

CONSISTENT COMMUNITY DETECTION IN MULTI-LAYER NETWORKS WITH PERSONALIZED EDGE DIFFERENTIAL PRIVACY

Yaoming Zhen^a, Shirong Xu^b, AND Junhui Wang^c

Department of Statistical Sciences, University of Toronto^a;

Department of Statistics and Data Science, University of California, Los Angeles^b

Department of Statistics, The Chinese University of Hong Kong^c

Supplementary Material

This Supplementary Material contains all necessary lemmas and technical proofs of the main paper.

S1 Technical proofs

Proof of Lemma 1. For any $(i, j, l) \in [n] \times [n] \times [L]$, we have

$$\begin{aligned} & \sup_{1 \leq i < j \leq n, l \in [L]} \sup_{\tilde{x} \in \{0,1\}} \sup_{x, x' \in \{0,1\}} \frac{P(\mathcal{M}(\mathcal{A}_{i,j,l}) = \tilde{x} | \mathcal{A}_{i,j,l} = x)}{P(\mathcal{M}(\mathcal{A}_{i,j,l}) = \tilde{x} | \mathcal{A}_{i,j,l} = x')} \\ &= \max\left\{1, \frac{\theta}{1-\theta}, \frac{1-\theta}{\theta}\right\} = \exp(\epsilon), \end{aligned}$$

where the last inequality follows from the assumption that $\theta = (1 + \exp(-\epsilon))^{-1}$. ■

Proof of Lemma 2. As in the proof of Lemma 1, $\theta_{i,j} = (1 + \exp(-\epsilon_{i,j}))^{-1}$.

The decomposition (4) immediately yields that $\epsilon_{i,j} = \log \frac{1+f_j f_j}{1-f_i f_j}$ for $1 \leq i < j \leq n$. ■

Proof of Lemma 3. Combining the fact that $\mathbb{E}(\mathbf{A}_{i,j}^{(l)}) = d_i d_j \mathbf{B}_{c_i, c_j}^{(l)}$ and the definition of flipping mechanism, we have

$$\begin{aligned} \mathbb{E}(\mathcal{M}_{\theta_{i,j}}(\mathbf{A}_{i,j}^{(l)})) &= \mathbb{E}(\mathcal{M}_{\theta_{i,j}}(\mathbf{A}_{i,j}^{(l)} | \mathbf{A}_{i,j}^{(l)} = 1) d_i d_j \mathbf{B}_{c_i, c_j}^{(l)} \\ &\quad + \mathbb{E}(\mathcal{M}_{\theta_{i,j}}(\mathbf{A}_{i,j}^{(l)} | \mathbf{A}_{i,j}^{(l)} = 0) (1 - d_i d_j \mathbf{B}_{c_i, c_j}^{(l)})) \\ &= \theta_{i,j} d_i d_j \mathbf{B}_{c_i, c_j}^{(l)} + (1 - \theta_{i,j}) (1 - d_i d_j \mathbf{B}_{c_i, c_j}^{(l)}) \\ &= (2\theta_{i,j} - 1) d_i d_j \mathbf{B}_{c_i, c_j}^{(l)} + (1 - \theta_{i,j}). \\ &= f_i f_j d_i d_j \mathbf{B}_{c_i, c_j}^{(l)} + \frac{1}{2} (1 - f_i f_j). \end{aligned}$$

This completes the proof. ■

Proof of Lemma 4. We first define $\tilde{\mathbf{U}}$ such that $\tilde{\mathbf{U}}_{i,:} = \mathbf{U}_{i,:} / \|\mathbf{U}_{i,:}\|_2$. By the definition of $\mathbf{U}$, we have

$$\mathbf{U}_{i,:} = f_i d_i / \sqrt{\gamma c_i^*} \mathbf{O}_{c_i^*, :}.$$

By the fact that $\|\mathbf{O}_{c_i^*, :}\| = 1$, we further have $\tilde{\mathbf{U}} = \mathbf{O}_{c_i^*, :}$. Clearly, $\tilde{\mathbf{U}}$ has only K distinct rows, and each corresponding to one community. It immediately follows that $\tilde{\mathbf{U}}_{i,:} = \tilde{\mathbf{U}}_{j,:}$ if $c_i^* = c_j^*$. Further, as $\mathbf{O}$ is an orthogonal matrix, $\tilde{\mathbf{U}}_{i,:}$ and $\tilde{\mathbf{U}}_{j,:}$ are perpendicular to each other if $c_i^* \neq c_j^*$ and $\|\tilde{\mathbf{U}}_{i,:}\| = 1$. We thus conclude that $\|\tilde{\mathbf{U}}_{i,:} - \tilde{\mathbf{U}}_{j,:}\| = \sqrt{2}$ if $c_i^* \neq c_j^*$, for $i, j \in [n]$. This completes

