

GENERALIZED S-ESTIMATORS FOR LINEAR MIXED EFFECTS MODELS

Inna Chervoneva and Mark Vishnyakov

Thomas Jefferson University

Supplementary Material

This supplementary material provides proofs of Theorems 1-3, derivation of the fixed point equations for $\widehat{\boldsymbol{\theta}}_M$, and Supplementary Figures 1-6.

1 Proof of Theorem 1.

Using Lagrange multipliers to solve the constrained minimization problem, one minimizes

$$L = D - \lambda g(\mathbf{y}, \boldsymbol{\xi}) = \sum_{i=1}^M v_i \ln [\det(\mathbf{V}_i)] - \lambda \left[\sum_{i=1}^M \rho(d_i) - b_M \right].$$

By taking derivatives one obtains the system of $q + l + 1$ estimating equations,

$$\begin{aligned} \frac{\partial L}{\partial \boldsymbol{\beta}} &= \frac{\partial g}{\partial \boldsymbol{\beta}} = \mathbf{0} \\ \frac{\partial L}{\partial \boldsymbol{\theta}} &= \frac{\partial D}{\partial \boldsymbol{\theta}} - \lambda \frac{\partial g}{\partial \boldsymbol{\theta}} = \mathbf{0} \\ \frac{\partial L}{\partial \lambda} &= g(\mathbf{y}, \boldsymbol{\xi}) = \mathbf{0} \end{aligned} \tag{S1}$$

where

$$\frac{\partial g}{\partial \boldsymbol{\beta}} = \frac{1}{2} \sum_{i=1}^M u(d_i) \mathbf{A}'_i \mathbf{V}_i^{-1} (\mathbf{y}_i - \mathbf{A}_i \boldsymbol{\beta}) \tag{S2}$$

$$\frac{\partial D}{\partial \boldsymbol{\theta}} = \sum_{j=1}^l e_j^l \otimes \sum_{i=1}^M v_i \text{tr} \left[\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right] \tag{S3}$$

$$\frac{\partial g}{\partial \boldsymbol{\theta}} = \sum_{j=1}^l e_j^l \otimes \sum_{i=1}^M \left[-\frac{1}{2} \sum_{i=1}^M u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right], \tag{S4}$$

e_j^l is j^{th} canonical $l \times 1$ vector and $u_{m_i}(d) = \frac{1}{d} \frac{\partial}{\partial d} \rho_{m_i}(d)$ is the weighting function corresponding to the ρ_{m_i} -function. System (S1) implies that λ must satisfy the equation $\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \boldsymbol{\theta}} - \lambda \mathbf{1}'_{l \times 1} \frac{\partial g}{\partial \boldsymbol{\theta}} = 0$, where $\mathbf{1}_{l \times 1}$ is $l \times 1$ vector of ones, and, therefore,

$$\lambda = \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \boldsymbol{\theta}} \right) \left(\mathbf{1}'_{l \times 1} \frac{\partial g}{\partial \boldsymbol{\theta}} \right)^{-1}.$$

Then, for any j , equation $\frac{\partial \mathbf{L}}{\partial \theta_j} = 0$ in (S1) may be written as

$$\left(\mathbf{1}'_{l \times 1} \frac{\partial g}{\partial \boldsymbol{\theta}} \right) \frac{\partial D}{\partial \theta_j} - \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \boldsymbol{\theta}} \right) \frac{\partial g}{\partial \theta_j} = 0. \quad (\text{S5})$$

Denoting $\mathbf{r}_i = \mathbf{V}_i^{-\frac{1}{2}} (\mathbf{y}_i - \mathbf{A}_i \boldsymbol{\beta})$ and using (S4), we can write (S5) as

$$\sum_{i=1}^M u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i = 0, \text{ where } \mathbf{F}_{ji} = \left(\sum_{k=1}^l \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \frac{\partial D}{\partial \theta_j} - \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \boldsymbol{\theta}} \right) \frac{\partial \mathbf{V}_i}{\partial \theta_j}, \quad (\text{S6})$$

and \mathbf{F}_{ji} is the same as in Theorem 1. Hence, system (S1) may be re-written as system (2.10)-(2.12) of $q + l + 1$ equations in $q + l$ unknowns. System (2.10)-(2.12) is conditionally unbiased under the true model parameters $[\boldsymbol{\beta}'_0, \boldsymbol{\theta}'_0]'$ and normal distribution. This is trivial for (2.12). For (2.10), this easily follows from each summation term being an odd function of \mathbf{r}_i and each \mathbf{r}_i having a symmetric distribution around zero under the true model (e.g. Lemma 1 in Kackar and Harville, 1981). Finally, equations (2.11) are conditionally unbiased if $v_i = m_i^{-1} E_{d^2 \sim \chi^2_{m_i}} [d^2 u(d)]$. It is enough to show that under the true normal model, for any $j = 1, \dots, l$,

$$\begin{aligned} & \frac{\partial D}{\partial \theta_j} \sum_{i=1}^M E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\sum_{k=1}^l \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] \\ &= \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \boldsymbol{\theta}} \right) \sum_{i=1}^M E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] \end{aligned} \quad (\text{S7})$$

