

## Robust Estimation of Dispersion Parameter in Discretely Observed Diffusion Processes

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### Supplementary Material

In this supplementary note, some technical lemmas and the proofs of Theorems 2.1, 2.2, and 2.3 are provided. In particular, we deal with the case of  $\alpha > 0$  and refer to Kessler (1997) for the case of  $\alpha = 0$ . In what follows, we denote  $a_{i-1} = a(X_{t_{i-1}^n})$ ,  $b_{i-1}(\sigma) = b(X_{t_{i-1}^n}, \sigma)$  and  $c_{i-1}(\sigma) = b(X_{t_{i-1}^n}, \sigma)^2$  for notational simplicity. Also, we drop  $k_0$  and  $k_1$  in  $V_{n,i}^\alpha(\sigma; k_0, k_1)$  without confusion. Further, let

$$\begin{aligned} Z_i &= \frac{1}{\sqrt{h_n}}(W_{t_i^n} - W_{t_{i-1}^n}), \\ \Delta_i &= \int_{t_{i-1}^n}^{t_i^n} \{a(X_s) - a(X_{t_{i-1}^n})\} ds + \int_{t_{i-1}^n}^{t_i^n} \{b(X_s, \sigma_0) - b(X_{t_{i-1}^n}, \sigma_0)\} dW_s, \\ \Delta_{1,i} &= \Delta_i - \frac{1}{2}b_{i-1}(\sigma_0)\partial_x b_{i-1}(\sigma_0)(Z_i^2 - 1)h_n, \end{aligned} \tag{1}$$

provided  $\partial_x b$  exists. We shall use the relation  $A_n \lesssim B_n$ , where  $A_n$  and  $B_n$  are nonnegative, to mean that  $A_n \leq CB_n$  for some constant  $C > 0$ . For example,  $A \lesssim 1$  means that  $A$  is bounded by some constant  $C$ .

## S1 Proof of Theorem 2.1

**Lemma 1.** *Suppose that A1 and A3 holds. Then, for  $k \geq 1$ ,*

$$\max_{i \leq n} E|\Delta_i|^{2k} \lesssim h_n^{2k}. \tag{S1.1}$$

*If, in addition,  $\partial_x^2 a$  and  $\partial_x^3 b$  exist and belong to  $\mathcal{P}$ , then*

$$\max_{i \leq n} E|\Delta_{1,i}|^{2k} \lesssim h_n^{3k}. \tag{S1.2}$$

*Proof.* By Lemma 6 of Kessler (1997), we have that for  $t_{i-1}^n \leq t \leq t_i^n$ ,

$$E|X_t - X_{t_{i-1}^n}|^k \lesssim h_n^{k/2}. \tag{S1.3}$$

Using this and Theorem 6.3 in Friedman (1975), one can prove (S1.1).

In order to show (S1.2), note that

$$\int_{t_{i-1}^n}^{t_i^n} \int_{t_{i-1}^n}^s dW_z dW_s = \frac{h_n}{2}(Z_i^2 - 1).$$

Then, by (1.17) in Chapter 5 of Kloeden and Platen (1999), we can write that

$$\begin{aligned}
\Delta_{1,i} &= \int_{t_{i-1}^n}^{t_i^n} \int_{t_{i-1}^n}^s L_\sigma a(X_z) dz ds + \int_{t_{i-1}^n}^{t_i^n} \int_{t_{i-1}^n}^s L^\sigma a(X_z) dW_z ds \\
&\quad + \int_{t_{i-1}^n}^{t_i^n} \int_{t_{i-1}^n}^s L_\sigma b(X_z, \sigma) dz dW_s + \int_{t_{i-1}^n}^{t_i^n} \int_{t_{i-1}^n}^s \int_{t_{i-1}^n}^z L_\sigma L^\sigma b(X_u, \sigma) du dW_z dW_s \\
&\quad + \int_{t_{i-1}^n}^{t_i^n} \int_{t_{i-1}^n}^s \int_{t_{i-1}^n}^z L^\sigma L^\sigma b(X_u, \sigma) dW_u dW_z dW_s \\
&:= D_{1,i} + D_{2,i} + D_{3,i} + D_{4,i} + D_{5,i},
\end{aligned}$$

where  $L_\sigma$  is the one given in (2.4) and  $L^\sigma = b(x, \sigma) \partial_x$ . We first deal with  $E|D_{4,i}|^{2k}$ . It follows from Theorem 6.3 in Friedman (1975) and Jensen's inequality that

$$\begin{aligned}
E|D_{4,i}|^{2k} &\lesssim h_n^{k-1} \int_{t_{i-1}^n}^{t_i^n} E \left\{ \int_{t_{i-1}^n}^s \int_{t_{i-1}^n}^z L_\sigma L^\sigma b(X_u, \sigma) du dW_z \right\}^{2k} ds \\
&\lesssim h_n^{k-1} \int_{t_{i-1}^n}^{t_i^n} h_n^{k-1} \int_{t_{i-1}^n}^s E \left\{ \int_{t_{i-1}^n}^z L_\sigma L^\sigma b(X_u, \sigma) du \right\}^{2k} dz ds \\
&\lesssim h_n^{k-1} \int_{t_{i-1}^n}^{t_i^n} h_n^{k-1} \int_{t_{i-1}^n}^s h_n^{2k-1} \int_{t_{i-1}^n}^z E|L_\sigma L^\sigma b(X_u, \sigma)|^{2k} du dz ds \lesssim h_n^{4k}. \quad (\text{S1.4})
\end{aligned}$$

Analogous to the above, it can be shown that

$$E|D_{1,i}|^{2k} \lesssim h_n^{4k}, \quad E|D_{2,i}|^{2k} \lesssim h_n^{3k}, \quad E|D_{3,i}|^{2k} \lesssim h_n^{3k} \quad \text{and} \quad E|D_{5,i}|^{2k} \lesssim h_n^{3k},$$

which together with (S1.4) yield (S1.2).  $\square$

**Lemma 2.** *Suppose that **A1** and **A3** hold. Then, if  $f : \mathbb{R} \times \Theta \rightarrow \mathbb{R}$  belongs to  $\mathcal{P}$  and  $nh_n^p \rightarrow 0$  for some  $p > 1$ ,*

$$\sup_{\sigma \in \Theta} \max_{i \leq n} |f^j(X_{t_{i-1}^n}, \sigma) Z_i^k \Delta_i^l h_n^m| = o(1) \quad \text{a.s.},$$

where  $j, k, l \in \{0, 1, 2, \dots\}$  and  $m > -l$ .

