
Supplement on “Structural Testing of High-Dimensional Correlation Matrices”

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This supplement consists of the detailed proofs of Lemma 1, Theorems 1, 2, 3, 5, and Corollary 3 in “Structural Testing of High-Dimensional Correlation Matrices”. We also present the simulation results when the observations are generated from the Gamma population.

S1 Detailed proof of Lemma 1

This section is devoted to presenting the detailed proof of Lemma 1. Let $\mathbf{e}_\ell$ denote the ℓ th column of $p \times p$ dimensional identity matrix $\mathbf{I}_p$, $\mathbf{\Gamma} = [\text{diag}(\mathbf{\Sigma})]^{-1/2} \mathbf{\Sigma}^{1/2}$, and $\mathbf{C} = (c_{ij})_{i,j=1}^p$ be a $p \times p$ dimensional matrix with $c_{ij} = 2r_{ij}^3 + \beta_w r_{ij} \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2$. In order to simplify the expressions of mean and variance-covariance functions in the CLT given by Lemma 1, we define the following quantites

$$\begin{aligned} f_1(\mathbf{R}, i, j) &= 2r_{ij}^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2, \\ f_2(\mathbf{R}, \mathbf{D}_2) &= 2\text{tr}(\mathbf{R} \mathbf{D}_2)^2 + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4, \\ f_3(\mathbf{R}, \mathbf{D}_2) &= 2\text{tr}(\mathbf{R} \mathbf{D}_2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4, \\ f_4(\mathbf{R}, \mathbf{D}_1, i) &= 2\mathbf{e}_i^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_i + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k, \end{aligned}$$

$$\begin{aligned}
f_5(\mathbf{R}, \mathbf{D}_2) &= 2\text{tr}(\mathbf{R}\mathbf{D}_2)^3 + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k, \\
f_6(\mathbf{R}, \mathbf{D}_2, i) &= 2\mathbf{e}_i^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{R} \mathbf{e}_i + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k, \\
f_7(\mathbf{R}, \mathbf{D}_2, i) &= 2\mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_i + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k, \\
f_8(\mathbf{R}, \mathbf{D}_1, \mathbf{D}_2) &= 2\text{tr}(\mathbf{R}\mathbf{D}_1(\mathbf{R}\mathbf{D}_2)^2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k, \\
f_9(\mathbf{R}, \mathbf{D}_1, \mathbf{D}_2) &= 2\text{tr}(\mathbf{R}\mathbf{D}_1 \mathbf{R} \mathbf{D}_2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k.
\end{aligned}$$

Lemma 1. *Under Assumptions A-B-C, if the spectral norms of $\mathbf{D}_1$ and $\mathbf{D}_2$ are uniformly bounded in p , then we have*

$$[\Lambda_n(\mathbf{D}_1, \mathbf{D}_2)]^{-1/2} \begin{pmatrix} \text{tr}(\widehat{\mathbf{R}}_n \mathbf{D}_1) - \nu_1(\mathbf{D}_1) \\ \text{tr}(\widehat{\mathbf{R}}_n \mathbf{D}_2 \widehat{\mathbf{R}}_n \mathbf{D}_2) - \nu_2(\mathbf{D}_2) \end{pmatrix} \xrightarrow{d} N_2((0, 0)^T, \mathbf{I}_2),$$

where

$$\begin{aligned}
\nu_1(\mathbf{D}_1) &= \text{tr}(\mathbf{R}\mathbf{D}_1) - n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_1) + \beta_w \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell \right] \\
&\quad + \frac{1}{4n} \text{tr}(\mathbf{C}\mathbf{D}_1) + \frac{3}{4n} \left[2\text{tr}(\mathbf{R}\mathbf{D}_1) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right], \\
\nu_2(\mathbf{D}_2) &= \frac{n-3}{n^2} \text{tr}^2(\mathbf{R}\mathbf{D}_2) - 4n^{-1} \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_k \\
&\quad + \frac{n-3}{n} \text{tr}(\mathbf{R}\mathbf{D}_2)^2 + \beta_w n^{-1} \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k)^2 \\
&\quad - 2\beta_w n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
&\quad - 2\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
&\quad - 2\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^2 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
&\quad + \frac{3}{2n} f_2(\mathbf{R}, \mathbf{D}_2) + \frac{3\text{tr}(\mathbf{R}\mathbf{D}_2)}{2n^2} f_3(\mathbf{R}, \mathbf{D}_2)
\end{aligned}$$

$$\begin{aligned}
& + \frac{1}{2n} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{e}_j \mathbf{e}_i^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
& + \frac{1}{2n} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
& + \frac{\text{tr}(\mathbf{R} \mathbf{D}_2)}{2n^2} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R} \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
& + \frac{1}{2n} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{D}_2 \mathbf{R} \mathbf{e}_j f_1(\mathbf{R}, i, j),
\end{aligned}$$

$$\text{and } \Lambda_n(\mathbf{D}_1, \mathbf{D}_2) = \begin{pmatrix} \sigma_{11}(\mathbf{D}_1) & \sigma_{12}(\mathbf{D}_1, \mathbf{D}_2) \\ \sigma_{21}(\mathbf{D}_1, \mathbf{D}_2) & \sigma_{22}(\mathbf{D}_2) \end{pmatrix} \text{ with}$$

$$\begin{aligned}
\sigma_{11}(\mathbf{D}_1) &= n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&+ n^{-1} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_i \mathbf{e}_j^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
&- 2n^{-1} \sum_{i=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_i f_4(\mathbf{R}, \mathbf{D}_1, i), \\
\sigma_{22}(\mathbf{D}_2) &= 4n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_2)^4 + \beta_w \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&+ 4[n^{-1}\text{tr}(\mathbf{R} \mathbf{D}_2)]^2 n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_2)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&+ 4[n^{-1}\text{tr}(\mathbf{R} \mathbf{D}_2)]^2 + 8n^{-2}\text{tr}(\mathbf{R} \mathbf{D}_2) f_5(\mathbf{R}, \mathbf{D}_2) \\
&+ 4n^{-1} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_i \mathbf{e}_j^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
&+ 8n^{-1}\text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_i \mathbf{e}_j^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
&+ 4[n^{-1}\text{tr}(\mathbf{R} \mathbf{D}_2)]^2 n^{-1} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_i \mathbf{e}_j^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
&- 8n^{-1} \sum_{i=1}^p \mathbf{e}_i^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_i f_6(\mathbf{R}, \mathbf{D}_2, i)
\end{aligned}$$

$$\begin{aligned}
& -8n^{-1}\text{tr}(\mathbf{R}\mathbf{D}_2)n^{-1}\sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_i f_6(\mathbf{R}, \mathbf{D}_2, i) \\
& -8n^{-1}\text{tr}(\mathbf{R}\mathbf{D}_2)n^{-1}\sum_{i=1}^p \mathbf{e}_i^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_i f_7(\mathbf{R}, \mathbf{D}_2, i) \\
& -8[n^{-1}\text{tr}(\mathbf{R}\mathbf{D}_2)]^2 n^{-1}\sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_i f_7(\mathbf{R}, \mathbf{D}_2, i), \\
\sigma_{12}(\mathbf{D}_1, \mathbf{D}_2) = & 2n^{-1}f_8(\mathbf{R}, \mathbf{D}_1, \mathbf{D}_2) + 2n^{-2}\text{tr}(\mathbf{R}\mathbf{D}_2)f_9(\mathbf{R}, \mathbf{D}_1, \mathbf{D}_2) \\
& -2n^{-1}\sum_{i=1}^p \mathbf{e}_i^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_i f_4(\mathbf{R}, \mathbf{D}_1, i) \\
& -2n^{-1}\text{tr}(\mathbf{R}\mathbf{D}_2)n^{-1}\sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_i f_4(\mathbf{R}, \mathbf{D}_1, i) \\
& -2n^{-1}\sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_1 \mathbf{e}_i f_6(\mathbf{R}, \mathbf{D}_2, i) \\
& -2n^{-1}\text{tr}(\mathbf{R}\mathbf{D}_2)n^{-1}\sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_1 \mathbf{e}_i f_7(\mathbf{R}, \mathbf{D}_2, i) \\
& +2n^{-1}\sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_1 \mathbf{e}_i \mathbf{e}_j^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_j f_1(\mathbf{R}, i, j) \\
& +2n^{-1}\text{tr}(\mathbf{R}\mathbf{D}_2)n^{-1}\sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R}\mathbf{D}_1 \mathbf{e}_i \mathbf{e}_j^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_j f_1(\mathbf{R}, i, j), \\
\sigma_{21}(\mathbf{D}_1, \mathbf{D}_2) = & \sigma_{12}(\mathbf{D}_1, \mathbf{D}_2).
\end{aligned}$$

Recall that the definitions of the sample covariance and correlation matrices are

$$\widehat{\Sigma}_n = n^{-1} \sum_{k=1}^n (\mathbf{x}_k - \bar{\mathbf{x}})(\mathbf{x}_k - \bar{\mathbf{x}})^T, \quad \widehat{\mathbf{R}}_n = [\text{diag}(\widehat{\Sigma}_n)]^{-1/2} \widehat{\Sigma}_n [\text{diag}(\widehat{\Sigma}_n)]^{-1/2}.$$

Note that the correlation coefficient is invariant under shifting and positive scaling of random variables. Let $\mathbf{y}_k = \mathbf{\Gamma} \mathbf{w}_k$, $\bar{\mathbf{y}} = n^{-1} \sum_{k=1}^n \mathbf{y}_k$, and

$$\widehat{\Sigma}_n^* = n^{-1} \sum_{k=1}^n (\mathbf{y}_k - \bar{\mathbf{y}})(\mathbf{y}_k - \bar{\mathbf{y}})^T,$$

then under Assumption A, that is, $\mathbf{x}_k$ has the independent component structure, we have

$$\widehat{\mathbf{R}}_n = [\text{diag}(\widehat{\Sigma}_n^*)]^{-1/2} \widehat{\Sigma}_n^* [\text{diag}(\widehat{\Sigma}_n^*)]^{-1/2}.$$

Although a slight abuse of notions, for simplicity, in the following we still use $\widehat{\Sigma}_n$ to represent the sample covariance matrix generated by $\mathbf{y}_1, \dots, \mathbf{y}_n$, that is, we denote

$$\widehat{\Sigma}_n = n^{-1} \sum_{k=1}^n (\mathbf{y}_k - \bar{\mathbf{y}})(\mathbf{y}_k - \bar{\mathbf{y}})^T.$$

in the following sections.

S1.1 Truncation, centralization and rescaling

For $k = 1, \dots, n$, denote $\check{\mathbf{w}}_k = (\check{w}_{1k}, \dots, \check{w}_{pk})^T$ with $\check{w}_{\ell k} = w_{\ell k} I\{|w_{\ell k}| < \eta_n \sqrt{n}\}$, where $\eta_n = (\log n)^{-(1+\epsilon)/2}$, $\check{\mathbf{y}}_k = \mathbf{\Gamma} \check{\mathbf{w}}_k$, $\check{\bar{\mathbf{y}}} = n^{-1} \sum_{k=1}^n \check{\mathbf{y}}_k$,

$$\check{\Sigma}_n = n^{-1} \sum_{k=1}^n (\check{\mathbf{y}}_k - \check{\bar{\mathbf{y}}})(\check{\mathbf{y}}_k - \check{\bar{\mathbf{y}}})^T \quad \text{and} \quad \check{\mathbf{R}}_n = [\text{diag}(\check{\Sigma}_n)]^{-1/2} \check{\Sigma}_n [\text{diag}(\check{\Sigma}_n)]^{-1/2}.$$

According to the truncation lemma (Lemma 2) in Nouredine (2009), we have

$$P(\widehat{\mathbf{R}}_n \neq \check{\mathbf{R}}_n, i.o.) = 0.$$

Define $\tilde{\mathbf{w}}_k = (\tilde{w}_{1k}, \dots, \tilde{w}_{pk})^T$ with $\tilde{w}_{\ell k} = (\check{w}_{\ell k} - E\check{w}_{\ell k}) / \sqrt{E(\check{w}_{\ell k} - E\check{w}_{\ell k})^2}$. Also define $\tilde{\mathbf{y}}_k$, $\tilde{\Sigma}_n$, and $\tilde{\mathbf{R}}_n$ as the analogues of $\mathbf{y}_k$, $\bar{\mathbf{y}}$, $\widehat{\Sigma}_n$, and $\widehat{\mathbf{R}}_n$ with $\mathbf{w}_k$ replaced by $\tilde{\mathbf{w}}_k$. Based on (1.6) in Yin et al. (2023), for any $1 \leq \ell \leq p, 1 \leq k \leq n$, we have

$$|E\check{w}_{\ell k}| = o(n^{-3/2}), \quad E(\check{w}_{\ell k} - E\check{w}_{\ell k})^2 - 1 = o(n^{-1}), \quad (\text{S1.1})$$

$$(\sqrt{E(\check{w}_{\ell k} - E\check{w}_{\ell k})^2} - 1) / \sqrt{E(\check{w}_{\ell k} - E\check{w}_{\ell k})^2} = o(n^{-1}). \quad (\text{S1.2})$$

Since $\mathbf{R} = \mathbf{\Gamma} \mathbf{\Gamma}^T$, we have

$$\begin{aligned} \|\check{\Sigma}_n - \tilde{\Sigma}_n\| &= \|\mathbf{\Gamma} (n^{-1} \sum_{k=1}^n \check{\mathbf{w}}_k \check{\mathbf{w}}_k^T - n^{-1} \sum_{k=1}^n \tilde{\mathbf{w}}_k \tilde{\mathbf{w}}_k^T) \mathbf{\Gamma}^T - \mathbf{\Gamma} (\check{\bar{\mathbf{w}}} \check{\bar{\mathbf{w}}}^T - \tilde{\bar{\mathbf{w}}} \tilde{\bar{\mathbf{w}}}^T) \mathbf{\Gamma}^T\| \\ &\leq \|\mathbf{R}\| \left(\left\| n^{-1} \sum_{k=1}^n \check{\mathbf{w}}_k \check{\mathbf{w}}_k^T - n^{-1} \sum_{k=1}^n \tilde{\mathbf{w}}_k \tilde{\mathbf{w}}_k^T \right\| + \|\check{\bar{\mathbf{w}}} \check{\bar{\mathbf{w}}}^T - \tilde{\bar{\mathbf{w}}} \tilde{\bar{\mathbf{w}}}^T\| \right), \end{aligned}$$

where $\|\cdot\|$ denotes the spectral norm of a matrix. According to (1.7) in Yin et al. (2023),