To shorten the notation, let us denote $\mathbf{x}_i = [u(d_i)]^{1/2} \mathbf{r}_i$ and $\mathbf{W}_i = \sum_{k=1}^l \frac{\partial \mathbf{V}_i}{\partial \theta_k}$.

$$\text{Then, } \sum_{i=1}^M E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\sum_{k=1}^l \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] =$$

$$\sum_{i=1}^M E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} \left[\mathbf{x}_i' \mathbf{V}_i^{-1/2} \mathbf{W}_i \mathbf{V}_i^{-1/2} \mathbf{x}_i \right] = \sum_{i=1}^M \text{tr} \left[\mathbf{V}_i^{-1/2} \mathbf{W}_i \mathbf{V}_i^{-1/2} \mathbf{Cov}(\mathbf{x}_i) \right].$$

For symmetric $\rho(d)$ with symmetric $u(d)$, under the true model, \mathbf{x}_i has a symmetric distribution and $E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})}[\mathbf{x}_i] = \mathbf{0}$. Then $\mathbf{Cov}(\mathbf{x}_i) = E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} [u(d_i) \mathbf{r}_i \mathbf{r}_i']$ is a diagonal matrix with equal diagonal elements $E[z_1^2 u(z_1^2 + \dots + z_{m_i}^2)]$, where z_1, \dots, z_{m_i} are i.i.d. $N(0, 1)$, because off diagonal elements of matrix $u(d_i) \mathbf{r}_i \mathbf{r}_i'$ are asymmetric functions in each component of vector \mathbf{r}_i . Hence,

$$\begin{aligned} \sum_{i=1}^M E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} \left[u(d_i) \mathbf{r}_i' \mathbf{V}_i^{-1/2} \mathbf{W}_i \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] &= \sum_{i=1}^M m_i^{-1} E_{d^2 \sim \chi_{m_i}^2} [d^2 u(d)] \text{tr} [\mathbf{V}_i^{-1} \mathbf{W}_i], \\ \sum_{i=1}^M E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} \left[u(d_i) \mathbf{r}_i' \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] &= \sum_{i=1}^M m_i^{-1} E_{d^2 \sim \chi_{m_i}^2} [d^2 u(d)] \text{tr} \left[\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right], \end{aligned}$$

and $v_i = E_{d^2 \sim \chi_{m_i}^2} [d^2 u(d)] m_i^{-1}$ implies that both sides of (S7) are equal to

$$\left(\sum_{i=1}^M v_i \text{tr} [\mathbf{V}_i^{-1} \mathbf{W}_i] \right) \sum_{i=1}^M v_i \text{tr} \left[\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right]. \quad \square$$

2 Proof of Theorem 2.

The GS-estimator may be viewed as a GEE estimator with estimating equations (3.1)-(3.2). The following regularity conditions (Yuan and Jennrich, 1998) imply existence of a consistent sequence of GS-estimators:

C1 With probability 1, $\Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi}_0) \rightarrow \mathbf{0}$, where $\boldsymbol{\xi}_0$ is the true value of $\boldsymbol{\xi}$.

C2 There is a neighborhood of $\boldsymbol{\xi}_0$, $N(\boldsymbol{\xi}_0)$, such that with probability 1, all $\Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi})$ are continuously differentiable and Jacobians $\frac{\partial}{\partial \boldsymbol{\xi}} \Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi})$ converge uniformly to a nonstochastic limit, which is nonsingular at $\boldsymbol{\xi}_0$.

C1 follows from Lemma 5.3 in Shao (2003) under conditions

C3 For each $h_i(\mathbf{y}_i) = \sup_{\boldsymbol{\xi} \in \Theta} \|\Psi_i(\mathbf{y}_i, \boldsymbol{\xi})\|$, $i = 1, \dots, M$, there exists $\delta > 0$ such that $\sup_i E |h_i(\mathbf{y}_i)|^{1+\delta} < \infty$ and $\sup_i E \|\mathbf{y}_i\|^\delta < \infty$.

C4 for any $c > 0$ and any sequence $\{\mathbf{x}_i\} \in \mathcal{R}^{m_i}$ such that $\|\mathbf{x}_i\| \leq c$, the sequence of functions $\{g_i(\boldsymbol{\xi}) = \boldsymbol{\Psi}_i(\mathbf{x}_i, \boldsymbol{\xi})\}$ is equicontinuous on any open subset of Θ .