*Proof.* Due to **A3** and (S1.1), we have that for any  $\epsilon, \kappa > 0$ ,

$$\begin{aligned}
\sum_{n=1}^{\infty} P\left(\max_{i \leq n} |(1 + |X_{t_{i-1}^n}|)^C Z_i^k \Delta_i^l h_n^m| > \epsilon\right) &\leq \sum_{n=1}^{\infty} \frac{n}{\epsilon^\kappa} \max_{i \leq n} E|(1 + |X_{t_{i-1}^n}|)^C Z_i^k \Delta_i^l h_n^m|^\kappa \\
&\lesssim \sum_{i=1}^{\infty} nh_n^{(l+m)\kappa} = \sum_{n=1}^{\infty} o(n^{1-(l+m)\kappa/p}).
\end{aligned}$$

By choosing  $\kappa$  such that  $1 - (l+m)\kappa/p < -1$ , the lemma is asserted.  $\square$

**Lemma 3.** *Suppose that **A1**–**A3** hold; let  $f(x, \sigma) \in \mathcal{P} : \mathbb{R} \times \Theta \rightarrow \mathbb{R}$  be differentiable with respect to  $x$  and  $\sigma$ . Then, if  $\partial_x f$ ,  $\partial_\sigma f$  and  $g : \mathbb{R} \rightarrow \mathbb{R}$  belong to  $\mathcal{P}$  and  $nh_n^p \rightarrow 0$  for some  $p > 1$ ,*

$$\sup_{\sigma \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n g(Z_i) f(X_{t_{i-1}^n}, \sigma) - E(g(Z)) \int f(x, \sigma) d\mu_0(x) \right| = o(1) \quad \text{a.s.}$$

*Proof.* First, we show that for each  $\sigma \in \Theta$ ,

$$\frac{1}{n} \sum_{i=1}^n f(X_{t_{i-1}^n}, \sigma) \xrightarrow{a.s.} \int f(x, \sigma) d\mu_0(x) \quad \text{as } n \rightarrow \infty. \quad (\text{S1.5})$$

In view of ergodic property, we have

$$\frac{1}{nh_n} \int_0^{nh_n} f(X_s, \sigma) ds \xrightarrow{a.s.} \int f(x, \sigma) d\mu_0(x) \quad \text{as } n \rightarrow \infty. \quad (\text{S1.6})$$

By using Jensen's inequality, Cauchy's inequality and (S1.3), we have that for any  $r > 0$ ,

$$\begin{aligned} & E \left| \frac{1}{n} \sum_{i=1}^n \frac{1}{h_n} \int_{t_{i-1}^n}^{t_i^n} \{f(X_{t_{i-1}^n}, \sigma) - f(X_s, \sigma)\} ds \right|^{2r} \\ & \leq \frac{1}{nh_n} \sum_{i=1}^n \int_{t_{i-1}^n}^{t_i^n} \left\{ E |X_{t_{i-1}^n} - X_s|^{4r} \right\}^{\frac{1}{2}} \left\{ \int_0^1 E |\partial_x f(X_s + u(X_{t_{i-1}^n} - X_s), \sigma)|^{4r} du \right\}^{\frac{1}{2}} ds \\ & = O(h_n^r) = o(n^{-r/p}). \end{aligned}$$

Hence, it follows that for any  $\epsilon > 0$  and  $r > p$ ,

$$\sum_{n=1}^{\infty} P \left( \left| \frac{1}{n} \sum_{i=1}^n \frac{1}{h_n} \int_{t_{i-1}^n}^{t_i^n} \{f(X_{t_{i-1}^n}, \sigma) - f(X_s, \sigma)\} ds \right| > \epsilon \right) \leq \sum_{n=1}^{\infty} o(n^{-r/p}) < \infty,$$

which together with (S1.6) establishes (S1.5).

Next, let  $S_i(\sigma) := \{g(Z_i) - E(g(Z))\}f(X_{t_{i-1}^n}, \sigma)$  and  $\mathcal{G}_i^n := \sigma\{W_s : s \leq t_i^n\}$ . Then,  $\{(S_i(\sigma), \mathcal{G}_i^n)\}_{i=1}^n$  becomes a martingale difference, and thus, using Burkholder's inequality and Jensen's inequality, we have that for any  $r > 1$ ,

$$E \left| \frac{1}{n} \sum_{i=1}^n S_i(\sigma) \right|^{2r} \leq \frac{1}{n^{r+1}} \sum_{i=1}^n E |S_i(\sigma)|^{2r} = O(n^{-r}).$$

Therefore, we have

$$\left| \frac{1}{n} \sum_{i=1}^n g(Z_i) f(X_{t_{i-1}^n}, \sigma) - E(g(Z)) \frac{1}{n} \sum_{i=1}^n f(X_{t_{i-1}^n}, \sigma) \right| = o(1) \quad a.s., \quad (\text{S1.7})$$

and so the point wise convergence in the lemma is asserted from (S1.5). To establish the uniform convergence, it suffices to show that  $\{n^{-1} \sum_{i=1}^n g(Z_i) f(X_{t_{i-1}^n}, \sigma)\}$  is equicontinuous with probability one. Observe that

$$\begin{aligned} & \sup_{\sigma_1, \sigma_2 \in \Theta} \frac{1}{|\sigma_1 - \sigma_2|} \frac{1}{n} \sum_{i=1}^n \left| g(Z_i) f(X_{t_{i-1}^n}, \sigma_1) - g(Z_i) f(X_{t_{i-1}^n}, \sigma_2) \right| \\ & \leq \sup_{\sigma \in \Theta} \frac{1}{n} \sum_{i=1}^n \left| g(Z_i) \partial_{\sigma} f(X_{t_{i-1}^n}, \sigma) \right| \lesssim \frac{1}{n} \sum_{i=1}^n g(Z_i) (1 + |X_{t_{i-1}^n}|^{2C}) \\ & = O(1) \quad a.s., \end{aligned}$$

where the convergence of the last line is due to (S1.5) and (S1.7). Hence, we can ensure the lemma by the compactness of  $\Theta$ .  $\square$