$\|n^{-1} \sum_{k=1}^n \tilde{\mathbf{w}}_k \tilde{\mathbf{w}}_k^T - n^{-1} \sum_{k=1}^n \tilde{\mathbf{w}}_k \tilde{\mathbf{w}}_k^T\| = o_{a.s.}(n^{-1})$. Note that

$$\begin{aligned}\bar{\tilde{\mathbf{w}}} &= \frac{1}{\sqrt{\mathbb{E}(\check{w}_{\ell k} - \mathbb{E}\check{w}_{\ell k})^2}} \bar{\tilde{\mathbf{w}}} - \frac{|\mathbb{E}\check{w}_{\ell k}|}{\sqrt{\mathbb{E}(\check{w}_{\ell k} - \mathbb{E}\check{w}_{\ell k})^2}} \mathbf{1}_p, \\ \bar{\tilde{\mathbf{w}}} - \bar{\tilde{\mathbf{w}}} &= \frac{\sqrt{\mathbb{E}(\check{w}_{\ell k} - \mathbb{E}\check{w}_{\ell k})^2} - 1}{\sqrt{\mathbb{E}(\check{w}_{\ell k} - \mathbb{E}\check{w}_{\ell k})^2}} \bar{\tilde{\mathbf{w}}} + \frac{|\mathbb{E}\check{w}_{\ell k}|}{\sqrt{\mathbb{E}(\check{w}_{\ell k} - \mathbb{E}\check{w}_{\ell k})^2}} \mathbf{1}_p,\end{aligned}$$

combining the above equations with (S1.1) and (S1.2), we obtain

$$\|\bar{\tilde{\mathbf{w}}} \bar{\tilde{\mathbf{w}}}^T - \bar{\tilde{\mathbf{w}}} \bar{\tilde{\mathbf{w}}}^T\| \leq \|\bar{\tilde{\mathbf{w}}}(\bar{\tilde{\mathbf{w}}} - \bar{\tilde{\mathbf{w}}})^T\| + \|(\bar{\tilde{\mathbf{w}}} - \bar{\tilde{\mathbf{w}}}) \bar{\tilde{\mathbf{w}}}^T\| = o_p(n^{-1}),$$

the last equality follows from the fact that $\bar{\tilde{\mathbf{w}}}^T \bar{\tilde{\mathbf{w}}} = O_p(1)$. Therefore, we have

$$\|\tilde{\Sigma}_n - \tilde{\Sigma}_n\| = o_p(n^{-1}). \quad (\text{S1.3})$$

Let γ_ℓ be the transpose of the ℓ th row of the matrix $\mathbf{\Gamma}$, it can be verified that

$$\begin{aligned}\max_{\ell=1,\dots,p} \|\mathbf{e}_\ell^T \tilde{\Sigma}_n \mathbf{e}_\ell - \mathbf{e}_\ell^T \tilde{\Sigma}_n \mathbf{e}_\ell\| &\leq \max_{\ell=1,\dots,p} \left(\|\gamma_\ell^T (n^{-1} \sum_{k=1}^n \tilde{\mathbf{w}}_k \tilde{\mathbf{w}}_k^T - n^{-1} \sum_{k=1}^n \tilde{\mathbf{w}}_k \tilde{\mathbf{w}}_k^T) \gamma_\ell\| \right. \\ &\quad \left. + \|\gamma_\ell^T (\bar{\tilde{\mathbf{w}}} \bar{\tilde{\mathbf{w}}}^T - \bar{\tilde{\mathbf{w}}} \bar{\tilde{\mathbf{w}}}^T) \gamma_\ell\| \right) = o_p(n^{-1}).\end{aligned} \quad (\text{S1.4})$$

Combining (S1.3) and (S1.4), we get

$$\begin{aligned}\|\check{\mathbf{R}}_n - \tilde{\mathbf{R}}_n\| &= \|\text{diag}^{-1/2}(\check{\Sigma}_n) \check{\Sigma}_n \text{diag}^{-1/2}(\check{\Sigma}_n) - \text{diag}^{-1/2}(\tilde{\Sigma}_n) \tilde{\Sigma}_n \text{diag}^{-1/2}(\tilde{\Sigma}_n)\| \\ &\leq \|(\text{diag}^{-1/2}(\check{\Sigma}_n) - \text{diag}^{-1/2}(\tilde{\Sigma}_n)) \check{\Sigma}_n \text{diag}^{-1/2}(\check{\Sigma}_n)\| \\ &\quad + \|\text{diag}^{-1/2}(\tilde{\Sigma}_n) (\check{\Sigma}_n - \tilde{\Sigma}_n) \text{diag}^{-1/2}(\tilde{\Sigma}_n)\| \\ &\quad + \|\text{diag}^{-1/2}(\tilde{\Sigma}_n) \tilde{\Sigma}_n (\text{diag}^{-1/2}(\check{\Sigma}_n) - \text{diag}^{-1/2}(\tilde{\Sigma}_n))\| = o_p(n^{-1}).\end{aligned}$$

It follows that

$$|\text{tr}(\check{\mathbf{R}}_n \mathbf{D}_1) - \text{tr}(\tilde{\mathbf{R}}_n \mathbf{D}_1)| = \sum_{\ell=1}^p |\lambda_\ell((\check{\mathbf{R}}_n - \tilde{\mathbf{R}}_n) \mathbf{D}_1)| \leq p \|\check{\mathbf{R}}_n - \tilde{\mathbf{R}}_n\| \|\mathbf{D}_1\| = o_p(1).$$

Similarly, we have $|\text{tr}(\check{\mathbf{R}}_n \mathbf{D}_2 \check{\mathbf{R}}_n \mathbf{D}_2) - \text{tr}(\tilde{\mathbf{R}}_n \mathbf{D}_2 \tilde{\mathbf{R}}_n \mathbf{D}_2)| = o_p(1)$. Therefore, without loss of generality, in the following proof, we assume that $w_{\ell k}$ satisfies that $|w_{\ell k}| \leq \eta_n \sqrt{n}$, $\mathbb{E} w_{\ell k} = 0$ and $\mathbb{E} w_{\ell k}^2 = 1$.

S1.2 Some primary steps

We decompose $\text{tr}(\hat{\mathbf{R}}_n \mathbf{D}_1)$ and $\text{tr}(\hat{\mathbf{R}}_n \mathbf{D}_2 \hat{\mathbf{R}}_n \mathbf{D}_2)$ into the sum of several terms in this section, and then analyze the terms in the following sections. Since for $\ell = 1, \dots, p$,

$$\hat{\sigma}_{\ell\ell} = n^{-1} \sum_{k=1}^n (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k)^2 - (n^{-1} \sum_{k=1}^n \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k)^2,$$

then we have

$$\max_{1 \leq \ell \leq p} |\hat{\sigma}_{\ell\ell} - 1| \leq \max_{1 \leq \ell \leq p} \left| n^{-1} \sum_{k=1}^n (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k)^2 - 1 \right| + \max_{1 \leq \ell \leq p} \left| n^{-1} \sum_{k=1}^n \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k \right|^2.$$

By Lemma 5.7 in Yin and Ma (2022), we have

$$\max_{1 \leq \ell \leq p} \left| n^{-1} \sum_{k=1}^n (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k)^2 - 1 \right| = O_p(n^{-1/2}).$$

For any $\epsilon > 0$, by Markov's inequality, we get

$$P \left(\max_{1 \leq \ell \leq p} n^{1/3} \left| n^{-1} \sum_{k=1}^n \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k \right|^2 > \epsilon \right) \leq \epsilon^{-2} \sum_{\ell=1}^p n^{2/3} \mathbb{E} \left| n^{-1} \sum_{k=1}^n \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{w}_k \right|^4 = o(1).$$

It follows that

$$\max_{1 \leq \ell \leq p} |\hat{\sigma}_{\ell\ell} - 1| = o_p(n^{-1/3}). \quad (\text{S1.5})$$

Moreover, by the Taylor expansion, we have

$$(\hat{\sigma}_{\ell\ell})^{-1/2} = 1 - \frac{1}{2}(\hat{\sigma}_{\ell\ell} - 1) + \frac{3}{8}(\hat{\sigma}_{\ell\ell} - 1)^2 - \frac{5}{16}(\sigma_{\ell\ell}^*)^{-7/2}(\hat{\sigma}_{\ell\ell} - 1)^3, \quad (\text{S1.6})$$

where $\sigma_{\ell\ell}^*$ lies between 1 and $\hat{\sigma}_{\ell\ell}$. Since $\sigma_{\ell\ell}^* \geq \min(1, \hat{\sigma}_{\ell\ell})$ for all ℓ , based on (S1.5), we have

$$\max_{1 \leq \ell \leq p} |(\sigma_{\ell\ell}^*)^{-7/2}(\hat{\sigma}_{\ell\ell} - 1)^3| = o_p(n^{-1}). \quad (\text{S1.7})$$

Combining (S1.6) and (S1.7), we have

$$(\hat{\sigma}_{\ell\ell})^{-1/2} = 1 - \frac{1}{2}(\hat{\sigma}_{\ell\ell} - 1) + \frac{3}{8}(\hat{\sigma}_{\ell\ell} - 1)^2 + o_p(n^{-1}), \quad (\text{S1.8})$$

where $o_p(n^{-1})$ is uniform for all $\ell = 1, \dots, p$. Since

$$\begin{aligned} \hat{\mathbf{R}}_n &= [\text{diag}(\hat{\Sigma}_n)]^{-1/2} \hat{\Sigma}_n [\text{diag}(\hat{\Sigma}_n)]^{-1/2} \\ &= \hat{\Sigma}_n + [\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p] \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) + \hat{\Sigma}_n [\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p], \end{aligned}$$

where $\text{diag}(\hat{\Sigma}_n) = \text{diag}(\hat{\sigma}_{11}, \dots, \hat{\sigma}_{pp})$. Then by (S1.5) and (S1.8), we have

$$\begin{aligned} \text{tr}(\hat{\mathbf{R}}_n \mathbf{D}_1) &= \text{tr}(\hat{\Sigma}_n \mathbf{D}_1) + \text{tr}[(\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) \mathbf{D}_1] \\ &\quad + \text{tr}[\hat{\Sigma}_n (\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_1] \\ &= \text{tr}(\hat{\Sigma}_n \mathbf{D}_1) - \text{tr}[(\text{diag}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \mathbf{D}_1] + \frac{3}{4} \text{tr}[(\text{diag}(\hat{\Sigma}_n) - \mathbf{I}_p)^2 \hat{\Sigma}_n \mathbf{D}_1] \\ &\quad + \frac{1}{4} \text{tr}[(\text{diag}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n (\text{diag}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_1] + o_p(1) \end{aligned} \quad (\text{S1.9})$$

and

$$\begin{aligned} &\text{tr}(\hat{\mathbf{R}}_n \mathbf{D}_2 \hat{\mathbf{R}}_n \mathbf{D}_2) \quad (\text{S1.10}) \\ &= \text{tr} \left\{ [\hat{\Sigma}_n \mathbf{D}_2 + (\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) \mathbf{D}_2 + \hat{\Sigma}_n (\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2]^2 \right\} \\ &= \text{tr}(\hat{\Sigma}_n \mathbf{D}_2 \hat{\Sigma}_n \mathbf{D}_2) + 2 \text{tr}[(\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) \mathbf{D}_2 \hat{\Sigma}_n \mathbf{D}_2] \\ &\quad + 2 \text{tr}[(\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2 \hat{\Sigma}_n \mathbf{D}_2 \hat{\Sigma}_n] \\ &\quad + \text{tr}[(\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) \mathbf{D}_2 (\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) \mathbf{D}_2] \\ &\quad + 2 \text{tr}[(\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2 (\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \text{diag}^{-1/2}(\hat{\Sigma}_n) \mathbf{D}_2 \hat{\Sigma}_n] \\ &\quad + \text{tr}[(\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2 \hat{\Sigma}_n (\text{diag}^{-1/2}(\hat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2 \hat{\Sigma}_n] \\ &= \text{tr}(\hat{\Sigma}_n \mathbf{D}_2 \hat{\Sigma}_n \mathbf{D}_2) - 2 \text{tr}[(\text{diag}(\hat{\Sigma}_n) - \mathbf{I}_p) \hat{\Sigma}_n \mathbf{D}_2 \hat{\Sigma}_n \mathbf{D}_2] \\ &\quad + \frac{3}{2} \text{tr}[(\text{diag}(\hat{\Sigma}_n) - \mathbf{I}_p)^2 \hat{\Sigma}_n \mathbf{D}_2 \hat{\Sigma}_n \mathbf{D}_2] \end{aligned}$$

$$\begin{aligned}
& + \frac{1}{2} \text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n (\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2] \\
& + \frac{1}{2} \text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \mathbf{D}_2 (\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n] \\
& + \frac{1}{2} \text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_2 (\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_2] + o_p(1).
\end{aligned}$$

S1.3 Analysis of the constant order terms

In this section, we prove that the last two terms in (S1.9) and the last four terms in (S1.10) converge in probability to their expectations, which are of the constant order. Let $\mathbf{r}_i = n^{-1/2} \mathbf{w}_i$ for $i = 1, \dots, n$, then we have

$$\begin{aligned}
& \text{Etr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p)^2 \widehat{\Sigma}_n \mathbf{D}_1] \\
& = \sum_{i=1}^n \sum_{j=1}^n \sum_{h=1}^n \sum_{k=1}^p \text{E}[\mathbf{e}_k^T \Gamma \mathbf{r}_i \mathbf{r}_i^T \Gamma^T \mathbf{D}_1 \mathbf{e}_k (\mathbf{e}_k^T \Gamma \mathbf{r}_j \mathbf{r}_j^T \Gamma^T \mathbf{e}_k - n^{-1}) \\
& \quad \times (\mathbf{e}_k^T \Gamma \mathbf{r}_h \mathbf{r}_h^T \Gamma^T \mathbf{e}_k - n^{-1})] + o(1) \\
& = \sum_{i=1}^n \sum_{j=1}^n \sum_{h=1}^n \sum_{k=1}^p \text{E}[\mathbf{r}_i^T \Gamma^T \mathbf{D}_1 \mathbf{e}_k \mathbf{e}_k^T \Gamma \mathbf{r}_i (\mathbf{r}_j^T \Gamma^T \mathbf{e}_k \mathbf{e}_k^T \Gamma \mathbf{r}_j - n^{-1}) \\
& \quad \times (\mathbf{r}_h^T \Gamma^T \mathbf{e}_k \mathbf{e}_k^T \Gamma \mathbf{r}_h - n^{-1})] + o(1) \\
& = \frac{(n-1)}{n^2} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_\ell^T \Gamma^T \mathbf{e}_k \mathbf{e}_k^T \Gamma \mathbf{e}_\ell)^2 \right] \\
& \quad + \sum_{i=1}^n \sum_{k=1}^p \text{E}[\mathbf{r}_i^T \Gamma^T \mathbf{D}_1 \mathbf{e}_k \mathbf{e}_k^T \Gamma \mathbf{r}_i (\mathbf{r}_i^T \Gamma^T \mathbf{e}_k \mathbf{e}_k^T \Gamma \mathbf{r}_i - n^{-1})^2] + o(1) \\
& = \frac{(n-1)}{n^2} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \Gamma \mathbf{e}_\ell)^4 \right] + o(1),
\end{aligned}$$

where the last equality is from (9.9.6) of Bai and Silverstein (2010). We can also prove that

$$\text{Var}\{\text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p)^2 \widehat{\Sigma}_n \mathbf{D}_1]\} = o(1).$$

From Chebyshev's inequality,

$$\text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p)^2 \widehat{\Sigma}_n \mathbf{D}_1] \tag{S1.11}$$

$$= \frac{(n-1)}{n^2} \left[2\text{tr}(\mathbf{R}\mathbf{D}_1) + \beta_w \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R}\mathbf{D}_1 \mathbf{e}_\ell \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^4 \right] + o_p(1).$$

Similarly,

$$\text{tr}[(\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \widehat{\boldsymbol{\Sigma}}_n (\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \mathbf{D}_1] \quad (\text{S1.12})$$

$$= \frac{(n-1)}{n^2} \text{tr}(\mathbf{C}\mathbf{D}_1) + o_p(1),$$

$$\text{tr}[(\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p)^2 \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_2 \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_2] \quad (\text{S1.13})$$

$$= \frac{(n-1)(n-2)}{n^3} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2 \mathbf{R}\mathbf{D}_2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R}\mathbf{D}_2 \mathbf{R}\mathbf{D}_2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right] \\ + \frac{(n-1)\text{tr}(\mathbf{R}\mathbf{D}_2)}{n^3} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right] + o_p(1),$$

$$\text{tr}[(\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \widehat{\boldsymbol{\Sigma}}_n (\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \mathbf{D}_2 \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_2] \quad (\text{S1.14})$$

$$= \frac{(n-1)(n-2)}{n^3} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{e}_j \mathbf{e}_i^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1),$$

$$\text{tr}[(\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \mathbf{D}_2 (\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_2 \widehat{\boldsymbol{\Sigma}}_n] \quad (\text{S1.15})$$

$$= \frac{(n-1)(n-2)}{n^3} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\ + \frac{(n-1)\text{tr}(\mathbf{R}\mathbf{D}_2)}{n^3} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R} \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1),$$

$$\text{tr}[(\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_2 (\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_2] \quad (\text{S1.16})$$

$$= \frac{(n-1)(n-2)}{n^3} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_j^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_i \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1),$$

where $\mathbf{C} = (c_{ij})_{i,j=1}^p$ is a $p \times p$ dimensional matrix whose (i, j) th element is given by

$$c_{ij} = 2r_{ij}^3 + \beta_w r_{ij} \sum_{\ell=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_\ell)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_\ell)^2,$$

and r_{ij} is the (i, j) th element of the population correlation matrix $\mathbf{R}$.