Under assumptions **R1-R4**, conditions **C3** and **C4** are satisfied. To show that **C2** holds, we compute the Jacobian $\frac{\partial}{\partial \boldsymbol{\xi}} \boldsymbol{\Psi}_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi})$, which is partitioned as

$$\frac{\partial}{\partial \boldsymbol{\xi}} \boldsymbol{\Psi}_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi}) = \begin{bmatrix} \frac{\partial}{\partial \boldsymbol{\beta}'} \boldsymbol{\Psi}_{\mathbf{M}}^{\beta}(\mathbf{y}, \boldsymbol{\xi}) & \frac{\partial}{\partial \boldsymbol{\theta}'} \boldsymbol{\Psi}_{\mathbf{M}}^{\beta}(\mathbf{y}, \boldsymbol{\xi}) \\ \frac{\partial}{\partial \boldsymbol{\beta}'} \boldsymbol{\Psi}_{\mathbf{M}}^{\theta}(\mathbf{y}, \boldsymbol{\xi}) & \frac{\partial}{\partial \boldsymbol{\theta}'} \boldsymbol{\Psi}_{\mathbf{M}}^{\theta}(\mathbf{y}, \boldsymbol{\xi}) \end{bmatrix}$$

where $\boldsymbol{\Psi}_{\mathbf{M}}^{\beta}(\mathbf{y}, \boldsymbol{\xi}) = \frac{1}{M} \sum_{i=1}^M \boldsymbol{\Psi}_{\beta i}(\mathbf{y}_i, \boldsymbol{\xi})$, $\boldsymbol{\Psi}_{\mathbf{M}}^{\theta}(\mathbf{y}, \boldsymbol{\xi}) = \frac{1}{M} \sum_{i=1}^M \boldsymbol{\Psi}_{\theta i}(\mathbf{y}_i, \boldsymbol{\xi})$ and

$$\frac{\partial}{\partial \boldsymbol{\beta}'} \boldsymbol{\Psi}_{i\beta}(\mathbf{y}_i, \boldsymbol{\xi}) = -u(d_i) \mathbf{A}'_i \mathbf{V}_i^{-1} \mathbf{A}_i - u'(d_i) d_i^{-1} \mathbf{A}'_i \mathbf{V}_i^{-1/2} \mathbf{r}_i \mathbf{r}'_i \mathbf{V}_i^{-1/2} \mathbf{A}_i$$

$$\begin{aligned} \frac{\partial}{\partial \theta_j} \boldsymbol{\Psi}_{i\beta}(\mathbf{y}_i, \boldsymbol{\xi}) &= u(d_i) \mathbf{A}'_i \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \\ &\quad + u'(d_i) d_i^{-1} \left(\mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right) \mathbf{A}'_i \mathbf{V}_i^{-1/2} \mathbf{r}_i \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial \boldsymbol{\beta}'} \boldsymbol{\Psi}_{i\theta j}(\mathbf{y}_i, \boldsymbol{\xi}) &= -2u'(d_i) d_i^{-1} \mathbf{A}'_i \mathbf{V}_i^{-1/2} \mathbf{r}_i \left[\mathbf{r}'_i \mathbf{V}_i^{-1/2} \mathbf{F}_{ji} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] \\ &\quad - 2u(d_i) \mathbf{A}'_i \mathbf{V}_i^{-1} \mathbf{F}_{ji} \mathbf{V}_i^{-1/2} \mathbf{r}_i - \frac{c_j}{2} u(d_i) \mathbf{A}'_i \mathbf{V}_i^{-1/2} \mathbf{r}_i \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial \boldsymbol{\theta}'} \boldsymbol{\Psi}_{\theta}(\mathbf{y}, \boldsymbol{\xi}) &= \left(\mathbf{1}'_{l \times 1} \frac{\partial g}{\partial \boldsymbol{\theta}} \right) \frac{\partial^2 D}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} + \frac{\partial D}{\partial \boldsymbol{\theta}} \left(\mathbf{1}'_{l \times 1} \frac{\partial^2 g}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \right) \\ &\quad - \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \boldsymbol{\theta}} \right) \frac{\partial^2 g}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} - \frac{\partial g}{\partial \boldsymbol{\theta}} \left(\mathbf{1}'_{l \times 1} \frac{\partial^2 D}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} \right) + \mathbf{c} \frac{\partial g}{\partial \boldsymbol{\theta}'}, \end{aligned}$$

$$\frac{\partial^2 D}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} = \sum_{j=1}^l \sum_{k=1}^l e_j^l \otimes (e_k^l)' \otimes \sum_{i=1}^M v_i \left\{ \text{tr} \left(\mathbf{V}_i^{-1} \frac{\partial^2 \mathbf{V}_i}{\partial \theta_j \partial \theta_k} - \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right) \right\},$$

$$\begin{aligned} \frac{\partial^2 g}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'} &= \sum_{j=1}^l \sum_{k=1}^l e_j^l \otimes (e_k^l)' \otimes \\ &\quad \sum_{i=1}^M \left\{ -u'(d_i) d_i^{-1} \left(\mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right. \\ &\quad \left. - u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} - \frac{\partial^2 \mathbf{V}_i}{\partial \theta_j \partial \theta_k} + \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \mathbf{V}_i^{-1/2} \mathbf{r}'_i \right\} \end{aligned}$$

