

**FEATURE-WEIGHTED ELASTIC NET: USING  
FEATURES OF FEATURES” FOR  
BETTER PREDICTION**

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**Supplementary Material**

The online supplementary materials provide (i) details on an alternative algorithm with  $\theta$  as a parameter, (ii) the proof for Theorem 1, (iii) details on the simulation study in Section 5, and (iv) details on the simulation study in Section 7.

## S1 Alternative algorithm with $\theta$ as a parameter

Assume that  $\mathbf{y}$  and the columns of  $\mathbf{X}$  are centered so that  $\hat{\beta}_0 = 0$  and we can ignore the intercept term in the rest of the discussion. If we consider  $\theta$  as an argument of the objective function, then we wish to solve

$$\begin{aligned} (\hat{\beta}, \hat{\theta}) &= \operatorname{argmin}_{\beta, \theta} J_{\lambda, \alpha}(\beta, \theta) \\ &= \operatorname{argmin}_{\beta, \theta} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda \sum_{j=1}^p w_j(\theta) \left[ \alpha |\beta_j| + \frac{1 - \alpha}{2} \beta_j^2 \right]. \end{aligned}$$

$J$  is not jointly convex  $\beta$  and  $\theta$ , so reaching a global minimum is a difficult task. Instead, we content ourselves with reaching a local minimum. A reasonable approach for doing so is to alternate between optimizing  $\beta$  and  $\theta$ : the steps are outlined in Algorithm 2.

Unfortunately, Algorithm 2 is slow due to repeated solving of the elastic net problem in Step 2(b)ii for each  $\lambda_i$ . The algorithm does not take advantage of the fact that once  $\alpha$  and  $\theta$  are fixed, the elastic net problem can be solved quickly for an entire path of  $\lambda$  values. We have also found that Algorithm 2 does not predict as well as Algorithm 1 in our simulations.

## S2 Proof of Theorem 1

For the moment, consider the more general penalty factor  $w_j(\theta) = \frac{\sum_{\ell=1}^p f(\mathbf{z}_\ell^T \theta)}{p f(\mathbf{z}_j^T \theta)}$ ,

where  $f$  is some function with range  $[0, +\infty)$ . (Fwelnet makes the choice

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**Algorithm 2** *Minimizing the fwelnet objective function via alternating minimization*

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1. Select a value of  $\alpha \in [0, 1]$  and a sequence of  $\lambda$  values  $\lambda_1 > \dots > \lambda_m$ .
  2. For  $i = 1, \dots, m$ :
    - (a) Initialize  $\beta^{(0)}(\lambda_i)$  at the elastic net solution for  $\lambda_i$ . Initialize  $\theta^{(0)} = \mathbf{0}$ .
    - (b) For  $k = 0, 1, \dots$  until convergence:
      - i. Fix  $\beta = \beta^{(k)}$ , update  $\theta^{(k+1)}$  via gradient descent. That is, set 
$$\Delta\theta = \frac{\partial J_{\lambda_i, \alpha}}{\partial \theta} \Big|_{\beta=\beta^{(k)}, \theta=\theta^{(k)}} \quad \text{and update } \theta^{(k+1)} = \theta^{(k)} - \eta \Delta\theta,$$
 where  $\eta$  is the step size computed via backtracking line search to ensure that  $J_{\lambda_i, \alpha}(\beta^{(k)}, \theta^{(k+1)}) < J_{\lambda_i, \alpha}(\beta^{(k)}, \theta^{(k)})$ .
      - ii. Fix  $\theta = \theta^{(k+1)}$ , update  $\beta^{(k+1)}$  by solving the elastic net with updated penalty factors  $w_j(\theta^{(k+1)})$ .
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$f(x) = e^x$ .)

First note that if feature  $j$  belongs to group  $k$ , then  $\mathbf{z}_j^T \theta = \theta_k$ , and its penalty factor is

$$w_j(\theta) = \frac{\sum_{\ell=1}^p f(\mathbf{z}_\ell^T \theta)}{p f(\mathbf{z}_j^T \theta)} = \frac{\sum_{\ell=1}^p f(\theta_\ell)}{p f(\theta_k)} = \frac{\sum_{\ell=1}^K p_\ell f(\theta_\ell)}{p f(\theta_k)},$$

where  $p_\ell$  denotes the number of features in group  $\ell$ . Letting  $v_k = \frac{f(\theta_k)}{\sum_{\ell=1}^K p_\ell f(\theta_\ell)}$  for  $k = 1, \dots, K$ , minimizing the fwelnet objective function (3.2) over  $\beta$  and  $\theta$  reduces to

$$\text{minimize}_{\beta, \theta} \quad \frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \frac{\lambda}{p} \sum_{k=1}^K \frac{1}{v_k} \left[ \alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2 \right].$$

