

---

**MM ALGORITHMS FOR VARIANCE COMPONENT ESTIMATION  
AND SELECTION IN LOGISTIC LINEAR MIXED MODEL**

Liuyi Hu<sup>1</sup>, Wenbin Lu<sup>1</sup>, Jin Zhou<sup>2</sup> and Hua Zhou<sup>3</sup>

<sup>1</sup>North Carolina State University, <sup>2</sup>University of Arizona, and <sup>3</sup>University of California, Los Angeles

**Supplementary Material**

**S1 Derivation of approximated log-likelihood in (3.2)**

$$L_{LA}(\beta, \sigma) = h(u^* | \beta, \sigma^2) - \frac{1}{2} \ln \det \nabla^2 \{-h(u^* | \beta, \sigma^2)\},$$

where

$$h(u^* | \beta, \sigma^2) = \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \sum_{i=1}^m q_i \ln \sigma_i^2 - \frac{n}{2} \ln 2\pi - \frac{1}{2} \sum_{i=1}^m \frac{\|u_i^*\|_2^2}{\sigma_i^2}.$$

The gradient and Hessian of  $h(u | \beta, \sigma)$  at  $u = u^*$  are

$$\begin{aligned} \nabla_u h(u | \beta, \sigma^2)|_{u=u^*} &= Z^T (y - p^*) - \begin{pmatrix} \sigma_1^{-2} u_1^* \\ \vdots \\ \sigma_m^{-2} u_m^* \end{pmatrix}, \\ \nabla_u^2 h(u | \beta, \sigma^2)|_{u=u^*} &= -\{Z^T W^* Z + \text{blkdiag}(\sigma_1^{-2} I_{q_1}, \dots, \sigma_m^{-2} I_{q_m})\}, \end{aligned}$$

where  $p^* = (p_1^*, \dots, p_n^*)^T$  with  $p_j^* = e^{\eta_j^*} / (1 + e^{\eta_j^*})$  and  $W^* = \text{diag}(w^*)$  is a

diagonal matrix with entries

$$w_j^* = p_j^*(1 - p_j^*) = \frac{e^{\eta_j^*}}{(1 + e^{\eta_j^*})^2}.$$

Therefore,

$$\begin{aligned} L_{\text{LA}}(\beta, \sigma) &= \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \sum_{i=1}^m q_i \ln \sigma_i^2 - \frac{n}{2} \ln 2\pi - \frac{1}{2} \sum_{i=1}^m \frac{\|u_i^*\|_2^2}{\sigma_i^2} \\ &\quad - \frac{1}{2} \ln \det \{Z^T W^* Z + \text{blkdiag}(\sigma_1^{-2} I_{q_1}, \dots, \sigma_m^{-2} I_{q_m})\}. \end{aligned} \quad (\text{S1.1})$$

Using the matrix determinant lemma, we have

$$\begin{aligned} &\ln \det \{Z^T W^* Z + \text{blkdiag}(\sigma_1^{-2} I_{q_1}, \dots, \sigma_m^{-2} I_{q_m})\} \\ &= \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) + \ln \det (\text{blkdiag}(\sigma_1^{-2} I_{q_1}, \dots, \sigma_m^{-2} I_{q_m})) + \ln \det W^* \\ &= \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) - \sum_{i=1}^m q_i \ln \sigma_i^2 + \ln \det W^*. \end{aligned} \quad (\text{S1.2})$$

Substitute (S1.2) to (S1.1) gives

$$\begin{aligned} L_{\text{LA}}(\beta, \sigma) &= \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \sum_{i=1}^m \frac{\|u_i^*\|_2^2}{\sigma_i^2} \\ &\quad - \frac{1}{2} \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) - \frac{1}{2} \ln \det W^* + \text{constant term}, \end{aligned}$$

where the constant term equals  $-\frac{n}{2} \ln 2\pi$ .

