

NETWORK GARCH MODEL

Jing Zhou¹, Dong Li², Rui Pan³, and Hansheng Wang⁴

¹*Renmin University of China*, ²*Tsinghua University*,

³*Central University of Finance and Economics*, and ⁴*Peking University*

Supplementary Material

S.1 PROOF OF THEOREM 1

S.2 PROOF OF THEOREM 2

S1. Proof of Theorem 1

Since $\{\mathbf{B}_t\}$ is i.i.d., by the (1.4) in Kesten and Spitzer (1984) or Theorem 3.2 in Bougerol and Picard (1992), it suffices to prove that $\rho(E(\mathbf{B}_t)) < 1$, where $\rho(C)$ is the spectral radius of matrix C . Note that $E(\mathbf{B}_t) = (\beta_0 + \alpha_0)\mathbf{I}_N + \lambda_0\mathbf{D}^{-1}A$. Suppose that μ is any arbitrary eigenvalue of $\mathbf{D}^{-1}A$. Since $\mathbf{D}^{-1}A = (a_{ij}/d_i)$, by the Gershgorin circle theorem on the eigenvalues of matrices and by the definition of d_i , we can get that $|\mu| \leq 1$. Thus, the largest eigenvalue in modular of $E(\mathbf{B}_t)$ is smaller than $\alpha_0 + \lambda_0 + \beta_0$, i.e., $\rho(E(\mathbf{B}_t)) \leq \alpha_0 + \lambda_0 + \beta_0 < 1$. Therefore, the result holds.

S2. Proof of Theorem 2

To prove the main results conveniently, we first denote some notations and give some facts. Let K and ρ be generic constants taking many different values with $K > 0$ and $\rho \in (0, 1)$ in what following.

By the expressions of $\tilde{\sigma}_{it}^2(\theta)$ and $\sigma_{it}^2(\theta)$ in (2.6), we have the following facts

$$\begin{aligned}
 \tilde{\sigma}_{it}^2(\theta) &= \sum_{k=1}^t \beta^{k-1} \left\{ \omega + \alpha y_{i,t-k}^2 + \lambda d_i^{-1} \sum_{j \neq i} a_{ij} y_{j,t-k}^2 \right\}, \quad t \geq 1, \\
 \sigma_{it}^2(\theta) &= \sum_{k=1}^{\infty} \beta^{k-1} \left\{ \omega + \alpha y_{i,t-k}^2 + \lambda d_i^{-1} \sum_{j \neq i} a_{ij} y_{j,t-k}^2 \right\}, \\
 \sigma_{it}^2(\theta) &= \tilde{\sigma}_{it}^2(\theta) + \beta^t \sigma_{i0}^2(\theta), \\
 \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} &= \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} + t \beta^{t-1} \sigma_{i0}^2(\theta) \mathbf{e} + \beta^t \frac{\partial \sigma_{i0}^2(\theta)}{\partial \theta}, \\
 \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \theta \partial \theta'} &= \frac{\partial^2 \tilde{\sigma}_{it}^2(\theta)}{\partial \theta \partial \theta'} + t(t-1) \beta^{t-2} \sigma_{i0}^2(\theta) \mathbf{e} \mathbf{e}' \\
 &\quad + 2t \beta^{t-1} \frac{\partial \sigma_{i0}^2(\theta)}{\partial \theta} \mathbf{e}' + \beta^t \frac{\partial^2 \sigma_{i0}^2(\theta)}{\partial \theta \partial \theta'},
 \end{aligned} \tag{S2.1}$$

where $\mathbf{e} = (0, 0, 0, 1)'$, and

$$\begin{aligned}
 \sup_{\theta \in \Theta} \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta_k} &\leq K, \\
 \sup_{\theta \in \Theta} \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} &\leq K \sum_{j=1}^{\infty} \rho^{js} \left\{ \sup_{\theta \in \Theta} \sigma_{i,t-j}^2(\theta) \right\}^s,
 \end{aligned} \tag{S2.2}$$

where $\theta_k \in \{\omega, \alpha, \lambda\}$, by the inequality $x/(1+x) \leq x^s$ for $x > 0$ and

$s \in (0, 1]$.

Lemma 1. *If Assumptions 1–2 hold, then*

(i). $\lim_{T \rightarrow \infty} \sup_{\theta \in \Theta} |L(\theta) - \tilde{L}(\theta)| = 0$ a.s..

