

## REGULARIZING LASSO: A CONSISTENT VARIABLE SELECTION METHOD

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### Supplementary Material

The supplementary material is organized as the follows. In Section S1, we provide additional information regarding the simulation studies in Section 5.1. In Section S2, we give lemmas for establishing asymptotic results when the precision matrix is sparse. The proofs of all lemmas and theorems are given in Section S3.

## S1 Additional Tables of Simulation Results

Table 1 provides the average computational time (in minutes) for the eight methods under the simulation settings. SIS clearly requires the least computational effort, whereas RLASSO as well as Scout require much longer computational time. But all methods except RLASSO(CLIME) can be computed under a reasonable amount of time for  $p = 5000$  and  $n = 100$ . RLASSO(CLIME) takes much longer because of inverting a matrix of 5000 dimension. However, 790.8 minutes of computation may still be acceptable. In an unreported simulation with  $p = 2000$  and the same other settings, the average computational time for RLASSO(CLIME) is 46.7 minutes.

Finally, because the estimation of  $\Sigma$  is an important step in RLASSO, we provide the average Frobenius norms of estimated  $\Sigma$  and  $\Omega$  in the simulation in Table 2. It is clear that  $\mathbf{S}$  is not a good estimator of  $\Sigma$  in terms of the Frobenius norm,  $\hat{\Sigma}$  by thresholding (or  $\hat{\Omega}$  by CLIME) is a good estimator when  $\Sigma$  (or  $\Omega$ ) is sparse but not so good when  $\Sigma$  (or  $\Omega$ ) is not sparse.

## S2 Lemmas

**Lemma S1.** *Assume conditions (C1)-(C2) and (C3''), for any  $\lambda_n \rightarrow 0$ , there exist positive constants  $C_{14}, C_{15}, C_{16}$  such that*

$$P\left(\|\tilde{\beta} - \beta\|_{\infty} > t\right) \leq 8 \exp\left(-C_{14}n[t/s_h r_q]^{1-\frac{2}{q}}\right) + 4p^2 \exp\left(-C_{15}nt^2/s_h^2\right) + 8p \exp(-C_{16}nt^2)$$

for any  $0 < t < 8M^{1-h}s_h$ .

Table 1: Average computational time of various methods (in minutes)  
 $p = 5000$  and  $n = 100$

	Model 1	Model 2	Model 3	Model 4
RLASSO(AT)	6.6	2.1	4.1	3.6
RLASSO(CLIME)	790.8	700.1	651.8	758.6
RLASSO(GLASSO)	11.0	1.7	7.6	8.2
Scout(1,1)	9.6	1.5	5.8	6.3
LASSO	1.5	0.6	1.2	1.4
LASSO+T	2.8	0.8	1.2	1.9
SLSE+T	4.2	1.5	2.3	2.0
SIS	0.6	0.4	0.6	0.5

Table 2: The average Frobenius norms of estimated  $\Sigma$  and  $\Omega$

		Model 1	Model 2	Model 3	Model 4
$\ \mathbf{S} - \Sigma\ _F$	$\mathbf{S} = \mathbf{X}'\mathbf{X}/n$	618.17	2877.25	2495.43	570.22
$\ \hat{\Sigma} - \Sigma\ _F$	RLASSO(AT)	50.88	40.71	385.38	40.17
$\ \hat{\Omega} - \Omega\ _F$	RLASSO(CLIME)	41.02	273.01	49.99	42.03

**Lemma S2.** Assume conditions (C1'), (C2) and (C3'''), for any  $\lambda_n \rightarrow 0$ , there exist positive constants  $C_{21}$ ,  $C_{22}$ ,  $C_{23}$  and  $C_{24}$  such that

$$P\left(\|\tilde{\beta} - \beta\|_\infty > t\right) \leq C_{21} \left[ \exp\left(-C_{22}n[t/r_q s_h]^{1-q}\right) + n^{-\frac{l-1-\tau}{2}} \right] \\ + C_{23}p^2 s_h^{2l} t^{-2l} n^{-l} + C_{24}pt^{-2l} n^{-l},$$

for any  $0 < t < 8M^{1-h} s_h$ .