Thus, under assumption **R5**, $\Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi})$ is continuously differentiable. Assumptions **R1-R6** and above arguments imply that conditions 6-8 in Yuan and Jennrich (1998) are satisfied, and the main Theorem in Yuan (1997) implies that convergence of $\frac{\partial}{\partial \boldsymbol{\xi}} \Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi})$ to $\Phi(\boldsymbol{\xi})$ is uniform. Hence condition **C2** is satisfied as well. Then existence of a consistent sequence of GS-estimators follows from the main result of Yuan and Jennrich (1998). \square

3 Proof of Theorem 3.

Under assumptions **R1-R7**, the multivariate Lindenber-Feller Central Limit Theorem (e.g. Mukhopadhyay, 2009) gives that $M^{1/2} \Psi_M(\mathbf{y}, \boldsymbol{\xi}_0) \rightarrow \mathbf{N}(0, \boldsymbol{\Omega}(\boldsymbol{\xi}_0))$ in distribution if the Lindenber's condition is satisfied. The proof of Theorem 5.14 in Shao (2003) implies that the Lindenber's condition follows if $\sup_i E \|\Psi_i(\mathbf{y}_i, \boldsymbol{\xi}_0)\|^{2+\delta} < \infty$ for some $\delta > 0$. The latter condition is satisfied under the assumptions **R1-R4**. Thus, assumptions 1-3 in Yuan and Jennrich (1998) are satisfied, and Theorem 4 in Yuan and Jennrich (1998) implies that $M^{-1/2} (\hat{\boldsymbol{\xi}}_M \rightarrow \boldsymbol{\xi}_0) \rightarrow \mathbf{N}(0, \boldsymbol{\Sigma})$ where $\boldsymbol{\Sigma} = [\Phi(\boldsymbol{\xi}_0)]^{-1} \boldsymbol{\Omega}(\boldsymbol{\xi}_0) [\Phi(\boldsymbol{\xi}_0)']^{-1}$,

$$\boldsymbol{\Omega}(\boldsymbol{\xi}_0) = \lim_{M \rightarrow \infty} \{\boldsymbol{\Omega}_M(\boldsymbol{\xi}_0)\}, \quad \boldsymbol{\Omega}_M(\boldsymbol{\xi}_0) = \frac{1}{M} \sum_{i=1}^M E[\Psi_i(\mathbf{y}_i, \boldsymbol{\xi}_0) \Psi_i(\mathbf{y}_i, \boldsymbol{\xi}_0)'],$$

$$\Phi(\boldsymbol{\xi}_0) = \lim_{M \rightarrow \infty} \{\Phi_M(\boldsymbol{\xi}_0)\}, \quad \Phi_M(\boldsymbol{\xi}_0) = E \left[\frac{\partial}{\partial \boldsymbol{\xi}} \Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi}) \Big|_{\boldsymbol{\xi}=\boldsymbol{\xi}_0} \right].$$

First, we compute $\Phi_M(\boldsymbol{\xi}_0) = E \left[\frac{\partial}{\partial \boldsymbol{\xi}} \Psi_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi}) \Big|_{\boldsymbol{\xi}=\boldsymbol{\xi}_0} \right]$. Under the true model parameters $\boldsymbol{\xi}_0 = [\boldsymbol{\beta}'_0, \boldsymbol{\theta}'_0]'$, vectors $\mathbf{r}_i = \mathbf{V}_i^{-1/2} (\mathbf{y}_i - \mathbf{A}_i \boldsymbol{\beta})$ have multivariate standard normal distribution, and Lemma 5.1 in Lopuhaa (1989) implies that $d_i = (\mathbf{r}'_i \mathbf{r}_i)^{1/2}$ and $d_i^{-1} \mathbf{r}_i$ are independent, $E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} [d_i^{-1} \mathbf{r}_i] = 0$, and $Cov(d_i^{-1} \mathbf{r}_i) = E_{\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})} [d_i^{-2} \mathbf{r}_i \mathbf{r}'_i] = m_i^{-1} \mathbf{I}$. Hence,

$$E \left[\frac{\partial}{\partial \boldsymbol{\beta}'} \Psi_{i\beta}(\mathbf{y}_i, \boldsymbol{\xi}) \right] = -E[u(d_i)] \mathbf{A}'_i \mathbf{V}_i^{-1} \mathbf{A}_i - E[u'(d_i) d_i] \mathbf{A}'_i \mathbf{V}_i^{-1/2} (m_i^{-1} \mathbf{I}) \mathbf{V}_i^{-1/2} \mathbf{A}_i,$$

$$\text{so that } E \left[\frac{\partial}{\partial \boldsymbol{\beta}'} \Psi_{\mathbf{M}}^{\boldsymbol{\beta}}(\mathbf{y}, \boldsymbol{\xi}) \right] = -\frac{1}{M} \sum_{i=1}^M E[u(d_i) + m_i^{-1} u'(d_i) d_i] \mathbf{A}'_i \mathbf{V}_i^{-1} \mathbf{A}_i.$$