(ii). $E|\ell_t(\theta_0)| < \infty$, and $E\ell_t(\theta) \geq E\ell_t(\theta_0)$, where the equality holds if and only if $\theta = \theta_0$.

(iii). Any $\theta \neq \theta_0$ has a neighbourhood $V(\theta)$ such that

$$\liminf_{T \rightarrow \infty} \inf_{\theta^* \in V(\theta)} \tilde{L}(\theta^*) > E\ell_t(\theta_0).$$

PROOF. The proof is similar to that of Theorem 2.1 in Francq and Zakoïan (2004), and thus it is omitted. □

Lemma 2. *If the conditions in Theorem 2 hold, then*

- (i). $\sqrt{NT} \left\| \frac{\partial \tilde{L}(\theta_0)}{\partial \theta} - \frac{\partial L(\theta_0)}{\partial \theta} \right\| = o(1)$ a.s.;
- (ii). $\sup_{\|\theta - \theta_0\| < \eta} \left\| \frac{\partial^2 \tilde{L}(\theta)}{\partial \theta \partial \theta'} - \frac{\partial^2 L(\theta_0)}{\partial \theta \partial \theta'} \right\| = O_p(\eta)$;
- (iii). $\frac{\partial^2 L(\theta_0)}{\partial \theta \partial \theta'} \rightarrow \Sigma$;
- (iv). $\sqrt{NT} \frac{\partial L(\theta_0)}{\partial \theta} \rightarrow_d \mathcal{N}(0, (\kappa_4 - 1)\Sigma)$.

PROOF. (i). A simple calculation yields

$$\begin{aligned} \sqrt{NT} \left\| \frac{\partial \tilde{L}(\theta)}{\partial \theta} - \frac{\partial L(\theta)}{\partial \theta} \right\| &\leq \frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \left\| \frac{1}{\tilde{\sigma}_{it}^2(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \\ &\quad + \frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \left\| \frac{y_{it}^2}{\tilde{\sigma}_{it}^4(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{y_{it}^2}{\sigma_{it}^4(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\|. \end{aligned}$$

First,

$$\begin{aligned} &\frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \left\| \frac{1}{\tilde{\sigma}_{it}^2(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \\ &\leq \frac{K}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \left\| \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \\ &\quad + \frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \left\| [\sigma_{it}^2(\theta) - \tilde{\sigma}_{it}^2(\theta)] \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\|. \end{aligned}$$

By the facts in (S2.1), we have

$$\begin{aligned} &\frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \sup_{\theta \in \Theta} \left\| \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \\ &\leq \frac{1}{\sqrt{N}} \sum_{i=1}^N \sup_{\theta \in \Theta} \sigma_{i0}^2(\theta) \left(\frac{1}{\sqrt{T}} \sum_{t=1}^T t \rho^{t-1} \right) \\ &\quad + \frac{1}{\sqrt{N}} \sum_{i=1}^N \sup_{\theta \in \Theta} \left\| \frac{\partial \sigma_{i0}^2(\theta)}{\partial \theta} \right\| \left(\frac{1}{\sqrt{T}} \sum_{t=1}^T \rho^t \right) \end{aligned}$$

$\rightarrow 0$ a.s.

since $\sup_{\theta \in \Theta} \|\partial \sigma_{i0}^2(\theta) / \partial \theta\| < \infty$ a.s. Further,

$$\begin{aligned} & \frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \sup_{\theta \in \Theta} \left\| [\sigma_{it}^2(\theta) - \tilde{\sigma}_{it}^2(\theta)] \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \\ & \leq \frac{1}{\sqrt{N}} \sum_{i=1}^N \sup_{\theta \in \Theta} \sigma_{i0}^2(\theta) \left(\frac{1}{\sqrt{T}} \sum_{t=1}^T \rho^t \sup_{\theta \in \Theta} \left\| \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \right) \rightarrow 0 \quad \text{a.s.} \end{aligned}$$

since, by (S2.2),

$$\sum_{t=1}^{\infty} P \left(\rho^t \sup_{\theta \in \Theta} \left\| \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| > \epsilon \right) \leq \sum_{t=1}^{\infty} \frac{\rho^t}{\epsilon} E \sup_{\theta \in \Theta} \left\| \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| < \infty,$$

for any $\epsilon > 0$. Thus,

$$\frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \sup_{\theta \in \Theta} \left\| \frac{1}{\tilde{\sigma}_{it}^2(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \rightarrow 0 \quad \text{a.s.}$$