### S3 Proofs

*Proof of Lemma 1.* From (1),

$$\hat{\beta}_{M_j} - \beta_{M_j} = \frac{1}{n} \sum_{i=1}^n x_{ij} \left( \mu + \sum_{k=1}^p \beta_k x_{ik} + \sigma_i \epsilon_i \right) - \sum_{k=1}^p \mathbf{E}(x_j x_k) \beta_k \\ = \sum_{k=1}^p \left[ \frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} - \mathbf{E}(x_j x_k) \right] \beta_k + \frac{1}{n} \sum_{i=1}^n \sigma_i x_{ij} \epsilon_i + \frac{1}{n} \sum_{i=1}^n \mu x_{ij}.$$

Then,

$$P\left(|\hat{\beta}_{M_j} - \beta_{M_j}| > t\right) \leq P\left(\sum_{k=1}^p \left| \frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} - \mathbf{E}(x_j x_k) \right| |\beta_k| > \frac{t}{3}\right) \\ + P\left(\left| \frac{1}{n} \sum_{i=1}^n \sigma_i x_{ij} \epsilon_i \right| > \frac{t}{3}\right) + P\left(\left| \frac{1}{n} \sum_{i=1}^n \mu x_{ij} \right| > \frac{t}{3}\right) \quad (\text{S3.1})$$

Under condition (C1), applying Lemma 1 of Cai and Liu (2011) to  $\pm(x_{ij}x_{ik} - \mathbb{E}(x_{ij}x_{ik}))$  gives that there exist  $D_1 > 0$  and  $D_2 > 0$  such that

$$\max_{jk} P \left( \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right| > t \right) \leq 2 \exp(-D_1nt^2),$$

for all  $0 < t \leq D_2$ . Then, under condition (C2), it follows by Bonferroni inequality that there exist  $C_1 > 0$  and  $C_3 > 0$  such that, for any  $1 \leq j \leq p$ ,

$$\begin{aligned} & P \left( \sum_{k=1}^p \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right| |\beta_k| > \frac{t}{3} \right) \\ & \leq P \left( M^{1-h} s_h \max_k \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_{ij}x_{ik}) \right| > \frac{t}{3} \right) \\ & \leq p \cdot \max_{jk} P \left( \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right| > \frac{t}{3M^{1-h}s_h} \right) \\ & \leq 2p \exp(-C_1nt^2/s_h^2), \end{aligned} \tag{S3.2}$$

for all  $0 < t \leq C_3s_h$ . Similarly,

$$\max_{1 \leq j \leq p} P \left( \left| \frac{1}{n} \sum_{i=1}^n \sigma_i x_{ij} \epsilon_i \right| > \frac{t}{3} \right) \leq 2 \exp(-C_2nt^2), \tag{S3.3}$$

$$\max_{1 \leq j \leq p} P \left( \left| \frac{1}{n} \sum_{i=1}^n \mu x_{ij} \right| > \frac{t}{3} \right) \leq 2 \exp(-C_2nt^2), \tag{S3.4}$$

for some  $C_2 > 0$ . Therefore,

$$P \left( |\hat{\beta}_{M_j} - \beta_{M_j}| > t \right) \leq 2p \exp(-C_1nt^2/s_h^2) + 4 \exp(-C_2nt^2). \tag{S3.5}$$

Then, Lemma 1 follows by

$$P \left( \|\hat{\beta}_M - \beta_M\|_\infty > t \right) \leq \sum_{j=1}^p P \left( |\hat{\beta}_{M_j} - \beta_{M_j}| > t \right).$$

□

**Proof of Lemma 2.** By Karush-Kuhn-Tucker conditions, the solution  $\tilde{\beta}$  to (8) satisfies that

$$\hat{\Sigma} \tilde{\beta} - \hat{\beta}_M = -\lambda_n \mathbf{Z}, \tag{S3.6}$$

where  $\mathbf{Z}$  has the form of

$$\mathbf{Z} = \begin{cases} 1, & \text{if } \tilde{\beta}_j > 0; \\ -1, & \text{if } \tilde{\beta}_j < 0; \\ \in [-1, 1], & \text{if } \tilde{\beta}_j = 0. \end{cases} \tag{S3.7}$$