**Lemma 4.** *Suppose that **A1–A4** and **A6(i)** hold; let  $f(x, \sigma) \in \mathcal{P} : \mathbb{R} \times \Theta \rightarrow \mathbb{R}$  be differentiable with respect to  $x$  and  $\sigma$ . If  $\partial_x f$  and  $\partial_\sigma f$  belong to  $\mathcal{P}$  and  $nh_n^p \rightarrow 0$  for some  $p > 1$ , then for any  $k \in \{0, 1, 2, \dots\}$ ,*

$$\sup_{\sigma \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n f(X_{t_{i-1}^n}, \sigma) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2} - \frac{(2k)!}{2^k k!} \int \left( 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right)^{-\frac{2k+1}{2}} f(x, \sigma) d\mu_0(x) \right| = o(1) \quad a.s..$$

*Proof.* Let

$$\begin{aligned} T_i(\sigma) &:= f(X_{t_{i-1}^n}, \sigma) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2} - E \left[ f(X_{t_{i-1}^n}, \sigma) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2} \middle| \mathcal{G}_{i-1}^n \right] \\ &= f(X_{t_{i-1}^n}, \sigma) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2} - \frac{(2k)!}{2^k k!} \left( 1 + \alpha \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} \right)^{-\frac{2k+1}{2}} f(X_{t_{i-1}^n}, \sigma). \end{aligned}$$

For any  $r \in \{1, 2, \dots\}$ , since  $\max_{i \leq n} \sup_{\sigma \in \Theta} E|T_i(\sigma)|^{2r} \lesssim 1$  by **A3**, it follows from Burkholder's inequality and Jensen's inequality that

$$E \left| \frac{1}{n} \sum_{i=1}^n T_i(\sigma) \right|^{2r} \leq \frac{1}{n^{r+1}} \sum_{i=1}^n E|T_i(\sigma)|^{2r} = O(n^{-r}). \quad (\text{S1.8})$$

The point wise convergence is therefore proved by (S1.5) and (S1.8). Moreover, since

$$\begin{aligned} & \sup_{\sigma_1, \sigma_2 \in \Theta} \frac{1}{|\sigma_1 - \sigma_2|} \frac{1}{n} \sum_{i=1}^n \left| f(X_{t_{i-1}^n}, \sigma_1) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_1)} Z_i^2} - f(X_{t_{i-1}^n}, \sigma_2) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_2)} Z_i^2} \right| \\ & \leq \sup_{\sigma \in \Theta} \frac{1}{n} \sum_{i=1}^n \left| \partial_\sigma f(X_{t_{i-1}^n}, \sigma) + \frac{\alpha}{2} \frac{c_{i-1}(\sigma_0) \partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^2} Z_i^2 f(X_{t_{i-1}^n}, \sigma) \right| Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_2)} Z_i^2} \\ & \lesssim \frac{1}{n} \sum_{i=1}^n (1 + |X_{t_{i-1}^n}|^{2C}) (1 + Z_i^2) Z_i^{2k} = O(1) \quad a.s., \end{aligned}$$

$\{n^{-1} \sum_{i=1}^n f(X_{t_{i-1}^n}, \sigma) Z_i^{2k} e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2}\}$  is equicontinuous with probability one, and thus the uniform convergence is established.  $\square$

**Lemma 5.** *Suppose that **A1–A4** and **A6(i)** with some  $2k$  hold. For any  $k_0 \in \{1, \dots, k\}$  and  $k_1 \in \{0\} \cup \mathbb{N}$ , if  $nh_n^p \rightarrow 0$  for some  $p > 1$ , then*

$$\sup_{\sigma \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n V_{n,i}^\alpha(\sigma) - V(\sigma) \right| = o(1) \quad a.s.,$$

where

$$V(\sigma) := \int \frac{1}{c(x, \sigma)^{\frac{\alpha}{2}}} \left\{ \frac{1}{\sqrt{1 + \alpha}} - \left( 1 + \frac{1}{\alpha} \right) \left( 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right)^{-\frac{1}{2}} \right\} d\mu_0(x).$$

*Proof.* Since

$$X_{t_i^n} = X_{t_{i-1}^n} + a_{i-1} h_n + b_{i-1}(\sigma_0) Z_i \sqrt{h_n} + \Delta_i, \quad (\text{S1.9})$$

we can express that

$$\begin{aligned} X_{t_i^n} - r_{k_0}(h_n, X_{t_{i-1}^n}, \sigma) &= b_{i-1}(\sigma_0)Z_i\sqrt{h_n} + \Delta_i - \sum_{j=2}^{k_0} \frac{h_n^j}{j!} L_\sigma^i X_{t_{i-1}^n} \\ &:= b_{i-1}(\sigma_0)Z_i\sqrt{h_n} + J_i(\sigma), \end{aligned} \quad (\text{S1.10})$$

and therefore

$$\begin{aligned} &\frac{(X_{t_i^n} - r_{k_0}(h_n, X_{t_{i-1}^n}, \sigma))^2}{c_{i-1}(\sigma)h_n} \left\{ 1 + \sum_{j=1}^{k_1} h_n^j d_j(X_{t_{i-1}^n}, \sigma) \right\} \\ &= \left\{ \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2 + \frac{2b_{i-1}(\sigma_0)Z_i J_i(\sigma)\sqrt{h_n} + J_i(\sigma)^2}{c_{i-1}(\sigma)h_n} \right\} \left\{ 1 + \sum_{j=1}^{k_1} h_n^j d_j(X_{t_{i-1}^n}, \sigma) \right\} \\ &:= \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2 + K_i(\sigma). \end{aligned} \quad (\text{S1.11})$$

