

OPTIMAL MODEL AVERAGING BASED ON GENERALIZED METHOD OF MOMENTS

Supplementary Material

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The following two lemmas will be used in the proofs of Proposition 1 and Theorem 1, respectively.

Lemma 1 (*Stein, 1981*) *Let $a \sim \text{Normal}(0, 1)$ and $g(a) : \mathcal{R} \rightarrow \mathcal{R}$ be an indefinite integral of the Lebesgue measurable function $\dot{g}(a)$. Thus, $\dot{g}(a)$ is the derivative of $g(a)$. Suppose that $E|\dot{g}(a)| < \infty$. Then we have $E\{\dot{g}(a)\} = E\{ag(a)\}$.*

Lemma 2 (*Zhang, 2010; Gao et al., 2019*) *Let*

$$\tilde{\mathbf{w}} = \operatorname{argmin}_{\mathbf{w} \in \mathcal{W}} \{L(w) + a_n(w) + b_n\},$$

where $a_n(\mathbf{w})$ is a term related to \mathbf{w} and b_n is a term unrelated to \mathbf{w} . If

$$\sup_{\mathbf{w} \in \mathcal{W}} |a_n(\mathbf{w})|/L^*(\mathbf{w}) = o_p(1), \quad \sup_{\mathbf{w} \in \mathcal{W}} |L(\mathbf{w}) - L^*(\mathbf{w})|/L^*(\mathbf{w}) = o_p(1),$$

and there exists a positive constant c and a positive integer N such that when

$n \geq N$, $\inf_{\mathbf{w} \in \mathcal{W}} L^*(\mathbf{w}) \geq c > 0$ almost surely, then $L(\tilde{\mathbf{w}})/\inf_{\mathbf{w} \in \mathcal{W}} L(\mathbf{w}) \rightarrow 1$ in probability.

S.1 Proof of Proposition 1

Let $f(\cdot)$ be a function with $f[\sqrt{n}\{\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)\}] = \sqrt{n}\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \sqrt{n}\hat{\boldsymbol{\mu}}$.

It is seen that

$$\begin{aligned}
 R(\mathbf{w}) & \tag{S.1} \\
 &= E \left([\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)]^T \boldsymbol{\Omega} [\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right) \\
 &= E \left([\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \hat{\boldsymbol{\mu}} + \hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)]^T \boldsymbol{\Omega} [\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \hat{\boldsymbol{\mu}} + \hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right) \\
 &= E \left([\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \hat{\boldsymbol{\mu}}]^T \boldsymbol{\Omega} [\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \hat{\boldsymbol{\mu}}] \right) + E \left[[\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)]^T \boldsymbol{\Omega} [\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right] \\
 &\quad + 2E \left([\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \hat{\boldsymbol{\mu}}]^T \boldsymbol{\Omega} [\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right) \tag{S.2}
 \end{aligned}$$

and

$$\begin{aligned}
 & E \left([\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \hat{\boldsymbol{\mu}}]^T \boldsymbol{\Omega} [\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right) \\
 &= n^{-1} E \left([\sqrt{n}\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \sqrt{n}\hat{\boldsymbol{\mu}}]^T \boldsymbol{\Omega} \sqrt{n}[\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right) \\
 &= n^{-1} E \left(f[\sqrt{n}\{\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)\}]^T \boldsymbol{\Omega} \sqrt{n}[\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)] \right) \\
 &= n^{-1} \left[E \left\{ f(\boldsymbol{\pi})^T \boldsymbol{\Omega} \boldsymbol{\pi} \right\} + o(1) \right] \\
 &= n^{-1} \left[E \left(\text{trace} \left\{ \frac{\partial f(\boldsymbol{\pi})}{\partial \boldsymbol{\pi}^T} \boldsymbol{\Omega} \mathbf{V} \right\} \right) + o(1) \right] \\
 &= n^{-1} \left[E \left(\text{trace} \left[\frac{\partial(\sqrt{n}\boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\} - \sqrt{n}\hat{\boldsymbol{\mu}})}{\partial \sqrt{n}\{\hat{\boldsymbol{\mu}} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0)\}^T} \boldsymbol{\Omega} \mathbf{V} \right] \right) + o(1) \right]
 \end{aligned}$$

$$\begin{aligned}
&= n^{-1}E \left(\text{trace} \left[\frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\mu}}^T} \boldsymbol{\Omega} \mathbf{V} \right] \right) - n^{-1} \text{trace}(\boldsymbol{\Omega} \mathbf{V}) + o(n^{-1}) \\
&= n^{-1}E \left(\text{trace} \left[\sum_{m=1}^M w_m \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^T} \boldsymbol{\Pi}_m^T \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^T} \boldsymbol{\Omega} \mathbf{V} \right] \right) - n^{-1} \text{trace}(\boldsymbol{\Omega} \mathbf{V}) + o(n^{-1}),
\end{aligned}$$

where the third, fourth and fifth steps are from Lemma 1 and Conditions (C.1)-(C.2). The above two formulas imply (3.6). This completes the proof.