Combining (S1.9) and (S1.11)-(S1.12), we have

$$\text{tr}(\widehat{\mathbf{R}}_n \mathbf{D}_1) = \text{tr}(\widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_1) - \text{tr}[(\text{diag}(\widehat{\boldsymbol{\Sigma}}_n) - \mathbf{I}_p) \widehat{\boldsymbol{\Sigma}}_n \mathbf{D}_1] + \frac{1}{4n} \text{tr}(\mathbf{C}\mathbf{D}_1) \quad (\text{S1.17})$$

$$+\frac{3}{4n}\left[2\text{tr}(\mathbf{R}\mathbf{D}_1)+\beta_w\sum_{k=1}^p\mathbf{e}_k^T\mathbf{R}\mathbf{D}_1\mathbf{e}_k\sum_{\ell=1}^p(\mathbf{e}_k^T\mathbf{\Gamma}\mathbf{e}_\ell)^4\right]+o_p(1).$$

Also combining (S1.10) and (S1.13)-(S1.16), we obtain

$$\begin{aligned} & \text{tr}(\widehat{\mathbf{R}}_n\mathbf{D}_2\widehat{\mathbf{R}}_n\mathbf{D}_2) \\ = & \text{tr}(\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2)-2\text{tr}[(\text{diag}(\widehat{\mathbf{\Sigma}}_n)-\mathbf{I}_p)\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2] \\ & +\frac{3}{2n}\left[2\text{tr}(\mathbf{R}\mathbf{D}_2\mathbf{R}\mathbf{D}_2)+\beta_w\sum_{k=1}^p\mathbf{e}_k^T\mathbf{R}\mathbf{D}_2\mathbf{R}\mathbf{D}_2\mathbf{e}_k\sum_{\ell=1}^p(\mathbf{e}_k^T\mathbf{\Gamma}\mathbf{e}_\ell)^4\right] \\ & +\frac{3\text{tr}(\mathbf{R}\mathbf{D}_2)}{2n^2}\left[2\text{tr}(\mathbf{R}\mathbf{D}_2)+\beta_w\sum_{k=1}^p\mathbf{e}_k^T\mathbf{R}\mathbf{D}_2\mathbf{e}_k\sum_{\ell=1}^p(\mathbf{e}_k^T\mathbf{\Gamma}\mathbf{e}_\ell)^4\right] \\ & +\frac{1}{2n}\sum_{i=1}^p\sum_{j=1}^p\mathbf{e}_i^T\mathbf{R}\mathbf{e}_j\mathbf{e}_i^T\mathbf{D}_2\mathbf{R}\mathbf{D}_2\mathbf{e}_j\left[2(\mathbf{e}_i^T\mathbf{R}\mathbf{e}_j)^2+\beta_w\sum_{k=1}^p(\mathbf{e}_i^T\mathbf{\Gamma}\mathbf{e}_k)^2(\mathbf{e}_j^T\mathbf{\Gamma}\mathbf{e}_k)^2\right] \\ & +\frac{1}{2n}\sum_{i=1}^p\sum_{j=1}^p\mathbf{e}_i^T\mathbf{D}_2\mathbf{e}_j\mathbf{e}_i^T\mathbf{R}\mathbf{D}_2\mathbf{R}\mathbf{e}_j\left[2(\mathbf{e}_i^T\mathbf{R}\mathbf{e}_j)^2+\beta_w\sum_{k=1}^p(\mathbf{e}_i^T\mathbf{\Gamma}\mathbf{e}_k)^2(\mathbf{e}_j^T\mathbf{\Gamma}\mathbf{e}_k)^2\right] \\ & +\frac{\text{tr}(\mathbf{R}\mathbf{D}_2)}{2n^2}\sum_{i=1}^p\sum_{j=1}^p\mathbf{e}_i^T\mathbf{D}_2\mathbf{e}_j\mathbf{e}_i^T\mathbf{R}\mathbf{e}_j\left[2(\mathbf{e}_i^T\mathbf{R}\mathbf{e}_j)^2+\beta_w\sum_{k=1}^p(\mathbf{e}_i^T\mathbf{\Gamma}\mathbf{e}_k)^2(\mathbf{e}_j^T\mathbf{\Gamma}\mathbf{e}_k)^2\right] \\ & +\frac{1}{2n}\sum_{i=1}^p\sum_{j=1}^p\mathbf{e}_i^T\mathbf{R}\mathbf{D}_2\mathbf{e}_j\mathbf{e}_j^T\mathbf{R}\mathbf{D}_2\mathbf{e}_i\left[2(\mathbf{e}_i^T\mathbf{R}\mathbf{e}_j)^2+\beta_w\sum_{k=1}^p(\mathbf{e}_i^T\mathbf{\Gamma}\mathbf{e}_k)^2(\mathbf{e}_j^T\mathbf{\Gamma}\mathbf{e}_k)^2\right]+o_p(1). \end{aligned} \tag{S1.18}$$

S1.4 Analysis of the main terms

In this section, we derive the CLT for

$$\begin{pmatrix} \text{tr}(\widehat{\mathbf{\Sigma}}_n\mathbf{D}_1)-\text{tr}[(\text{diag}(\widehat{\mathbf{\Sigma}}_n)-\mathbf{I}_p)\widehat{\mathbf{\Sigma}}_n\mathbf{D}_1] \\ \text{tr}(\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2)-2\text{tr}[(\text{diag}(\widehat{\mathbf{\Sigma}}_n)-\mathbf{I}_p)\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2\widehat{\mathbf{\Sigma}}_n\mathbf{D}_2] \end{pmatrix}.$$

According to the expression of $\widehat{\mathbf{\Sigma}}_n$,

$$\widehat{\mathbf{\Sigma}}_n=\mathbf{\Xi}_n+\mathbf{\Delta}_n=\frac{n-1}{n}\sum_{j=1}^n\mathbf{\Gamma}\mathbf{r}_j\mathbf{r}_j^T\mathbf{\Gamma}^T-\frac{1}{n}\sum_{j\neq k}\mathbf{\Gamma}\mathbf{r}_j\mathbf{r}_k^T\mathbf{\Gamma}^T,$$

we have

$$\text{tr}(\widehat{\mathbf{\Sigma}}_n\mathbf{D}_1)-\text{Etr}(\widehat{\mathbf{\Sigma}}_n\mathbf{D}_1)=\text{tr}(\mathbf{\Xi}_n\mathbf{D}_1)-\text{Etr}(\mathbf{\Xi}_n\mathbf{D}_1)+o_p(1),$$

$$\begin{aligned}
& \text{tr}[(\text{diag}(\widehat{\Xi}_n) - \mathbf{I}_p)\widehat{\Xi}_n\mathbf{D}_1] - \text{Etr}[(\text{diag}(\widehat{\Xi}_n) - \mathbf{I}_p)\widehat{\Xi}_n\mathbf{D}_1] \\
&= \text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_1] - \text{Etr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_1] + o_p(1), \\
& \text{tr}(\widehat{\Xi}_n\mathbf{D}_2\widehat{\Xi}_n\mathbf{D}_2) - \text{Etr}(\widehat{\Xi}_n\mathbf{D}_2\widehat{\Xi}_n\mathbf{D}_2) = \text{tr}(\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2) - \text{Etr}(\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2) + o_p(1), \\
& \text{tr}[(\text{diag}(\widehat{\Xi}_n) - \mathbf{I}_p)\widehat{\Xi}_n\mathbf{D}_2\widehat{\Xi}_n\mathbf{D}_2] - \text{Etr}[(\text{diag}(\widehat{\Xi}_n) - \mathbf{I}_p)\widehat{\Xi}_n\mathbf{D}_2\widehat{\Xi}_n\mathbf{D}_2] \\
&= \text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2] - \text{Etr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2] + o_p(1).
\end{aligned}$$

Let $(E_j - E_{j-1})x = E_j(x) - E_{j-1}(x)$, where x is a random variable and $E_j(\cdot)$ denotes the conditional expectation with respect to the σ -field generated by $\mathbf{r}_1, \dots, \mathbf{r}_j$, then

$$\begin{aligned}
& \text{tr}(\Xi_n\mathbf{D}_1) - \text{Etr}(\Xi_n\mathbf{D}_1) = \sum_{j=1}^n (E_j - E_{j-1})\text{tr}(\Xi_n\mathbf{D}_1), \\
& \text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_1] - \text{Etr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_1] \\
&= \sum_{j=1}^n (E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_1], \\
& \text{tr}(\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2) - \text{Etr}(\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2) = \sum_{j=1}^n (E_j - E_{j-1})\text{tr}(\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2), \\
& \text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2] - \text{Etr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2] \\
&= \sum_{j=1}^n (E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2].
\end{aligned}$$

It is not difficult to prove that the martingale difference sequences

$$\begin{aligned}
& \{(E_j - E_{j-1})\text{tr}(\Xi_n\mathbf{D}_1), j = 1, \dots, n\}, \\
& \{(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_1], j = 1, \dots, n\}, \\
& \{(E_j - E_{j-1})\text{tr}(\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2), j = 1, \dots, n\}, \\
& \{(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n\mathbf{D}_2\Xi_n\mathbf{D}_2], j = 1, \dots, n\}
\end{aligned}$$

satisfy the Lindeberg-type conditions, that is,

$$\begin{aligned}
& \sum_{j=1}^n E[(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_1)]^2 \delta_{\{|(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_1)| > \varepsilon\}} \rightarrow 0, \\
& \sum_{j=1}^n E\{(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_1]\}^2 \delta_{\{|(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_1]| > \varepsilon\}} \rightarrow 0, \\
& \sum_{j=1}^n E[(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2)]^2 \delta_{\{|(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2)| > \varepsilon\}} \rightarrow 0, \\
& \sum_{j=1}^n E\{(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2]\}^2 \delta_{\{|(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2]| > \varepsilon\}} \rightarrow 0,
\end{aligned}$$

where $\delta_{\{\cdot\}}$ represents the indicator function, then the proof is omitted. We will calculate

$$\begin{aligned}
\sigma_{11}(\mathbf{D}_1) &= \sigma_{110}(\mathbf{D}_1) + \sigma_{220}(\mathbf{D}_1) - 2\sigma_{120}(\mathbf{D}_1), \\
\sigma_{22}(\mathbf{D}_2) &= \sigma_{330}(\mathbf{D}_2) + 4\sigma_{440}(\mathbf{D}_2) - 4\sigma_{340}(\mathbf{D}_2), \\
\sigma_{12}(\mathbf{D}_1, \mathbf{D}_2) &= \sigma_{130}(\mathbf{D}_1, \mathbf{D}_2) - 2\sigma_{140}(\mathbf{D}_1, \mathbf{D}_2) - \sigma_{230}(\mathbf{D}_1, \mathbf{D}_2) + 2\sigma_{240}(\mathbf{D}_1, \mathbf{D}_2),
\end{aligned}$$

where

$$\begin{aligned}
\sigma_{110}(\mathbf{D}_1) &= \sum_{j=1}^n E_{j-1} [(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_1)]^2, \\
\sigma_{220}(\mathbf{D}_1) &= \sum_{j=1}^n E_{j-1} \{(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_1]\}^2, \\
\sigma_{120}(\mathbf{D}_1) &= \sum_{j=1}^n E_{j-1} \{[(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_1)][(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_1]]\}, \\
\sigma_{330}(\mathbf{D}_2) &= \sum_{j=1}^n E_{j-1} [(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2)]^2, \\
\sigma_{440}(\mathbf{D}_2) &= \sum_{j=1}^n E_{j-1} \{(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2]\}^2, \\
\sigma_{340}(\mathbf{D}_2) &= \sum_{j=1}^n E_{j-1} \{[(E_j - E_{j-1})\text{tr}(\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2)] \\
&\quad \times [(E_j - E_{j-1})\text{tr}[(\text{diag}(\Xi_n) - \mathbf{I}_p)\Xi_n \mathbf{D}_2 \Xi_n \mathbf{D}_2]]\},
\end{aligned}$$

and

$$\begin{aligned}
\sigma_{130}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_1)] [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)] \}, \\
\sigma_{140}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_1)] \\
&\quad \times [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2]] \}, \\
\sigma_{230}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_1]] \\
&\quad \times [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)] \}, \\
\sigma_{240}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_1]] \\
&\quad \times [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2]] \}.
\end{aligned}$$

Based on (1.15) of Bai and Silverstein (2004), after tedious calculations, we obtain

$$\begin{aligned}
\sigma_{110}(\mathbf{D}_1) &= \sum_{j=1}^n \mathbf{E}_{j-1} [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_1)]^2 \\
&= \sum_{j=1}^n \mathbf{E} [\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{r}_j - n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_1)]^2 + o_p(1) \\
&= n^{-1} \left[2 \text{tr}(\mathbf{R} \mathbf{D}_1)^2 + \beta_w \sum_{\ell=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell)^2 \right] + o_p(1), \\
\sigma_{220}(\mathbf{D}_1) &= \sum_{j=1}^n \mathbf{E}_{j-1} \{ (\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_1] \}^2 \\
&= \sum_{j=1}^n \mathbf{E}_{j-1} \left[\sum_{\ell=1}^p \sum_{i=1}^{j-1} \mathbf{r}_i^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_i (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \right. \\
&\quad \left. + \frac{n-j}{n} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \right]^2 + o_p(1) \\
&= n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_2} \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1), \\
\sigma_{120}(\mathbf{D}_1) &= \sum_{j=1}^n \mathbf{E}_{j-1} \{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_1)] [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_1]] \}
\end{aligned}$$

$$\begin{aligned}
&= \sum_{j=1}^n \mathbb{E}_{j-1} \left\{ \left[\sum_{\ell=1}^p \sum_{i=1}^{j-1} \mathbf{r}_i^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_i (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \right. \right. \\
&\quad \left. \left. + \frac{n-j}{n} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \right] [\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{r}_j - n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_1)] \right\} + o_p(1) \\
&= n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \left[2 \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \right] + o_p(1).
\end{aligned}$$

Therefore, we have

$$\begin{aligned}
\sigma_{11}(\mathbf{D}_1) &= n^{-1} \left[2 \text{tr}(\mathbf{R} \mathbf{D}_1)^2 + \beta_w \sum_{\ell=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell)^2 \right] \\
&\quad + n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_2} \left[2 (\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&\quad - 2n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \left[2 \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \right] + o_p(1).
\end{aligned} \tag{S1.19}$$