Note that by Lemma 2,

$$\max_{i \leq n} \sup_{\sigma \in \Theta} |K_i(\sigma)| = o(1) \quad a.s.. \quad (\text{S1.12})$$

Let

$$U_i^\alpha(\sigma) := \frac{1}{c_{i-1}(\sigma)^{\frac{\alpha}{2}}} \left[ \frac{1}{\sqrt{1+\alpha}} - \left(1 + \frac{1}{\alpha}\right) \exp\left(-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2\right) \right]. \quad (\text{S1.13})$$

Then, we have that

$$\begin{aligned} |V_{n,i}^\alpha(\sigma) - U_i^\alpha(\sigma)| &= \frac{1}{c_{i-1}(\sigma)^{\frac{\alpha}{2}}} \left| \frac{1}{\sqrt{1+\alpha}} \sum_{j=1}^{k_1} h_n^j d_j^\alpha(X_{t_{i-1}^n}, \sigma) \right. \\ &\quad \left. - \left(1 + \frac{1}{\alpha}\right) \left[ \left\{ 1 + \sum_{j=1}^{k_1} h_n^j d_j^\alpha(X_{t_{i-1}^n}, \sigma) \right\} e^{-\frac{\alpha}{2} K_i(\sigma)} - 1 \right] e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2} \right| \\ &\lesssim h_n (1 + |X_{t_{i-1}^n}|)^C (1 + e^{-\frac{\alpha}{2} K_i(\sigma)} + |e^{-\frac{\alpha}{2} K_i(\sigma)} - 1|), \end{aligned}$$

and thus it follows from Lemma 2 and (S1.12) that

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n \sup_{\sigma \in \Theta} |V_{n,i}^\alpha(\sigma) - U_i^\alpha(\sigma)| &\lesssim (1 + e^{\max_{i \leq n} \sup_{\sigma \in \Theta} \frac{\alpha}{2} |K_i(\sigma)|}) \max_{i \leq n} (1 + |X_{t_{i-1}^n}|)^C h_n \\ &\quad + \max_{i \leq n} \sup_{\sigma \in \Theta} |K_i(\sigma)| e^{\max_{i \leq n} \sup_{\sigma \in \Theta} \frac{\alpha}{2} |K_i(\sigma)|} \\ &= o(1) \quad a.s.. \end{aligned} \quad (\text{S1.14})$$

Moreover, by Lemmas 3 and 4, we have

$$\sup_{\sigma \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n U_i^\alpha(\sigma) - V(\sigma) \right| = o(1) \quad a.s..$$

Combining the above and (S1.14), we obtain the lemma.  $\square$

**Proof of Theorem 2.1.** Note that

$$\left(\frac{1}{x}\right)^{\alpha/2} \left\{ \frac{1}{\sqrt{1+\alpha}} - \left(1 + \frac{1}{\alpha}\right) \left(1 + \alpha \frac{x_0}{x}\right)^{-1/2} \right\}, \quad x > 0, x_0 > 0$$

is minimized only at  $x = x_0$ . Then,  $V(\sigma)$  has a unique minimum at  $\sigma = \sigma_0$  by **A5**. Since  $\Theta$  is compact and the contrast function converges uniformly to  $V(\sigma)$  with probability one, we can establish the theorem by standard arguments.  $\square$

## S2 Proof of Theorem 2.2

**Lemma 6.** Suppose that **A1**, **A3–A4** and **A6(ii)** with some  $2k$  hold. For any  $k_0 \in \{1, \dots, k\}$  and  $k_1 \in \{0\} \cup \mathbb{N}$ , if  $nh_n^2 \rightarrow 0$ , then

$$\left| \frac{1}{\sqrt{n}} \sum_{i=1}^n \partial_\sigma V_{n,i}^\alpha(\sigma_0) - \frac{1}{\sqrt{n}} \sum_{i=1}^n \partial_\sigma U_i^\alpha(\sigma_0) \right| = o_P(1).$$

*Proof.* Let  $R_{1,i-1}(\sigma)$  be such that

$$\frac{\partial}{\partial \sigma} \left[ \frac{1}{c_{i-1}(\sigma)^{\frac{\alpha}{2}}} \left\{ 1 + \sum_{j=1}^{k_1} h_n^j d_j^\alpha(X_{t_{i-1}^n}, \sigma) \right\} \right] = -\frac{\alpha}{2} \frac{\partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} + h_n R_{1,i-1}(\sigma),$$

then, from (S1.11), we have

$$\begin{aligned} & \partial_\sigma V_{n,i}^\alpha(\sigma) \\ = & \left\{ -\frac{\alpha}{2} \frac{\partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} + h_n R_{1,i-1}(\sigma) \right\} \left\{ \frac{1}{\sqrt{1+\alpha}} - \left(1 + \frac{1}{\alpha}\right) \exp\left(-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2 - \underbrace{\frac{\alpha}{2} K_i(\sigma)}_{(*)}\right) \right\} \\ & + \frac{1+\alpha}{2} \frac{1}{c_{i-1}(\sigma)^{\frac{\alpha}{2}}} \left\{ 1 + \sum_{j=1}^{k_1} h_n^j d_j^\alpha(X_{i-1}, \sigma) \right\} \left\{ -\frac{c_{i-1}(\sigma_0) \partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^2} Z_i^2 + \partial_\sigma K_i(\sigma) \right\} \\ & \times \exp\left(-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2 - \underbrace{\frac{\alpha}{2} K_i(\sigma)}_{(*)}\right). \end{aligned}$$