Further, $E \left[\frac{\partial}{\partial \theta_j} \Psi_{i\beta}(\mathbf{y}_i, \boldsymbol{\xi}) \right] = \mathbf{0}$ since $\frac{\partial}{\partial \theta_j} \Psi_{i\beta}(\mathbf{y}_i, \boldsymbol{\xi})$ is asymmetric function

in each component of $\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})$. Similarly, $E \left[\frac{\partial}{\partial \beta} \Psi_{i\theta_j}(\mathbf{y}_i, \boldsymbol{\xi}) \right] = \mathbf{0}$ since $\frac{\partial}{\partial \beta} \Psi_{i\theta_j}(\mathbf{r}_i, \mathbf{A}_i, \boldsymbol{\xi})$ is asymmetric function in each component of $\mathbf{r}_i \sim \mathbf{N}_{m_i}(\mathbf{0}, \mathbf{I})$.

$$\begin{aligned} \text{Finally, } E \left[\frac{\partial}{\partial \theta'} M \Psi_{\theta}(\mathbf{y}, \boldsymbol{\xi}) \right] &= \left(\mathbf{1}'_{l \times 1} E \left[\frac{\partial g}{\partial \theta} \right] \right) \frac{\partial^2 D}{\partial \theta \partial \theta'} - \frac{\partial D}{\partial \theta} \left(\mathbf{1}'_{l \times 1} E \left[\frac{\partial^2 g}{\partial \theta \partial \theta'} \right] \right) \\ &- \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \theta} \right) E \left[\frac{\partial^2 g}{\partial \theta \partial \theta'} \right] - E \left[\frac{\partial g}{\partial \theta} \right] \left(\mathbf{1}'_{l \times 1} \frac{\partial^2 D}{\partial \theta \partial \theta'} \right) + \mathbf{c} E \left[\frac{\partial g}{\partial \theta'} \right], \end{aligned}$$

where $E \left[\frac{\partial g}{\partial \theta} \right] = \frac{\partial D}{\partial \theta}$ because $v_i = E_{d^2 \sim \chi_{m_i}^2} [d^2 u(d)] m_i^{-1}$. Hence,

$$E \left[\frac{\partial}{\partial \theta'} M \Psi_{\theta}(\mathbf{y}, \boldsymbol{\xi}) \right] = - \left(\mathbf{1}'_{l \times 1} \frac{\partial D}{\partial \theta} \mathbf{I} - \frac{\partial D}{\partial \theta} \mathbf{1}'_{l \times 1} \right) \left(E \left[\frac{\partial^2 g}{\partial \theta \partial \theta'} \right] - \frac{\partial^2 D}{\partial \theta \partial \theta'} \right) + \mathbf{c} \frac{\partial D}{\partial \theta'},$$

where $E \left[\frac{\partial^2 g}{\partial \theta \partial \theta'} \right] =$

$$\begin{aligned} E \left[\sum_{j=1}^l \sum_{k=1}^l e_j^l \otimes (e_k^l)' \otimes \sum_{i=1}^M \left\{ -u'(d_i) d_i^{-1} \left(\mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right. \right. \\ \left. \left. - u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} - \frac{\partial^2 \mathbf{V}_i}{\partial \theta_j \partial \theta_k} + \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \mathbf{V}_i^{-1/2} \mathbf{r}'_i \right\} \right]. \end{aligned}$$

Again using Lemma 5.1 in Lopuhaa (1989),

$$\begin{aligned} E \left[u'(d_i) d_i^{-1} \left(\mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] = \\ \frac{E[u'(d_i) d_i^3]}{m_i(m_i+2)} \text{tr} \left(\left[\left(\mathbf{V}_i^{-\frac{1}{2}} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-\frac{1}{2}} \right)' \otimes \left(\mathbf{V}_i^{-\frac{1}{2}} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-\frac{1}{2}} \right) \right] \mathbf{H}_{m_i} \right), \end{aligned}$$

where $\mathbf{H}_{m_i} = [\mathbf{I} + \mathbf{K}_{m_i, m_i} + \text{vec}(\mathbf{I}) \text{vec}(\mathbf{I})']$, \mathbf{K}_{m_i, m_i} is a $(m_i^2 \times m_i^2)$ -block matrix with (i, j) -th block $K(i, j)$ of size $(m_i \times m_i)$, which has 1 at entry (i, j) and 0 otherwise ($K(i, j)_{ij} = \delta_{ij}$). Also, with the arguments similar to the above,

$$\begin{aligned} E \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} - \frac{\partial^2 \mathbf{V}_i}{\partial \theta_j \partial \theta_k} + \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \mathbf{V}_i^{-1/2} \mathbf{r}'_i \right] = \\ m_i^{-1} E \left[u(d_i) d_i^2 \right] \text{tr} \left[\mathbf{V}_i^{-1} \left(\frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} - \frac{\partial^2 \mathbf{V}_i}{\partial \theta_j \partial \theta_k} + \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \right], \text{ so that,} \end{aligned}$$