the proof. ■

Lemma 1. *Under the conditions of Lemma 4, there exists an orthogonal matrix $\mathbf{O}^{(1)} \in \mathbb{R}^{K \times K}$, such that the left singular vectors corresponding to the non-zero singular values of $\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})$ is $\mathbf{U}\mathbf{O}^{(1)}$.*

Proof of Lemma 1. Since the columns of $\mathbf{V}$ and $\mathbf{U}$ are all orthonormal vectors, $\mathbf{V} \otimes \mathbf{U}$ is also a column orthogonal matrix. Note that the Tucker decomposition of $\tilde{\mathcal{P}}$ implies that

$$\mathcal{M}_1(\tilde{\mathcal{P}}) = \mathbf{U}\mathcal{M}_1(\mathcal{C})(\mathbf{V} \otimes \mathbf{U})^T.$$

Let $\mathbf{O}_{\mathcal{M}_1(\mathcal{C})}\Sigma_{\mathcal{M}_1(\mathcal{C})}\mathbf{V}_{\mathcal{M}_1(\mathcal{C})}^T$ be the singular value decomposition of $\mathcal{M}_1(\mathcal{C})$, it then follows that

$$\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U}) = (\mathbf{U}\mathbf{O}_{\mathcal{M}_1(\mathcal{C})})\Sigma_{\mathcal{M}_1(\mathcal{C})}\mathbf{V}_{\mathcal{M}_1(\mathcal{C})}^T.$$

Therefore, $(\mathbf{U}\mathbf{O}_{\mathcal{M}_1(\mathcal{C})})\Sigma_{\mathcal{M}_1(\mathcal{C})}\mathbf{V}_{\mathcal{M}_1(\mathcal{C})}^T$ is the singular value decomposition of $\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})$. The desired result immediately follows by taking $\mathbf{O}^{(1)} = \mathbf{O}_{\mathcal{M}_1(\mathcal{C})}$. ■

Lemma 2. *Under Assumptions A, it holds true that*

$$\sigma_k(\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})) \asymp ns_n\bar{\psi}\sqrt{L}, \text{ for } k \in [K],$$

where $\bar{\psi} = \frac{1}{n} \sum_{i=1}^n f_i^2 d_i^2$.

Proof of Lemma 2. The Tucker decomposition of $\tilde{\mathcal{P}}$ implies that

$$\sigma_k(\mathcal{M}_1(\tilde{\mathcal{P}})(V \otimes U)) = \sigma_k(U\mathcal{M}_1(\mathcal{C})(V \otimes U)^T(V \otimes U)) = \sigma_k(\mathcal{M}_1(\mathcal{C})),$$

where the last equality follows from the fact that U has orthonormal columns and hence does not affect the singular values. It then follows from the Tucker decomposition of $\mathcal{B} \times_1 \Gamma \times_2 \Gamma$ that

$$\sigma_k(\mathcal{M}_1(\mathcal{C})) = \sigma_k(\mathcal{O}\mathcal{M}_1(\mathcal{C})(V \otimes \mathcal{O})^T) = \sigma_k(\mathcal{M}_1(\mathcal{B} \times_1 \Gamma \times_2 \Gamma)).$$

Let $\mathbf{F}^{(l)} = \Gamma\mathcal{B}_{:, :, l}\Gamma$, for $l \in [L]$, and then we have $\mathcal{M}_1(\mathcal{B} \times_1 \Gamma \times_2 \Gamma) = ([\mathbf{F}^{(1)}, \dots, \mathbf{F}^{(L)}])$. With this, it follows that

$$\sigma_k(\mathcal{M}_1(\tilde{\mathcal{P}})(V \otimes U)) = \sigma_k([\mathbf{F}^{(1)}, \dots, \mathbf{F}^{(L)}]) = \left(\lambda_k \left(\sum_{l=1}^L \mathbf{F}^{(l)} (\mathbf{F}^{(l)})^T \right) \right)^{1/2},$$

where $\lambda_k(\cdot)$ denotes the k -th largest eigenvalue of a symmetric matrix. By Assumptions A, we have

$$\lambda_k(\mathbf{F}^{(l)} (\mathbf{F}^{(l)})^T) \asymp (ns_n \bar{\psi})^2, \text{ for } l \in [L].$$

