

## A NOTE ON MINIMAL ORTHOGONAL PLANS

Fang Dong

*National University of Singapore*

*Abstract:* Sufficient conditions are given for plans, which permit the uncorrelated estimation of all main effects and some specified interaction effects, to have a minimal number of experimental runs.

*Key words and phrases:* Fractional factorial design, orthogonal main effect plan, resolution-4 plan.

### 1. Introduction

Orthogonal main effect plans (OMEPE's) are plans which allow for the uncorrelated estimation of all main effects, provided that all interactions can be validly assumed to be negligible. In practice, such an assumption may not always be realistic. When an experimenter is not prepared to assume the absence of *all* interactions, a plan which permits estimation of all main effects and all or some specified interaction effects is required. When experimental runs are expensive or time consuming, it is important to know the minimal number of experimental runs necessary to construct such a plan for a given number of factors having a specified number of levels. This problem has been considered by Webb (1968) for resolution-4 plans for  $2^n$  experiments, Margolin (1969a) for resolution-4 plans for  $2^n 3^m$  experiments, Margolin (1969b) for general resolution-4 plans, and Jacroux (1992) for orthogonal main effect plans. For resolution-4, -5 and general resolution- $r$  plans, readers are referred to Box and Hunter (1961) and Webb (1965). Briefly, a resolution-4 plan is a plan which permits the estimation of all main effects under the assumption that interactions involving three or more factors are negligible, and a resolution-5 plan is a plan which permits the estimation of all main effects and all two factor interactions under the assumption that interactions involving three or more factors are negligible.

In this article, plans which permit the uncorrelated estimation of all main effects and all  $i$ -factor interaction effects among  $r$  ( $i < r$ ) specific factors are considered. They include the OMEPE's (when  $r = 2$ ) and some of the class one compromise plans (when  $r = 3$ ) introduced in Addelman (1962b). These plans are more flexible and easier to construct than the orthogonal resolution-4 or -5 plans. They can usually include more factors than an orthogonal resolution-4

or -5 plan with the same number of experimental runs. Sufficient conditions for these plans to have a minimal number of experimental runs are derived.

## 2. Main Results

For a given plan, the following notations will be used throughout this article:  $N$  = total number of observations to be taken;  $n(p_1, \dots, p_m; i_1, \dots, i_m)$  = number of observations to be taken at the  $i_1$ th level of factor  $p_1, \dots$ , the  $i_m$ th level of factor  $p_m$ . A plan is minimal if it has a minimal number of experimental runs.

Assume that a plan is to be constructed with  $k \geq r + 1$  factors which permits the uncorrelated estimation of all main effects and all  $i$ -factor ( $i < r$ ) interaction effects among factors  $p_1, \dots, p_r$ . It will be called an  $O(p_1, \dots, p_r)$  plan. It is an OMEP when  $r = 2$  and a class one compromise plan of Addelman (1962b) when  $r = 3$ . When  $r > 2$ , an  $O(p_1, \dots, p_r)$  plan is weaker than an orthogonal resolution- $(r + 1)$  plan. It is obvious that an orthogonal resolution-4 plan is an  $O(p_1, p_2, p_3)$  plan for any three of its factors, and an orthogonal resolution-5 plan is an  $O(p_1, p_2, p_3, p_4)$  plan for any four of its factors. An  $O(p_1, \dots, p_r)$  plan can usually include more factors than an orthogonal resolution- $(r + 1)$  plan with the same number of experimental runs. For example, using the lower bound given in Margolin (1969b), we know that a minimal resolution-4 plan (orthogonal or not) for a  $4^6$  experiment has 64 runs whereas an  $O(p_1, p_2, p_3)$  plan with the same number of runs can include 12 factors of four levels each (Addelman (1962b)). Another advantage for  $O(p_1, \dots, p_r)$  plans is that they remain as  $O(p_1, \dots, p_r)$  plans under collapsing (Addelman (1962a)); hence making it very easy to construct  $O(p_1, \dots, p_r)$  plans from existing  $O(p_1, \dots, p_r)$  plans and from orthogonal resolution- $(r + 1)$  plans.