Since the martingale differences of $\text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)$ and $\text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2]$ are

$$\begin{aligned}
&(\mathbb{E}_j - \mathbb{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2) \\
&= \frac{2(n-j)}{n} (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{r}_j - n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{D}_2)) \\
&\quad + 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{r}_j - n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2)) \\
&\quad + 2 \sum_{k=1}^{j-1} (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{r}_k \mathbf{r}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{r}_j - n^{-1} \mathbf{r}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{r}_k) + o_p(1)
\end{aligned} \tag{S1.20}$$

and

$$\begin{aligned}
&(\mathbb{E}_j - \mathbb{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2] \\
&= \sum_{\ell=1}^p \sum_{i=1}^{j-1} \sum_{h=1}^{j-1} \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_i \mathbf{r}_i^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{r}_h \mathbf{r}_h^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{e}_\ell (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \\
&\quad + \frac{(n-j)(n-j-1)}{n^2} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_\ell (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \\
&\quad + \frac{(n-j) \text{tr}(\mathbf{R} \mathbf{D}_2)}{n^2} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_\ell (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1}) \\
&\quad + \frac{(n-j)}{n} \sum_{\ell=1}^p \sum_{h=1}^{j-1} \mathbf{r}_h^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_h (\mathbf{r}_j^T \mathbf{\Gamma}^T \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{r}_j - n^{-1})
\end{aligned} \tag{S1.21}$$

$$+\frac{(n-j)}{n}\sum_{\ell=1}^p\sum_{h=1}^{j-1}\mathbf{r}_h^T\mathbf{\Gamma}^T\mathbf{D}_2\mathbf{e}_\ell\mathbf{e}_\ell^T\mathbf{R}\mathbf{D}_2\mathbf{\Gamma}\mathbf{r}_h(\mathbf{r}_j^T\mathbf{\Gamma}^T\mathbf{e}_\ell\mathbf{e}_\ell^T\mathbf{\Gamma}\mathbf{r}_j-n^{-1})+o_p(1),$$

respectively, then we have

$$\begin{aligned}\sigma_{330}(\mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)]^2 \\ &= 4n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2)^4 + \beta_w \sum_{\ell=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell)^2 \right] \\ &\quad + 4[n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)]^2 n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2)^2 + \beta_w \sum_{\ell=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell)^2 \right] \\ &\quad + 4[n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)^2]^2 + 8n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2)^3 \right. \\ &\quad \left. + \beta_w \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \right] + o_p(1), \\ \sigma_{440}(\mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \left\{ (\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2] \right\}^2 \\ &= n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_{\ell_2} \\ &\quad \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\ &\quad + 2n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_{\ell_2} \\ &\quad \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\ &\quad + [n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)]^2 n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell_2} \\ &\quad \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1), \\ \sigma_{340}(\mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \left\{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)] \right. \\ &\quad \left. \times [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2]] \right\} \\ &= 2n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_\ell \end{aligned}$$

$$\begin{aligned}
& \times \left[2\mathbf{e}_\ell^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{R}\mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R}\mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& + 2n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_\ell \\
& \times \left[2\mathbf{e}_\ell^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{R}\mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R}\mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& + 2n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_\ell \\
& \times \left[2\mathbf{e}_\ell^T \mathbf{R}\mathbf{D}_2 \mathbf{R}\mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& + 2[n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)]^2 n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_\ell \\
& \times \left[2\mathbf{e}_\ell^T \mathbf{R}\mathbf{D}_2 \mathbf{R}\mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] + o_p(1).
\end{aligned}$$

Therefore, we have

$$\begin{aligned}
\sigma_{22}(\mathbf{D}_2) &= 4n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2)^4 + \beta_w \sum_{\ell=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R}\mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell)^2 \right] \tag{S1.22} \\
&+ 4[n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)]^2 n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2)^2 + \beta_w \sum_{\ell=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell)^2 \right] \\
&+ 4[n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)]^2 + 8n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2)^3 \right. \\
&\left. + \beta_w \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R}\mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \right] \\
&+ 4n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_{\ell_2} \\
&\times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R}\mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&+ 8n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T (\mathbf{R}\mathbf{D}_2)^2 \mathbf{e}_{\ell_2} \\
&\times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R}\mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&+ 4[n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2)]^2 n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T \mathbf{R}\mathbf{D}_2 \mathbf{e}_{\ell_2}
\end{aligned}$$

$$\begin{aligned}
& \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
& - 8n^{-1} \sum_{\ell=1}^p \mathbf{e}_{\ell}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_{\ell} \\
& \times \left[2\mathbf{e}_{\ell}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{R} \mathbf{e}_{\ell} + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& - 8n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell} \\
& \times \left[2\mathbf{e}_{\ell}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{R} \mathbf{e}_{\ell} + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& - 8n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_{\ell}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_{\ell} \\
& \times \left[2\mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_{\ell} + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& - 8[n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2)]^2 n^{-1} \sum_{\ell=1}^p \mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell} \\
& \times \left[2\mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_{\ell} + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] + o_p(1).
\end{aligned}$$

Based on (S1.20) and (S1.21), after complicated calculation, we get

$$\begin{aligned}
\sigma_{130}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \left\{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_1)] [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)] \right\} \\
&= 2n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1 (\mathbf{R} \mathbf{D}_2)^2) + \beta_w \sum_{\ell=1}^p \mathbf{e}_{\ell}^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_{\ell} \mathbf{e}_{\ell}^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_{\ell} \right] \\
&\quad + 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{D}_2) + \beta_w \sum_{\ell=1}^p \mathbf{e}_{\ell}^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_{\ell} \mathbf{e}_{\ell}^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_{\ell} \right] + o_p(1), \\
\sigma_{140}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n \mathbf{E}_{j-1} \left\{ [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_1)] \right. \\
&\quad \times [(\mathbf{E}_j - \mathbf{E}_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2]] \left. \right\} \\
&= n^{-1} \sum_{\ell=1}^p \mathbf{e}_{\ell}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_{\ell} \left[2\mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_{\ell} + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \right] \\
&\quad + n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell} \left[2\mathbf{e}_{\ell}^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_{\ell} + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \right] + o_p(1),
\end{aligned}$$

$$\begin{aligned}
\sigma_{230}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n E_{j-1} \{ [(E_j - E_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_1]] \\
&\quad \times [(E_j - E_{j-1}) \text{tr}(\mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2)] \} \\
&= 2n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \left[2\mathbf{e}_\ell^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
&\quad + 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \\
&\quad \times \left[2\mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] + o_p(1), \\
\sigma_{240}(\mathbf{D}_1, \mathbf{D}_2) &= \sum_{j=1}^n E_{j-1} \{ [(E_j - E_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_1]] \\
&\quad \times [(E_j - E_{j-1}) \text{tr}[(\text{diag}(\mathbf{\Xi}_n) - \mathbf{I}_p) \mathbf{\Xi}_n \mathbf{D}_2 \mathbf{\Xi}_n \mathbf{D}_2]] \} \\
&= n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_{\ell_2} \\
&\quad \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&\quad + n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell_2} \\
&\quad \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1).
\end{aligned}$$

Thus, we have

$$\begin{aligned}
\sigma_{12}(\mathbf{D}_1, \mathbf{D}_2) &= 2n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1 (\mathbf{R} \mathbf{D}_2)^2) + \beta_w \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \right] \quad (\text{S1.23}) \\
&\quad + 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{D}_2) + \beta_w \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_\ell^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \right] \\
&\quad - 2n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_\ell \left[2\mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \right] \\
&\quad - 2n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \left[2\mathbf{e}_\ell^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
&\quad - 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_\ell
\end{aligned}$$

$$\begin{aligned}
& \times \left[2\mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& - 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \\
& \times \left[2\mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_\ell + \beta_w \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& + 2n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T (\mathbf{R} \mathbf{D}_2)^2 \mathbf{e}_{\ell_2} \\
& \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
& + 2n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{\ell_1=1}^p \sum_{\ell_2=1}^p \mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_{\ell_1} \mathbf{e}_{\ell_2}^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_{\ell_2} \\
& \times \left[2(\mathbf{e}_{\ell_1}^T \mathbf{R} \mathbf{e}_{\ell_2})^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_{\ell_1}^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_{\ell_2}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] + o_p(1).
\end{aligned}$$

According to the CLT for martingale difference sequence, we obtain that

$$\Lambda_n^{-1/2} \begin{pmatrix} \text{tr}(\widehat{\Sigma}_n \mathbf{D}_1) - \text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_1] - \mu_{01}(\mathbf{D}_1) \\ \text{tr}(\widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2) - 2\text{tr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2] - \mu_{02}(\mathbf{D}_2) \end{pmatrix} \quad (\text{S1.24})$$

is asymptotically distributed as the standard bivariate normal distribution with

$$\Lambda_n = \begin{pmatrix} \sigma_{11}(\mathbf{D}_1) & \sigma_{12}(\mathbf{D}_1, \mathbf{D}_2) \\ \sigma_{21}(\mathbf{D}_1, \mathbf{D}_2) & \sigma_{22}(\mathbf{D}_2) \end{pmatrix},$$

where $\sigma_{11}(\mathbf{D}_1), \sigma_{22}(\mathbf{D}_2), \sigma_{12}(\mathbf{D}_1, \mathbf{D}_2)$ are given in (S1.19)-(S1.22)-(S1.23) and

$$\begin{aligned}
\mu_{01}(\mathbf{D}_1) &= \text{Etr}(\widehat{\Sigma}_n \mathbf{D}_1) - \text{Etr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_1], \\
\mu_{02}(\mathbf{D}_2) &= \text{Etr}(\widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2) - 2\text{Etr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2].
\end{aligned}$$

Because

$$\begin{aligned}
& \text{Etr}(\widehat{\Sigma}_n \mathbf{D}_1) - \text{Etr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_1] \\
&= \text{tr}(\mathbf{R} \mathbf{D}_1) - n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1) + \beta_w \sum_{k=1}^p \sum_{\ell=1}^p \mathbf{e}_k^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \right] + o(1),
\end{aligned}$$

$$\begin{aligned}
& \text{Etr}(\widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2) \\
&= \frac{(n-1)}{n^2} \text{tr}^2(\mathbf{R} \mathbf{D}_2) + \frac{(n-1)}{n} \text{tr}(\mathbf{R} \mathbf{D}_2)^2 + \beta_w n^{-1} \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k)^2 + o(1), \\
& \text{Etr}[(\text{diag}(\widehat{\Sigma}_n) - \mathbf{I}_p) \widehat{\Sigma}_n \mathbf{D}_2 \widehat{\Sigma}_n \mathbf{D}_2] \\
&= n^{-2} \text{tr}^2(\mathbf{R} \mathbf{D}_2) + n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2)^2 + 2n^{-1} \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_k \\
& \quad + \beta_w n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
& \quad + \beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
& \quad + \beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^2 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell + o(1),
\end{aligned}$$

then we have

$$\begin{aligned}
\mu_{01}(\mathbf{D}_1) &= \text{tr}(\mathbf{R} \mathbf{D}_1) - n^{-1} \left[2\text{tr}(\mathbf{R} \mathbf{D}_1) + \beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p \mathbf{e}_k^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \right] + o(1), \\
\mu_{02}(\mathbf{D}_2) &= \frac{n-3}{n^2} \text{tr}^2(\mathbf{R} \mathbf{D}_2) - 4n^{-1} \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_k \\
& \quad + \frac{n-3}{n} \text{tr}(\mathbf{R} \mathbf{D}_2)^2 + \beta_w n^{-1} \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k)^2 \\
& \quad - 2\beta_w n^{-1} \text{tr}(\mathbf{R} \mathbf{D}_2) n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
& \quad - 2\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
& \quad - 2\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^2 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell + o(1).
\end{aligned}$$

By (S1.17)-(S1.18)-(S1.24), we obtain that

$$\begin{pmatrix} \sigma_{11}(\mathbf{D}_1) & \sigma_{12}(\mathbf{D}_1, \mathbf{D}_2) \\ \sigma_{21}(\mathbf{D}_1, \mathbf{D}_2) & \sigma_{22}(\mathbf{D}_2) \end{pmatrix}^{-1/2} \begin{pmatrix} \text{tr}(\widehat{\mathbf{R}}_n \mathbf{D}_1) - \nu_1(\mathbf{D}_1) \\ \text{tr}(\widehat{\mathbf{R}}_n \mathbf{D}_2 \widehat{\mathbf{R}}_n \mathbf{D}_2) - \nu_2(\mathbf{D}_2) \end{pmatrix} \xrightarrow{d} N(\mathbf{0}_2, \mathbf{I}_2),$$

where $\mathbf{0}_2 = (0, 0)^T$, $\mathbf{I}_2$ is the 2×2 dimensional identity matrix,

$$\begin{aligned}
\nu_1(\mathbf{D}_1) &= \text{tr}(\mathbf{R}\mathbf{D}_1) - n^{-1} \left[2\text{tr}(\mathbf{R}\mathbf{D}_1) + \beta_w \sum_{k=1}^p \sum_{\ell=1}^p \mathbf{e}_k^T \mathbf{D}_1 \mathbf{\Gamma} \mathbf{e}_\ell (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \right] \\
&\quad + \frac{1}{4n} \text{tr}(\mathbf{C}\mathbf{D}_1) + \frac{3}{4n} \left[2\text{tr}(\mathbf{R}\mathbf{D}_1) + \beta_w \sum_{\ell=1}^p \mathbf{e}_\ell^T \mathbf{R} \mathbf{D}_1 \mathbf{e}_\ell \sum_{k=1}^p (\mathbf{e}_\ell^T \mathbf{\Gamma} \mathbf{e}_k)^4 \right], \\
\nu_2(\mathbf{D}_2) &= \frac{n-3}{n^2} \text{tr}^2(\mathbf{R}\mathbf{D}_2) - 4n^{-1} \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_k \\
&\quad + \frac{n-3}{n} \text{tr}(\mathbf{R}\mathbf{D}_2)^2 + \beta_w n^{-1} \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_k)^2 \\
&\quad - 2\beta_w n^{-1} \text{tr}(\mathbf{R}\mathbf{D}_2) n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
&\quad - 2\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
&\quad - 2\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^2 \mathbf{e}_k^T \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{\Gamma} \mathbf{e}_\ell \\
&\quad + \frac{3}{2n} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2 \mathbf{R} \mathbf{D}_2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right] \\
&\quad + \frac{3\text{tr}(\mathbf{R}\mathbf{D}_2)}{2n^2} \left[2\text{tr}(\mathbf{R}\mathbf{D}_2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right] \\
&\quad + \frac{1}{2n} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{e}_j \mathbf{e}_i^T \mathbf{D}_2 \mathbf{R} \mathbf{D}_2 \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&\quad + \frac{1}{2n} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{R} \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&\quad + \frac{\text{tr}(\mathbf{R}\mathbf{D}_2)}{2n^2} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R} \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
&\quad + \frac{1}{2n} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_j \mathbf{e}_j^T \mathbf{R} \mathbf{D}_2 \mathbf{e}_i \left[2(\mathbf{e}_i^T \mathbf{R} \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right].
\end{aligned}$$

S2 Proof of Theorem 1

Note that under H_0 , $\boldsymbol{\theta} = \mathbf{A}^{-1}\mathbf{a}$ with $\mathbf{a} = (\text{tr}[(\mathbf{R} - \mathbf{J}_0)\mathbf{J}_1], \dots, \text{tr}[(\mathbf{R} - \mathbf{J}_0)\mathbf{J}_K])^T$. By Lemma 1, for $k = 1, \dots, K$, we have

$$\text{tr}(\widehat{\mathbf{R}}_n \mathbf{J}_k) - \text{tr}(\mathbf{R} \mathbf{J}_k) = O_p(1).$$

Since $\hat{a}_k = \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{J}_0)\mathbf{J}_k]$ and $a_k = \text{tr}[(\mathbf{R} - \mathbf{J}_0)\mathbf{J}_k]$, then we have $p^{-1}(\hat{a}_k - a_k) = o_p(1)$. It follows that

$$\hat{\boldsymbol{\theta}} - \boldsymbol{\theta} = \mathbf{A}^{-1}(\hat{\mathbf{a}} - \mathbf{a}) = (p^{-1}\mathbf{A})^{-1}p^{-1}(\hat{\mathbf{a}} - \mathbf{a}) = O_p(n^{-1}) = o_p(1).$$

Then the proof of Theorem 1 is completed.