Further, it follows from Weyl's inequality that

$$\lambda_k \left(\sum_{l=1}^L \mathbf{F}^{(l)} (\mathbf{F}^{(l)})^T \right) \asymp n^2 s_n^2 \bar{\psi}^2 L, \text{ for } k \in [K].$$

The desired result then follows immediately. $\blacksquare$

Lemma 3. Denote $\delta_n = \|\mathcal{M}_1(\tilde{\mathcal{A}})(V \otimes U) - \mathcal{M}_1(\tilde{\mathcal{P}})(V \otimes U)\|$. Then there exists an orthogonal matrix $\mathcal{O}^{(2)}$ such that

$$\|\hat{U} - U\mathcal{O}^{(2)}\|_F \leq \frac{2\sqrt{2} \left(2\sigma_1(\mathcal{M}_1(\tilde{\mathcal{P}})(V \otimes U)) + \delta_n \right) \delta_n}{\sigma_K^2(\mathcal{M}_1(\tilde{\mathcal{P}})(V \otimes U))}.$$

Proof of Lemma 3. By Lemma 1, we know that there exists an orthogonal matrix $\mathbf{O}^{(1)}$ such that, $\mathbf{U}\mathbf{O}^{(1)}$ are the left singular vectors corresponding to the non-zero singular values of $\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})$. Applying similar argument in Lemma 1 to $\tilde{\mathcal{A}}$, there exists an orthogonal matrix $\mathbf{O}^{(3)} \in \mathbb{R}^{K \times K}$ such that $\hat{\mathbf{U}}\mathbf{O}^{(3)}$ are the left singular vectors corresponding to the first K leading singular values of $\mathcal{M}_1(\tilde{\mathcal{A}})(\mathbf{V} \otimes \mathbf{U})$. By Theorem 3 in [2], there exists an orthogonal matrix $\mathbf{O}^{(4)}$ such that

$$\begin{aligned} \|\hat{\mathbf{U}} - \mathbf{U}\mathbf{O}^{(1)}\mathbf{O}^{(4)}(\mathbf{O}^{(3)})^T\|_F &= \|\hat{\mathbf{U}}\mathbf{O}^{(3)} - \mathbf{U}\mathbf{O}^{(1)}\mathbf{O}^{(4)}\|_F \\ &\leq \frac{2\sqrt{2}\left(2\sigma_1(\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})) + \delta_n\right)\delta_n}{\sigma_K^2(\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U}))}. \end{aligned}$$

The desired result then follows immediately by taking $\mathbf{O}^{(2)} = \mathbf{O}^{(1)}\mathbf{O}^{(4)}(\mathbf{O}^{(3)})^T$.

■

Lemma 4. Let δ_n be defined in Lemma 3. Under Assumptions A to D, it holds true that

$$\delta_n = O_p\left(\sqrt{n\varphi_n \log n}\right),$$

where $\varphi_n = 1 - \min_{i \in [n]} f_i + 4s_n$.

Proof of Lemma 4. Note that

$$\left\|\mathcal{M}_1(\tilde{\mathcal{A}} - \tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})\right\| = \left\|\sum_{l=1}^L \sum_{i \leq j} (\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)}) \mathbf{E}^{(i,j)} \mathbf{V}_{l,:}^T \otimes \mathbf{U}\right\|, \quad (\text{S1.1})$$

where $\mathbf{E}^{(i,j)}$ is the square matrix with 1 in the (i, j) -th and (j, i) -th entries

and 0 otherwise. Clearly, $(\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)})\mathbf{E}^{(i,j)}(\mathbf{V}_{l,:}^T \otimes \mathbf{U})$ are independent zero-mean random matrices, for $i, j \in [n]$, $l \in [L]$.