S.2 Proof of Proposition 2

It is implied by (2.5) that

$$\frac{\partial \{\widehat{\boldsymbol{\mu}} - \boldsymbol{\mu}(\boldsymbol{\Pi}_m^T \widehat{\boldsymbol{\theta}}_m)\}^T \boldsymbol{\Omega} \{\widehat{\boldsymbol{\mu}} - \boldsymbol{\mu}(\boldsymbol{\Pi}_m^T \widehat{\boldsymbol{\theta}}_m)\}}{\partial \widehat{\boldsymbol{\theta}}_m} = \mathbf{0}, \quad (\text{S.3})$$

which is

$$\mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \left\{ \widehat{\boldsymbol{\mu}} - \boldsymbol{\mu}(\boldsymbol{\Pi}_m^T \widehat{\boldsymbol{\theta}}_m) \right\} = \mathbf{0}. \quad (\text{S.4})$$

Taking derivative of the both sides of (S.4) with respect to $\widehat{\boldsymbol{\mu}}^T$, we have

$$\sum_{\tau=1}^{d_m} \mathbf{A}_\tau(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \left\{ \widehat{\boldsymbol{\mu}} - \boldsymbol{\mu}(\boldsymbol{\Pi}_m^T \widehat{\boldsymbol{\theta}}_m) \right\} \frac{\partial \widehat{\theta}_{m,\tau}}{\partial \widehat{\boldsymbol{\mu}}^T} + \mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \quad (\text{S.5})$$

$$- \sum_{\tau=1}^{d_m} \mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \frac{\partial \boldsymbol{\mu}(\boldsymbol{\Pi}_m^T \widehat{\boldsymbol{\theta}}_m)}{\partial \widehat{\theta}_{m,\tau}} \frac{\partial \widehat{\theta}_{m,\tau}}{\partial \widehat{\boldsymbol{\mu}}^T} = \mathbf{0}. \quad (\text{S.6})$$

From the definitions of \mathbf{D}_m and \mathbf{B}_m in (3.11) and (3.12), Equation (S.5) is simplified to

$$\mathbf{D}_m \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^T} + \mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} - \mathbf{B}_m \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^T} = \mathbf{0},$$

which implies

$$(\mathbf{D}_m - \mathbf{B}_m)^\top (\mathbf{D}_m - \mathbf{B}_m) \frac{\partial \hat{\boldsymbol{\theta}}_m}{\partial \hat{\boldsymbol{\mu}}^\top} = -(\mathbf{D}_m - \mathbf{B}_m)^\top \mathbf{A}(\hat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega},$$

which, along with the condition that $(\mathbf{D}_m - \mathbf{B}_m)^\top (\mathbf{D}_m - \mathbf{B}_m)$ is invertible, implies (3.13). This completes the proof.

S.3 Proofs of (3.16), (3.17), (3.20) and (3.21)

Let $\hat{\mathbf{B}}_m = \mathbf{A}(\hat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \mathbf{A}^\top(\hat{\boldsymbol{\theta}}_m)$. Then, we have

$$\begin{aligned} & \text{trace} \left[\sum_{m=1}^M w_m \frac{\partial \boldsymbol{\mu}\{\hat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \hat{\boldsymbol{\theta}}(\mathbf{w})^\top} \boldsymbol{\Pi}_m^\top \frac{\partial \hat{\boldsymbol{\theta}}_m}{\partial \hat{\boldsymbol{\mu}}^\top} \boldsymbol{\Omega} \hat{\mathbf{V}} \right] \\ &= \text{trace} \left\{ \sum_{m=1}^M w_m \frac{\mathbf{X}^\top \mathbf{X}}{n} \boldsymbol{\Pi}_m^\top (\hat{\mathbf{B}}_m^\top \hat{\mathbf{B}}_m)^{-1} \hat{\mathbf{B}}_m^\top \mathbf{A}(\hat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \boldsymbol{\Omega} \hat{\mathbf{V}} \right\} \\ &= \hat{\sigma}^2 \text{trace} \left\{ \sum_{m=1}^M w_m \mathbf{A}(\hat{\boldsymbol{\theta}}_m)^\top (\hat{\mathbf{B}}_m^\top \hat{\mathbf{B}}_m)^{-1} \hat{\mathbf{B}}_m^\top \mathbf{A}(\hat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \right\} \\ &= \hat{\sigma}^2 \text{trace} \left\{ \sum_{m=1}^M w_m \mathbf{A}(\hat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \mathbf{A}(\hat{\boldsymbol{\theta}}_m)^\top (\hat{\mathbf{B}}_m^\top \hat{\mathbf{B}}_m)^{-1} \hat{\mathbf{B}}_m^\top \right\} \\ &= \hat{\sigma}^2 \text{trace} \left\{ \sum_{m=1}^M w_m \hat{\mathbf{B}}_m (\hat{\mathbf{B}}_m^\top \hat{\mathbf{B}}_m)^{-1} \hat{\mathbf{B}}_m^\top \right\} \\ &= \hat{\sigma}^2 \sum_{m=1}^M w_m d_m, \end{aligned} \tag{S.7}$$

where the first step is from (3.13)-(3.15) and the second step is from (3.14)-(3.15). Hence, (3.16) is proved.