We now introduce a method of constructing  $O(p_1, \dots, p_r)$  plans from existing  $O(p_1, \dots, p_r)$  plans. It is related to some of the methods given in Addelman (1962a) and Jacroux (1992). Let  $c_j, j = 1, \dots, k$  be  $k$  positive integers and  $d_0$  be an  $O(p_1, \dots, p_r)$  plan with  $N$  experimental runs for  $k \geq r + 1$  factors in which factor  $j$  has  $u_j$  levels,  $j = 1, \dots, k$ . Then an  $O(p_1, \dots, p_r)$  plan  $d$  with  $N \prod_{j=1}^k c_j$  experimental runs for  $k$  factors in which factor  $j$  has  $v_j \leq c_j u_j$  levels,  $j = 1, \dots, k$ , can be constructed as follows: Denote the levels of factor  $j$  of  $d_0$  by  $0, 1, \dots, u_j - 1$ ,  $j = 1, \dots, k$ . From  $d_0$ , we obtain the plans  $d_{i_1, \dots, i_k}$ ,  $i_j = 0, 1, \dots, c_j - 1$ ,  $j = 1, \dots, k$ , by replacing the levels of factor  $j$  in  $d_0$  with levels  $i_j u_j, i_j u_j + 1, \dots, i_j u_j + (u_j - 1)$  for  $j = 1, \dots, k$ . Let  $d_c$  be the plan obtained by combining the runs in  $\{d_{i_1, \dots, i_k}, i_j = 0, 1, \dots, c_j - 1, j = 1, \dots, k\}$ , and let  $d$  be the plan obtained from  $d_c$  by collapsing the levels in  $d_c$  to the required number of levels in  $d$  (for the collapsing method, see Addelman (1962a)), then  $d$  is the desired plan.

From Table 1 in Addelman (1962b), we know that for experiments  $2^4$ ,  $3^7$ ,  $4^{12}$  and  $5^{19}$ , there exist  $O(p_1, p_2, p_3)$  plans in 8, 27, 64 and 125 runs respectively. Using the construction method mentioned above, many  $O(p_1, \dots, p_r)$  plans can be constructed. For example, by the  $O(p_1, p_2, p_3)$  plan in 8 runs, an  $O(p_1, p_2, p_3)$  plan  $d$  can be constructed for experiment  $4 \cdot 3 \cdot 2^2$  in 32 runs where  $p_1, p_2$  are the factors with 4 and 3 levels and  $p_3$  is a factor with 2 levels. Let  $d_0$  be the  $O(p_1, p_2, p_3)$  plan in 8 runs. We obtain the plans  $d_{i_1, i_2}$ ,  $i_j = 0, 1$ ,  $j = 1, 2$ , by replacing the levels of factor  $j$  in  $d_0$  with levels  $2i_j, 2i_j + 1$  for  $j = 1, 2$ . Let  $d_c$  be the plan obtained by combining the runs in  $\{d_{i_1, i_2}, i_j = 0, 1, j = 1, 2\}$ , then  $d$  can be obtained from  $d_c$  by collapsing the levels in  $d_c$  to the required number of levels in  $d$ . Similarly, there exist  $O(p_1, p_2, p_3)$  plans, for example, for experiment  $5 \cdot 3 \cdot 2^5$  in 54 runs where  $p_1, p_2$  are the factors with 5 and 3 levels and  $p_3$  is a factor with 2 levels; for experiment  $4 \cdot 3^2 \cdot 2^9$  in 64 runs where  $p_1$  is the factor with 4 levels and  $p_2, p_3$  are the factors with 3 levels; and for experiment  $5 \cdot 4^2 \cdot 3^4 \cdot 2^{12}$  in 125 runs where  $p_1$  is the factor with 5 levels and  $p_2, p_3$  are the factors with 4 levels. Using Theorem 1 which will be stated and proved shortly, we see that all these plans are minimal  $O(p_1, p_2, p_3)$  plans.