Next, we proceed to verify the required conditions of the matrix Bernstein inequality (Theorem 1.6 in [1]) in order to provide an probabilistic upper bound for (S1.1). For each $i, j \in [n]$ and $l \in [L]$, we have

$$\left\| (\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)})\mathbf{E}^{(i,j)}\mathbf{V}_{l,:}^T \otimes \mathbf{U} \right\| \leq \begin{cases} \left\| (\mathbf{V}_{l,:} \otimes \mathbf{U}_{j,:}, \mathbf{V}_{l,:} \otimes \mathbf{U}_{i,:}) \right\|, & \text{if } i < j, \\ \left\| (\mathbf{V}_{l,:} \otimes \mathbf{U}_{i,:}) \right\|, & \text{if } i = j. \end{cases} \quad (\text{S1.2})$$

The right-hand side of (S1.2) can be further upper bounded by $\sqrt{2}\|\mathbf{V}_{l,:}\| \max\{\|\mathbf{U}_{i,:}\|, \|\mathbf{U}_{j,:}\|\}$, where the coefficient $\sqrt{2}$ is due to the case that $i < j$ and $c_i^* = c_j^*$. Note that

$$\begin{aligned} \sigma_{L_0}(\mathcal{M}_3(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O})) &= \sigma_{L_0}(\mathbf{V} \mathcal{M}_3(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O})) \\ &= \sigma_{L_0}(\mathcal{M}_3(\mathbf{B})(\mathbf{\Gamma} \otimes \mathbf{\Gamma})) = \Omega(n\bar{\psi}\sqrt{L}s_n), \end{aligned}$$

where the last equality follows from Assumption D. Therefore, for any $l \in [L]$, we have

$$\begin{aligned} \|\mathbf{V}_{l,:}\|_{\sigma_{L_0}}(\mathcal{M}_3(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O})) &\leq \|(\mathbf{V}_{l,:})^T \mathcal{M}_3(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O})\| \\ &= \|(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O} \times_3 \mathbf{V})_{:,l}\|_F \\ &= \|(\mathbf{B} \times_1 \mathbf{\Gamma} \times_2 \mathbf{\Gamma})_{:,l}\|_F = O(n\bar{\psi}s_n). \end{aligned}$$

Combining the above two bounds implies that

$$\|\mathbf{V}_{l,:}\| = \frac{\|\mathbf{V}_{l,:}\|_{\sigma_{L_0}} \left(\mathcal{M}_3(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O}) \right)}{\sigma_{L_0} \left(\mathcal{M}_3(\mathbf{C} \times_1 \mathbf{O} \times_2 \mathbf{O}) \right)} = O\left(\frac{1}{\sqrt{L}}\right).$$

Thus, it follows from Assumption A and B that

$$\left\| (\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)}) \mathbf{E}^{(i,j)} \mathbf{V}_{l,:}^T \otimes \mathbf{U} \right\| \leq \sqrt{\frac{2C_1}{L \min\{n_{c^*i}, n_{c_j^*}\}}} \leq C_2 \sqrt{\frac{1}{nL}},$$

for some absolute constant C_2 . We next proceed to bound the second-order central moment. Denote

$$\begin{aligned} \text{Var}_1 &= \sum_{l=1}^L \sum_{i \leq j} \mathbb{E} \left(\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)} \right)^2 \mathbf{E}^{(i,j)} (\mathbf{V}_{l,:}^T \otimes \mathbf{U}) (\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T \mathbf{E}^{(i,j)}, \\ \text{Var}_2 &= \sum_{l=1}^L \sum_{i \leq j} \mathbb{E} \left(\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)} \right)^2 \left(\mathbf{V}_{l,:}^T \otimes \mathbf{U} \right)^T \mathbf{E}^{(i,j)} \mathbf{E}^{(i,j)} (\mathbf{V}_{l,:}^T \otimes \mathbf{U}). \end{aligned}$$

In what follows, we proceed to bound $\|\text{Var}_1\|$ and $\|\text{Var}_2\|$, separately, and then obtain an upper bound for $\max\{\|\text{Var}_1\|, \|\text{Var}_2\|\}$.