Simple algebra from (S3.6) yields

$$\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta} = \boldsymbol{\Sigma}^{-1}[\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M - \lambda_n \mathbf{Z} - (\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})\boldsymbol{\beta} - (\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})(\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta})]. \quad (\text{S3.8})$$

Hence,

$$\|\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}\|_\infty \leq v_p(\|\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M\|_\infty + \|\lambda_n \mathbf{Z}\|_\infty + \|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})\boldsymbol{\beta}\|_\infty + \|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})(\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta})\|_\infty).$$

Equivalently,

$$\frac{1}{v_p}\|\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}\|_\infty - \|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})(\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta})\|_\infty \leq \|\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M\|_\infty + \|\lambda_n \mathbf{Z}\|_\infty + \|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})\boldsymbol{\beta}\|_\infty$$

Then, by  $\|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})(\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta})\|_\infty \leq \|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 \|\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}\|_\infty$ , it holds that

$$\begin{aligned} & P(\|\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}\|_\infty > t) \\ &= P\left(\|\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}\|_\infty > t, \|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 \leq \frac{1}{2v_p}\right) + P\left(\|\tilde{\boldsymbol{\beta}} - \boldsymbol{\beta}\|_\infty > t, \|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 > \frac{1}{2v_p}\right) \\ &\leq P\left(\|\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M\|_\infty + \|\lambda_n \mathbf{Z}\|_\infty + \|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})\boldsymbol{\beta}\|_\infty > \frac{t}{2v_p}, \|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 \leq \frac{1}{2v_p}\right) \\ &\quad + P\left(\|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 > \frac{1}{2v_p}\right) \\ &\leq P\left(\|\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M\|_\infty > \frac{t}{6v_p}\right) + P\left(\|\lambda_n \mathbf{Z}\|_\infty > \frac{t}{6v_p}\right) + P\left(\|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})\boldsymbol{\beta}\|_\infty > \frac{t}{6v_p}\right) \\ &\quad + P\left(\|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 > \frac{1}{2v_p}\right) \\ &:= I + II + III + IV \end{aligned}$$

By Lemma 1, there exist positive constants  $C_4$  and  $C_5$  such that

$$I \leq 2p^2 \exp(-C_4 nt^2 / (s_h v_p)^2) + 4p \exp(-C_5 nt^2 / v_p^2). \quad (\text{S3.9})$$

By the choice of  $\lambda_n$ ,  $II = 0$ , when  $n$  is sufficiently large. For  $III$ , under condition (C2),  $\|(\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma})\boldsymbol{\beta}\|_\infty \leq \|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 \|\boldsymbol{\beta}\|_\infty \leq M \|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1$ . Then, it follows from Theorem 1(i) of Cai and Liu (2011) that,

$$III + IV \leq 2P\left(\|\hat{\boldsymbol{\Sigma}} - \boldsymbol{\Sigma}\|_1 > \frac{t}{6Mv_p}\right) \leq C_6 n^{-1/2} p^{-(\delta-2)} (r_q v_p / t)^{1/(1-q)}, \quad (\text{S3.10})$$

for some  $C_6 > 0$ . This completes the proof of the lemma.  $\square$

**Proof of Theorem 1.** Note that  $\widehat{\mathcal{M}}_\beta = \mathcal{M}_{\tilde{\beta}, t_n}$  and

$$\begin{aligned}
& P\left(\mathcal{M}_{\beta, a_n t_n} \subset \widehat{\mathcal{M}}_\beta\right) \\
&= 1 - P\left(\cup_{j:|\beta_j| > a_n t_n} \left\{|\tilde{\beta}_j| \leq t_n\right\}\right) \\
&\geq 1 - P\left(\cup_{j:|\beta_j| > a_n t_n} \left\{|\tilde{\beta}_j - \beta_j| > (a_n - 1)t_n\right\}\right) \\
&\geq 1 - P\left(\|\tilde{\beta} - \beta\|_\infty > (a_n - 1)t_n\right) \\
&\geq 1 - O\left[\exp\left(-C_7(\log n)^{-2}n^{1-2\alpha_1-2\alpha_3-2\eta}\right) + (1/p)^{\delta-2}\left((\log n)(1/n)^{\frac{1-q}{2}-\alpha_2-\alpha_3-\eta}\right)^{1/(1-q)}\right]
\end{aligned}$$