From (3.14) and (3.16), we have

$$C(\mathbf{w})$$

S.3 Proofs of (3.16), (3.17), (3.20) and (3.21)

$$\begin{aligned}
&= [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}]^\top \boldsymbol{\Omega} [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}] \\
&\quad + 2n^{-1} \text{trace} \left[\sum_{m=1}^M w_m \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top} \boldsymbol{\Pi}_m^\top \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^\top} \boldsymbol{\Omega} \widehat{\mathbf{V}} \right] \\
&= [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}]^\top \boldsymbol{\Omega} [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}] + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m \\
&= n^{-1} \{ \mathbf{X}^\top \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \mathbf{X}^\top \mathbf{y} \}^\top (\mathbf{X}^\top \mathbf{X})^{-1} \{ \mathbf{X}^\top \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \mathbf{X}^\top \mathbf{y} \} + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m \\
&= n^{-1} \left\{ \boldsymbol{\theta}(\mathbf{w})^\top \mathbf{X}^\top \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) + \mathbf{y}^\top \mathbf{X} (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} - 2\mathbf{y}^\top \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) \right\} + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m \\
&= n^{-1} \| \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \mathbf{y} \|^2 + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m - \mathbf{y}^\top \{ \mathbf{I}_n - \mathbf{X} (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \} \mathbf{y},
\end{aligned}$$

which is (3.17).

The proof of (3.20) is exactly the same as that of (3.16). For (3.21),

$$\begin{aligned}
&C(\mathbf{w}) \\
&= [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}]^\top \boldsymbol{\Omega} [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}] \\
&\quad + 2n^{-1} \text{trace} \left[\sum_{m=1}^M w_m \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top} \boldsymbol{\Pi}_m^\top \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^\top} \boldsymbol{\Omega} \widehat{\mathbf{V}} \right] \\
&= [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}]^\top \boldsymbol{\Omega} [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}] + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m \\
&= n^{-1} \{ \mathbf{Z}^\top \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \mathbf{Z}^\top \mathbf{y} \}^\top (\mathbf{Z}^\top \mathbf{Z})^{-1} \{ \mathbf{Z}^\top \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \mathbf{Z}^\top \mathbf{y} \} + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m \\
&= n^{-1} \left\{ \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top \mathbf{X}^\top \mathbf{P}_Z \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) + \mathbf{y}^\top \mathbf{P}_Z \mathbf{y} - 2\mathbf{y}^\top \mathbf{P}_Z \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) \right\} + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m \\
&= n^{-1} \| \mathbf{P}_Z \mathbf{X} \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \mathbf{y} \|^2 + 2n^{-1} \widehat{\sigma}^2 \sum_{m=1}^M w_m d_m - \mathbf{y}^\top (\mathbf{I}_n - \mathbf{P}_Z) \mathbf{y}.
\end{aligned}$$

Hence, (3.21) is proved.

S.4 Examples where Conditions (C.3)-(C.5) and (C.7) are satisfied

We first consider the example with the linear regression candidate models, which are described in Remark 1 detailedly. In this example, $\mathbf{V} = \sigma^2 E(\mathbf{X}_i \mathbf{X}_i^\top)$, $\widehat{\mathbf{V}} = \widehat{\sigma}^2 \mathbf{X}^\top \mathbf{X}/n$, $\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}/\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top |_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\tilde{\boldsymbol{\theta}}_{\mathbf{w}}} = \mathbf{X}^\top \mathbf{X}/n$, and

$$\begin{aligned} \tilde{\boldsymbol{\theta}}_m &= (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{y} \\ &= (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top (\mathbf{X} \boldsymbol{\theta} + \boldsymbol{\epsilon}) \\ &= (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{X} \boldsymbol{\theta} + (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \boldsymbol{\epsilon}. \end{aligned}$$

Therefore, when $\mathbf{X}^\top \mathbf{X}/n$ converges to a positive definite matrix, $\mathbf{X}^\top \boldsymbol{\epsilon}/n = o_p(1)$ and $\widehat{\sigma}^2 - \sigma^2 = o_p(1)$, Conditions (C.3)-(C.5) and (C.7) are satisfied in this example.

Second, we consider the example with linear regression models with instrumental variables, which are described in Remark 2 detailedly. In this example, $\boldsymbol{\Omega} = (\mathbf{Z}^\top \mathbf{Z}/n)^{-1}$, $\mathbf{V} = \sigma^2 E(\mathbf{Z}_i \mathbf{Z}_i^\top)$ with \mathbf{Z}_i^\top being the i^{th} row of \mathbf{Z} , $\widehat{\mathbf{V}} = \widehat{\sigma}^2 \mathbf{Z}^\top \mathbf{Z}/n$, $\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}/\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top |_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\tilde{\boldsymbol{\theta}}_{\mathbf{w}}} = \mathbf{Z}^\top \mathbf{X}/n$, and

$$\begin{aligned} \tilde{\boldsymbol{\theta}}_m &= (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \mathbf{y} \\ &= (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z (\mathbf{X} \boldsymbol{\theta} + \boldsymbol{\epsilon}) \\ &= (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \mathbf{X} \boldsymbol{\theta} + (\boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \mathbf{X} \boldsymbol{\Pi}_m^\top)^{-1} \boldsymbol{\Pi}_m \mathbf{X}^\top \mathbf{P}_Z \boldsymbol{\epsilon}. \end{aligned}$$