Let $\mathbf{Q} = \mathbf{F}\mathbf{D}$ and then we can verify that $(\mathbf{V}_{l,:}^T \otimes \mathbf{U})(\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T =$

$\|\mathbf{V}_{l,:}\|^2 \mathbf{Q} \mathbf{Z} \mathbf{\Gamma}^{-2} \mathbf{Z}^T \mathbf{Q}$. Thus,

$$\begin{aligned}
 \text{Var}_1 &= \sum_{l=1}^L \|\mathbf{V}_{l,:}\|^2 \sum_{i \leq j} \mathbb{E}(\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)})^2 \mathbf{E}^{(i,j)} \mathbf{Q} \mathbf{Z} \mathbf{\Gamma}^{-2} \mathbf{Z}^T \mathbf{Q} \mathbf{E}^{(i,j)} \\
 &= \sum_{l=1}^L \|\mathbf{V}_{l,:}\|^2 \sum_{i \leq j} (f_i f_j \mathbf{P}_{i,j,l} + \frac{1 - f_i f_j}{2}) (\frac{1 + f_i f_j}{2} - f_i f_j \mathbf{P}_{i,j,l}) \\
 &\quad \times \mathbf{E}^{(i,j)} \mathbf{Q} \mathbf{Z} \mathbf{\Gamma}^{-2} \mathbf{Z}^T \mathbf{Q} \mathbf{E}^{(i,j)} \\
 &= \sum_{l=1}^L \|\mathbf{V}_{l,:}\|^2 \sum_{i \leq j} \left(\frac{1}{4} - \frac{f_i^2 f_j^2}{4} (1 - 2\mathbf{P}_{i,j,l})^2 \right) \mathbf{E}^{(i,j)} \mathbf{Q} \mathbf{Z} \mathbf{\Gamma}^{-2} \mathbf{Z}^T \mathbf{Q} \mathbf{E}^{(i,j)} \\
 &\preceq \sum_{l=1}^L \frac{\varphi_n \|\mathbf{V}_{l,:}\|^2}{4} \sum_{i \leq j} \mathbf{E}^{(i,j)} \mathbf{Q} \mathbf{Z} \mathbf{\Gamma}^{-2} \mathbf{Z}^T \mathbf{Q} \mathbf{E}^{(i,j)} = \sum_{l=1}^L \frac{\varphi_n \|\mathbf{V}_{l,:}\|^2}{4} (K \mathbf{I}_n + \mathbf{G}_c),
 \end{aligned}$$

where $\varphi_n = 1 - \min_{i \in [n]} f + 4s_n$, $\mathbf{I}_n$ is the n -dimensional identity matrix, and $\mathbf{G}_c$ is a $n \times n$ matrix such that $(\mathbf{G}_c)_{ii} = 0$ for $i \in [n]$, $(\mathbf{G}_c)_{ij} = (f_i f_j d_i d_j) \gamma_{c_i^*}^{-1}$ if $c_i^* = c_j^*$ and 0 otherwise for $i \neq j$. Herein, the partial order $\preceq$ between two matrix $\mathbf{M}^{(1)}$ and $\mathbf{M}^{(2)}$ is defined as $\mathbf{M}^{(1)} \preceq \mathbf{M}^{(2)}$ if and only if $\mathbf{M}^{(2)} - \mathbf{M}^{(1)}$ is positive semi-definite. This leads to

$$\|\text{Var}_1\| \leq \sum_{l=1}^L \frac{\varphi_n \|\mathbf{V}_{l,:}\|^2}{4} \left(K + \max_{i \in [n]} \sqrt{\frac{f_i^2 d_i^2 n_{c_i^*}}{\gamma_{c_i^*}}} \right) \leq \sum_{l=1}^L \frac{\varphi_n \|\mathbf{V}_{l,:}\|^2}{4} (K + \sqrt{C_1}),$$

where the first inequality follows from the triangle inequality, Gershgorin circle theorem and Cauchy-Schwarz inequality, and the second inequality follows from Assumption B.