Further, let  $\tilde{U}_i^\alpha(\sigma)$  be the term without  $(*)$  in the RHS of the above equation. We first show that  $|\sum_{i=1}^n \{\partial_\sigma V_{n,i}^\alpha(\sigma_0) - \tilde{U}_i^\alpha(\sigma)\}| = o_P(\sqrt{n})$ . For this, we decompose  $\partial_\sigma V_{n,i}^\alpha(\sigma_0) - \tilde{U}_i^\alpha(\sigma_0)$  into  $I_i + II_i + III_i$ , where

$$\begin{aligned} I_i &= \frac{1+\alpha}{2} \frac{\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} (1 - e^{-\frac{\alpha}{2} K_i(\sigma_0)}), \\ II_i &= \left[ \frac{1+\alpha}{2} \frac{1}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}}} \left\{ \frac{\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)} Z_i^2 - \partial_\sigma K_i(\sigma_0) \right\} \sum_{j=1}^{k_1} h_n^j d_j^\alpha(X_{i-1}, \sigma_0) \right. \\ & \quad \left. + \left(1 + \frac{1}{\alpha}\right) h_n R_{1,i-1}(\sigma_0) \right] e^{-\frac{\alpha}{2} Z_i^2} (1 - e^{-\frac{\alpha}{2} K_i(\sigma_0)}), \\ III_i &= -\frac{1+\alpha}{2} \frac{1}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}}} \partial_\sigma K_i(\sigma_0) e^{-\frac{\alpha}{2} Z_i^2} (1 - e^{-\frac{\alpha}{2} K_i(\sigma_0)}). \end{aligned}$$

To deal with  $n^{-1/2} \sum_{i=1}^n I_i$ , we note that

$$\begin{aligned} \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} K_i(\sigma_0) &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{2\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1} b_{i-1}(\sigma_0)} (Z_i^3 - Z_i) e^{-\frac{\alpha}{2} Z_i^2} \frac{\Delta_i}{\sqrt{h_n}} \\ &+ \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} R_{2,i}(\sigma_0) \\ &:= \sum_{i=1}^n (I_i^a + I_i^b), \end{aligned} \quad (\text{S2.1})$$

where

$$R_{2,i}(\sigma) = K_i(\sigma) - \frac{2b_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)} \frac{Z_i \Delta_i}{\sqrt{h_n}}.$$

By simple calculations and (S1.1), we have

$$\sup_{\sigma \in \Theta} \max_{i \leq n} \left\{ E |R_{2,i}(\sigma)|^{2k} \vee E |\partial_\sigma R_{2,i}(\sigma)|^{2k} \right\} \lesssim h_n^{2k}. \quad (\text{S2.2})$$

Observe that from (1),

$$\begin{aligned} \sum_{i=1}^n I_i^a &= \frac{\sqrt{h_n}}{\sqrt{n}} \sum_{i=1}^n \frac{\partial_\sigma c_{i-1}(\sigma_0) \partial_\sigma b_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} (Z_i^3 - Z_i) (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} \\ &+ \frac{1}{\sqrt{nh_n}} \sum_{i=1}^n \frac{2\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1} b_{i-1}(\sigma_0)} (Z_i^3 - Z_i) e^{-\frac{\alpha}{2} Z_i^2} \Delta_{1,i}. \end{aligned} \quad (\text{S2.3})$$

Since

$$\frac{\sqrt{h_n}}{\sqrt{n}} \sum_{i=1}^n \frac{\partial_\sigma c_{i-1}(\sigma_0) \partial_\sigma b_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} E \left[ (Z_i^3 - Z_i) (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} \middle| \mathcal{G}_{i-1}^n \right] = 0 \quad (\text{S2.4})$$

and, by Lemma 2,

$$\begin{aligned} \frac{h_n}{n} \sum_{i=1}^n \left\{ \frac{\partial_\sigma c_{i-1}(\sigma_0) \partial_\sigma b_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} \right\}^2 E \left[ (Z_i^3 - Z_i)^2 (Z_i^2 - 1)^2 e^{-\alpha Z_i^2} \middle| \mathcal{G}_{i-1}^n \right] \\ \lesssim \max_{i \leq n} (1 + |X_{i-1}|)^C h_n = o(1) \quad a.s., \end{aligned} \quad (\text{S2.5})$$

according to Lemma 9 of Genon-Catalot and Jacod (1993), the first term of the righthand side of (S2.3) converges to zero in probability. Also, by using Cauchy's inequality, Lemma 1 and (S2.2), it can be shown that

$$\frac{1}{\sqrt{nh_n}} \sum_{i=1}^n E \left| \frac{2\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1} b_{i-1}(\sigma_0)} (Z_i^3 - Z_i) e^{-\frac{\alpha}{2} Z_i^2} \Delta_{1,i} \right| \lesssim \sqrt{nh_n}, \quad (\text{S2.6})$$

$$E \sum_{i=1}^n |I_i^b| \lesssim \sqrt{nh_n} \quad (\text{S2.7})$$

and

$$\begin{aligned} & \frac{1}{\sqrt{n}} \sum_{i=1}^n E \left| \frac{\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} K_i(\sigma_0)^2 \right| \\ & \lesssim \frac{1}{\sqrt{n} h_n} \sum_{i=1}^n E \left| (1 + |X_{i-1}|)^C (Z_i^2 - 1) Z_i^2 \Delta_i^2 \right| + \frac{1}{\sqrt{n}} \sum_{i=1}^n E \left| (1 + |X_{i-1}|)^C (Z_i^2 - 1) R_{2,i}(\sigma_0)^2 \right| \\ & \lesssim \sqrt{n} h_n. \end{aligned} \quad (\text{S2.8})$$

Since

$$1 - e^{-\frac{\alpha}{2} K_i(\sigma_0)} = \frac{\alpha}{2} K_i(\sigma_0) - \frac{\alpha^2}{8} K_i(\sigma_0)^2 e^{\zeta_i},$$

where  $|\zeta_i| \leq \alpha |K_i(\sigma_0)|/2$ , by combining (S2.4)-(S2.8) and the fact that  $e^{\max_i |K_i(\sigma_0)|} = O(1)$  a.s., it follows that

$$\begin{aligned} \frac{1}{\sqrt{n}} \sum_{i=1}^n I_i &= \frac{\alpha(1+\alpha)}{4} \sum_{i=1}^n (I_i^a + I_i^b) - \frac{\alpha^2(1+\alpha)}{16} \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial_\sigma c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} (Z_i^2 - 1) e^{-\frac{\alpha}{2} Z_i^2} K_i(\sigma_0)^2 e^{\zeta_i} \\ &= o_P(1). \end{aligned} \quad (\text{S2.9})$$