S3 Proof of Theorem 2

In this section, we derive the limiting distributions of T_{1n} and T_{2n} . Let

$$\mathbf{R}_P = \mathbf{J}_0 + \theta_{1P}\mathbf{J}_1 + \dots + \theta_{KP}\mathbf{J}_K,$$

where $\boldsymbol{\theta}_P = (\theta_{1P}, \dots, \theta_{KP})^T = \mathbf{A}^{-1}\mathbf{a}_P$ and $\mathbf{a}_P = (a_{1P}, \dots, a_{KP})^T$ with $a_{kP} = \text{tr}[(\mathbf{R} - \mathbf{J}_0)\mathbf{J}_k]$ for $k = 1, \dots, K$. After simple calculation, we obtain

$$\begin{aligned} \widehat{\mathbf{R}}_0^{-1} &= \mathbf{R}_P^{-1} - \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \widehat{\mathbf{R}}_0^{-1} \mathbf{J}_k \mathbf{R}_P^{-1} \\ &= \mathbf{R}_P^{-1} - \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) (\widehat{\mathbf{R}}_0^{-1} - \mathbf{R}_P^{-1} + \mathbf{R}_P^{-1}) \mathbf{J}_k \mathbf{R}_P^{-1} \\ &= \mathbf{R}_P^{-1} - \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1} - \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) (\widehat{\mathbf{R}}_0^{-1} - \mathbf{R}_P^{-1}) \mathbf{J}_k \mathbf{R}_P^{-1} \\ &= \mathbf{R}_P^{-1} - \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1} + \sum_{k_1=1}^K \sum_{k_2=1}^K (\hat{\theta}_{k_1} - \theta_{k_1P}) (\hat{\theta}_{k_2} - \theta_{k_2P}) \widehat{\mathbf{R}}_0^{-1} \mathbf{J}_{k_1} \mathbf{R}_P^{-1} \mathbf{J}_{k_2} \mathbf{R}_P^{-1}. \end{aligned} \tag{S3.25}$$

By (S3.25) and $\hat{\theta}_k = \theta_{kP} + O_p(n^{-1})$ for $k = 1, \dots, K$, we have

$$\begin{aligned}
T_{1n} &= \text{tr}[(\hat{\mathbf{R}}_n \hat{\mathbf{R}}_0^{-1} - \mathbf{I}_p)^2] \\
&= \text{tr}\{[\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p + \hat{\mathbf{R}}_n(\hat{\mathbf{R}}_0^{-1} - \mathbf{R}_P^{-1})]^2\} \\
&= \text{tr}[(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)^2] + 2\text{tr}[(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)\hat{\mathbf{R}}_n(\hat{\mathbf{R}}_0^{-1} - \mathbf{R}_P^{-1})] \\
&\quad + \text{tr}[\hat{\mathbf{R}}_n(\hat{\mathbf{R}}_0^{-1} - \mathbf{R}_P^{-1})\hat{\mathbf{R}}_n(\hat{\mathbf{R}}_0^{-1} - \mathbf{R}_P^{-1})] \\
&= \text{tr}[(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)^2] + 2 \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1}) \\
&\quad - 2 \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1}) \\
&\quad + 2 \sum_{k_1=1}^K \sum_{k_2=1}^K (\hat{\theta}_{k_1} - \theta_{k_1P})(\hat{\theta}_{k_2} - \theta_{k_2P}) \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \hat{\mathbf{R}}_n \hat{\mathbf{R}}_0^{-1} \mathbf{J}_{k_1} \mathbf{R}_P^{-1} \mathbf{J}_{k_2} \mathbf{R}_P^{-1}) \\
&\quad - 2 \sum_{k_1=1}^K \sum_{k_2=1}^K (\hat{\theta}_{k_1} - \theta_{k_1P})(\hat{\theta}_{k_2} - \theta_{k_2P}) \text{tr}(\hat{\mathbf{R}}_n \hat{\mathbf{R}}_0^{-1} \mathbf{J}_{k_1} \mathbf{R}_P^{-1} \mathbf{J}_{k_2} \mathbf{R}_P^{-1}) \\
&\quad + \sum_{k_1=1}^K \sum_{k_2=1}^K (\hat{\theta}_{k_1} - \theta_{k_1P})(\hat{\theta}_{k_2} - \theta_{k_2P}) \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_{k_1} \mathbf{R}_P^{-1} \hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_{k_2} \mathbf{R}_P^{-1}) \\
&\quad - 2 \sum_{k_1, k_2, k_3=1}^K \prod_{\ell=1}^3 (\hat{\theta}_{k_\ell} - \theta_{k_\ell P}) \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_{k_1} \mathbf{R}_P^{-1} \hat{\mathbf{R}}_n \hat{\mathbf{R}}_0^{-1} \mathbf{J}_{k_2} \mathbf{R}_P^{-1} \mathbf{J}_{k_3} \mathbf{R}_P^{-1}) \\
&\quad + \sum_{k_1, k_2, k_3, k_4=1}^K \prod_{\ell=1}^4 (\hat{\theta}_{k_\ell} - \theta_{k_\ell P}) \text{tr}(\hat{\mathbf{R}}_n \hat{\mathbf{R}}_0^{-1} \mathbf{J}_{k_1} \mathbf{R}_P^{-1} \mathbf{J}_{k_2} \mathbf{R}_P^{-1} \hat{\mathbf{R}}_n \hat{\mathbf{R}}_0^{-1} \mathbf{J}_{k_3} \mathbf{R}_P^{-1} \mathbf{J}_{k_4} \mathbf{R}_P^{-1}) \\
&= \text{tr}[(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)^2] + 2 \sum_{k=1}^K p(\hat{\theta}_k - \theta_{kP}) p^{-1} \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1}) \\
&\quad - 2 \sum_{k=1}^K p(\hat{\theta}_k - \theta_{kP}) p^{-1} \text{tr}(\hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \hat{\mathbf{R}}_n \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1}) + o_p(1).
\end{aligned}$$

From the proof of Lemma 1, we get

$$p^{-1} \text{tr}(\hat{\mathbf{R}}_n \mathbf{M}_1) = p^{-1} \text{tr}(\mathbf{R} \mathbf{M}_1) + o_p(1),$$

$$p^{-1} \text{tr}(\hat{\mathbf{R}}_n \mathbf{M}_1 \hat{\mathbf{R}}_n \mathbf{M}_2) = y_n p^{-1} \text{tr}(\mathbf{R} \mathbf{M}_1) p^{-1} \text{tr}(\mathbf{R} \mathbf{M}_2) + p^{-1} \text{tr}(\mathbf{R} \mathbf{M}_1 \mathbf{R} \mathbf{M}_2) + o_p(1),$$

where $\mathbf{M}_1$ and $\mathbf{M}_2$ are $p \times p$ dimensional non-random symmetric matrices with uniformly bounded spectral norms. Hence, we have

$$\begin{aligned}
T_{1n} &= \text{tr}[(\widehat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)^2] + 2 \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \text{tr}(\mathbf{R} \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1}) \\
&\quad - 2 \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) [n^{-1} \text{tr}(\mathbf{R} \mathbf{R}_P^{-1}) \text{tr}(\mathbf{R} \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1}) + \text{tr}(\mathbf{R} \mathbf{R}_P^{-1} \mathbf{R} \mathbf{R}_P^{-1} \mathbf{J}_k \mathbf{R}_P^{-1})] + o_p(1) \\
&= \text{tr}[(\widehat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)^2] - 2n^{-1} \text{tr}(\mathbf{R} \mathbf{R}_P^{-1}) \text{tr}(\widehat{\mathbf{R}}_0 \mathbf{R}_P^{-1} \mathbf{R} \mathbf{R}_P^{-1}) + 2n^{-1} \text{tr}^2(\mathbf{R} \mathbf{R}_P^{-1}) \\
&\quad - 2 \text{tr}[\widehat{\mathbf{R}}_0 \mathbf{R}_P^{-1} \mathbf{R} \mathbf{R}_P^{-1} (\mathbf{R} \mathbf{R}_P^{-1} - \mathbf{I}_p)] + 2 \text{tr}[\mathbf{R} \mathbf{R}_P^{-1} (\mathbf{R} \mathbf{R}_P^{-1} - \mathbf{I}_p)] + o_p(1) \\
&= \text{tr}[(\widehat{\mathbf{R}}_n \mathbf{R}_P^{-1} - \mathbf{I}_p)^2] - 2 \text{tr}(\widehat{\mathbf{R}}_0 \mathbf{H}) + 2n^{-1} \text{tr}^2(\mathbf{R} \mathbf{R}_P^{-1}) + 2 \text{tr}[\mathbf{R} \mathbf{R}_P^{-1} (\mathbf{R} \mathbf{R}_P^{-1} - \mathbf{I}_p)] + o_p(1) \\
&= \text{tr}(\widehat{\mathbf{R}}_n \mathbf{R}_P^{-1})^2 - 2 \text{tr}(\widehat{\mathbf{R}}_n \mathbf{R}_P^{-1}) + p - 2 \text{tr}(\mathbf{J}_0 \mathbf{H}) - 2 \text{tr}(\widehat{\mathbf{R}}_n \mathbf{B}_P) \\
&\quad + 2 \text{tr}(\mathbf{J}_0 \mathbf{B}_P) + 2n^{-1} \text{tr}^2(\mathbf{R} \mathbf{R}_P^{-1}) + 2 \text{tr}[\mathbf{R} \mathbf{R}_P^{-1} (\mathbf{R} \mathbf{R}_P^{-1} - \mathbf{I}_p)] + o_p(1),
\end{aligned} \tag{S3.26}$$

where

$$\begin{aligned}
\mathbf{H} &= n^{-1} \text{tr}(\mathbf{R} \mathbf{R}_P^{-1}) \mathbf{R}_P^{-1} \mathbf{R} \mathbf{R}_P^{-1} + \mathbf{R}_P^{-1} \mathbf{R} \mathbf{R}_P^{-1} (\mathbf{R} \mathbf{R}_P^{-1} - \mathbf{I}_p), \\
\mathbf{B}_P &= \sum_{k=1}^K h_{kP} \mathbf{J}_k, \quad h_{kP} = (\text{tr}(\mathbf{J}_1 \mathbf{H}), \dots, \text{tr}(\mathbf{J}_K \mathbf{H})) \mathbf{A}^{-1} \mathbf{e}_k,
\end{aligned}$$

and $\mathbf{e}_k$ denotes the k th column of the identity matrix $\mathbf{I}_K$. Due to the fact that

$$\begin{aligned}
\text{tr}(\mathbf{J}_0 \mathbf{H}) &= \text{tr}(\mathbf{R}_P \mathbf{H}) - \text{tr}(\mathbf{R} \mathbf{B}_P) + \text{tr}(\mathbf{J}_0 \mathbf{B}_P), \\
\text{tr}(\mathbf{R}_P \mathbf{H}) &= n^{-1} \text{tr}^2(\mathbf{R} \mathbf{R}_P^{-1}) + \text{tr}[\mathbf{R} \mathbf{R}_P^{-1} (\mathbf{R} \mathbf{R}_P^{-1} - \mathbf{I}_p)],
\end{aligned}$$

we obtain

$$T_{1n} = \text{tr}(\widehat{\mathbf{R}}_n \mathbf{R}_P^{-1})^2 - 2 \text{tr}[\widehat{\mathbf{R}}_n (\mathbf{R}_P^{-1} + \mathbf{B}_P)] + 2 \text{tr}(\mathbf{R} \mathbf{B}_P) + p + o_p(1).$$

By Lemma 1 and the Delta method, we have

$$\sigma_{1n}^{-1} (T_{1n} - \mu_{1n}) \xrightarrow{d} N(0, 1),$$

where

$$\begin{aligned}\mu_{1n} &= \nu_2(\mathbf{R}_P^{-1}) - 2\nu_1(\mathbf{R}_P^{-1} + \mathbf{B}_P) + 2\text{tr}(\mathbf{R}\mathbf{B}_P) + p, \\ \sigma_{1n}^2 &= \sigma_{22}(\mathbf{R}_P^{-1}) + 4\sigma_{11}(\mathbf{R}_P^{-1} + \mathbf{B}_P) - 4\sigma_{12}(\mathbf{R}_P^{-1} + \mathbf{B}_P, \mathbf{R}_P^{-1}).\end{aligned}$$

Note that $\widehat{\mathbf{R}}_0 - \mathbf{R}_P = \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \mathbf{J}_k$. From $\hat{\theta}_k - \theta_{kP} = O_p(n^{-1})$, we get

$$\begin{aligned}T_{2n} &= \text{tr}[(\widehat{\mathbf{R}}_n - \widehat{\mathbf{R}}_0)^2] \\ &= \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P)^2] - 2\text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P)(\widehat{\mathbf{R}}_0 - \mathbf{R}_P)] + \text{tr}[(\widehat{\mathbf{R}}_0 - \mathbf{R}_P)^2] \\ &= \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P)^2] - 2 \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P) \mathbf{J}_k] + o_p(1) \\ &= \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P)^2] - 2 \sum_{k=1}^K (\hat{\theta}_k - \theta_{kP}) \text{tr}[(\mathbf{R} - \mathbf{R}_P) \mathbf{J}_k] + o_p(1) \\ &= \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P)^2] - 2\text{tr}(\widehat{\mathbf{R}}_n \tilde{\mathbf{B}}_P) + 2\text{tr}[(\mathbf{R} - \mathbf{R}_P) \mathbf{R}_P] \\ &\quad - 2\text{tr}[(\mathbf{R} - \mathbf{R}_P) \mathbf{J}_0] + 2\text{tr}(\mathbf{J}_0 \tilde{\mathbf{B}}_P) + o_p(1),\end{aligned}$$

where $\tilde{\mathbf{B}}_P = \sum_{k=1}^K \tilde{h}_{kP} \mathbf{J}_k$ and $\tilde{h}_{kP} = (\text{tr}[(\mathbf{R} - \mathbf{R}_P) \mathbf{J}_1], \dots, \text{tr}[(\mathbf{R} - \mathbf{R}_P) \mathbf{J}_K]) \mathbf{A}^{-1} \mathbf{e}_k$. Due to the fact that $\text{tr}[(\mathbf{R} - \mathbf{R}_P)(\mathbf{R}_P - \mathbf{J}_0)] = \text{tr}(\mathbf{R} \tilde{\mathbf{B}}_P) - \text{tr}(\mathbf{J}_0 \tilde{\mathbf{B}}_P)$, we have

$$\begin{aligned}T_{2n} &= \text{tr}[(\widehat{\mathbf{R}}_n - \mathbf{R}_P)^2] - 2\text{tr}(\widehat{\mathbf{R}}_n \tilde{\mathbf{B}}_P) + 2\text{tr}(\mathbf{R} \tilde{\mathbf{B}}_P) + o_p(1) \\ &= \text{tr}(\widehat{\mathbf{R}}_n^2) - 2\text{tr}[\widehat{\mathbf{R}}_n(\mathbf{R}_P + \tilde{\mathbf{B}}_P)] + \text{tr}(\mathbf{R}_P^2) + 2\text{tr}(\mathbf{R} \tilde{\mathbf{B}}_P) + o_p(1).\end{aligned}\tag{S3.27}$$

By Lemma 1 and the Delta method, we have

$$\sigma_{2n}^{-1}(T_{2n} - \mu_{2n}) \xrightarrow{d} N(0, 1),$$

where

$$\begin{aligned}\mu_{2n} &= \nu_2(\mathbf{I}_p) - 2\nu_1(\mathbf{R}_P + \tilde{\mathbf{B}}_P) + \text{tr}(\mathbf{R}_P^2) + 2\text{tr}(\mathbf{R} \tilde{\mathbf{B}}_P), \\ \sigma_{2n}^2 &= \sigma_{22}(\mathbf{I}_p) + 4\sigma_{11}(\mathbf{R}_P + \tilde{\mathbf{B}}_P) - 4\sigma_{12}(\mathbf{R}_P + \tilde{\mathbf{B}}_P, \mathbf{I}_p).\end{aligned}$$