$$\begin{aligned} E \left[\frac{\partial^2 g}{\partial \theta \partial \theta'} \right] &= \sum_{j=1}^l \sum_{k=1}^l e_j^l \otimes (e_k^l)' \otimes \sum_{i=1}^M \left\{ -\frac{E[u'(d_i) d_i^3]}{m_i(m_i+2)} \times \right. \\ &\quad \left. \text{tr} \left(\left[\left(\mathbf{V}_i^{-\frac{1}{2}} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-\frac{1}{2}} \right)' \otimes \left(\mathbf{V}_i^{-\frac{1}{2}} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-\frac{1}{2}} \right) \right] \mathbf{H}_{m_i} \right) \right\} \\ &\quad + v_i \text{tr} \left[\mathbf{V}_i^{-1} \left(-2 \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} + \frac{\partial^2 \mathbf{V}_i}{\partial \theta_j \partial \theta_k} \right) \right] \end{aligned}$$

Hence, $E \left[\frac{\partial}{\partial \boldsymbol{\xi}} \boldsymbol{\Psi}_{\mathbf{M}}(\mathbf{y}, \boldsymbol{\xi}_0) \Big|_{\boldsymbol{\xi}=\boldsymbol{\xi}_0} \right] = \text{Diag}(\Phi_{\beta\beta}, \Phi_{\theta\theta})$, where

$$\begin{aligned} \Phi_{\beta\beta} &= -\frac{1}{M} \sum_{i=1}^M \mathbf{A}'_i \mathbf{V}_i^{-1} \mathbf{A}_i E [u(d_i) + m_i^{-1} u'(d_i) d_i] \\ \Phi_{\theta\theta} &= -\frac{1}{M} \left(\mathbf{1}'_{l \times 1} \left\{ \sum_{j=1}^l e_j^l \otimes \sum_{i=1}^M v_i \text{tr} \left[\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right] \right\} \mathbf{I} - \frac{\partial D}{\partial \boldsymbol{\theta}} \mathbf{1}'_{l \times 1} \right) \times \\ &\sum_{j=1}^l \sum_{k=1}^l e_j^l \otimes (e_k^l)' \otimes \sum_{i=1}^M \left\{ \frac{E [u'(d_i) d_i^3]}{m_i (m_i + 2)} \text{tr} \left(\left[\left(\mathbf{V}_i^{-\frac{1}{2}} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \mathbf{V}_i^{-\frac{1}{2}} \right)' \otimes \left(\mathbf{V}_i^{-\frac{1}{2}} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-\frac{1}{2}} \right) \right] \mathbf{H}_{m_i} \right) \right. \\ &\left. + v_i \text{tr} \left(\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_k} \right) \right\} + \frac{1}{M} \mathbf{c} \sum_{j=1}^l (e_j^l)' \otimes \sum_{i=1}^M v_i \text{tr} \left[\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right]. \end{aligned}$$

Further, $\boldsymbol{\Omega}_M(\boldsymbol{\xi}_0) = \frac{1}{M} \sum_{i=1}^M E [\boldsymbol{\Psi}_i(\mathbf{y}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_i(\mathbf{y}_i, \boldsymbol{\xi}_0)']$ may be written as

$$\boldsymbol{\Omega}_M(\boldsymbol{\xi}_0) = \begin{bmatrix} \frac{1}{M} \sum_{i=1}^M E [\boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] & \frac{1}{M} \sum_{i=1}^M E [\boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] \\ \frac{1}{M} \sum_{i=1}^M E [\boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] & \frac{1}{M} \sum_{i=1}^M E [\boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] \end{bmatrix},$$

where $\sum_{i=1}^M E [\boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] = \sum_{i=1}^M m_i^{-1} E [u^2(d_i) d_i^2] \mathbf{A}'_i \mathbf{V}_i^{-1} \mathbf{A}_i$ and

$\sum_{i=1}^M E [\boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] = \mathbf{0}$ because $\boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \mathbf{A}_i, \boldsymbol{\xi})$ is asymmetric

and $\boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \mathbf{A}_i, \boldsymbol{\xi})'$ is symmetric function in each component of \mathbf{r}_i .

Similarly, $\sum_{i=1}^M E [\boldsymbol{\Psi}_{i\beta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0)'] = \mathbf{0}$.

Finally, $\boldsymbol{\Omega}_M^\theta(\boldsymbol{\xi}_0) = \sum_{i=1}^M E [\boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0) \boldsymbol{\Psi}_{i\theta}(\mathbf{r}_i, \boldsymbol{\xi}_0)']$ has elements

$$\begin{aligned} [\boldsymbol{\Omega}_M^\theta(\boldsymbol{\xi}_0)]_{jk} &= E \left[\sum_{i=1}^M u^2(d_i) \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ki} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \right] \\ &+ E \left[c_j \sum_{i=1}^M u(d_i) [\rho(d_i) - b_{m_i}] \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \right] + \\ &+ E \left[c_k \sum_{i=1}^M u(d_i) [\rho(d_i) - b_{m_i}] \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ki} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \right] + \\ &+ E \left[c_j c_k \sum_{i=1}^M [\rho(d_i) - b_{m_i}]^2 \right], \text{ where } \mathbf{F}_{ji} = \mathbf{F}_{ji}(\boldsymbol{\theta}) \text{ is given in (2.13)}. \end{aligned}$$