by Lemma 2 and the choice of  $a_n$  and  $t_n$ . Similarly,

$$\begin{aligned}
& P(\widehat{\mathcal{M}}_\beta \subset \mathcal{M}_{\beta, t_n/a_n}) \\
&= P\left(\cap_{j:|\beta_j| \leq t_n/a_n} \left\{|\tilde{\beta}_j| \leq t_n\right\}\right) \\
&\geq 1 - P\left(\cup_{j:|\beta_j| \leq t_n/a_n} \left\{|\tilde{\beta}_j - \beta_j| > (1 - a_n^{-1})t_n\right\}\right) \\
&\geq 1 - P\left(\|\tilde{\beta} - \beta\|_\infty > (1 - a_n^{-1})t_n\right) \\
&= 1 - O\left[\exp\left(-C_7(\log n)^{-2}n^{1-2\alpha_1-2\alpha_3-2\eta}\right) + (1/p)^{\delta-2}\left((\log n)(1/n)^{\frac{1-q}{2}-\alpha_2-\alpha_3-\eta}\right)^{1/(1-q)}\right].
\end{aligned}$$

This completes the proof of the first part of Theorem 1. In particular, if we choose  $h = 0$ ,

$$\begin{aligned}
& P(\mathcal{M}_\beta \subset \widehat{\mathcal{M}}_\beta) \\
&= 1 - P\left(\cup_{j \in \mathcal{M}_\beta} \left\{|\tilde{\beta}_j| \leq t_n\right\}\right) \\
&\geq 1 - P\left(\cup_{j \in \mathcal{M}_\beta} \left\{|\beta_j| - |\tilde{\beta}_j - \beta_j| \leq t_n\right\}\right) \tag{S3.11} \\
&\geq 1 - P\left(\cup_{j \in \mathcal{M}_\beta} \left\{|\tilde{\beta}_j - \beta_j| \geq t_n/2\right\}\right) \\
&= 1 - O\left[\exp\left(-C_8n^{1-2\alpha_1-2\alpha_3-2\eta}\right) + (1/p)^{\delta-2}(1/n)^{(\frac{1-q}{2}-\alpha_2-\alpha_3-\eta)/(1-q)}\right],
\end{aligned}$$

since under (C5), by the choice of  $t_n$ ,  $\min_{j \in \mathcal{M}_\beta} |\beta_j| > \frac{3}{2}t_n$  for large enough  $n$ .

On the other hand,

$$\begin{aligned}
& P(\widehat{\mathcal{M}}_\beta \subset \mathcal{M}_\beta) \\
&= 1 - P\left(\cup_{j \notin \mathcal{M}_\beta} \left\{|\tilde{\beta}_j| > t_n\right\}\right) \\
&= 1 - P\left(\cup_{j \notin \mathcal{M}_\beta} \left\{|\tilde{\beta}_j - \beta_j| > t_n\right\}\right) \tag{S3.12} \\
&= 1 - O\left[\exp\left(-C_8n^{1-2\alpha_1-2\alpha_3-2\eta}\right) + (1/p)^{\delta-2}(1/n)^{(\frac{1-q}{2}-\alpha_2-\alpha_3-\eta)/(1-q)}\right].
\end{aligned}$$

(S3.11) and (S3.12) together prove the theorem.  $\square$

**Proof of Lemma 3.** The proof is analogous to that of Lemma 1. Under condition (C1'), it holds that

$$\mathbb{E} |x_{ij}x_{ik} - \mathbb{E}(x_jx_k)|^{2l} \leq 2^{2l-1} \left[ \mathbb{E}|x_{ij}x_{ik}|^{2l} + (\mathbb{E}|x_{ij}x_{ik}|)^{2l} \right] = O(1).$$

Then, by Chebyshev Inequality and Theorem 2 in Whittle (1960).

$$\max_{jk} P \left( \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right| > t \right) \leq t^{-2l} \mathbb{E} \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right|^{2l} = O(t^{-2l}n^{-l}).$$