S4 Proof of Theorem 3

Note that under the null hypothesis H_0 , $\mathbf{R}_P$ degenerates to $\mathbf{R}_0$, in this case, $\mathbf{H} = y_n \mathbf{R}_0^{-1}$ and $\mathbf{B}_P = y_n \mathbf{B}$, where $\mathbf{B} = \sum_{k=1}^K h_k \mathbf{J}_k$ and $h_k = (\text{tr}(\mathbf{J}_1 \mathbf{R}_0^{-1}), \dots, \text{tr}(\mathbf{J}_K \mathbf{R}_0^{-1})) \mathbf{A}^{-1} \mathbf{e}_k$. Based on the conclusion (a) in Theorem 2, we have $\sigma_{10}^{-1}(T_{1n} - \mu_{10}) \xrightarrow{d} N(0, 1)$, where

$$\mu_{10} = \nu_2(\mathbf{R}_0^{-1}) - 2\nu_1(\mathbf{C}_1) + 2y_n \text{tr}(\mathbf{R}_0 \mathbf{B}) + p,$$

$$\sigma_{10}^2 = \sigma_{22}(\mathbf{R}_0^{-1}) + 4\sigma_{11}(\mathbf{C}_1) - 4\sigma_{12}(\mathbf{C}_1, \mathbf{R}_0^{-1}),$$

$\mathbf{C}_1 = \mathbf{R}_0^{-1} + y_n \mathbf{B}$ and $\nu_1(\cdot), \nu_2(\cdot), \sigma_{11}(\cdot), \sigma_{12}(\cdot, \cdot), \sigma_{22}(\cdot)$ are defined in Lemma 1. After calculation and simplification, we get

$$\begin{aligned} \mu_{10} = & py_n - 3y_n^2 - 7y_n + \beta_w y_n - 0.5n^{-1} \text{tr}(\mathbf{C}_0 \mathbf{C}_1) \\ & - (2y_n + 4)\beta_w n^{-1} \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{R}_0^{-1} \mathbf{\Gamma} \mathbf{e}_\ell \\ & + (1.5y_n + 2)n^{-1} \left[2p + \beta_w \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right] \\ & + (0.5y_n + 1)n^{-1} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R}_0 \mathbf{e}_j \mathbf{e}_i^T \mathbf{R}_0^{-1} \mathbf{e}_j \\ & \quad \times \left[2(\mathbf{e}_i^T \mathbf{R}_0 \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\ & + 2n^{-1} \left[2\text{tr}(\mathbf{R}_0 \mathbf{C}_1) + \beta_w \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{C}_1 \mathbf{\Gamma} \mathbf{e}_\ell \right] \\ & - 1.5n^{-1} \left[2\text{tr}(\mathbf{R}_0 \mathbf{C}_1) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R}_0 \mathbf{C}_1 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right], \end{aligned}$$

and

$$\begin{aligned} \sigma_{10}^2 = & 4y_n^2 - 4(2 + \beta_w)y_n(1 + y_n)^2 \\ & + 4(1 + y_n)^2 n^{-1} \left[2\text{tr}(\mathbf{R}_0^2) + \beta_w \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \end{aligned}$$

$$\begin{aligned}
& +4n^{-1} \left[2\text{tr}(\mathbf{R}_0 \mathbf{C}_1)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{C}_1 \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
& +4n^{-1} \sum_{i=1}^p \sum_{j=1}^p \mathbf{e}_i^T \mathbf{R}_0 \mathbf{C}_1 \mathbf{e}_i \mathbf{e}_j^T \mathbf{R}_0 \mathbf{C}_1 \mathbf{e}_j \left[2(\mathbf{e}_i^T \mathbf{R}_0 \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right] \\
& -8n^{-1} \sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}_0 \mathbf{C}_1 \mathbf{e}_i \left[2\mathbf{e}_i^T \mathbf{R}_0 \mathbf{C}_1 \mathbf{R}_0 \mathbf{e}_i + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{C}_1 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& +8(1+y_n)n^{-1} \left[2\text{tr}(\mathbf{R}_0^2 \mathbf{C}_1) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{\Gamma} \mathbf{e}_k \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{C}_1 \mathbf{\Gamma} \mathbf{e}_k \right] \\
& -8(1+y_n)n^{-1} \sum_{i=1}^p \mathbf{e}_i^T \mathbf{R}_0 \mathbf{C}_1 \mathbf{e}_i \left[2\mathbf{e}_i^T \mathbf{R}_0^2 \mathbf{e}_i + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 \mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{\Gamma} \mathbf{e}_k \right],
\end{aligned}$$

where $\mathbf{C}_0$ is a $p \times p$ matrix with (i, j) th element being $c_{0ij} = 2r_{0ij}^3 + \beta_w r_{0ij} \sum_{\ell=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_\ell)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_\ell)^2$.

Similarly, under the null hypothesis H_0 , we have $\tilde{\mathbf{B}}_P = \mathbf{0}_{p \times p}$. By the conclusion (b) in Theorem 2, we get $\sigma_{20}^{-1}(T_{2n} - \mu_{20}) \xrightarrow{d} N(0, 1)$, where

$$\begin{aligned}
\mu_{20} &= \nu_2(\mathbf{I}_p) - 2\nu_1(\mathbf{R}_0) + \text{tr}(\mathbf{R}_0^2), \\
\sigma_{20}^2 &= \sigma_{22}(\mathbf{I}_p) + 4\sigma_{11}(\mathbf{R}_0) - 4\sigma_{12}(\mathbf{R}_0, \mathbf{I}_p).
\end{aligned}$$

After calculation and simplification, we obtain that

$$\begin{aligned}
\mu_{20} &= py_n + y_n^2 + n^{-1} \text{tr}(\mathbf{R}_0^2) + \beta_w n^{-1} \sum_{k=1}^p (\mathbf{e}_k^T \mathbf{\Gamma}^T \mathbf{\Gamma} \mathbf{e}_k)^2 - 0.5n^{-1} \text{tr}(\mathbf{C}_0 \mathbf{R}_0) \\
& -2n^{-1} \left[2\text{tr}(\mathbf{R}_0^2) + \beta_w \sum_{k=1}^p \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^3 \mathbf{e}_k^T \mathbf{R}_0 \mathbf{\Gamma} \mathbf{e}_\ell \right] \\
& +0.5n^{-1} \left[2\text{tr}(\mathbf{R}_0^2) + \beta_w \sum_{k=1}^p \mathbf{e}_k^T \mathbf{R}_0^2 \mathbf{e}_k \sum_{\ell=1}^p (\mathbf{e}_k^T \mathbf{\Gamma} \mathbf{e}_\ell)^4 \right] \\
& +n^{-1} \sum_{i=1}^p \sum_{j=1}^p (\mathbf{e}_i^T \mathbf{R}_0 \mathbf{e}_j)^2 \left[2(\mathbf{e}_i^T \mathbf{R}_0 \mathbf{e}_j)^2 + \beta_w \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2 \right],
\end{aligned}$$

and $\sigma_{20}^2 = 4[n^{-1} \text{tr}(\mathbf{R}_0^2)]^2$.

S5 Proof of Theorem 5

The proof of Theorem 5 is similar to those of Theorem 1 in Cai et al. (2013) and Theorem 2.1 in Yang (2020). Throughout this section, we denote the constants that do not depend on p, n by C and may vary from one expression to the next. Before presenting the proof, we list some technical lemmas that will be used.

S5.1 Technical lemmas

The first lemma is on the large deviations for $\hat{\eta}_{ij}$.

Lemma 2 (Lemma 4.3 of Yang (2020)). *Under assumptions F or F^* , there exists some constant $C > 0$ such that for any $M > 0$ and $\epsilon > 0$,*

$$P\left(\max_{1 \leq i < j \leq p} |\hat{\eta}_{ij} - \eta_{ij}| \geq C \frac{\varepsilon_n}{\log p}\right) = O(p^{-M} + n^{-\epsilon/8}),$$

where $\varepsilon_n = (\log p)^{3/2}/n^{1/2}$ if assumption F holds and $\varepsilon_n = (\log p)^{-1}$ if assumption F^* holds.

The second lemma is an extension of the Bernstein inequality.

Lemma 3 (Lemma 8 of Xiao and Wu (2013)). *Let $X, X_1, \dots, X_n$ be i.i.d. random variables with mean zero and unit variance. Assume that for some $0 < \alpha < 1$,*

$$\mathbb{E}(|X|^{3(1-\alpha)} e^{t|X|^\alpha}) \leq A, \quad \text{for all } 0 \leq t < T.$$

Let $S_n = X_1 + \dots + X_n$. If $x^{1-\alpha} \geq 2A/T^2$, then we have

$$P(S_n \geq x) \leq \exp\left\{-\frac{x^2}{2(n + x^{2-\alpha}/T)}\right\} + nP(X \geq x).$$

The next lemma is the classical Bonferroni's inequality.

Lemma 4 (Lemma 1 of Cai et al. (2013)). *Let $B = \bigcup_{t=1}^p B_t$. For any $k < [p/2]$, we have*

$$\sum_{t=1}^{2k} (-1)^{t-1} E_t \leq P(B) \leq \sum_{t=1}^{2k-1} (-1)^{t-1} E_t,$$

where $E_t = \sum_{1 \leq i_1 < \dots < i_t \leq p} P(B_{i_1} \cap \dots \cap B_{i_t})$.

Suppose that $(N_i)_{i \in \mathcal{I}_n}$ is a Gaussian random vector whose entries have zero mean and unit variance, where $\mathcal{I}_n$ is an index set with cardinality $|\mathcal{I}_n| = s_n$. Let $\Sigma_n = (\sigma_{ij})_{i,j \in \mathcal{I}_n}$ be the covariance matrix of $(N_i)_{i \in \mathcal{I}_n}$. Assume that $s_n \rightarrow \infty$ as $n \rightarrow \infty$.

Consider either of the following two conditions.

(C1) For any sequence $\{b_n\}$ such that $b_n \rightarrow \infty$, $\gamma(n, b_n) = o(1/\log b_n)$;

and $\limsup_{n \rightarrow \infty} \gamma_n < 1$.

(C2) For any sequence $\{b_n\}$ such that $b_n \rightarrow \infty$, $\gamma(n, b_n) = o(1)$;

$\sum_{i \neq j \in \mathcal{I}_n} \sigma_{ij}^2 = O(s_n^{2-\delta})$ for some $\delta > 0$; and $\limsup_{n \rightarrow \infty} \gamma_n < 1$,

where

$$\gamma(n, b_n) := \sup_{i \in \mathcal{I}_n} \sup_{\mathcal{A} \subset \mathcal{I}_n, |\mathcal{A}|=b_n} \inf_{j \in \mathcal{A}} |\sigma_{ij}| \quad \text{and} \quad \gamma_n := \sup_{i,j \in \mathcal{I}_n; i \neq j} |\sigma_{ij}|.$$

Lemma 5 (Lemma 7 of Xiao and Wu (2013)). *Assume that either (C1) or (C2) holds. For a fixed $z \in \mathbb{R}$ and a sequence $\{z_n\}$ satisfying $z_n^2 = 2 \log s_n - \log \log s_n - \log \pi + 2z + o(1)$, denote*

$$A'_i = \{|N_i| > z_n\} \quad \text{and} \quad Q'_d = \sum_{\mathcal{A} \subset \mathcal{I}_n, |\mathcal{A}|=d} P\left(\bigcap_{i \in \mathcal{A}} A'_i\right),$$

then for all $d \geq 1$, it holds that

$$\lim_{n \rightarrow \infty} Q'_d = \frac{e^{-dz}}{d!}.$$

S5.2 The detailed proof

Without loss of generality, we assume that $\boldsymbol{\mu} = \mathbf{0}$ and $\sigma_{ii} = 1$ for $i = 1, \dots, p$. We divide the proof of Theorem 5 into two steps:

Step 1: Effects of estimated variances and means. Denote

$$\widetilde{M}_{n,1} = \max_{1 \leq i < j \leq p} \frac{n(\hat{r}_{ij} - r_{ij})^2}{\eta_{ij}}.$$

Under the event $\{|\hat{\eta}_{ij}/\eta_{ij} - 1| \leq C\varepsilon_n/\log p\}$, we have

$$|\widetilde{M}_n - \widetilde{M}_{n,1}| \leq C\widetilde{M}_n \frac{\varepsilon_n}{\log p}.$$

Thus, by Lemma 2, it suffices to show that for any $t \in \mathbb{R}$,

$$P\left(\widetilde{M}_{n,1} - 4\log p + \log \log p \leq t\right) \rightarrow \exp\left(-\frac{1}{\sqrt{8\pi}} \exp\left(-\frac{t}{2}\right)\right).$$

Let $\widetilde{\Sigma}_n$ be the non-centralized sample covariance matrix, that is, $\widetilde{\Sigma}_n = (\tilde{\sigma}_{ij})_{i,j=1}^p = n^{-1} \sum_{k=1}^n \mathbf{x}_k \mathbf{x}_k^T$, and $\widetilde{\mathbf{R}}_n = (\tilde{r}_{ij})_{i,j=1}^p$ be the sample correlation matrix corresponding to $\widetilde{\Sigma}_n$.

Denote

$$\widetilde{M}_{n,2} = \max_{1 \leq i < j \leq p} \frac{n(\tilde{r}_{ij} - r_{ij})^2}{\eta_{ij}}.$$

Using Lemma 3 and the Bernstein inequality, we can show that

$$\max_{1 \leq i \leq p} |\bar{x}_i| = O_p\left(\sqrt{\log p/n}\right), \quad (\text{S5.28})$$

$$\max_{1 \leq i \leq j \leq p} |\tilde{\sigma}_{ij} - \sigma_{ij}| = O_p\left(\sqrt{\log p/n}\right). \quad (\text{S5.29})$$

From (S5.28), (S5.29), and the first order Taylor expansion of the 3 variate function $x(yz)^{-1/2}$ for $x \in \mathbb{R}$ and $y, z > 0$ (see equation (5) in Cai and Zhang (2016)),

$$\frac{\hat{x}}{(\hat{y}\hat{z})^{1/2}} = \frac{x}{(yz)^{1/2}} + \frac{\hat{x} - x}{(yz)^{1/2}} - \frac{x}{(yz)^{1/2}} \left(\frac{\hat{y} - y}{2y} + \frac{\hat{z} - z}{2z} \right) + o(\hat{x} - x) + o(\hat{y} - y) + o(\hat{z} - z),$$

we get

$$\begin{aligned}\tilde{r}_{ij} - r_{ij} &= \frac{1}{n} \sum_{k=1}^n \left[x_{ik}x_{jk} - \frac{r_{ij}}{2} (x_{ik}^2 + x_{jk}^2) \right] + o_p(n^{-1/2}), \\ \hat{r}_{ij} - r_{ij} &= \frac{1}{n} \sum_{k=1}^n \left[(x_{ik} - \bar{x}_i)(x_{jk} - \bar{x}_j) - \frac{r_{ij}}{2} ((x_{ik} - \bar{x}_i)^2 + (x_{jk} - \bar{x}_j)^2) \right] + o_p(n^{-1/2}).\end{aligned}\tag{S5.30}$$

It follows that

$$\hat{r}_{ij} - \tilde{r}_{ij} = \bar{x}_i\bar{x}_j - \frac{r_{ij}}{2} (\bar{x}_i^2 + \bar{x}_j^2) + o_p(n^{-1/2}).$$