Next, we turn to bound the spectral norm of Var_2 . Note that

$$\begin{aligned}
 \text{Var}_2 &= \sum_{l=1}^L \sum_{i \leq j} \mathbb{E}(\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)})^2 (\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T \mathbf{I}^{(i,j)} (\mathbf{V}_{l,:}^T \otimes \mathbf{U}) \\
 &= \sum_{l=1}^L (\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T \left(\sum_{i \leq j} \mathbb{E}(\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)})^2 \mathbf{I}^{(i,j)} \right) (\mathbf{V}_{l,:}^T \otimes \mathbf{U}) \\
 &= \sum_{l=1}^L (\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T \left(\sum_{i \leq j} \left(\frac{1}{4} - \frac{f_i^2 f_j^2}{4} (1 - 2\mathcal{P}_{i,j,l})^2 \right) \mathbf{I}^{(i,j)} \right) (\mathbf{V}_{l,:}^T \otimes \mathbf{U}) \\
 &\preceq \sum_{l=1}^L \frac{n\varphi_n}{4} (\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T (\mathbf{V}_{l,:}^T \otimes \mathbf{U}),
 \end{aligned}$$

where $\mathbf{I}^{(i,j)}$ is the diagonal matrix with the (i, i) -th and (j, j) -th entries being 1 and all other entries being zero. Hence,

$$\|\text{Var}_2\| \leq \left\| \sum_{l=1}^L \frac{n\varphi_n}{4} (\mathbf{V}_{l,:}^T \otimes \mathbf{U})^T (\mathbf{V}_{l,:}^T \otimes \mathbf{U}) \right\| \leq \sum_{l=1}^L \frac{n\varphi_n \|\mathbf{V}_{l,:}\|^2}{4}.$$

Since $n \gg (K + C_1)$, the variance condition can be met by the fact that

$$\sigma^2 = \max\{\|\text{Var}_1\|, \|\text{Var}_2\|\} \leq \sum_{l=1}^L n\varphi_n \|\mathbf{V}_{l,:}\|^2 / 4.$$

in [1] yields that, for any $t > 0$,

$$\begin{aligned}
 &\mathbb{P}\left(\left\| \sum_{l=1}^L \sum_{i \leq j} (\tilde{\mathbf{A}}_{i,j}^{(l)} - \tilde{\mathbf{P}}_{i,j}^{(l)}) \mathbf{E}^{(i,j)} (\mathbf{V}_{l,:}^T \otimes \mathbf{U}) \right\| \geq t \right) \\
 &\leq (n + KL_0) \exp \left\{ - \frac{6t^2}{3 \sum_{l=1}^L n\varphi_n \|\mathbf{V}_{l,:}\|^2 + \frac{4C_2}{\sqrt{Ln}} t} \right\}.
 \end{aligned}$$

By the upper bound of $\|\mathbf{V}_{l,:}\|$, there exists an absolute constant C_3 , such

that for any $t > 0$,

$$\mathbb{P}\left(\left\| \mathcal{M}_1(\tilde{\mathcal{A}} - \tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U}) \right\| \geq t \right) \leq \exp \left\{ 2 \log n - \frac{C_3 t^2}{n\varphi_n + \frac{t}{\sqrt{nL}}} \right\}.$$

Taking $t = \sqrt{\frac{8}{C_3} \varphi_n n \log n}$, then with probability at least $1 - n^{-2}$, we have

$$\left\| \mathcal{M}_1(\tilde{\mathcal{A}} - \tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U}) \right\| < \sqrt{\frac{8}{C_3} \varphi_n n \log n}. \quad (\text{S1.3})$$

This completes the proof. $\blacksquare$

Proof of Theorem 1. Let $\hat{\mathbf{U}} \in \mathbb{R}^{n \times K}$ be any matrix having orthonormal columns such that the column space of $\hat{\mathbf{U}}$ is the same as the one that spanned by the first K leading left singular vectors of $\mathcal{M}_1(\tilde{\mathcal{A}})$. Then the $(1 + \tau)$ -optimal K -medians algorithm is applied to estimate assignment matrix and spectral embedding centers, which finds a pair of solution $(\hat{\mathbf{Z}}, \hat{\mathbf{W}})$ such that

$$\|\hat{\mathbf{Z}}\hat{\mathbf{W}} - \hat{\mathbf{U}}\|_{2,1} \leq (1 + \tau) \min_{\mathbf{Z} \in \Delta, \mathbf{W} \in \mathbb{R}^{K \times K}} \|\mathbf{Z}\mathbf{W} - \hat{\mathbf{U}}\|_{2,1}.$$