Using results from Neudecker (1969) and Lemma 5.1 in Lopuhaa (1989)),

$$\begin{aligned}
& E \left[\sum_{i=1}^{\mathbf{M}} u^2(d_i) \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ki} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \right] = \\
& \sum_{i=1}^{\mathbf{M}} E \left[u^2(d_i) d_i^4 \right] E \left[\text{vec} \left(\frac{\mathbf{r}_i \mathbf{r}'_i}{d_i^2} \right)' \left\{ \left(\mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ki} \mathbf{V}_i^{-\frac{1}{2}} \right)' \otimes \left(\mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \right) \right\} \text{vec} \left(\frac{\mathbf{r}_i \mathbf{r}'_i}{d_i^2} \right) \right] = \\
& \sum_{i=1}^{\mathbf{M}} \frac{E[u^2(d_i) d_i^4]}{m_i(m_i+2)} \text{tr} \left\{ \left[\left(\mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ki} \mathbf{V}_i^{-\frac{1}{2}} \right)' \otimes \left(\mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \right) \right] [\mathbf{I} + \mathbf{K}_{m_i} + \text{vec}(\mathbf{I}) \text{vec}(\mathbf{I})'] \right\},
\end{aligned}$$

$$\text{and } E \left[\sum_{i=1}^{\mathbf{M}} u(d_i) [\rho(d_i) - b_{m_i}] \left(\mathbf{r}'_i \mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \mathbf{r}_i \right) \right] =$$

$$\sum_{i=1}^{\mathbf{M}} m_i^{-1} E \left[u(d_i) d_i^2 (\rho(d_i) - b_{m_i}) \right] \text{tr} \left[\mathbf{V}_i^{-1} \mathbf{F}_{ji} \right], \text{ so that}$$

$$\begin{aligned}
[\boldsymbol{\Omega}_M(\boldsymbol{\xi}_0)]_{jk} &= c_j c_k \sum_{i=1}^{\mathbf{M}} E \left[(\rho(d_i) - b_{m_i})^2 \right] \\
&+ \sum_{i=1}^{\mathbf{M}} m_i^{-1} E \left[u(d_i) d_i^2 (\rho(d_i) - b_{m_i}) \right] \text{tr} \left[\mathbf{V}_i^{-1} (c_j \mathbf{F}_{ji} + c_k \mathbf{F}_{ki}) \right] \\
&+ \sum_{i=1}^{\mathbf{M}} \frac{E[u^2(d_i) d_i^4]}{m_i(m_i+2)} \text{tr} \left\{ \left[\left(\mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ki} \mathbf{V}_i^{-\frac{1}{2}} \right)' \otimes \left(\mathbf{V}_i^{-\frac{1}{2}} \mathbf{F}_{ji} \mathbf{V}_i^{-\frac{1}{2}} \right) \right] \mathbf{H}_{m_i} \right\}. \square
\end{aligned}$$

4 Derivation of fixed point equations for $\widehat{\boldsymbol{\theta}}_M$.

Linear covariance structure of \mathbf{V}_i ($\mathbf{V}_i = \sum_{g=1}^l \theta_g \mathbf{L}_{ig}$, $i = 1, \dots, M$) implies that $\forall i$,

$$\frac{\partial \mathbf{V}_i}{\partial \theta_j} = \mathbf{L}_{ij} \text{ and } \sum_{j=1}^l \theta_j \frac{\partial \mathbf{V}_i}{\partial \theta_j} = \mathbf{V}_i \text{ so that}$$

$$\text{tr} \left(\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right) = \text{tr} \left(\mathbf{V}_i^{-1} \mathbf{L}_{ij} \mathbf{V}_i^{-1} \mathbf{V}_i \right) = \sum_{s=1}^l \theta_s \text{tr} \left[\mathbf{V}_i^{-1} \mathbf{L}_{ij} \mathbf{V}_i^{-1} \mathbf{L}_{is} \right]$$

Hence, $\text{tr} \left(\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \boldsymbol{\theta}} \right) = \mathbf{Q}_i(\boldsymbol{\theta}) \boldsymbol{\theta}$, where $\mathbf{Q}_i(\boldsymbol{\theta})$ the $l \times l$ matrix with entries $[\mathbf{Q}_i(\boldsymbol{\theta})]_{jk} = \text{tr} \left[\mathbf{V}_i^{-1} \mathbf{L}_{ij} \mathbf{V}_i^{-1} \mathbf{L}_{ik} \right]$. Respectively, $\text{tr} \left(\mathbf{V}_i^{-1} \frac{\partial \mathbf{V}_i}{\partial \theta_j} \right) = [\mathbf{Q}_i(\boldsymbol{\theta})]_j \cdot \theta_j$, where $[\mathbf{Q}_i(\boldsymbol{\theta})]_j$ stands for j^{th} row of matrix $\mathbf{Q}_i(\boldsymbol{\theta})$.