Now note that

$$\partial_\sigma K_i(\sigma) = -2 \frac{b_{i-1}(\sigma_0) \partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^2} \frac{Z_i \Delta_i}{\sqrt{h_n}} + \partial_\sigma R_{2,i}(\sigma).$$

Using this equation, (S2.2) and a similar method as for (S2.9), we can show that  $\sum_{i=1}^n II_i = o_P(\sqrt{n})$  and  $\sum_{i=1}^n III_i = o_P(\sqrt{n})$ . Hence, we have

$$\left| \frac{1}{\sqrt{n}} \sum_{i=1}^n \partial_\sigma V_{n,i}^\alpha(\sigma_0) - \frac{1}{\sqrt{n}} \sum_{i=1}^n \tilde{U}_i^\alpha(\sigma_0) \right| = o_P(1). \quad (\text{S2.10})$$

Next, letting

$$R_{3,i}(\sigma) := \tilde{U}_i^\alpha(\sigma) - \partial_\sigma U_i^\alpha(\sigma) - \frac{1+\alpha}{2} \frac{1}{c_{i-1}(\sigma)^{\frac{\alpha}{2}}} \partial_\sigma K_i(\sigma) e^{-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2},$$

then it can be seen that  $|R_{3,i}(\sigma)| \lesssim h_n (1 + |X_{i-1}|)^C (1 + Z_i^2 + \partial_\sigma K_i(\sigma))$ . Using again (1) and Lemma 9 of Genon-Catalot and Jacod (1993), we can establish that

$$\left| \frac{1}{\sqrt{n}} \sum_{i=1}^n \tilde{U}_i^\alpha(\sigma_0) - \frac{1}{\sqrt{n}} \sum_{i=1}^n \partial_\sigma U_i^\alpha(\sigma_0) \right| = o_P(1),$$

which together with (S2.10) completes the proof.  $\square$

**Lemma 7.** *Suppose that **A1–A4** and **A6(ii)** with some  $2k$  hold. If  $nh_n^p \rightarrow 0$  for some  $p > 1$ , then*

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n \partial_\sigma U_i^\alpha(\sigma_0) \xrightarrow{d} N(0, \Sigma_\alpha^*),$$

where

$$\Sigma_\alpha^* = \frac{1}{4} \left\{ 2 \frac{(1+\alpha)^2(1+2\alpha^2)}{(1+2\alpha)^2 \sqrt{1+2\alpha}} - \frac{\alpha^2}{1+\alpha} \right\} \int \frac{(\partial_\sigma c(x, \sigma_0))^2}{c(x, \sigma_0)^{\alpha+2}} d\mu_0(x). \quad (\text{S2.11})$$



*Proof.* According to Theorems 3.2 and 3.4 in Hall and Heyde (1980), it suffices to verify that

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n E \left[ \partial_{\sigma} U_i^{\alpha}(\sigma_0) \middle| \mathcal{G}_{i-1}^n \right] \xrightarrow{P} 0, \quad (\text{S2.12})$$

$$\frac{1}{n} \sum_{i=1}^n E \left[ (\partial_{\sigma} U_i^{\alpha}(\sigma_0))^2 \middle| \mathcal{G}_{i-1}^n \right] \xrightarrow{P} \Sigma_{\alpha}^*, \quad (\text{S2.13})$$

$$\frac{1}{n^2} \sum_{i=1}^n E \left[ (\partial_{\sigma} U_i^{\alpha}(\sigma_0))^4 \middle| \mathcal{G}_{i-1}^n \right] \xrightarrow{P} 0. \quad (\text{S2.14})$$

Note that

$$\partial_{\sigma} U_i^{\alpha}(\sigma_0) = \frac{1}{2} \frac{\partial_{\sigma} c_{i-1}(\sigma_0)}{c_{i-1}(\sigma_0)^{\frac{\alpha}{2}+1}} \left\{ (1+\alpha)(1-Z_i^2) e^{-\frac{\alpha}{2} Z_i^2} - \frac{\alpha}{\sqrt{1+\alpha}} \right\}. \quad (\text{S2.15})$$

Then, it follows from the fact that  $E[e^{-\frac{\alpha}{2} Z_i^2} | \mathcal{G}_{i-1}^n] = (1+\alpha)^{-\frac{1}{2}}$  and  $E[Z_i^2 e^{-\frac{\alpha}{2} Z_i^2} | \mathcal{G}_{i-1}^n] = (1+\alpha)^{-\frac{3}{2}}$  that

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n E \left[ \partial_{\sigma} U_i^{\alpha}(\sigma_0) \middle| \mathcal{G}_{i-1}^n \right] = 0,$$

which yields (S2.12). Next, by (S2.15), we have that

$$\frac{1}{n} \sum_{i=1}^n E \left[ (\partial_{\sigma} U_i^{\alpha}(\sigma_0))^2 \middle| \mathcal{G}_{i-1}^n \right] = \frac{1}{4} \left\{ 2 \frac{(1+\alpha)^2(1+2\alpha^2)}{(1+2\alpha)^2 \sqrt{1+2\alpha}} - \frac{\alpha^2}{1+\alpha} \right\} \frac{1}{n} \sum_{i=1}^n \frac{(\partial_{\sigma} c_{i-1}(\sigma_0))^2}{c_{i-1}(\sigma_0)^{\alpha+2}},$$

where we have used the fact that  $E[(1-Z_i^2)^2 e^{-\alpha Z_i^2} | \mathcal{G}_{i-1}^n] = 2(1+2\alpha^2)(1+2\alpha)^{-\frac{5}{2}}$  and  $E[(1-Z_i^2) e^{-\frac{\alpha}{2} Z_i^2} | \mathcal{G}_{i-1}^n] = \alpha(1+\alpha)^{-\frac{3}{2}}$ . Hence, (S2.13) is asserted by Lemma 3. Finally, using again Lemma 3 and the fact that

$$\frac{1}{n^2} \sum_{i=1}^n E \left[ (\partial_{\sigma} U_i^{\alpha}(\sigma_0))^4 \middle| \mathcal{G}_{i-1}^n \right] \lesssim \frac{1}{n^2} \sum_{i=1}^n (1 + |X_{i-1}|^{2C}),$$

