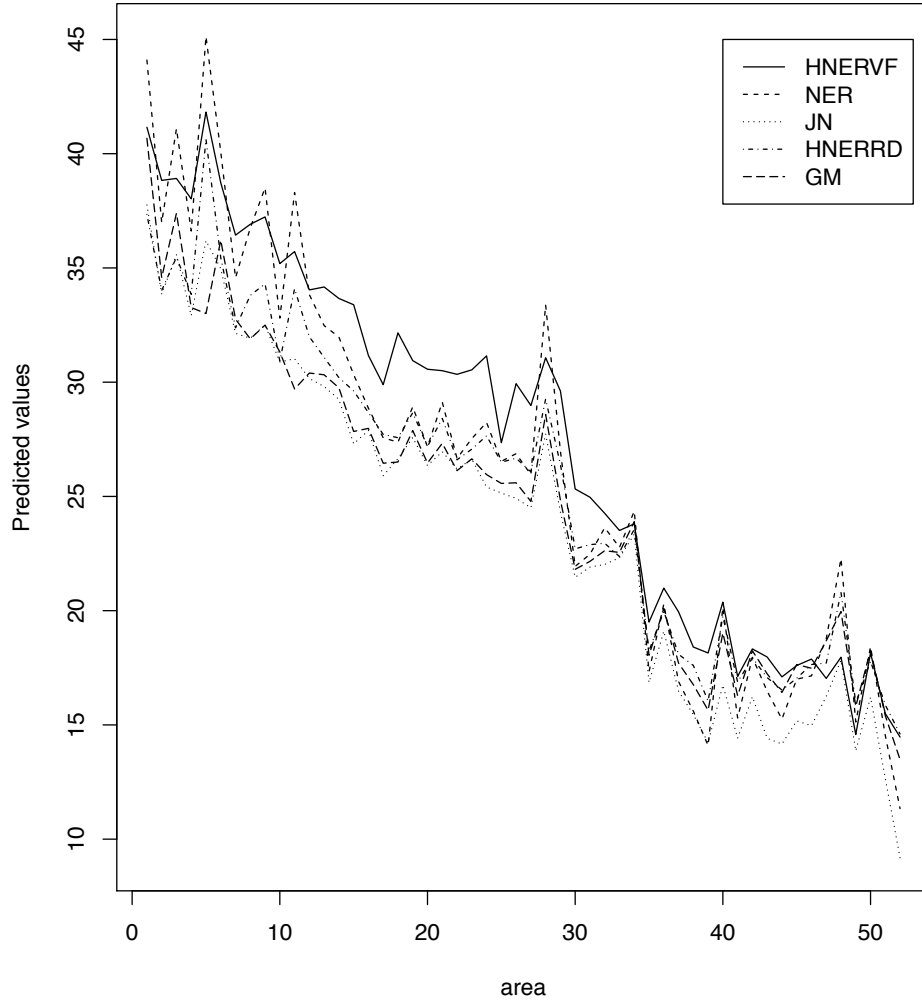


Figure 1: Predicted Values of  $\mu_i$  from Each Model.