Thus, from (S5.29), we get

$$\left| \widetilde{M}_{n,1} - \widetilde{M}_{n,2} \right| \leq Cn \max_{1 \leq i \leq p} |\bar{x}_i|^4 + Cn^{1/2} \widetilde{M}_{n,2}^{1/2} \max_{1 \leq i \leq p} |\bar{x}_i|^2 + o_p(1),$$

which together with (S5.28) implies that we only need to prove that for any $t \in \mathbb{R}$,

$$P \left(\widetilde{M}_{n,2} - 4 \log p + \log \log p \leq t \right) \rightarrow \exp \left(-\frac{1}{\sqrt{8\pi}} \exp \left(-\frac{t}{2} \right) \right).$$

Step 2: Truncation. From (S5.29) and (S5.30), we get $\widetilde{M}_{n,2} = \max_{\alpha \in \mathcal{I}_n} Q_\alpha^2 + o_p(1)$, where

$$Q_\alpha = \frac{1}{\sqrt{n\eta_{ij}}} \sum_{k=1}^n V_{k\alpha},$$

and

$$V_{k\alpha} = x_{ik}x_{jk} - \frac{r_{ij}}{2} (x_{ik}^2 + x_{jk}^2).$$

Let $\hat{V}_{k\alpha} = V_{k\alpha} I\{|V_{k\alpha}| \leq \tau_n\} - \mathbb{E}V_{k\alpha} I\{|V_{k\alpha}| \leq \tau_n\}$, where $\tau_n = \eta^{-1}8 \log(p+n)$ if assumption

F holds and $\tau_n = \sqrt{n}/(\log p)^8$ if assumption F* holds. Denote

$$\hat{Q}_\alpha = \frac{1}{\sqrt{n\eta_{ij}}} \sum_{k=1}^n \hat{V}_{k\alpha}.$$

Note that $|V_{k\alpha}| \leq x_{ik}^2 + x_{jk}^2$ for $k = 1, \dots, n$. If assumption F holds, then we have

$$\begin{aligned}\max_{\alpha \in \mathcal{I}_n} \frac{1}{\sqrt{n\eta_{ij}}} \sum_{k=1}^n \mathbb{E}|V_{k\alpha}| I\{|V_{k\alpha}| \geq \tau_n\} &\leq C\sqrt{n} \max_{\alpha \in \mathcal{I}_n} \mathbb{E}|V_{1\alpha}| I\{|V_{1\alpha}| \geq \tau_n\} \\ &\leq C\sqrt{n}(p+n)^{-2} \max_{\alpha \in \mathcal{I}_n} \mathbb{E}|V_{1\alpha}| \exp(\eta|V_{1\alpha}|/4) \leq C\sqrt{n}(p+n)^{-2},\end{aligned}$$

and if assumption F^* holds, then we have

$$\begin{aligned} \max_{\alpha \in \mathcal{I}_n} \frac{1}{\sqrt{n\eta_{ij}}} \sum_{k=1}^n \mathbb{E}|V_{k\alpha}| I\{|V_{k\alpha}| \geq \tau_n\} &\leq C\sqrt{n} \max_{\alpha \in \mathcal{I}_n} \mathbb{E}|V_{1\alpha}| I\{|V_{1\alpha}| \geq \tau_n\} \\ &\leq C\sqrt{n} \max_{\alpha \in \mathcal{I}_n} \mathbb{E}|V_{1\alpha}|^{2+2\gamma_0+\epsilon/2} / \tau_n^{1+2\gamma_0+\epsilon/2} \leq Cn^{-\gamma_0-\epsilon/8}. \end{aligned}$$

Thus, we get

$$\begin{aligned} &P\left(\max_{\alpha \in \mathcal{I}_n} |\hat{Q}_\alpha - \hat{Q}_\alpha| \geq (\log p)^{-1}\right) \\ &\leq P\left(\max_{\alpha \in \mathcal{I}_n} \frac{1}{\sqrt{n\eta_{ij}}} \sum_{k=1}^n |V_{k\alpha}| I\{|V_{k\alpha}| \geq \tau_n\} - \mathbb{E}V_{k\alpha} I\{|V_{k\alpha}| \geq \tau_n\}| \geq (\log p)^{-1}\right) \\ &\leq P\left(\max_{\alpha \in \mathcal{I}_n} \max_{1 \leq k \leq n} |V_{k\alpha}| \geq \tau_n\right) \leq \sum_{k=1}^n P\left(\max_{\alpha \in \mathcal{I}_n} |V_{k\alpha}| \geq \tau_n\right) \\ &=nP\left(\max_{\alpha \in \mathcal{I}_n} |V_{1\alpha}| \geq \tau_n\right) \leq np \max_{1 \leq i \leq p} P(x_{i1}^2 \geq \tau_n/2) \\ &=O(p^{-1} + n^{-\epsilon/8}). \end{aligned} \tag{S5.31}$$

Note that

$$\left| \max_{\alpha \in \mathcal{I}_n} Q_\alpha^2 - \max_{\alpha \in \mathcal{I}_n} \hat{Q}_\alpha^2 \right| \leq 2 \max_{\alpha \in \mathcal{I}_n} |\hat{Q}_\alpha| \max_{\alpha \in \mathcal{I}_n} |Q_\alpha - \hat{Q}_\alpha| + \max_{\alpha \in \mathcal{I}_n} |Q_\alpha - \hat{Q}_\alpha|^2. \tag{S5.32}$$

Based on (S5.31) and (S5.32), it suffices to prove that for any $t \in \mathbb{R}$,

$$P\left(\max_{\alpha \in \mathcal{I}_n} \hat{Q}_\alpha^2 - 4 \log p + \log \log p \leq t\right) \rightarrow \exp\left(-\frac{1}{\sqrt{8\pi}} \exp\left(-\frac{t}{2}\right)\right). \tag{S5.33}$$

Let q be the cardinality of $\mathcal{I}_n$ and $y_p = t + 4 \log p - \log \log p$. We arrange the two dimensional indices $\{(i, j) : 1 \leq i < j \leq p\}$ in any ordering and set them as $\{(i_m, j_m) : 1 \leq m \leq q\}$. Denote $V_{km} = x_{i_mk}x_{j_mk} - \frac{r_{imjm}}{2}(x_{i_mk}^2 + x_{j_mk}^2)$, $\eta_m = \eta_{i_mj_m}$, and $Q_m = \frac{1}{\sqrt{n\eta_m}} \sum_{k=1}^n V_{km}$.

By Lemma 4, for any integer s with $0 < s < q/2$, we have

$$\begin{aligned} &\sum_{d=1}^{2s} (-1)^{d-1} \sum_{1 \leq m_1 < \dots < m_d \leq q} P\left(\bigcap_{j=1}^d E_{m_j}\right) \leq P\left(\max_{1 \leq m \leq q} \hat{Q}_m^2 \geq y_p\right) \\ &\leq \sum_{d=1}^{2s-1} (-1)^{d-1} \sum_{1 \leq m_1 < \dots < m_d \leq q} P\left(\bigcap_{j=1}^d E_{m_j}\right), \end{aligned} \tag{S5.34}$$

where $E_{m_j} = \{\hat{Q}_{m_j}^2 \geq y_p\}$. Let $\tilde{V}_{km} = \hat{V}_{km}/(\eta_m)^{1/2}$ for $1 \leq m \leq q$ and $\mathbf{V}_k = (\tilde{V}_{km_1}, \dots, \tilde{V}_{km_d})$ for $1 \leq k \leq n$. Denote $|\mathbf{a}|_{\min} = \min_{1 \leq i \leq d} |a_i|$ for any vector $\mathbf{a} \in \mathbb{R}^d$. Then we have

$$P\left(\bigcap_{j=1}^d E_{m_j}\right) = P\left(\left|n^{-1/2} \sum_{k=1}^n \mathbf{V}_k\right|_{\min} \geq y_p^{1/2}\right).$$

By Lemma 9 in Xiao and Wu (2013), we get

$$\begin{aligned} P\left(\left|n^{-1/2} \sum_{k=1}^n \mathbf{V}_k\right|_{\min} \geq y_p^{1/2}\right) &\leq P(|\mathbf{N}_d|_{\min} \geq y_p^{1/2} - \epsilon_n(\log p)^{-1/2}) \\ &\quad + c_{1,d} \exp\left(-\frac{n^{1/2}\epsilon_n}{c_{2,d}\tau_n(\log p)^{1/2}}\right), \end{aligned} \quad (\text{S5.35})$$

where $c_{j,d} = c_j d^{5/2}$, $j = 1, 2$, $c_j > 0$ are constants, $\mathbf{N}_d = (N_{m_1}, \dots, N_{m_d})$ is a d -dimensional Gaussian vector with zero mean and $\text{Cov}(\mathbf{N}_d) = \text{Cov}(\mathbf{V}_1)$. Let

$$\epsilon_n = \begin{cases} (\log p)^{3/2}/n^{3/10}, & \text{if assumption F holds,} \\ (\log p)^{-1/2}, & \text{if assumption F* holds,} \end{cases}$$

then we have $\epsilon_n \rightarrow 0$ and for any $M > 0$,

$$c_{1,d} \exp\left(-\frac{n^{1/2}\epsilon_n}{c_{2,d}\tau_n(\log p)^{1/2}}\right) = O(p^{-M}). \quad (\text{S5.36})$$

It follows from (S5.34), (S5.35), and (S5.36) that

$$P\left(\max_{1 \leq m \leq q} \hat{Q}_m^2 \geq y_p\right) \leq \sum_{d=1}^{2s-1} (-1)^{d-1} \sum_{1 \leq m_1 < \dots < m_d \leq q} P(|\mathbf{N}_d|_{\min} \geq y_p^{1/2} - \epsilon_n(\log p)^{-1/2}) + o(1). \quad (\text{S5.37})$$

Similarly, we also have

$$P\left(\max_{1 \leq m \leq q} \hat{Q}_m^2 \geq y_p\right) \geq \sum_{d=1}^{2s} (-1)^{d-1} \sum_{1 \leq m_1 < \dots < m_d \leq q} P(|\mathbf{N}_d|_{\min} \geq y_p^{1/2} + \epsilon_n(\log p)^{-1/2}) - o(1). \quad (\text{S5.38})$$

For $0 < k, \ell < q$, denote $C_{k\ell} = 1/(\eta_{m_k}\eta_{m_\ell})^{1/2}$. By some elementary calculations, we get if assumptions F holds,

$$\max_{k,\ell} |\text{Cov}(N_{m_k}, N_{m_\ell}) - C_{k\ell} \text{Cov}(V_{1m_k}, V_{1m_\ell})| \leq C(p+n)^{-1}, \quad (\text{S5.39})$$

and if assumption F^* holds,

$$\max_{k,\ell} |\text{Cov}(N_{m_k}, N_{m_\ell}) - C_{k\ell} \text{Cov}(V_{1m_k}, V_{1m_\ell})| \leq C \tau_n^{-2\gamma_0 - \epsilon/2}. \quad (\text{S5.40})$$

Based on assumption E (or E^*), (S5.39), and (S5.40), we know that the covariance matrix of $(N_1, \dots, N_q)$ satisfies assumption E (or E^*). Thus, by Lemma 5, we get

$$\lim_{n \rightarrow \infty} \sum_{1 \leq m_1 < \dots < m_d \leq q} P(|\mathbf{N}_d|_{\min} \geq y_p^{1/2} \pm \epsilon_n (\log p)^{-1/2}) = \frac{1}{d!} \left(\frac{1}{\sqrt{8\pi}} \exp\left(-\frac{t}{2}\right) \right)^d. \quad (\text{S5.41})$$

Submitting (S5.41) into (S5.37) and (S5.38), we get

$$\limsup_{n \rightarrow \infty} P\left(\max_{1 \leq m \leq q} \hat{Q}_m^2 \geq y_p\right) \leq \sum_{d=1}^{2s-1} (-1)^{d-1} \frac{1}{d!} \left(\frac{1}{\sqrt{8\pi}} \exp\left(-\frac{t}{2}\right) \right)^d$$

and

$$\liminf_{n \rightarrow \infty} P\left(\max_{1 \leq m \leq q} \hat{Q}_m^2 \geq y_p\right) \geq \sum_{d=1}^{2s} (-1)^{d-1} \frac{1}{d!} \left(\frac{1}{\sqrt{8\pi}} \exp\left(-\frac{t}{2}\right) \right)^d$$

for any positive integer s . Letting $s \rightarrow \infty$, we prove that (S5.33) holds.

S6 Proof of Corollary 3

Note that $\mathbf{R}_0 = (r_{0ij})_{i,j=1}^p$ represents the structured population correlation matrix under the null hypothesis H_0 . Let

$$M_{n,1} = \max_{1 \leq i < j \leq p} \frac{n(\hat{r}_{ij} - r_{0ij})^2}{\hat{\eta}_{ij}}.$$

By Theorem 5, we get for any $t \in \mathbb{R}$,

$$P(M_{n,1} - 4 \log p + \log \log p \leq t) \rightarrow \exp\left(-\frac{1}{\sqrt{8\pi}} \exp\left(-\frac{t}{2}\right)\right).$$

Under the null hypothesis H_0 , we have

$$\eta_{ij} = \text{Var}\left(\frac{(x_{i1} - \mu_i)(x_{j1} - \mu_j)}{(\sigma_{ii}\sigma_{jj})^{1/2}} - \frac{r_{0ij}}{2} \left(\frac{(x_{i1} - \mu_i)^2}{\sigma_{ii}} + \frac{(x_{j1} - \mu_j)^2}{\sigma_{jj}}\right)\right),$$

and

$$n^{-1} \sum_{k=1}^n \frac{(x_{ik} - \bar{x}_i)(x_{jk} - \bar{x}_j)}{(\hat{\sigma}_{ii}\hat{\sigma}_{jj})^{1/2}} - \frac{\hat{r}_{0ij}}{2} \left(\frac{(\hat{x}_{ik} - \bar{x}_i)^2}{\hat{\sigma}_{ii}} + \frac{(x_{jk} - \bar{x}_j)^2}{\hat{\sigma}_{jj}} \right) = \hat{r}_{ij} - \hat{r}_{0ij},$$

then η_{ij} can be estimated by

$$\hat{\eta}_{0ij} = n^{-1} \sum_{k=1}^n \left[\frac{(x_{ik} - \bar{x}_i)(x_{jk} - \bar{x}_j)}{(\hat{\sigma}_{ii}\hat{\sigma}_{jj})^{1/2}} - \frac{\hat{r}_{0ij}}{2} \left(\frac{(x_{ik} - \bar{x}_i)^2}{\hat{\sigma}_{ii}} + \frac{(x_{jk} - \bar{x}_j)^2}{\hat{\sigma}_{jj}} \right) - (\hat{r}_{ij} - \hat{r}_{0ij}) \right]^2.$$

Next, we prove that $\max_{1 \leq i < j \leq p} |\hat{\eta}_{0ij} - \eta_{ij}| = o_p(1)$. After some elementary calculations, we get

$$\hat{\eta}_{0ij} = n^{-1} \sum_{k=1}^n [T_{kij} + (\hat{r}_{ij} - \hat{r}_{0ij})R_{kij}]^2,$$

where

$$T_{kij} = \frac{(x_{ik} - \bar{x}_i)(x_{jk} - \bar{x}_j)}{(\hat{\sigma}_{ii}\hat{\sigma}_{jj})^{1/2}} - \frac{\hat{r}_{ij}}{2} \left(\frac{(x_{ik} - \bar{x}_i)^2}{\hat{\sigma}_{ii}} + \frac{(x_{jk} - \bar{x}_j)^2}{\hat{\sigma}_{jj}} \right),$$

$$R_{kij} = \frac{(x_{ik} - \bar{x}_i)^2}{2\hat{\sigma}_{ii}} + \frac{(x_{jk} - \bar{x}_j)^2}{2\hat{\sigma}_{jj}} - 1.$$

Thus, we have

$$\hat{\eta}_{0ij} - \eta_{ij} = 2(\hat{r}_{ij} - \hat{r}_{0ij})n^{-1} \sum_{k=1}^n T_{kij}R_{kij} + (\hat{r}_{ij} - \hat{r}_{0ij})^2 n^{-1} \sum_{k=1}^n R_{kij}^2. \quad (\text{S6.42})$$