(S2.14) is verified.  $\square$

**Lemma 8.** *Suppose that A1–A4 and A6(ii) with some  $2k$  hold. For any  $k_0 \in \{1, \dots, k\}$  and  $k_1 \in \{0\} \cup \mathbb{N}$ , if  $nh_n^p \rightarrow 0$  for some  $p > 1$ , then*

$$\frac{1}{n} \sum_{i=1}^n \partial_{\sigma}^2 V_{n,i}^{\alpha}(\sigma_0) \xrightarrow{P} \frac{\alpha^2 + 2}{4(1+\alpha)^{\frac{3}{2}}} \int \frac{(\partial_{\sigma} c(x, \sigma_0))^2}{c(x, \sigma_0)^{\frac{\alpha}{2}+2}} d\mu_0(x).$$

*Proof.* Let

$$R_{4,i}(\sigma) := \left\{ 1 + \sum_{j=1}^{k_1} h_n^j d_j^{\alpha}(X_{i-1}, \sigma) \right\} \left\{ - \frac{c_{i-1}(\sigma_0) \partial_{\sigma} c_{i-1}(\sigma)}{c_{i-1}(\sigma)^2} Z_i^2 + \partial_{\sigma} K_i(\sigma) \right\} + \frac{c_{i-1}(\sigma_0) \partial_{\sigma} c_{i-1}(\sigma)}{c_{i-1}(\sigma)^2} Z_i^2$$

and

$$R_{5,i}(\sigma) := \frac{h_n}{\sqrt{1+\alpha}} R_{1,i-1}(\sigma) + (1+\alpha) \left\{ \frac{1}{2} \frac{R_{4,i}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}}} - \frac{h_n}{\alpha} R_{1,i-1}(\sigma) \right\} \exp \left( - \frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2 - \frac{\alpha}{2} K_i(\sigma) \right).$$

Then, we can rewrite

$$\begin{aligned}\partial_\sigma V_{n,i}^\alpha(\sigma) &= -\frac{1}{2} \frac{\alpha}{\sqrt{1+\alpha}} \frac{\partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} \\ &\quad + \frac{1+\alpha}{2} \left\{ \frac{\partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} - \frac{c_{i-1}(\sigma_0) \partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+2}} Z_i^2 \right\} \exp\left(-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2 - \frac{\alpha}{2} K_i(\sigma)\right) \\ &\quad + R_{5,i}(\sigma),\end{aligned}$$

and thus we have

$$\begin{aligned}&\partial_\sigma^2 V_{n,i}^\alpha(\sigma) \\ &= -\frac{1}{2} \frac{\alpha}{\sqrt{1+\alpha}} \left\{ \frac{\partial_\sigma^2 c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} - \left(1 + \frac{\alpha}{2}\right) \frac{(\partial_\sigma c_{i-1}(\sigma))^2}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+2}} \right\} \\ &\quad + \frac{1+\alpha}{2} \left[ \frac{\partial_\sigma^2 c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} - \left(1 + \frac{\alpha}{2}\right) \frac{(\partial_\sigma c_{i-1}(\sigma))^2}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+2}} - \left\{ \frac{\partial_\sigma^2 c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+2}} - \left(2 + \frac{\alpha}{2}\right) \frac{(\partial_\sigma c_{i-1}(\sigma))^2}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+3}} \right\} c_{i-1}(\sigma_0) Z_i^2 \right] \\ &\quad \times \exp\left(-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2\right) \underbrace{\left\{ 1 - \frac{\alpha}{2} K_i(\sigma) e^{\zeta_i(\sigma)} \right\}}_{(*)} \\ &\quad + \frac{\alpha(1+\alpha)}{4} \left\{ \frac{\partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+1}} - \frac{c_{i-1}(\sigma_0) \partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^{\frac{\alpha}{2}+2}} Z_i^2 \right\} \left\{ \frac{c_{i-1}(\sigma_0) \partial_\sigma c_{i-1}(\sigma)}{c_{i-1}(\sigma)^2} Z_i^2 - \underbrace{\partial_\sigma K_i(\sigma)}_{(*)} \right\} \\ &\quad \times \exp\left(-\frac{\alpha}{2} \frac{c_{i-1}(\sigma_0)}{c_{i-1}(\sigma)} Z_i^2\right) \underbrace{\left\{ 1 - \frac{\alpha}{2} K_i(\sigma) e^{\zeta_i(\sigma)} \right\}}_{(*)} + \partial_\sigma R_{5,i}(\sigma) \\ &:= \widehat{U}_i^\alpha(\sigma) + R_{6,i}(\sigma),\end{aligned}$$

where  $|\zeta_i(\sigma)| \leq \alpha K_i(\sigma)/2$  and  $\widehat{U}_i^\alpha(\sigma)$  denotes the term without  $(*)$  in the RHS of the above equation. Then, by some calculations, it can be shown that

$$\begin{aligned}|R_{4,i}(\sigma)| + |\partial_\sigma R_{4,i}(\sigma)| &\lesssim (1 + |X_{i-1}|)^C (h_n Z_i^2 + |\partial_\sigma K_i(\sigma)| + |\partial_\sigma^2 K_i(\sigma)|), \\ |\partial_\sigma R_{5,i}(\sigma)| &\lesssim (1 + |X_{i-1}|)^C (h_n + |R_{4,i}(\sigma)| + |\partial_\sigma R_{4,i}(\sigma)|) (1 + Z_i^2 + |\partial_\sigma K_i(\sigma)|) (1 + |K_i(\sigma)|) e^{|K_i(\sigma)|}\end{aligned}$$

and

$$|R_{6,i}(\sigma)| \lesssim (1 + |X_{i-1}|)^C (1 + Z_i^2) (1 + Z_i^2 + |\partial_\sigma K_i(\sigma)|) |K_i(\sigma)| e^{|K_i(\sigma)|} + |\partial_\sigma R_{5,i}(\sigma)|.$$