Therefore, by replacing (S3.2) with

$$\begin{aligned} & P \left( \sum_{k=1}^p \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right| |\beta_k| > \frac{t}{3} \right) \\ & \leq p \cdot \max_{jk} P \left( \left| \frac{1}{n} \sum_{i=1}^n x_{ij}x_{ik} - \mathbb{E}(x_jx_k) \right| > \frac{t}{3M^{1-h} s_h} \right) \\ & \leq C_9 p s_h^{2l} t^{-2l} n^{-l}, \end{aligned}$$

for some  $C_9 > 0$  and replacing (S3.3) and (S3.4) with

$$\begin{aligned} & \max_{1 \leq j \leq p} P \left( \left| \frac{1}{n} \sum_{i=1}^n \sigma_i x_{ij} \epsilon_i \right| > \frac{t}{3} \right) \leq \frac{C_{10}}{2} t^{-2l} n^{-l}, \\ & \max_{1 \leq j \leq p} P \left( \left| \frac{1}{n} \sum_{i=1}^n \mu x_{ij} \right| > \frac{t}{3} \right) \leq \frac{C_{10}}{2} t^{-2l} n^{-l}, \end{aligned}$$

for some  $C_{10} > 0$ , the rest of proof follows from (S3.1).  $\square$

**Proof of Lemma 4.** From Lemma 3, it holds that

$$P \left( \|\hat{\beta}_M - \beta_M\|_\infty > \frac{t}{6v_p} \right) \leq C_{11} p^2 s_h^{2l} v_p^{2l} t^{-2l} n^{-l} + C_{12} p v_p^{2l} t^{-2l} n^{-l}, \quad (\text{S3.13})$$

for some  $C_{11} > 0$  and  $C_{12} > 0$ .

Under condition (C3'), it follows from Theorem 1(ii) of Cai and Liu (2011) that, there exists  $C_{13} > 0$  that

$$P \left( \|\hat{\Sigma} - \Sigma\|_1 > \frac{t}{6Mv_p} \right) \leq C_{13} \left( n^{-1/2} p^{-(\delta-2)} (r_q v_p / t)^{1/(1-q)} + n^{-\frac{l-1-\tau}{2}} \right). \quad (\text{S3.14})$$

Replacing (S3.9) and (S3.10) with (S3.13) and (S3.14) and observing that by the choice of  $\lambda_n$ ,  $P(\lambda_n \mathbf{Z} > \frac{t}{6v_p}) = 0$ , when  $n$  is sufficiently large, the rest of the proof resembles the proof of Lemma 2.  $\square$

**Proof of Theorem 2.** The proof is the same as the proof of Theorem 1 by replacing results in Lemma 2 with results in Lemma 4.  $\square$

**Proof of Lemma S1.** From (S3.6),

$$\tilde{\beta} = \hat{\Omega}\hat{\beta}_M - \lambda_n\hat{\Omega}\mathbf{Z}.$$

Recall that,  $\beta = \Omega\beta_M$ . Hence,

$$\tilde{\beta} - \beta = \hat{\Omega}\hat{\beta}_M - \Omega\beta_M - \lambda_n\hat{\Omega}\mathbf{Z}.$$

Then,

$$P(\|\tilde{\beta} - \beta\|_\infty > t) \leq P(\|\hat{\Omega}\hat{\beta}_M - \Omega\beta_M\|_\infty > t/2) + P(\|\lambda_n\hat{\Omega}\mathbf{Z}\|_\infty > t/2). \quad (\text{S3.15})$$

Since

$$\hat{\Omega}\hat{\beta}_M - \Omega\beta_M = (\hat{\Omega} - \Omega)\hat{\beta}_M + \Omega(\hat{\beta}_M - \beta_M),$$

it holds that

$$P(\|\hat{\Omega}\hat{\beta}_M - \Omega\beta_M\|_\infty > t/2) \leq P(\|(\hat{\Omega} - \Omega)\hat{\beta}_M\|_\infty > t/4) + P(\|\Omega(\hat{\beta}_M - \beta_M)\|_\infty > t/4). \quad (\text{S3.16})$$