Define $\mathcal{S}_k = \{i : c_i^* = k, \|\hat{\mathbf{Z}}_{i,:}\hat{\mathbf{W}} - \tilde{\mathbf{U}}_{i,:}\mathbf{O}^{(2)}\| \geq \sqrt{2}/2\}$, where $\mathbf{O}^{(2)}$ is defined as in Lemma 3. It is easy to show that

$$\begin{aligned} \|\tilde{\mathbf{U}}\mathbf{O}^{(2)} - \hat{\mathbf{Z}}\hat{\mathbf{W}}\|_{2,1} &\leq \|\tilde{\mathbf{U}}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_{2,1} + \|\hat{\mathbf{U}} - \hat{\mathbf{Z}}\hat{\mathbf{W}}\|_{2,1} \\ &\leq (2 + \tau)\|\tilde{\mathbf{U}}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_{2,1}, \end{aligned}$$

where the last inequality follows from the $(1 + \tau)$ -optimality of $(\hat{\mathbf{Z}}, \hat{\mathbf{W}})$.

$$\|\tilde{\mathbf{U}}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_{2,1} = \sum_{i=1}^n \left\| \frac{\mathbf{U}_{i,:}}{\|\mathbf{U}_{i,:}\|} \mathbf{O}^{(2)} - \frac{\hat{\mathbf{U}}_{i,:}}{\|\hat{\mathbf{U}}_{i,:}\|} \right\| \leq 2 \sum_{i=1}^n \frac{\|\mathbf{U}_{i,:}\mathbf{O}^{(2)} - \hat{\mathbf{U}}_{i,:}\|}{\|\mathbf{U}_{i,:}\mathbf{O}^{(2)}\|},$$

where the inequality follows from the fact that $\left\| \frac{v_1}{\|v_1\|} - \frac{v_2}{\|v_2\|} \right\| \leq 2 \frac{\|v_1 - v_2\|}{\|v_1\|}$ for any two vectors with same dimension. By the Cauchy-Swartz inequality, we

further have

$$\begin{aligned}
 \|\tilde{\mathbf{U}}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_{2,1} &\leq 2\sqrt{\sum_{i=1}^n \|\mathbf{U}_{i,:}\mathbf{O}^{(2)} - \hat{\mathbf{U}}_{i,:}\|^2 \sum_{j=1}^n \frac{1}{\|\mathbf{U}_{j,:}\mathbf{O}^{(2)}\|^2}} \\
 &= 2\sqrt{\sum_{i=1}^n \|\mathbf{U}_{i,:}\mathbf{O}^{(2)} - \hat{\mathbf{U}}_{i,:}\|^2} \sqrt{\sum_{j=1}^n \frac{\gamma_{c_j^*}}{f_j^2 d_j^2}} \\
 &= 2\|\mathbf{U}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_F \sqrt{\sum_{j=1}^n \frac{\gamma_{c_j^*}}{f_j^2 d_j^2}} = 2\|\mathbf{U}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_F \sqrt{\sum_{k=1}^K n_k^2 v_k},
 \end{aligned}$$

where $v_k = n_k^{-2} \sum_{c_i^*=k} \gamma_k (f_i d_i)^{-2}$. Here it can be verified that v_k takes value in $[1, \infty)$ and higher privacy requirement of nodes within community k leads to larger value of v_k , which then leads to a slower convergence rate.

Next, we proceed to establish the connection between the Hamming error of $\hat{\mathbf{c}}$ and $\|\mathbf{U}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_F^2$. It is straightforward to verified that

$$\begin{aligned}
 Err(\hat{\mathbf{c}}, \mathbf{c}^*) &\leq \frac{1}{n} \sum_{k=1}^K |S_k| \\
 &\leq \frac{\sqrt{2}}{n} \|\tilde{\mathbf{U}}\mathbf{O}^{(2)} - \hat{\mathbf{Z}}\hat{\mathbf{W}}\|_{2,1} \\
 &\leq \frac{\sqrt{2}(4 + 2\tau)}{n} \|\mathbf{U}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_F \sqrt{\sum_{k=1}^K n_k^2 v_k}.
 \end{aligned}$$