Using the definition of the least common multiple and the greatest common divisor, the following result can be obtained easily.

**Lemma 1.** *Let  $x_1, \dots, x_n$  be  $n$  positive integers and  $x = \prod_{i=1}^n x_i$ . Then  $x = KL$  where  $K$  is the least common multiple for  $x/x_1, \dots, x/x_n$ , and  $L$  is the greatest common divisor for  $x_1, \dots, x_n$ .*

Now we state the main result.

**Theorem 1.** *suppose  $d$  is an  $O(p_1, \dots, p_r)$  plan with  $N$  experimental runs for  $k \geq r + 1$  factors in which factor  $i$  has  $s_i$  levels,  $i = 1, \dots, k$  and let  $s_e = \max_{j \neq p_1, \dots, p_r} s_j$ . If  $N = \prod_{i=1}^r t_i$  for  $t_1, \dots, t_r$  satisfying*

$$\prod_{i=1}^r t_i = \min \left\{ \prod_{i=1}^r x_i; x_i \geq s_{p_i}, i = 1, \dots, r, g(x_1, \dots, x_r) \geq s_e, \prod_{i=1}^r x_i < 2 \prod_{i=1}^r s_{p_i} \right\},$$

where  $g(x_1, \dots, x_r)$  is the greatest common divisor for  $x_1, \dots, x_r$ , then  $d$  is a minimal  $O(p_1, \dots, p_r)$  plan.

**Proof.** A necessary and sufficient condition for a plan to be an  $O(p_1, \dots, p_r)$  plan is that for any integers  $h_1, \dots, h_r \leq k$ ,  $h_i \neq h_j$  when  $i \neq j$ , at least  $r - 1$  of them belong to the set  $\{p_1, \dots, p_r\}$ , and  $i_j \leq s_{h_j}$ ,  $j = 1, \dots, r$ , we have  $N^{r-1}n(h_1, \dots, h_r; i_1, \dots, i_r) = \prod_{j=1}^r n(h_j; i_j)$  (for example, see Addelman (1962b)). Therefore for any  $h_1, \dots, h_{r-1}$  which belong to the set  $\{p_1, \dots, p_r\}$ , and  $i_j \leq s_{h_j}$ ,  $j = 1, \dots, r - 1$ , we have  $N^{r-2}n(h_1, \dots, h_{r-1}; i_1, \dots, i_{r-1}) = \prod_{j=1}^{r-1} n(h_j; i_j)$ . For an  $O(p_1, \dots, p_r)$  plan,  $n(p_1, \dots, p_r; i_1, \dots, i_r) \geq 1$  for any  $i_j \leq s_{p_j}$ ,  $j = 1, \dots, r$ .

Since  $N < 2 \prod_{i=1}^r s_{p_i}$ , there exist positive integers  $a_i$ ,  $i = 1, \dots, r$  such that  $N^{-(r-1)} \prod_{j=1}^r n(p_j; a_j) = n(p_1, \dots, p_r; a_1, \dots, a_r) = 1$ . Let

$$y_i = n(p_1, \dots, p_{i-1}, p_{i+1}, \dots, p_r; a_1, \dots, a_{i-1}, a_{i+1}, \dots, a_r), \quad i = 1, \dots, r.$$