Therefore, when $\mathbf{Z}^\top \mathbf{Z}/n$ converges to a positive definite matrices, $\mathbf{Z}^\top \mathbf{X}/n$

converges to a matrix with full column rank, $\mathbf{Z}^T \boldsymbol{\epsilon}/n = o_p(1)$ and $\widehat{\sigma}^2 - \sigma^2 = o_p(1)$, Conditions (C.3)-(C.5) and (C.7) are satisfied in this example.

S.5 Proof of Theorem 1

It is well-known that the following equalities are satisfied for any matrices

\mathbf{B}_1 and \mathbf{B}_2 with identical dimensions (see, for example, Li (1987)):

$$\lambda_{\max}(\mathbf{B}_1 + \mathbf{B}_2) \leq \lambda_{\max}(\mathbf{B}_1) + \lambda_{\max}(\mathbf{B}_2) \text{ and } \lambda_{\max}(\mathbf{B}_1 \mathbf{B}_2) \leq \lambda_{\max}(\mathbf{B}_1) \lambda_{\max}(\mathbf{B}_2) \quad (\text{S.8})$$

where the definition of $\lambda_{\max}(\cdot)$ is in Condition (C.5).

Now, we prove that uniformly for any $m \in \{1, \dots, M\}$,

$$\lambda_{\max} \left(\frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^T} \right) = O_p(1). \quad (\text{S.9})$$

Let $\mathbf{P}_{\text{BD}} = (\mathbf{D}_m - \mathbf{B}_m) \{(\mathbf{D}_m - \mathbf{B}_m)^T (\mathbf{D}_m - \mathbf{B}_m)\}^{-1} (\mathbf{D}_m - \mathbf{B}_m)^T$. By

(3.13), (S.8), the assumption that $(\mathbf{D}_m - \mathbf{B}_m)^T (\mathbf{D}_m - \mathbf{B}_m)$ is invertible,

and the truth that $\boldsymbol{\Omega}$ is a positive definite matrix, we have that uniformly

for $m \in \{1, \dots, M\}$,

$$\begin{aligned} & \lambda_{\max} \left(\frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^T} \right) \\ &= \lambda_{\max}^{1/2} \left(\frac{\partial \widehat{\boldsymbol{\theta}}_m^T}{\partial \widehat{\boldsymbol{\mu}}} \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^T} \right) \\ &= \lambda_{\max}^{1/2} \left(\boldsymbol{\Omega}^T \mathbf{A}(\widehat{\boldsymbol{\theta}}_m)^T (\mathbf{D}_m - \mathbf{B}_m) \{(\mathbf{D}_m - \mathbf{B}_m)^T (\mathbf{D}_m - \mathbf{B}_m)\}^{-2} (\mathbf{D}_m - \mathbf{B}_m)^T \mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \right) \\ &\leq \lambda_{\max}^{1/2} \left(\{(\mathbf{D}_m - \mathbf{B}_m)^T (\mathbf{D}_m - \mathbf{B}_m)\}^{-1} \right) \lambda_{\max}^{1/2} \left(\boldsymbol{\Omega}^T \mathbf{A}(\widehat{\boldsymbol{\theta}}_m)^T \mathbf{P}_{\text{BD}} \mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \boldsymbol{\Omega} \right) \end{aligned}$$

$$\begin{aligned}
 &\leq \lambda_{\max}^{1/2} \left(\{(\mathbf{D}_m - \mathbf{B}_m)^\top (\mathbf{D}_m - \mathbf{B}_m)\}^{-1} \right) \lambda_{\max}^{1/2} (\mathbf{P}_{\mathbf{BD}}) \lambda_{\max} \left(\mathbf{A}(\widehat{\boldsymbol{\theta}}_m) \right) \lambda_{\max}(\boldsymbol{\Omega}) \\
 &= O(1),
 \end{aligned} \tag{S.10}$$

hence, (S.9) is proved.

Let

$$\mathbf{H}_m = 2n^{-1} \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top} \boldsymbol{\Pi}_m^\top \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^\top} \boldsymbol{\Omega} \widehat{\mathbf{V}} \quad \text{and} \quad \mathbf{H}(\mathbf{w}) = \sum_{m=1}^M w_m \mathbf{H}_m.$$