Combined with Lemma 3 and 4, it follows that

$$\begin{aligned}
 Err(\hat{\mathbf{c}}, \mathbf{c}^*) &\leq \frac{\sqrt{2}(4+2\tau)}{n} \|\mathbf{U}\mathbf{O}^{(2)} - \hat{\mathbf{U}}\|_F \sqrt{\sum_{k=1}^K n_k^2 v_k} \\
 &\leq \frac{\sqrt{\sum_{k=1}^K n_k^2 v_k} (16+8\tau)}{n} \frac{(2\sigma_1(\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U})) + \delta_n) \delta_n}{\sigma_K^2(\mathcal{M}_1(\tilde{\mathcal{P}})(\mathbf{V} \otimes \mathbf{U}))} \\
 &= O_p\left(\sqrt{\sum_{k=1}^K v_k} \left(\frac{\sqrt{\varphi_n \log n}}{\sqrt{nLs_n\bar{\psi}}} + \frac{\varphi_n \log n}{ns_n^2 \bar{\psi}^2 L}\right)\right) \\
 &= O_p\left(\sqrt{\sum_{k=1}^K v_k} \frac{\sqrt{\varphi_n \log n}}{\sqrt{nLs_n\bar{\psi}}}\right).
 \end{aligned}$$

where the last equality follows from the lower of s_n in Assumption 3, yielding

that $\frac{\sqrt{\varphi_n \log n}}{\sqrt{nLs_n\bar{\psi}}}$ vanishes. This completes the proof. $\blacksquare$

Proof of Corollary 1. (1) Suppose $f_i f_j \asymp \alpha_n^2$, for $i, j \in [n]$, under the condition of Corollary 1, we have $v_k = \frac{\gamma_k}{n_k^2} \sum_{c_i^*=k} \frac{1}{f_i^2 d_i^2} \asymp \frac{n_k \alpha_n^2}{n_k^2} \cdot \frac{n_k}{\alpha_n^2} \asymp 1$, $\varphi_n = 1 - \alpha + 4s_n = O(1)$, and $\bar{\psi} = \frac{1}{n} \sum_{i=1}^n (f_i d_i)^2 \asymp \alpha_n^2$. By the result of theorem 1, we have

$$Err(\hat{\mathbf{c}}, \mathbf{c}^*) = O_P\left(\sqrt{\frac{\log n}{nLs_n^2 \alpha_n^4}}\right) = o_p(1).$$

(2) Under the condition of Corollary 1, we have

$$\begin{aligned}
 v_k &= \frac{\gamma_k}{n_k^2} \sum_{c_i^*=k} \frac{1}{f_i^2 d_i^2} \asymp \frac{(\alpha_n^2 \beta_n n_k + (1 - \beta_n) n_k) (\beta_n n_k / \alpha_n^2 + (1 - \beta_n) n_k)}{n_k^2} \\
 &\asymp (1 - \beta_n) (\beta_n / \alpha_n^2 + 1 - \beta_n).
 \end{aligned}$$

Therefore, $\sqrt{\sum_k v_k} \asymp \sqrt{(1 - \beta_n) (\beta_n / \alpha_n^2 + 1 - \beta_n)}$, and $\varphi_n = 1 - \alpha_n + 4s_n =$

$O(1)$. In addition,

$$\bar{\psi} = \frac{1}{n} \sum_{i=1}^n (f_i d_i)^2 = \frac{1}{n} \sum_{k=1}^K (\alpha_n^2 \beta_n n_k + (1 - \beta_n) n_k) = \Omega(1 - \beta_n).$$

Note that the condition $\frac{\beta_n}{\alpha_n^2(1-\beta_n)} \ll \frac{nLs_n^2}{\log n}$ and $\frac{\log n}{nLs_n^2} = o(1)$ imply that

$(1 - \beta_n)(\frac{\beta_n}{\alpha_n^2} + 1 - \beta_n) \ll \frac{nLs_n^2 \bar{\psi}^2}{\log n}$. By the result of Theorem 1, we have

$$Err(\hat{\mathbf{c}}, \mathbf{c}^*) = O_p\left(\sqrt{(1 - \beta_n)\left(\frac{\beta_n}{\alpha_n^2} + 1 - \beta_n\right)} \sqrt{\frac{\varphi_n \log n}{nLs_n^2 \bar{\psi}^2}}\right) = o_p(1).$$

■

Bibliography

[1] Joel A Tropp. User-friendly tail bounds for sums of random matrices.

Foundations of Computational Mathematics, 12(4):389–434, 2012.

[2] Yi Yu, Tengyao Wang, and Richard J Samworth. A useful variant of the

davis–kahan theorem for statisticians. *Biometrika*, 102(2):315–323, 2015.