For fixed  $\beta$ , we can explicitly determine the  $v_k$  values which minimize the expression above. By the Cauchy-Schwarz inequality,

$$\begin{aligned} & \frac{\lambda}{p} \sum_{k=1}^K \frac{1}{v_k} \left[ \alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2 \right] \\ &= \frac{\lambda}{p} \left( \sum_{k=1}^K \frac{1}{v_k} \left[ \alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2 \right] \right) \left( \sum_{k=1}^K p_k v_k \right) \\ &\geq \frac{\lambda}{p} \left( \sum_{k=1}^K \sqrt{p_k \left[ \alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2 \right]} \right)^2. \end{aligned} \quad (\text{S2.1})$$

Note that equality is attainable for (S2.1): letting  $a_k = \sqrt{\frac{[\alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2]}{p_k}}$ ,

equality occurs when there is some  $c \in \mathbb{R}$  such that

$$\begin{aligned} c \cdot \frac{1}{v_k} \left[ \alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2 \right] &= p_k v_k && \text{for all } k, \\ v_k &= \sqrt{c} a_k && \text{for all } k. \end{aligned}$$

Since  $\sum_{k=1}^K p_k v_k = 1$ , we have  $\sqrt{c} = \frac{1}{\sum_{k=1}^K p_k a_k}$ , giving  $v_k = \frac{a_k}{\sum_{k=1}^K p_k a_k}$  for all  $k$ . A solution for this is  $f(\theta_k) = a_k$  for all  $k$ , which is feasible for  $f$  having range  $[0, \infty)$ . (Note that if  $f$  only has range  $(0, \infty)$ , the connection still holds if  $\lim_{x \rightarrow -\infty} f(x) = 0$  or  $\lim_{x \rightarrow +\infty} f(x) = 0$ : the solution will just have  $\theta = +\infty$  or  $\theta = -\infty$ .)

Thus, the fwelnet solution is

$$\operatorname{argmin}_{\beta} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \frac{\lambda}{p} \left( \sum_{k=1}^K \sqrt{p_k} \left[ \alpha \|\beta^{(k)}\|_1 + \frac{1-\alpha}{2} \|\beta^{(k)}\|_2^2 \right] \right)^2. \quad (\text{S2.2})$$

When  $\alpha = 0$ , the penalty term is convex. Writing in constrained form, (S2.2) becomes minimizing  $\frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2$  subject to

$$\begin{aligned} \left( \sum_{k=1}^K \sqrt{p_k} \|\beta^{(k)}\|_2 \right)^2 &\leq C \text{ for some constant } C, \\ \sum_{k=1}^K \sqrt{p_k} \|\beta^{(k)}\|_2 &\leq \sqrt{C}. \end{aligned}$$

Converting back to Lagrange form again, there is some  $\lambda' \geq 0$  such that the fwelnet solution is

$$\operatorname{argmin}_{\beta} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\beta\|_2^2 + \lambda' \sum_{k=1}^K \sqrt{p_k} \|\beta^{(k)}\|_2.$$

## S3 Details on simulation study in Section 5

### S3.1 Setting 1: Noisy version of the true $\beta$

1. Set  $n = 100$ ,  $p = 50$ ,  $\beta \in \mathbb{R}^{50}$  with  $\beta_j = 2$  for  $j = 1, \dots, 5$ ,  $\beta_j = -1$  for  $j = 6, \dots, 10$ , and  $\beta_j = 0$  otherwise.
2. Generate  $x_{ij} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$  for  $i = 1, \dots, n$  and  $j = 1, \dots, p$ .
3. For each  $SNR_y \in \{0.5, 1, 2\}$  and  $SNR_Z \in \{0.5, 2, 10\}$ :
  - (a) Compute  $\sigma_y^2 = \left(\sum_{j=1}^p \beta_j^2\right) / SNR_y$ .
  - (b) Generate  $y_i = \sum_{j=1}^p x_{ij}\beta_j + \varepsilon_i$ , where  $\varepsilon_i \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma_y^2)$  for  $i = 1, \dots, n$ .
  - (c) Compute  $\sigma_Z^2 = \text{Var}(|\beta|) / SNR_Z$ .
  - (d) Generate  $z_j = |\beta_j| + \eta_j$ , where  $\eta_j \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma_Z^2)$ . Treat this as a column matrix to get  $\mathbf{Z} \in \mathbb{R}^{p \times 1}$ .