The first item in (S3.16) is bounded by

$$\begin{aligned} & P\left(\|\hat{\Omega} - \Omega\|_1\|\hat{\beta}_M\|_\infty > t/4\right) \\ & \leq P\left(\|\hat{\Omega} - \Omega\|_1\|\hat{\beta}_M\|_\infty > t/4 \cap \|\hat{\beta}_M - \beta_M\|_\infty \leq t/8\right) + P(\|\hat{\beta}_M - \beta_M\|_\infty > t/8) \\ & \leq P\left(\|\hat{\Omega} - \Omega\|_1\|\beta_M\|_\infty > t/8 \cap \|\hat{\beta}_M - \beta_M\|_\infty \leq t/8\right) \\ & \quad + P\left(\|\hat{\Omega} - \Omega\|_1\|\hat{\beta}_M - \beta_M\|_\infty > t/8 \cap \|\hat{\beta}_M - \beta_M\|_\infty \leq t/8\right) \\ & \quad + P\left(\|\hat{\beta}_M - \beta_M\|_\infty > t/8\right) \\ & \leq P\left(\|\hat{\Omega} - \Omega\|_1 > t/[8M^{1-h}s_h]\right) + P\left(\|\hat{\Omega} - \Omega\|_1 > 1\right) + P\left(\|\hat{\beta}_M - \beta_M\|_\infty > t/8\right). \end{aligned}$$

For the second item in (S3.16), it follows from the assumption  $\|\Omega\|_1 \leq M$  that

$$\begin{aligned} P(\|\Omega(\hat{\beta}_M - \beta_M)\|_\infty > t/4) & \leq P\left(\|\Omega\|_1\|\hat{\beta}_M - \beta_M\|_\infty > t/4\right) \\ & \leq P\left(\|\hat{\beta}_M - \beta_M\|_\infty > t/[4M]\right). \end{aligned}$$

Without loss of generality, assume  $M \geq 2$ . Then, for any  $0 < t < 8M^{1-h}s_h$ ,

$$\begin{aligned} P\left(\|\tilde{\beta} - \beta\|_\infty > t\right) & \leq 2P\left(\|\hat{\Omega} - \Omega\|_1 > t/[8M^{1-h}s_h]\right) \\ & \quad + 2P\left(\|\hat{\beta}_M - \beta_M\|_\infty > t/[4M]\right), \end{aligned} \quad (\text{S3.17})$$

since  $\|\lambda_n\hat{\Omega}\mathbf{Z}\|_\infty \leq |\lambda_n|\|\hat{\Omega}\|_1\|\mathbf{Z}\|_\infty \leq |\lambda_n|\|\hat{\Omega}\|_1 \leq 2|\lambda_n|\|\Omega\|_1 \leq 2M|\lambda_n|$ . By the choice of  $\lambda_n$ , when  $n$  is sufficiently large,  $P(\|\lambda_n\hat{\Omega}\mathbf{Z}\|_\infty > t/2) = 0$ .

Under (C1), it follows by Theorem 1(a) of Cai, Liu, and Luo (2011) that

$$P\left(\|\hat{\Omega} - \Omega\|_1 > t/[8M^{1-h}s_h]\right) \leq 4 \exp\left(-C_{14}n[t/s_h r_q]^{2/(1-q)}\right), \quad (\text{S3.18})$$

for some  $C_{14} > 0$ .

From Lemma 1, it holds that

$$P(\|\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M\|_\infty > t/(4M)) \leq 2p^2 \exp(-C_{15}nt^2/s_h^2) + 4p \exp(-C_{16}nt^2). \quad (\text{S3.19})$$

(S3.17), (S3.18) together with (S3.19) proves the lemma.  $\square$

**Proof of Lemma S2.** Under conditions (C1') and (C3'''), by Theorem 1(ii) of Cai, Liu, and Luo (2011).

$$P\left(\|\hat{\boldsymbol{\Omega}} - \boldsymbol{\Omega}\|_1 > \frac{t}{8M^{1-h} s_h}\right) \leq \frac{C_{21}}{2} \left[ \exp\left(-C_{22}n [t/s_h r_q]^{1-\frac{2}{q}}\right) + n^{-\frac{l-1-\tau}{2}} \right].$$