It is seen that

$$\begin{aligned}
 &C(\mathbf{w}) \\
 &= [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}]^\top \boldsymbol{\Omega} [\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \widehat{\boldsymbol{\mu}}] + \text{trace}\{\mathbf{H}(\mathbf{w})\} \\
 &= \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) + \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \right]^\top \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) + \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \right] \\
 &\quad + \text{trace}\{\mathbf{H}(\mathbf{w})\} \\
 &= L(\mathbf{w}) + 2 \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^\top \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} + \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}^\top \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} \\
 &\quad + \text{trace}\{\mathbf{H}(\mathbf{w})\} \\
 &= L(\mathbf{w}) + 2 \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]^\top \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} \\
 &\quad + 2 \left[\boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^\top \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} \\
 &\quad + \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}^\top \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} + \text{trace}\{\mathbf{H}(\mathbf{w})\},
 \end{aligned} \tag{S.11}$$

where the term $\{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}^\top \widehat{\boldsymbol{\Omega}} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}$ is unrelated to \mathbf{w} , and

$$L(\mathbf{w})$$

$$\begin{aligned}
 &= \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^T \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right] \\
 &= \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} + \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^T \boldsymbol{\Omega} \\
 &\quad \times \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} + \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right] \\
 &= L^*(\mathbf{w}) + \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]^T \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right] \\
 &\quad + 2 \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]^T \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]. \quad (\text{S.12})
 \end{aligned}$$

In addition, from Condition (C.6), we know that there exists a positive constant c and a positive integer N such that when $n \geq N$, $\inf_{\mathbf{w} \in \mathcal{W}} L^*(\mathbf{w}) \geq c > 0$ almost surely. Hence, by Lemma 2, to prove (4.2) it is sufficient to verify that

$$\sup_{\mathbf{w} \in \mathcal{W}} |L^*(\mathbf{w})^{-1} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]^T \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]| = o_p(1) \quad (\text{S.13})$$

$$\sup_{\mathbf{w} \in \mathcal{W}} |L^*(\mathbf{w})^{-1} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]^T \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]| = o_p(1) \quad (\text{S.14})$$

$$\sup_{\mathbf{w} \in \mathcal{W}} |L^*(\mathbf{w})^{-1} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right]^T \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}| = o_p(1), \quad (\text{S.15})$$

$$\sup_{\mathbf{w} \in \mathcal{W}} |L^*(\mathbf{w})^{-1} \left[\boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^T \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}| = o_p(1), \quad (\text{S.16})$$

and

$$\sup_{\mathbf{w} \in \mathcal{W}} |n^{-1} L^*(\mathbf{w})^{-1} \text{trace}\{\mathbf{H}(\mathbf{w})\}| = o_p(1). \quad (\text{S.17})$$

By Taylor's expansion, we obtain that

$$\begin{aligned}
& \left\| \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}\{\boldsymbol{\theta}^*(\mathbf{w})\} \right\|^2 \\
&= \left\| \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\tilde{\boldsymbol{\theta}}_{\mathbf{w}}} \{\widehat{\boldsymbol{\theta}}(\mathbf{w}) - \boldsymbol{\theta}^*(\mathbf{w})\} \right\|^2 \\
&\leq \lambda_{\max} \left[\frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\tilde{\boldsymbol{\theta}}_{\mathbf{w}}} \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}^\top}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\tilde{\boldsymbol{\theta}}_{\mathbf{w}}} \right] \left\| \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \boldsymbol{\theta}^*(\mathbf{w}) \right\|^2 \\
&\leq \lambda_{\max}^2 \left[\frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^\top} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\tilde{\boldsymbol{\theta}}_{\mathbf{w}}} \right] \left\| \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \boldsymbol{\theta}^*(\mathbf{w}) \right\|^2 \\
&= O_p(n^{-1}Mp), \tag{S.18}
\end{aligned}$$

where $\tilde{\boldsymbol{\theta}}_{\mathbf{w}}^*$ is a vector between $\widehat{\boldsymbol{\theta}}(\mathbf{w})$ and $\boldsymbol{\theta}^*(\mathbf{w})$ and can be related to \mathbf{w} , the third step is from (S.8), and the last step is from Conditions (C.4) and (C.5).

From (S.18) and Condition (C.6), we can obtain (S.13)-(S.14). From (S.18) and Conditions (C.1), (C.3) and (C.6), we can obtain (S.15). From Conditions (C.1), (C.3) and (C.6), we can obtain (S.16).

It is seen that

$$\begin{aligned}
& \text{trace}\{\mathbf{H}(\mathbf{w})\} \\
&\leq \max_{1 \leq m \leq M} \text{trace}(\mathbf{H}_m) \\
&= 2^{-1} \max_{1 \leq m \leq M} \text{trace}(\mathbf{H}_m + \mathbf{H}_m^\top) \\
&\leq 2^{-1} \max_{1 \leq m \leq M} \text{rank}(\mathbf{H}_m + \mathbf{H}_m^\top) \lambda_{\max}(\mathbf{H}_m + \mathbf{H}_m^\top) \\
&\leq 2 \max_{1 \leq m \leq M} \text{rank}(\mathbf{H}_m) \lambda_{\max}(\mathbf{H}_m)
\end{aligned}$$