## S2 Derivation of approximated log-likelihood in (3.8)

$$L_{\text{LA}}(\beta, \sigma) = h(u^* | \beta, \sigma) - \frac{1}{2} \ln \det \nabla^2 \{-h(u^* | \beta, \sigma^2)\},$$

where

$$h(u \mid \beta, \sigma^2) = \sum_j \{y_j \eta_j - \ln(1 + e^{\eta_j})\} - \frac{1}{2} \|u\|_2^2 - \frac{n}{2} \ln 2\pi.$$

The gradient and Hessian  $h(u \mid \beta, \sigma)$  at  $u = u^*$  are

$$\begin{aligned} \nabla_u h(u \mid \beta, \sigma)|_{u=u^*} &= D^T Z^T (y - p^*) - u^*, \\ \nabla_u^2 h(u \mid \beta, \sigma)|_{u=u^*} &= -(D^T Z^T W^* Z D + I_q), \end{aligned}$$

where  $p^* = (p_1^*, \dots, p_n^*)^T$  with  $p_j^* = e^{\eta_j^*} / (1 + e^{\eta_j^*})$  and  $W^* = \text{diag}(w^*)$  is a diagonal matrix with entries

$$w_j^* = p_j^* (1 - p_j^*) = \frac{e^{\eta_j^*}}{(1 + e^{\eta_j^*})^2}.$$

Using the matrix determinant lemma, we have

$$\ln \det (D^T Z^T W^* Z D + I_q) = \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) + \ln \det W^*.$$

Therefore,

$$\begin{aligned} &L_{\text{LA}}(\beta, \sigma) \\ &= \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \|u^*\|_2^2 - \frac{n}{2} \ln 2\pi - \frac{1}{2} \ln \det (D^T Z^T W^* Z D + I_q) \\ &= \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \|u^*\|_2^2 - \frac{1}{2} \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) \\ &\quad - \frac{1}{2} \ln \det W^* + \text{constant term}, \end{aligned}$$

where the constant term equals  $-\frac{n}{2} \ln 2\pi$ .

### S3 Proof of ascent property in (3.6)

*Proof.* From (3.2) the approximated log-likelihood is

$$L_{\text{LA}}(\beta, \sigma) = \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \sum_{i=1}^m \frac{\|u_i^*\|_2^2}{\sigma_i^2} - \frac{1}{2} \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) - \frac{1}{2} \ln \det W^* + \text{terms without } \beta, \sigma^2,$$

where  $u^*$  is the maximizer of  $h(u \mid \beta, \sigma)$ ,  $\eta^* = X\beta + Zu^*$  and  $W^* = \text{diag}(w^*)$

is a diagonal matrix with entries

$$w_j^* = p_j^*(1 - p_j^*) = \frac{e^{\eta_j^*}}{(1 + e^{\eta_j^*})^2}.$$

Thus

$$L_{\text{LA}}(\sigma \mid \beta, u^*) = -\frac{1}{2} \sum_{i=1}^m \frac{\|u_i^*\|_2^2}{\sigma_i^2} - \frac{1}{2} \ln \det \left( W^{*-1} + \sum_i \sigma_i^2 Z_i Z_i^T \right) + c,$$

where  $c = \sum_j \{y_j \eta_j^* - \ln(1 + e^{\eta_j^*})\} - \frac{1}{2} \ln \det W^* - \frac{n}{2} \ln 2\pi$  is a constant not involving  $\sigma$ .

The minorization (2.4) leads to the surrogate function of  $L_{\text{LA}}(\sigma \mid \beta, u^*)$

$$g(\sigma^2 \mid \sigma^{2(t)}) = -\frac{1}{2} \sum_{i=1}^m \frac{\|u_i^*\|_2^2}{\sigma_i^2} - \frac{1}{2} \sum_{i=1}^m \sigma_i^2 \text{tr} \left\{ \left( \sum_i \sigma_i^{2(t)} Z_i Z_i^T + W^{*-1} \right)^{-1} Z_i Z_i^T \right\} + c^{(t)},$$

where  $c^{(t)}$  is a constant irrelevant to optimization. Since  $\sigma^{2(t+1)} = (\sigma_1^{2(t+1)}, \dots, \sigma_m^{2(t+1)})$

with

$$\sigma_i^{2(t+1)} = \left[ \frac{\|u_i^*\|_2^2}{\text{tr} \left\{ Z_i^T (\sum_i \sigma_i^{2(t)} Z_i Z_i^T + W^{*-1})^{-1} Z_i \right\}} \right]^{\frac{1}{2}}$$

maximizes the surrogate function  $g(\sigma^2 \mid \sigma^{2(t)})$ , we have the following inequality satisfied

$$L_{LA}(\sigma^{(t+1)} \mid \beta, u^*) \geq g(\sigma^{2(t+1)} \mid \sigma^{2(t)}) \geq g(\sigma^{2(t)} \mid \sigma^{2(t)}) = L_{LA}(\sigma^{(t)} \mid \beta, u^*).$$

Therefore, the iterates possess the ascent property.  $\square$