Furthermore, in view of Lemma 2, one can check that

$$\max_{i \leq n} \sup_{\sigma \in \Theta} |R_{6,i}(\sigma)| = o(1) \quad a.s..$$

Therefore, it follows from Lemmas 3 and 4 that

$$\begin{aligned}
& \frac{1}{n} \sum_{i=1}^n \partial_\sigma^2 V_{n,i}^\alpha(\sigma) \\
& \xrightarrow{a.s.} \frac{1}{2} \int \left[ -\frac{\alpha}{\sqrt{1+\alpha}} + (1+\alpha) \left\{ 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right\}^{-\frac{1}{2}} - (1+\alpha) \left\{ 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right\}^{-\frac{3}{2}} \right] \frac{\partial_\sigma^2 c(x, \sigma)}{c(x, \sigma)^{\frac{\alpha}{2}+1}} d\mu_0(x) \\
& + \frac{1}{4} \int \left[ \frac{\alpha(2+\alpha)}{\sqrt{1+\alpha}} - (1+\alpha)(2+\alpha) \left\{ 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right\}^{-\frac{1}{2}} + (1+\alpha)(4+\alpha) \left\{ 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right\}^{-\frac{3}{2}} \right] \\
& \quad \times \frac{(\partial_\sigma c(x, \sigma))^2}{c(x, \sigma)^{\frac{\alpha}{2}+2}} d\mu_0(x) \\
& + \frac{\alpha(1+\alpha)}{4} \int \left\{ 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right\}^{-\frac{3}{2}} \left[ \frac{c(x, \sigma_0)(\partial_\sigma c(x, \sigma))^2}{c(x, \sigma)^{\frac{\alpha}{2}+3}} - 3 \left\{ 1 + \alpha \frac{c(x, \sigma_0)}{c(x, \sigma)} \right\}^{-1} \right. \\
& \quad \left. \times \frac{c(x, \sigma_0)^2 (\partial_\sigma c(x, \sigma))^2}{c(x, \sigma)^{\frac{\alpha}{2}+4}} \right] d\mu_0(x) \quad \text{uniformly in } \sigma,
\end{aligned}$$

which yields the lemma.  $\square$

**Proof of Theorem 2.2.** Let  $l_n^\alpha(\sigma) = n^{-1} \sum_{i=1}^n V_{n,i}^\alpha(\sigma)$ . Using Taylor's theorem, we have that

$$0 = \sqrt{n} \partial_\sigma l_n^\alpha(\sigma_0) + \sqrt{n} (\hat{\sigma}_n^\alpha - \sigma_0) \int_0^1 \partial_\sigma^2 l_n^\alpha(\sigma_0 + u(\hat{\sigma}_n^\alpha - \sigma_0)) du.$$

Note that by Lemmas 6 and 7,

$$\sqrt{n} \partial_\sigma l_n^\alpha(\sigma_0) \xrightarrow{d} N(0, \Sigma_\alpha^*), \quad (\text{S2.16})$$

where  $\Sigma_\alpha^*$  is the one given in (S2.11). Moreover, since the limit of  $\partial_\sigma^2 l_n^\alpha(\sigma)$  is continuous in  $\sigma$  by **A2**, **A4** and **A6(ii)**, it follows from Lemma 8 that

$$\int_0^1 \partial_\sigma^2 l_n^\alpha(\sigma_0 + u(\hat{\sigma}_n^\alpha - \sigma_0)) du \xrightarrow{a.s.} \frac{\alpha^2 + 2}{4(1+\alpha)^{\frac{3}{2}}} \int \frac{(\partial_\sigma c(x, \sigma_0))^2}{c(x, \sigma_0)^{\frac{\alpha}{2}+2}} d\mu_0(x). \quad (\text{S2.17})$$

Therefore, the theorem is established from (S2.16) and (S2.17).  $\square$

### S3 Proof of Theorem 2.3

**Proof of Theorem 2.3.** First, we note that Lemmas 1–4 also hold for the diffusion process of the form  $dX_t = a(X_t, \theta)dt + b(X_t, \sigma)dW_t$ . In this case,  $X_{t_i^n}$  can be written as

$$X_{t_i^n} = X_{t_{i-1}^n} + a(X_{t_{i-1}^n}, \theta_0)h_n + b_{i-1}(\sigma_0)Z_i \sqrt{h_n} + \Delta_i,$$

and thus, we have

$$\begin{aligned}
X_{t_i^n} - \mathbf{r}_{k_0}(h_n, X_{t_{i-1}^n}, \theta, \sigma) &= X_{t_i^n} - X_{t_{i-1}^n} - a(X_{t_{i-1}^n}, \theta_0)h_n - \sum_{j=2}^{k_0} \frac{h_n^j}{j!} L_{\theta, \sigma}^j X_{t_{i-1}^n} \\
&= b_{i-1}(\sigma_0)Z_i \sqrt{h_n} + \{a(X_{t_{i-1}^n}, \theta_0) - a(X_{t_{i-1}^n}, \theta)\}h_n + \Delta_i - \sum_{j=2}^{k_0} \frac{h_n^j}{j!} L_{\theta, \sigma}^j X_{t_{i-1}^n} \\
&:= b_{i-1}(\sigma_0)Z_i \sqrt{h_n} + J_i(\theta, \sigma),
\end{aligned}$$

where

$$\Delta_i = \int_{t_{i-1}^n}^{t_i^n} \{a(X_s, \theta_0) - a(X_{t_{i-1}^n}, \theta_0)\} ds + \int_{t_{i-1}^n}^{t_i^n} \{b(X_s, \sigma_0) - b(X_{t_{i-1}^n}, \sigma_0)\} dW_s.$$

Since  $J_i(\theta, \sigma)$  is  $o(h^{1-\kappa})$  almost surely for any  $\kappa > 0$ , it can be seen that (S1.12) also holds, which in turn yields Lemma 5 for the general diffusion case. That is,

$$\sup_{(\theta, \sigma) \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n V_{n,i}^\alpha(\theta, \sigma) - V(\sigma) \right| = o(1) \quad a.s.,$$

which establishes the theorem. □