Similarly, we can prove that

$$\frac{1}{\sqrt{NT}} \sum_{i=1}^N \sum_{t=1}^T \sup_{\theta \in \Theta} \left\| \frac{y_{it}^2}{\tilde{\sigma}_{it}^4(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \theta} - \frac{y_{it}^2}{\sigma_{it}^4(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \theta} \right\| \rightarrow 0 \quad \text{a.s.}$$

Therefore, (i) is proved.

(ii). We first show

$$\sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^2 \tilde{L}(\theta)}{\partial \beta^2} - \frac{\partial^2 L(\theta_0)}{\partial \beta^2} \right| = O_p(\eta).$$

Note that

$$\begin{aligned} \sup_{\|\theta-\theta_0\|<\eta} \left| \frac{\partial^2 \tilde{L}(\theta)}{\partial \beta^2} - \frac{\partial^2 L(\theta_0)}{\partial \beta^2} \right| &\leq \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{\partial^2 \tilde{\ell}_t(\theta)}{\partial \beta^2} - \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} \right| \\ &\quad + \frac{1}{T} \sum_{t=1}^T \sup_{\|\theta-\theta_0\|<\eta} \left| \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} - \frac{\partial^2 \ell_t(\theta_0)}{\partial \beta^2} \right|, \end{aligned}$$

and

$$\begin{aligned} \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} &= \frac{1}{N} \sum_{i=1}^N \left\{ \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} - \frac{1}{\sigma_{it}^4(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \right\} \\ &\quad - \frac{1}{N} \sum_{i=1}^N \left\{ \frac{y_{it}^2}{\sigma_{it}^4(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} - 2 \frac{y_{it}^2}{\sigma_{it}^6(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \right\}. \end{aligned}$$

Then

$$\begin{aligned} &\frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{\partial^2 \tilde{\ell}_t(\theta)}{\partial \beta^2} - \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} \right| \\ &\leq \frac{1}{N} \sum_{i=1}^N \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} - \frac{1}{\tilde{\sigma}_{it}^2(\theta)} \frac{\partial^2 \tilde{\sigma}_{it}^2(\theta)}{\partial \beta^2} \right| \\ &\quad + \frac{1}{N} \sum_{i=1}^N \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{1}{\sigma_{it}^4(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} - \frac{1}{\tilde{\sigma}_{it}^4(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \beta} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \beta} \right| \\ &\quad + \frac{1}{N} \sum_{i=1}^N \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{y_{it}^2}{\sigma_{it}^4(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} - \frac{y_{it}^2}{\tilde{\sigma}_{it}^4(\theta)} \frac{\partial^2 \tilde{\sigma}_{it}^2(\theta)}{\partial \beta^2} \right| \\ &\quad + \frac{2}{N} \sum_{i=1}^N \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{y_{it}^2}{\sigma_{it}^6(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} - \frac{y_{it}^2}{\tilde{\sigma}_{it}^6(\theta)} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \beta} \frac{\partial \tilde{\sigma}_{it}^2(\theta)}{\partial \beta} \right| \\ &:= A_1 + A_2 + A_3 + A_4. \end{aligned}$$

For A_1 , by the facts in (S2.1), we have

$$\begin{aligned}
A_1 &\leq \frac{1}{N} \sum_{i=1}^N \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{[\sigma_{it}^2(\theta) - \tilde{\sigma}_{it}^2(\theta)] \partial^2 \sigma_{it}^2(\theta)}{\tilde{\sigma}_{it}^2(\theta) \sigma_{it}^2(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} \right| \\
&\quad + \frac{K}{N} \sum_{i=1}^N \frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} - \frac{\partial^2 \tilde{\sigma}_{it}^2(\theta)}{\partial \beta^2} \right| \\
&\leq \frac{1}{N} \sum_{i=1}^N \sup_{\theta \in \Theta} \sigma_{i0}^2(\theta) \frac{1}{T} \sum_{t=1}^T \rho^t \sup_{\theta \in \Theta} \left| \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} \right| \\
&\quad + \frac{1}{N} \sum_{i=1}^N \sup_{\theta \in \Theta} \sigma_{i0}^2(\theta) \left(\frac{1}{T} \sum_{t=1}^T t^2 \rho^{t-2} \right) + \frac{2}{N} \sum_{i=1}^N \sup_{\theta \in \Theta} \left| \frac{\partial \sigma_{i0}^2(\theta)}{\partial \beta} \right| \left(\frac{1}{T} \sum_{t=1}^T t \rho^{t-1} \right) \\
&\quad + \frac{1}{N} \sum_{i=1}^N \sup_{\theta \in \Theta} \frac{\partial^2 \sigma_{i0}^2(\theta)}{\partial \beta^2} \left(\frac{1}{T} \sum_{t=1}^T \rho^t \right) \rightarrow 0 \quad \text{as } T \rightarrow \infty.
\end{aligned}$$