$$\begin{aligned}
 &\leq 2 \max_{1 \leq m \leq M} \text{rank}(\mathbf{H}_m) 2n^{-1} \max_{1 \leq m \leq M} \lambda_{\max} \left[\frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^{\text{T}}} \boldsymbol{\Pi}_m^{\text{T}} \frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^{\text{T}}} \boldsymbol{\Omega} \widehat{\mathbf{V}} \right] \\
 &\leq 4n^{-1} p \max_{1 \leq m \leq M} \lambda_{\max} \left[\frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^{\text{T}}} \right] \lambda_{\max}(\boldsymbol{\Pi}_m^{\text{T}}) \lambda_{\max} \left(\frac{\partial \widehat{\boldsymbol{\theta}}_m}{\partial \widehat{\boldsymbol{\mu}}^{\text{T}}} \right) \\
 &\quad \times \lambda_{\max}(\boldsymbol{\Omega}) \lambda_{\max}(\widehat{\mathbf{V}}) \\
 &= O_p(p/n), \tag{S.19}
 \end{aligned}$$

where the fourth and sixth steps use (S.8) and the last step uses (S.9) and Conditions (C.3) and (C.5). Now, by (S.19) and Condition (C.6), we can obtain (S.17). As stated in above (S.13), the optimality (4.2) is implied by (S.13)-(S.17) This completes the proof.

S.6 Proof of Theorem 2

Let

$$\mathbf{G}(\mathbf{w}) = \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}^{\text{T}}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\widetilde{\boldsymbol{\theta}}_{\mathbf{w}}^*} \boldsymbol{\Omega} \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})^{\text{T}}} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\widetilde{\boldsymbol{\theta}}_{\mathbf{w}}^*}$$

and

$$\mathbf{g}(\mathbf{w}) = \frac{\partial \boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\}^{\text{T}}}{\partial \widehat{\boldsymbol{\theta}}(\mathbf{w})} \Big|_{\widehat{\boldsymbol{\theta}}(\mathbf{w})=\widetilde{\boldsymbol{\theta}}_{\mathbf{w}}^*} \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \},$$

where $\widetilde{\boldsymbol{\theta}}_{\mathbf{w}}^*$ is defined following (S.18). It is seen that

$$\begin{aligned}
 &C(\mathbf{w}) \\
 &= \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) + \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \right]^{\text{T}} \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) + \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \right] \\
 &\quad + \text{trace}\{\mathbf{H}(\mathbf{w})\}
 \end{aligned}$$

$$\begin{aligned}
&= \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^T \boldsymbol{\Omega} \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right] \\
&\quad + 2 \left[\boldsymbol{\mu}\{\widehat{\boldsymbol{\theta}}(\mathbf{w})\} - \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) \right]^T \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} \\
&\quad + \text{trace}\{\mathbf{H}(\mathbf{w})\} + \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}^T \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} \\
&= \{ \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \boldsymbol{\theta}_0 \}^T \mathbf{G}(\mathbf{w}) \{ \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \boldsymbol{\theta}_0 \} + 2 \{ \widehat{\boldsymbol{\theta}}(\mathbf{w}) - \boldsymbol{\theta}_0 \}^T \mathbf{g}(\mathbf{w}) + \text{trace}\{\mathbf{H}(\mathbf{w})\} \\
&\quad + \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}^T \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}, \tag{S.20}
\end{aligned}$$

where the first step is from the second step of (S.11) and the last step is from Taylor's expansion. Recall that $\mathbf{w}_{\tilde{m}}$ is a weight vector in which the \tilde{m}^{th} component is one and the other are zeros. From (4.1), (S.19), Conditions (C.1) and (C.3), and the second step of (S.20), we have

$$C(\mathbf{w}_{\tilde{m}}) = \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \}^T \boldsymbol{\Omega} \{ \boldsymbol{\mu}_{\text{true}}(\boldsymbol{\theta}_0) - \widehat{\boldsymbol{\mu}} \} + O_p(n^{-1}p) = O_p(n^{-1}p) \tag{S.21}$$

From (S.19), Condition (C.1) and the third step of (S.20), we have

$$C(\widehat{\mathbf{w}}) = \{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \}^T \mathbf{G}(\widehat{\mathbf{w}}) \{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \} + 2 \{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \}^T \mathbf{g}(\widehat{\mathbf{w}}) + O_p(n^{-1}p).$$

Combining the above equations and $C(\widehat{\mathbf{w}}) \leq C(\mathbf{w}_{\tilde{m}})$, we have

$$\{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \}^T \mathbf{G}(\widehat{\mathbf{w}}) \{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \} + 2 \{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \}^T \mathbf{g}(\widehat{\mathbf{w}}) + O_p(n^{-1}p) \leq O_p(n^{-1}p),$$

from which and Condition (C.7), we further have

$$\begin{aligned}
\kappa_2 \|\widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0\|^2 &\leq -2 \{ \widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0 \}^T \mathbf{g}(\widehat{\mathbf{w}}) - O_p(n^{-1}p) + O_p(n^{-1}p) \\
&\leq 2 \|\widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0\| \|\mathbf{g}(\widehat{\mathbf{w}})\| + O_p(n^{-1}p), \tag{S.22}
\end{aligned}$$

by which, we further have

$$\left\{ \|\widehat{\boldsymbol{\theta}}(\widehat{\mathbf{w}}) - \boldsymbol{\theta}_0\| - \kappa_2^{-1} \|\mathbf{g}(\widehat{\mathbf{w}})\| \right\}^2 \leq \kappa_2^{-2} \|\mathbf{g}(\widehat{\mathbf{w}})\|^2 + O_p(n^{-1}p). \quad (\text{S.23})$$

From Conditions (C.1), (C.3) and (C.5), it is easily to obtain $\|\mathbf{g}(\widehat{\mathbf{w}})\| = O_p(n^{-1/2}p^{1/2})$, which along with (S.23), implies (4.3). This completes the proof.