### S3.2 Setting 2: Grouped data setting

1. Set  $n = 100$ ,  $p = 150$ .
2. For  $j = 1, \dots, p$  and  $k = 1, \dots, 15$ , set  $z_{jk} = 1$  if  $10(k-1) < j \leq 10k$ ,  $z_{jk} = 0$  otherwise.

3. Generate  $\beta \in \mathbb{R}^{150}$  with  $\beta_j = 3$  or  $\beta_j = -3$  with equal probability for  $j = 1, \dots, 10G$ ,  $\beta_j = 0$  otherwise.  $G = 1$  for the first scenario where the response depends on the first group only, and  $G = 4$  for the second scenario where it depends on the first 4 groups.
4. Generate  $x_{ij} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$  for  $i = 1, \dots, n$  and  $j = 1, \dots, p$ .
5. For each  $SNR_y \in \{0.5, 1, 2\}$ :
  - (a) Compute  $\sigma_y^2 = \left( \sum_{j=1}^p \beta_j^2 \right) / SNR_y$ .
  - (b) Generate  $y_i = \sum_{j=1}^p x_{ij} \beta_j + \varepsilon_i$ , where  $\varepsilon_i \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma_y^2)$  for  $i = 1, \dots, n$ .

### S3.3 Setting 3: Noise variables

1. Set  $n = 100$ ,  $p = 100$ ,  $\beta \in \mathbb{R}^{100}$  with  $\beta_j = 2$  for  $j = 1, \dots, 10$ , and  $\beta_j = 0$  otherwise.
2. Generate  $x_{ij} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$  for  $i = 1, \dots, n$  and  $j = 1, \dots, p$ .
3. For each  $SNR_y \in \{0.5, 1, 2\}$ :
  - (a) Compute  $\sigma_y^2 = \left( \sum_{j=1}^p \beta_j^2 \right) / SNR_y$ .
  - (b) Generate  $y_i = \sum_{j=1}^p x_{ij} \beta_j + \varepsilon_i$ , where  $\varepsilon_i \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma_y^2)$  for  $i = 1, \dots, n$ .

- (c) Generate  $z_{jk} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$  for  $j = 1, \dots, p$  and  $k = 1, \dots, 10$ . Append a column of ones to get  $\mathbf{Z} \in \mathbb{R}^{p \times 11}$ .

## S4 Details on simulation study in Section 7

1. Set  $n = 150$ ,  $p = 50$ .

2. Generate  $\beta_1 \in \mathbb{R}^{50}$  with

$$\beta_{1,j} = \begin{cases} 5 \text{ or } -5 \text{ with equal probability} & \text{for } j = 1, \dots, 5, \\ 2 \text{ or } -2 \text{ with equal probability} & \text{for } j = 6, \dots, 10, \\ 0 & \text{otherwise.} \end{cases}$$

3. Generate  $\beta_2 \in \mathbb{R}^{50}$  with

$$\beta_{2,j} = \begin{cases} 5 \text{ or } -5 \text{ with equal probability} & \text{for } j = 1, \dots, 5, \\ 2 \text{ or } -2 \text{ with equal probability} & \text{for } j = 11, \dots, 15, \\ 0 & \text{otherwise.} \end{cases}$$

4. Generate  $x_{ij} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$  for  $i = 1, \dots, n$  and  $j = 1, \dots, p$ .

5. Generate response 1,  $\mathbf{y}_1 \in \mathbb{R}^{150}$ , in the following way:

- (a) Compute  $\sigma_1^2 = \left( \sum_{j=1}^p \beta_{1,j}^2 \right) / 0.5$ .

- (b) Generate  $y_{1,i} = \sum_{j=1}^p x_{ij} \beta_{1,j} + \varepsilon_{1,i}$ , where  $\varepsilon_{1,i} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma_1^2)$  for  $i = 1, \dots, n$ .



6. Generate response 2,  $\mathbf{y}_2 \in \mathbb{R}^{150}$ , in the following way:

(a) Compute  $\sigma_2^2 = \left( \sum_{j=1}^p \beta_{2,j}^2 \right) / 1.5$ .

(b) Generate  $y_{2,i} = \sum_{j=1}^p x_{ij} \beta_{2,j} + \varepsilon_{2,i}$ , where  $\varepsilon_{2,i} \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma_2^2)$  for  $i = 1, \dots, n$ .