Similarly, we can prove that $A_i \rightarrow 0$ a.s. for $i = 2, 3, 4$. Thus,

$$\frac{1}{T} \sum_{t=1}^T \sup_{\theta \in \Theta} \left| \frac{\partial^2 \tilde{\ell}_t(\theta)}{\partial \beta^2} - \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} \right| \rightarrow 0 \quad \text{as } T \rightarrow \infty.$$

On the other hand, by the Taylor expansion, we have

$$E \sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} - \frac{\partial^2 \ell_t(\theta_0)}{\partial \beta^2} \right| \leq \eta E \sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^3 \ell_t(\theta)}{\partial \beta^3} \right|.$$

By a simple calculation, it follows that

$$\begin{aligned} \frac{\partial^3 \ell_t(\theta)}{\partial \beta^3} &= \frac{1}{N} \sum_{i=1}^N \left\{ 2 - \frac{6y_{it}^2}{\sigma_{it}^2(\theta)} \right\} \left[\frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \right]^3 \\ &+ \frac{1}{N} \sum_{i=1}^N \left\{ 1 - \frac{y_{it}^2}{\sigma_{it}^2(\theta)} \right\} \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial^3 \sigma_{it}^2(\theta)}{\partial \beta^3} \\ &+ \frac{1}{N} \sum_{i=1}^N \left\{ \frac{6y_{it}^2}{\sigma_{it}^2(\theta)} - 3 \right\} \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2}. \end{aligned}$$

Choosing small enough $\eta > 0$ and $\delta = (1 - \rho^s)/(2\rho^s)$, similar to (4.25) in

Francq and Zakoïan (2004), we have

$$\begin{aligned} \sup_{\|\theta - \theta_0\| < \eta} \frac{y_{it}^2}{\sigma_{it}^2(\theta)} &\leq K \varepsilon_{it}^2 \left\{ 1 + \sum_{k=0}^{\infty} (1 + \delta)^k \rho^{ks} \left(y_{i,t-k-1}^{2s} + d_i^{-s} \sum_{j \neq i} a_{ij} y_{j,t-k-1}^{2s} \right) \right\}, \\ \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial \sigma_{it}^2(\theta)}{\partial \beta} &\leq K \sum_{k=1}^{\infty} \rho^{ks} \left\{ \sup_{\theta \in \Theta} \sigma_{i,t-k}^2(\theta) \right\}^s, \\ \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial^2 \sigma_{it}^2(\theta)}{\partial \beta^2} &\leq K \sum_{j=1}^{\infty} \rho^{js} \left\{ \sup_{\theta \in \Theta} \sigma_{i,t-j}^2(\theta) \right\}^s \sup_{\theta \in \Theta} \frac{1}{\sigma_{i,t-j}^2(\theta)} \frac{\partial \sigma_{i,t-j}^2(\theta)}{\partial \beta}, \\ \frac{1}{\sigma_{it}^2(\theta)} \frac{\partial^3 \sigma_{it}^2(\theta)}{\partial \beta^3} &\leq K \sum_{j=1}^{\infty} \rho^{js} \left\{ \sup_{\theta \in \Theta} \sigma_{i,t-j}^2(\theta) \right\}^s \sup_{\theta \in \Theta} \frac{1}{\sigma_{i,t-j}^2(\theta)} \frac{\partial^2 \sigma_{i,t-j}^2(\theta)}{\partial \beta^2}, \end{aligned}$$

where $s \in (0, 1]$. By taking $s = 1/2$ and Hölder's inequality, it is not hard

to show that

$$E \sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^3 \ell_t(\theta)}{\partial \beta^3} \right| = O(1).$$