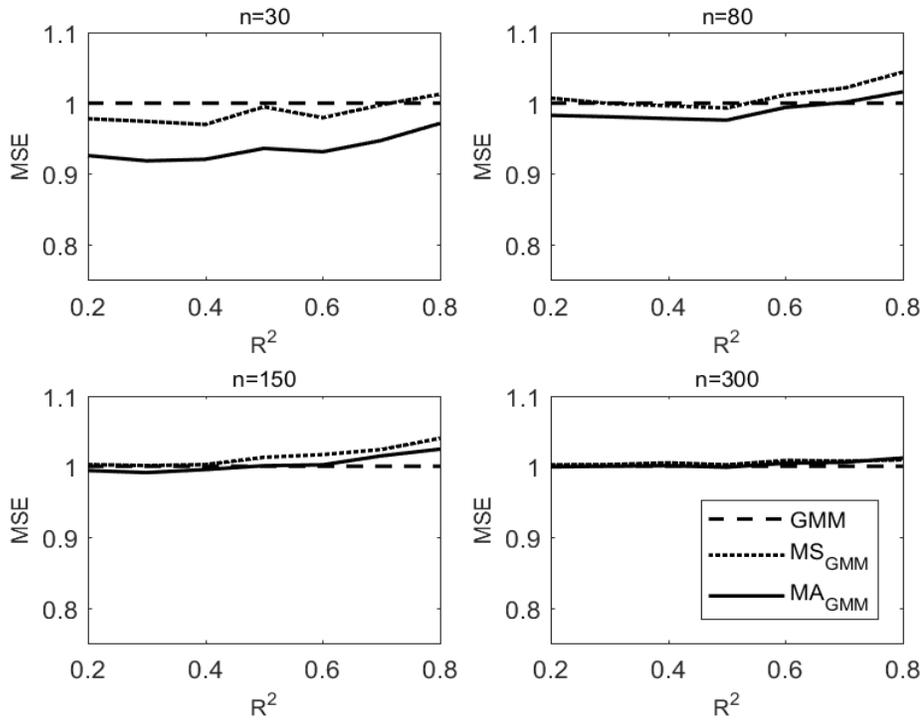
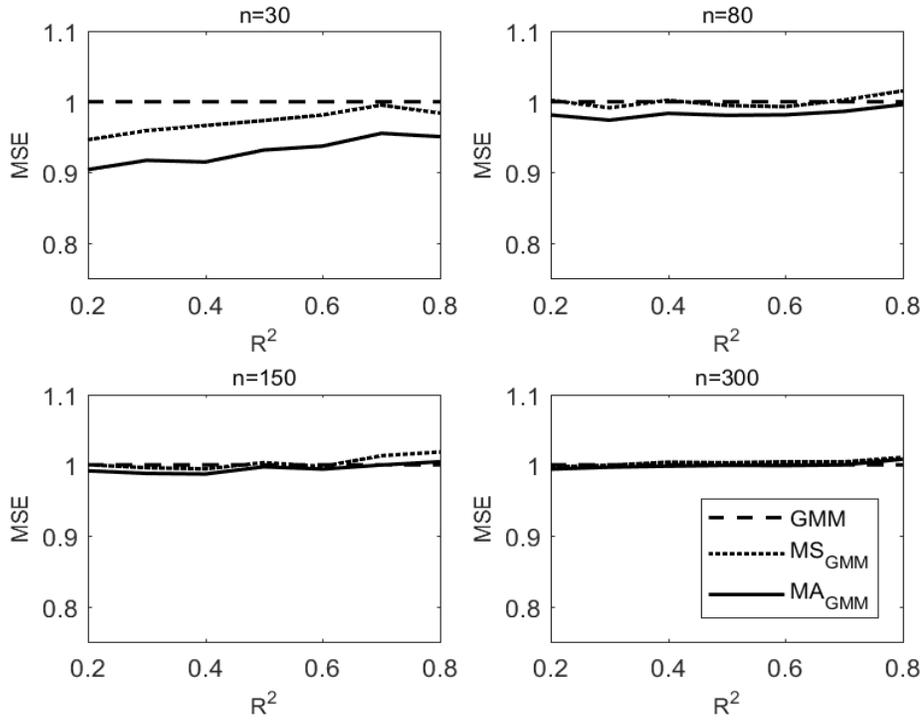


Figure S.1: MSE in simulation Design I, with $\widetilde{R}^2 = 0.5$.

Figure S.2: MSE in simulation Design I, with $\tilde{R}^2 = 0.8$.