Then  $y_i = \sum_{j=1}^{s_{p_i}} n(p_1, \dots, p_r; a_1, \dots, a_{i-1}, j, a_{i+1}, \dots, a_r) \geq s_{p_i}$ ,  $i = 1, \dots, r$  and  $\prod_{i=1}^r y_i = N^{-r(r-2)} (\prod_{j=1}^r n(p_j; a_j))^{r-1} = N^{-r(r-2)} N^{(r-1)^2} = N$ . Furthermore for any  $i \leq s_e$  and  $j \leq r$  we have

$$\begin{aligned} n(e; i) &= N^{-(r-1)} n(e; i) \prod_{l=1}^r n(p_l; a_l) \\ &= n(p_j; a_j) n(p_1, \dots, p_{j-1}, e, p_{j+1}, \dots, p_r; a_1, \dots, a_{j-1}, i, a_{j+1}, \dots, a_r). \end{aligned}$$

Hence  $n(e; i)$  is a multiple of  $n(p_j; a_j)$ ,  $j = 1, \dots, r$ . Therefore  $N = \sum_{i=1}^{s_e} n(e; i) \geq s_e K$  where  $K$  is the least common multiple for  $n(p_j; a_j)$ ,  $j = 1, \dots, r$ . By Lemma 1,  $L \geq s_e$  where  $L$  is the greatest common divisor for  $y_i$ ,  $i = 1, \dots, r$ . Since  $N = \prod_{i=1}^r t_i$  where  $t_i$ ,  $i = 1, \dots, r$  satisfy the condition in Theorem 1,  $d$  is minimal. This completes the proof.

The following result is a direct consequence of Theorem 1.

**Corollary 1.** *Let  $d$  and  $s_e$  be as in Theorem 1. If  $s_e^r = N < 2 \prod_{i=1}^r s_{p_i}$  and  $s_{p_i} \leq s_e$ ,  $i = 1, \dots, r$  then  $d$  is minimal.*

### 3. Discussion

In this article, we considered certain plans which permit the uncorrelated estimation of all main effects and some specified interaction effects. They are more flexible and easier to construct than orthogonal resolution-4 or -5 plans. Sufficient conditions are given for these plans to have a minimal number of experimental runs. This minimal number is useful to the experimenters for constructing minimal orthogonal designs. For example, for a  $10 \cdot 7 \cdot 4^2 \cdot 3^8$  experiment, a plan which permits the uncorrelated estimation of all main effects and two-factor interactions among the 10-, 7- and a 3-level factors requires at least 384 experimental runs. If this is considered too expensive or time consuming to perform, then either some factors should be dropped, or the levels of some factors should be redefined, or a non-orthogonal plan should be considered.

### Acknowledgement

I thank the associate editor and the referees for their helpful comments.

### References

- Addelman, S. (1962a). Orthogonal main-effect plans for asymmetrical factorial experiments. *Technometrics* 4, 21-46.

- Addelman, S. (1962b). Symmetrical and asymmetrical fractional factorial plans. *Technometrics* **4**, 47-58.
- Box, G. E. P. and Hunter, J. S. (1961). The  $2^{k-p}$  fractional factorial designs, part I. *Technometrics* **3**, 311-351.
- Jacroux, M. (1992). A note on the determination and construction of minimal orthogonal main-effect plans. *Technometrics* **34**, 92-96.
- Margolin, B. H. (1969a). Results on factorial designs of resolution-IV for the  $2^n$  and  $2^n \cdot 3^m$  series. *Technometrics* **11**, 431-444.
- Margolin, B. H. (1969b). Resolution-IV fractional factorial designs. *J. Roy. Statist. Soc. Ser.B* **31**, 514-523.
- Webb, S. (1965). *Design, testing and estimation in complex experimentation*: Part I. Expansible and contractible factorial designs and the application of linear programming to combinatorial problems. Aerospace Research Laboratories Technical Document 65-116.
- Webb, S. (1968). Non-orthogonal designs of even resolution. *Technometrics* **10**, 291-299.

Department of Mathematics, National University of Singapore, 10 Kent Ridge Crescent, Singapore 0511.

E-mail: matdongf@nus.sg

(Received December 1995; accepted June 1996)