Thus, we have

$$\frac{1}{T} \sum_{t=1}^T \sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^2 \ell_t(\theta)}{\partial \beta^2} - \frac{\partial^2 \ell_t(\theta_0)}{\partial \beta^2} \right| = O_p(1),$$

and in turn

$$\sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^2 \tilde{L}(\theta)}{\partial \beta^2} - \frac{\partial^2 L(\theta_0)}{\partial \beta^2} \right| = O_p(\eta).$$

Similarly, we can show that

$$\sup_{\|\theta - \theta_0\| < \eta} \left| \frac{\partial^2 \tilde{L}(\theta)}{\partial \theta_i \partial \theta_j} - \frac{\partial^2 L(\theta_0)}{\partial \theta_i \partial \theta_j} \right| = O_p(\eta),$$

where $\theta_i, \theta_j \in \{\omega, \alpha, \lambda, \beta\}$. Thus, (ii) holds.

(iii). By a simple calculation, it follows that

$$\frac{\partial^2 \ell_t(\theta_0)}{\partial \theta \partial \theta'} = \frac{1}{N} \sum_{i=1}^N \left\{ (1 - \varepsilon_{it}^2) \frac{1}{h_{it}} \frac{\partial^2 \sigma_{it}^2(\theta_0)}{\partial \theta \partial \theta'} + (2\varepsilon_{it}^2 - 1) \frac{1}{h_{it}^2} \frac{\partial \sigma_{it}^2(\theta_0)}{\partial \theta} \frac{\partial \sigma_{it}^2(\theta_0)}{\partial \theta'} \right\}.$$

Thus (iii) holds by the strong law of large numbers.

(iv). Note that

$$\sqrt{NT} \frac{\partial L(\theta_0)}{\partial \theta} = \frac{1}{\sqrt{NT}} \sum_{t=1}^T \sum_{i=1}^N (1 - \varepsilon_{it}^2) \frac{1}{h_{it}} \frac{\partial \sigma_{it}^2(\theta_0)}{\partial \theta}.$$

By the martingale central limit theorem, we have

$$\sqrt{NT} \frac{\partial L(\theta_0)}{\partial \theta} \rightarrow_d \mathcal{N}(0, (\kappa_4 - 1)\Sigma).$$

Proof of Theorem 2. By a standard compactness argument, using Lemma 1, we can complete the proof of Theorem 2(i) and then omit it.

Further, by the Taylor expansion and Lemma 2, we have

$$0 = \frac{\partial \tilde{L}(\hat{\theta})}{\partial \theta} = \frac{\partial \tilde{L}(\theta_0)}{\partial \theta} + \frac{\partial^2 \tilde{L}(\theta^*)}{\partial \theta \partial \theta'} (\hat{\theta} - \theta_0),$$

where θ^* lies in $\hat{\theta}$ and θ_0 satisfying $\|\theta^* - \theta_0\| \leq \|\hat{\theta} - \theta_0\|$. Thus,

$$\begin{aligned} \sqrt{NT}(\hat{\theta} - \theta_0) &= -\left(\frac{\partial^2 \tilde{L}(\theta^*)}{\partial \theta \partial \theta'}\right)^{-1} \sqrt{NT} \frac{\partial \tilde{L}(\theta_0)}{\partial \theta} \\ &= -\left(\frac{\partial^2 L(\theta_0)}{\partial \theta \partial \theta'}\right)^{-1} \sqrt{NT} \frac{\partial L(\theta_0)}{\partial \theta} + o_p(1) \\ &\rightarrow_d \mathcal{N}(0, (\kappa_4 - 1)\Sigma^{-1}). \end{aligned}$$

Bibliography

Bougerol, P. and Picard, N. (1992). Stationarity of GARCH processes and of some nonnegative time series. *J. Econometrics* **52**, 115–127.

Francq, C. and Zakoïan, J.-M. (2004). Maximum likelihood estimation of pure GARCH and ARMA-GARCH processes. *Bernoulli* **10**, 605–637.

BIBLIOGRAPHY

Kesten, H. and Spitzer, F. (1984). Convergence in distribution of products of random matrices, *Z. Wahrsch. Verw. Gebiete* 67, 363–386.