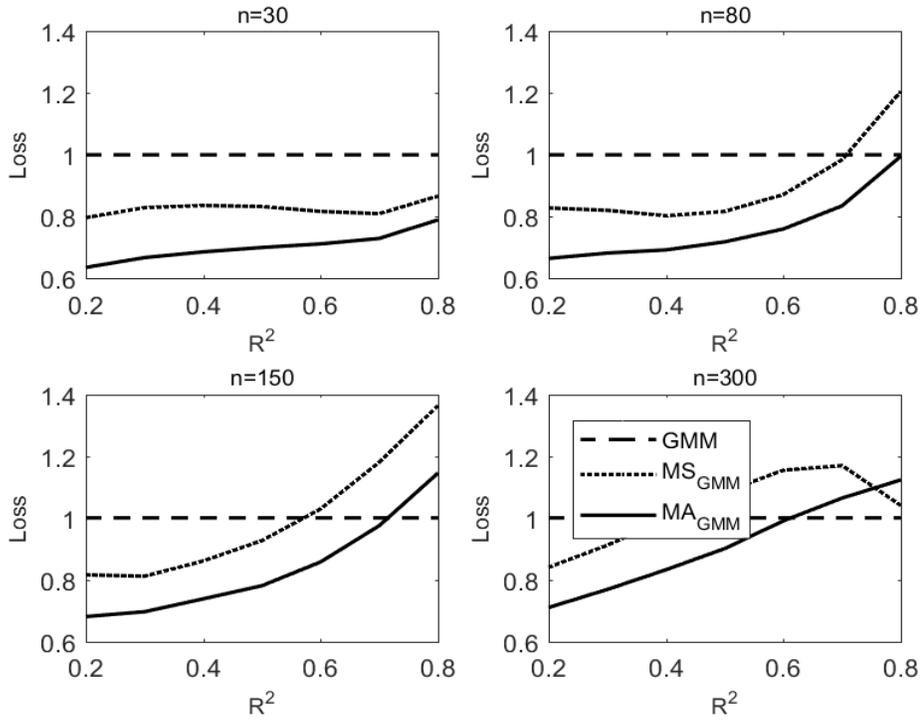
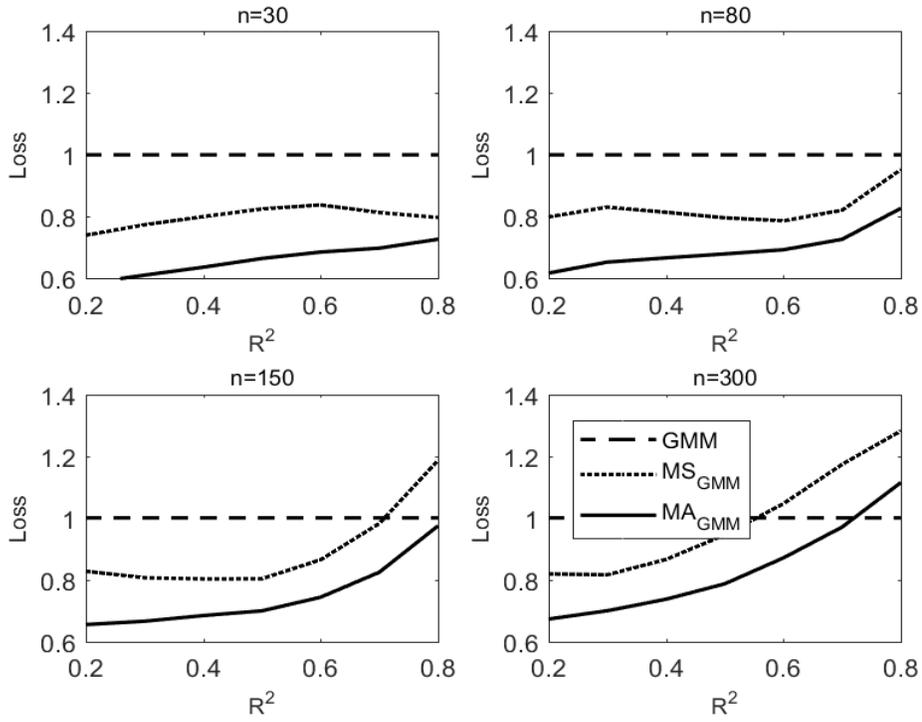


Figure S.3: Loss in simulation Design II, with $\tilde{R}^2 = 0.5$.

Figure S.4: Loss in simulation Design II, with $\tilde{R}^2 = 0.8$.

References

- GAO, Y., ZHANG, X., WANG, S., CHONG, T. T.-L. & ZOU, G. (2019). Frequentist model averaging for threshold models. *Annals of the Institute of Statistical Mathematics* **71**, 275–306.
- LI, K.-C. (1987). Asymptotic optimality for C_p, C_l , cross-validation and generalized cross-validation: Discrete index set. *The Annals of Statistics* **15**, 958–975.
- STEIN, C. M. (1981). Estimation of the mean of a multivariate normal distribution. *The Annals of Statistics* **153**, 1135–1151.
- ZHANG, X. (2010). Model averaging and its applications. *Ph.D. Thesis*, Academy of Mathematics and Systems Science, Chinese Academy of Sciences.