Note that $n^{-1} \sum_{k=1}^n R_{kij}^2 \leq \max_{1 \leq i \leq p} \hat{\vartheta}_{ii}/\hat{\sigma}_{ii}^2$, where $\hat{\vartheta}_{ii} = n^{-1} \sum_{k=1}^n ((x_{ik} - \bar{x}_i)^2 - \hat{\sigma}_{ii})^2$. By Lemma 3 of Cai et al. (2013), we have

$$P \left(\max_i |\hat{\vartheta}_{ii} - \vartheta_{ii}|/\sigma_{ii}^2 \geq C \frac{\varepsilon_n}{\log p} \right) = O(p^{-1} + n^{-\epsilon/8}),$$

where $\varepsilon_n = \max((\log p)^{1/6}/n^{1/2}, (\log p)^{-1})$. Therefore, we get

$$\max_{1 \leq i < j \leq p} n^{-1} \sum_{k=1}^n R_{kij}^2 = O_p(1). \quad (\text{S6.43})$$

Due to the conclusion that $\hat{\theta}_k = \theta_k + O_p(n^{-1})$ in the proof of Theorem 2, we have

$$\max_{1 \leq i < j \leq p} |\hat{r}_{0ij} - r_{0ij}| \leq \|\hat{\mathbf{R}}_0 - \mathbf{R}_0\| = o_p(n^{-1/2}). \quad (\text{S6.44})$$

From (S5.28) and (S5.29), we get $\max_{1 \leq i < j \leq p} |\hat{r}_{ij} - r_{0ij}| = o_p(1)$. It follows that

$$\max_{1 \leq i < j \leq p} |\hat{r}_{ij} - \hat{r}_{0ij}| = o_p(1). \quad (\text{S6.45})$$

Denote

$$M_{n,2} = \max_{1 \leq i < j \leq p} \frac{n(\hat{r}_{ij} - r_{0ij})^2}{\hat{\eta}_{0ij}}.$$

Combining (S6.42), (S6.43), and (S6.45), we have $\max_{1 \leq i < j \leq p} |\hat{\eta}_{0ij} - \hat{\eta}_{ij}| = o_p(1)$. It follows that

$$|M_{n,1} - M_{n,2}| \leq M_{n,1} \max_{1 \leq i < j \leq p} |\hat{\eta}_{ij}/\hat{\eta}_{0ij} - 1| = o_p(1).$$

From (S6.44), we get

$$|M_{n,2} - M_n| \leq Cn \max_{1 \leq i < j \leq p} |\hat{r}_{0ij} - r_{0ij}|^2 + Cn^{1/2} M_{n,2}^{1/2} \max_{1 \leq i < j \leq p} |\hat{r}_{0ij} - r_{0ij}| = o_p(1).$$

Based on the Slutsky's theorem, we derive that M_n converges to the Type I extreme value distribution under the null hypothesis H_0 . Thus, the proof of Corollary 3 is completed.

Table S1: Key notations used in the paper

Notation	Definition
p	the dimension
n	the sample size
y_n	the dimension-versus-sample-size ratio, defined as p/n
$\mathbf{x}_1, \dots, \mathbf{x}_n$	the <i>i.i.d.</i> sample
$\mathbf{\Sigma} = (\sigma_{ij})_{i,j=1}^p$	the population covariance matrix
$\mathbf{R} = (r_{ij})_{i,j=1}^p$	the population correlation matrix
$\text{diag}(\mathbf{\Sigma})$	the diagonal matrix formed by the diagonal elements of $\mathbf{\Sigma}$
$\mathbf{\Gamma}$	a $p \times p$ matrix, defined as $[\text{diag}(\mathbf{\Sigma})]^{-1/2} \mathbf{\Sigma}^{1/2}$
$\widehat{\mathbf{\Sigma}}_n = (\hat{\sigma}_{ij})_{i,j=1}^p$	the sample covariance matrix
$\widehat{\mathbf{R}}_n = (\hat{r}_{ij})_{i,j=1}^p$	the sample correlation matrix
β_w	the kurtosis, defined in Assumption A
$\mathbf{J}_0, \mathbf{J}_1, \dots, \mathbf{J}_K$	the basis matrices involved in H_0
$\mathbf{R}_0 = (r_{0ij})_{i,j=1}^p$	the population correlation matrix specified under H_0
$\widehat{\mathbf{R}}_0 = (\hat{r}_{0ij})_{i,j=1}^p$	the structured estimator of $\mathbf{R}$ under H_0
$\mathbf{A}$	a $K \times K$ matrix with the (i, j) th entry being $\text{tr}(\mathbf{J}_i \mathbf{J}_j)$
$\mathbf{I}_p$	the $p \times p$ identity matrix
$\mathbf{e}_i$	the i th column of the identity matrix, and its dimension is determined by the matrix it multiplies
$\mathbf{C}_0 = (c_{0ij})_{i,j=1}^p$	a matrix with $c_{0ij} = 2r_{0ij}^3 + \beta_w r_{0ij} \sum_{k=1}^p (\mathbf{e}_i^T \mathbf{\Gamma} \mathbf{e}_k)^2 (\mathbf{e}_j^T \mathbf{\Gamma} \mathbf{e}_k)^2$
h_k	a quantity, defined as $(\text{tr}(\mathbf{J}_1 \mathbf{R}_0^{-1}), \dots, \text{tr}(\mathbf{J}_K \mathbf{R}_0^{-1})) \mathbf{A}^{-1} \mathbf{e}_k$
$\mathbf{B}$	a $p \times p$ matrix, defined as $\sum_{k=1}^K h_k \mathbf{J}_k$
$\mathbf{C}_1$	a $p \times p$ matrix, defined as $\mathbf{R}_0^{-1} + y_n \mathbf{B}$

Table S2: Empirical size and power for the test statistics based on infinite norm (M_n), ratio-based quadratic norm (T_{1n}) and distance-based quadratic norm (T_{2n}) under scenarios 1–4, where n observations with dimension p are generated from the Gamma population.

n	p	Scenario 1			Scenario 2			Scenario 3			Scenario 4		
		M_n	T_{1n}	T_{2n}	M_n	T_{1n}	T_{2n}	M_n	T_{1n}	T_{2n}	M_n	T_{1n}	T_{2n}
Empirical size (%)													
100	50	4.1	5.1	3.8	3.6	4.9	3.7	3.9	4.5	4.3	3.7	4.6	4.7
	100	3.4	5.4	4.4	3.3	5.1	4.5	3.1	5.0	4.7	3.1	4.6	4.8
	300	3.2	5.7	4.6	3.1	5.1	4.6	2.6	5.3	4.8	2.2	4.6	4.6
	500	2.6	6.0	5.1	2.7	5.3	4.7	2.9	5.4	5.1	2.7	4.9	4.8
	1000	2.6	6.3	4.6	2.9	5.9	5.0	2.5	5.8	4.5	2.4	4.5	4.7
300	50	3.6	5.5	4.2	3.7	5.4	3.8	3.8	5.2	4.4	3.7	4.9	4.8
	100	3.5	5.3	4.7	3.5	5.2	4.3	3.3	5.2	4.6	3.6	4.8	5.0
	300	3.3	5.2	4.7	3.3	5.2	4.6	3.3	4.9	4.5	3.2	4.7	4.5
	500	3.1	5.3	4.4	2.8	4.8	4.7	2.9	5.0	4.4	3.1	4.7	4.6
	1000	3.0	5.7	5.1	2.7	5.7	5.0	2.7	5.2	4.9	2.4	5.2	5.2
Empirical power (%)													
100	50	5.6	100.0	11.6	6.8	19.3	54.4	73.6	4.9	9.1	5.0	22.4	14.1
	100	4.5	100.0	11.5	7.0	62.6	89.3	58.7	4.8	6.8	4.5	28.9	15.0
	300	3.2	100.0	11.5	7.4	99.9	100.0	35.0	5.9	4.9	2.7	57.2	15.3
	500	3.1	100.0	11.9	7.8	100.0	100.0	26.7	5.9	5.3	3.0	80.3	16.4
	1000	2.7	100.0	11.8	8.3	100.0	100.0	17.4	5.9	4.9	2.5	99.4	16.3
300	50	11.4	100.0	30.5	10.2	8.0	90.6	99.7	15.3	21.9	9.9	57.8	42.1
	100	8.4	100.0	31.1	11.2	35.1	99.8	99.3	6.2	11.9	7.4	67.0	43.5
	300	5.2	100.0	32.0	12.7	98.4	100.0	96.6	5.0	5.5	5.0	88.7	46.5
	500	4.7	100.0	31.5	12.6	100.0	100.0	93.9	5.3	5.1	4.3	97.1	46.5
	1000	3.9	100.0	32.5	12.8	100.0	100.0	89.0	5.9	5.1	3.2	100.0	47.9

Table S3: Empirical size and power for the test statistics based on infinite norm (M_n), ratio-based quadratic norm (T_{1n}) and distance-based quadratic norm (T_{2n}) under scenarios 5–8, where n observations with dimension p are generated from the Gamma population.

n	p	Scenario 5			Scenario 6			Scenario 7			Scenario 8		
		M_n	T_{1n}	T_{2n}	M_n	T_{1n}	T_{2n}	M_n	T_{1n}	T_{2n}	M_n	T_{1n}	T_{2n}
Empirical size (%)													
100	50	3.7	5.1	3.9	3.8	4.3	4.3	3.3	5.2	4.0	2.9	6.3	2.3
	100	3.2	5.0	4.6	3.2	4.5	4.4	3.6	5.3	4.4	2.9	5.6	3.1
	300	2.7	4.9	4.6	2.6	4.8	4.8	2.9	4.7	4.4	2.7	5.0	3.4
	500	2.9	5.3	4.9	2.9	4.8	4.9	2.9	5.3	5.1	2.7	5.3	3.3
	1000	2.3	4.9	4.7	2.6	4.5	4.6	2.8	4.9	5.2	2.7	5.1	2.7
300	50	3.5	5.3	4.0	3.7	4.9	4.9	3.6	5.5	4.0	2.8	6.1	2.6
	100	3.4	5.0	4.4	3.7	4.8	4.8	3.2	4.7	3.9	3.0	5.5	3.3
	300	3.2	4.8	4.4	3.2	5.0	4.9	3.0	5.1	4.7	2.9	5.1	3.7
	500	2.9	5.1	4.7	3.0	4.9	4.7	2.7	5.0	4.8	2.6	5.1	3.6
	1000	2.5	5.1	4.9	2.6	4.9	4.8	2.6	4.9	5.3	2.4	5.0	3.0
Empirical power (%)													
100	50	87.1	35.1	50.0	57.1	46.0	40.5	3.7	26.6	4.4	45.3	26.4	2.9
	100	75.3	46.7	75.8	56.6	42.8	36.3	3.6	44.9	4.7	34.2	23.4	3.2
	300	50.9	95.5	99.8	54.3	44.8	33.2	3.0	95.2	4.6	26.9	40.6	3.1
	500	39.8	99.8	100.0	52.7	48.0	32.2	3.0	100.0	5.4	25.7	60.6	3.0
	1000	27.1	100.0	100.0	48.8	59.7	32.0	2.7	100.0	5.5	24.8	91.2	2.1
300	50	99.9	83.9	94.1	98.4	94.5	91.8	4.0	45.9	4.9	97.5	61.2	5.6
	100	99.8	91.8	99.4	98.5	93.1	88.8	3.5	68.9	4.9	95.4	51.2	3.6
	300	99.4	100.0	100.0	98.0	92.4	85.4	3.1	99.3	5.9	93.1	65.3	2.7
	500	98.7	100.0	100.0	97.5	93.4	84.0	2.9	100.0	5.8	92.8	83.0	2.6
	1000	97.4	100.0	100.0	97.5	96.0	83.7	2.7	100.0	6.2	93.2	98.6	2.1

Table S4: Empirical size and power for the Tippett's minimum p -value test T_{tn} and the Cauchy combination test T_{cn} under scenarios 1–8, where n observations with dimension p are generated from the Gamma population.

Scenario		1		2		3		4		5		6		7		8	
n	p	T_{tn}	T_{cn}	T_{tn}	T_{cn}	T_{tn}	T_{cn}	T_{tn}	T_{cn}	T_{tn}	T_{cn}	T_{tn}	T_{cn}	T_{tn}	T_{cn}	T_{tn}	T_{cn}
Empirical size (%)																	
100	50	4.0	4.3	3.4	3.9	3.5	4.3	3.1	4.3	3.8	4.2	2.6	4.0	3.7	4.1	3.8	3.8
	100	4.1	4.3	3.9	4.1	3.8	4.3	2.7	4.1	3.8	4.3	2.4	3.7	4.0	4.4	3.4	3.6
	300	4.3	4.4	4.3	4.3	3.7	4.1	2.6	3.6	3.8	4.1	2.5	3.8	3.7	3.8	3.8	3.6
	500	4.4	4.3	4.1	4.1	4.0	4.3	3.1	4.1	3.8	4.0	2.9	4.0	3.9	4.1	3.7	3.7
	1000	4.5	4.5	4.6	4.5	4.4	4.2	2.6	3.6	3.9	3.9	2.5	3.8	4.1	4.1	3.1	3.0
300	50	4.0	4.4	3.8	4.1	3.7	4.3	3.2	4.3	3.7	4.2	3.0	4.2	3.8	4.2	3.8	3.9
	100	4.0	4.6	4.3	4.5	3.8	4.4	3.2	4.3	3.8	4.2	2.9	4.3	3.3	3.7	3.7	3.7
	300	3.9	4.1	4.0	4.0	3.6	4.0	2.9	4.0	3.6	4.0	2.7	4.0	3.8	4.2	3.7	3.4
	500	3.9	4.1	4.0	4.1	3.7	4.0	3.0	3.9	3.6	3.8	2.5	3.9	3.5	3.8	3.4	3.2
	1000	4.4	4.5	4.5	4.2	4.0	4.1	2.9	3.9	3.9	4.1	2.6	4.0	4.0	4.2	3.3	3.2
Empirical power (%)																	
100	50	99.9	99.9	51.7	53.7	61.4	61.7	14.2	17.2	84.4	86.8	59.9	66.2	18.3	18.7	41.2	42.4
	100	100.0	100.0	94.3	95.0	46.7	47.0	18.0	20.5	86.4	88.4	59.0	64.9	33.6	33.2	33.2	34.2
	300	100.0	100.0	100.0	100.0	26.6	26.9	40.9	41.4	99.7	99.8	59.0	64.0	90.7	90.2	39.9	40.7
	500	100.0	100.0	100.0	100.0	20.4	21.0	67.6	66.4	100.0	100.0	60.1	65.1	99.8	99.8	55.5	56.3
	1000	100.0	100.0	100.0	100.0	13.9	14.0	98.7	98.4	100.0	100.0	65.5	68.6	100.0	100.0	86.4	86.7
300	50	100.0	100.0	87.7	88.4	99.4	99.4	44.4	49.7	100.0	100.0	99.6	99.8	34.3	34.0	96.8	97.1
	100	100.0	100.0	99.9	100.0	98.2	98.2	52.2	56.7	100.0	100.0	99.6	99.8	55.5	55.1	94.0	94.4
	300	100.0	100.0	100.0	100.0	93.8	93.6	79.4	80.0	100.0	100.0	99.5	99.6	98.4	98.3	94.5	95.3
	500	100.0	100.0	100.0	100.0	90.2	90.1	93.3	92.9	100.0	100.0	99.5	99.7	100.0	100.0	96.7	97.2
	1000	100.0	100.0	100.0	100.0	83.7	83.6	99.9	99.9	100.0	100.0	99.6	99.8	100.0	100.0	99.7	99.8

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