From Lemma 3, it holds that

$$P(\|\hat{\boldsymbol{\beta}}_M - \boldsymbol{\beta}_M\|_\infty > t/(4M)) \leq \frac{C_{23}}{2} p^2 s_h^{2l} t^{-2l} n^{-l} + \frac{C_{24}}{2} pt^{-2l} n^{-l}.$$

The rest of proof follows by (S3.17).  $\square$

**Proof of Theorem 3 and Theorem 4.** By using results in Lemma S1 and Lemma S2, the proof is the same as that of Theorem 1.  $\square$

**Proof of (15).** Decompose  $\boldsymbol{\Sigma}$  as

$$\boldsymbol{\Sigma} = \begin{pmatrix} \mathbf{I}_{s_0} & \boldsymbol{\Sigma}'_{21} \\ \boldsymbol{\Sigma}_{21} & \mathbf{I}_{p-s_0} \end{pmatrix},$$

where  $\boldsymbol{\Sigma}_{21} = (\mathbf{B}' \mathbf{0})'$ .

$$\mathbf{S}_{21} \mathbf{S}_{11}^{-1} = \mathbf{S}_{21} (\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0}) + (\mathbf{S}_{21} - \boldsymbol{\Sigma}_{21}) \mathbf{I}_{s_0} + \boldsymbol{\Sigma}_{21}. \quad (\text{S3.20})$$

It is well known (e.g. see Bickel and Levina (2008)) that for normally distributed covariates,

$$\max_{1 \leq i, j \leq p} |s_{ij} - \rho_{ij}| = O_P\left(\sqrt{n^{-1} \log p}\right),$$

where  $s_{ij}$  is the  $(i, j)$ th element of  $\mathbf{S}$ . Then,

$$\|\mathbf{S}_{21} - \boldsymbol{\Sigma}_{21}\|_\infty = \max_{s_0 < i \leq p} \sum_{j=1}^{s_0} |s_{ij} - \rho_{ij}| = O_P\left(s_0 \sqrt{n^{-1} \log p}\right).$$

Hence,  $\|(\mathbf{S}_{21} - \boldsymbol{\Sigma}_{21}) \mathbf{I}_{s_0}\|_\infty = O_P\left(s_0 \sqrt{n^{-1} \log p}\right)$ .

Moreover,  $\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0} = \mathbf{S}_{11}^{-1} (\mathbf{I}_{s_0} - \mathbf{S}_{11})$ . Then,

$$\|\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0}\|_\infty \leq \|\mathbf{S}_{11}^{-1}\|_\infty \|\mathbf{I}_{s_0} - \mathbf{S}_{11}\|_\infty \leq (1 + \|\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0}\|_\infty) \|\mathbf{I}_{s_0} - \mathbf{S}_{11}\|_\infty.$$



Hence,  $\|\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0}\|_\infty \leq \|\mathbf{I}_{s_0} - \mathbf{S}_{11}\|_\infty / (1 - \|\mathbf{I}_{s_0} - \mathbf{S}_{11}\|_\infty) = O_P\left(s_0\sqrt{n^{-1}\log p}\right)$ . Then,

$$\begin{aligned} \|\mathbf{S}_{21}(\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0})\|_\infty &\leq \|\mathbf{S}_{21}\|_\infty \|\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0}\|_\infty \\ &\leq (\|\boldsymbol{\Sigma}_{21}\|_\infty + \|\mathbf{S}_{21} - \boldsymbol{\Sigma}_{21}\|_\infty) \|\mathbf{S}_{11}^{-1} - \mathbf{I}_{s_0}\|_\infty \\ &= O_P\left(s_0 + s_0\sqrt{n^{-1}\log p}\right) \cdot O_P\left(s_0\sqrt{n^{-1}\log p}\right) \\ &= O_P\left(s_0^2\sqrt{n^{-1}\log p}\right) \end{aligned}$$

Since  $\|\boldsymbol{\Sigma}_{21}\|_\infty \geq \|\mathbf{B}\|_\infty \geq 1 + 2\gamma$ , under the assumptions that  $n^{-1}\log p \rightarrow 0$  and  $s_0$  is fixed, it follows from (S3.20) that  $P(\|\mathbf{S}_{21}\mathbf{S}_{11}^{-1}\|_\infty \geq 1 + \gamma) \rightarrow 1$ .  $\square$

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