Under the assumption of linear covariance structure, $\mathbf{V}_i = \sum_{g=1}^l \theta_g \mathbf{L}_{ig}$, $i = 1, \dots, M$, each matrix \mathbf{F}_{ji} in (2.13) simplifies to

$$\mathbf{F}_{ji} = \left(\sum_{s=1}^l \mathbf{L}_{is} \right) \sum_{k=1}^M v_k \text{tr} \left(\mathbf{V}_k^{-1} \frac{\partial \mathbf{V}_k}{\partial \theta_j} \right) - \mathbf{L}_{ij} \sum_{k=1}^M v_k \text{tr} \left(\mathbf{V}_k^{-1} \sum_{s=1}^l \mathbf{L}_{is} \right), \text{ where}$$

$\sum_{k=1}^M v_k tr \left(\mathbf{V}_k^{-1} \sum_{s=1}^l \mathbf{L}_{is} \right)$ is a constant further denoted by S_{Ml} ,

$$S_{Ml} = \sum_{k=1}^M v_k tr \left(\mathbf{V}_k^{-1} \sum_{s=1}^l \frac{\partial \mathbf{V}_k}{\partial \theta_s} \right) = \sum_{k=1}^M v_k tr \left(\mathbf{V}_k^{-1} \sum_{s=1}^l \mathbf{L}_{is} \right).$$

Hence, \mathbf{F}_{ji} may be expressed as

$$\mathbf{F}_{ji} = \left(\sum_{s=1}^l \mathbf{L}_{is} \right) \sum_{k=1}^M v_k [\mathbf{Q}_k(\boldsymbol{\theta})]_j \cdot \boldsymbol{\theta} - S_{Ml} \mathbf{L}_{ij}.$$

Let us further denote by $\mathbf{Q}(\boldsymbol{\theta})$ the $l \times l$ matrix with entries

$$[\mathbf{Q}(\boldsymbol{\theta})]_{jk} = \sum_{i=1}^M v_i tr [\mathbf{V}_i^{-1} \mathbf{L}_{ij} \mathbf{V}_i^{-1} \mathbf{L}_{ik}]$$

Then $\mathbf{F}_{ji} = \left(\sum_{s=1}^l \mathbf{L}_{is} \right) [\mathbf{Q}(\boldsymbol{\theta})]_j \cdot \boldsymbol{\theta} - S_{Ml} \mathbf{L}_{ij}$, where $[\mathbf{Q}(\boldsymbol{\theta})]_j$ stands for j^{th} row of matrix $\mathbf{Q}(\boldsymbol{\theta})$, and system (3.2) is written in the matrix form as

$$\begin{aligned} & \frac{1}{M} \sum_{j=1}^l e_j^l \otimes \sum_{i=1}^M \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\left(\sum_{s=1}^l \mathbf{L}_{is} \right) [\mathbf{Q}(\boldsymbol{\theta})]_j \cdot \boldsymbol{\theta} - S_{Ml} \mathbf{L}_{ij} \right) \mathbf{V}_i^{-1/2} \mathbf{r}_i + c_j (\rho(d_i) - b_{m_i}) \right] = \\ & \frac{1}{M} \sum_{j=1}^l e_j^l \otimes \sum_{i=1}^M \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\left(\sum_{s=1}^l \mathbf{L}_{is} \right) [\mathbf{Q}(\boldsymbol{\theta})]_j \cdot \boldsymbol{\theta} \right) \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] + \\ & \frac{S_{Ml}}{M} \sum_{j=1}^l e_j^l \otimes \sum_{i=1}^M \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \mathbf{L}_{ij} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right] + \frac{1}{M} \sum_{j=1}^l e_j^l \otimes c_j \sum_{i=1}^M [(\rho(d_i) - b_{m_i})] = \mathbf{0}, \end{aligned}$$

where $\sum_{i=1}^M [(\rho(d_i) - b_{m_i})] = \sum_{i=1}^M \rho_{m_i}^\epsilon(d_i) - b_M = g(\mathbf{y}, \boldsymbol{\xi})$,

the third term is $l \times 1$ vector $M^{-1} g(\mathbf{y}, \boldsymbol{\xi}) \mathbf{c}$,

the second term is $l \times 1$ vector \mathbf{U}_M with entries

$$U_{Mj} = M^{-1} S_{Ml} \sum_{i=1}^M \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \mathbf{L}_{ij} \mathbf{V}_i^{-1/2} \mathbf{r}_i \right],$$

and the first term is written as

$$\frac{1}{M} \sum_{j=1}^l e_j^l \otimes [\mathbf{Q}(\boldsymbol{\theta})]_j \cdot \boldsymbol{\theta} \sum_{i=1}^M \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\left(\sum_{s=1}^l \mathbf{L}_{is} \right) \right) \mathbf{V}_i^{-1/2} \mathbf{r}_i \right].$$

Let us further denote

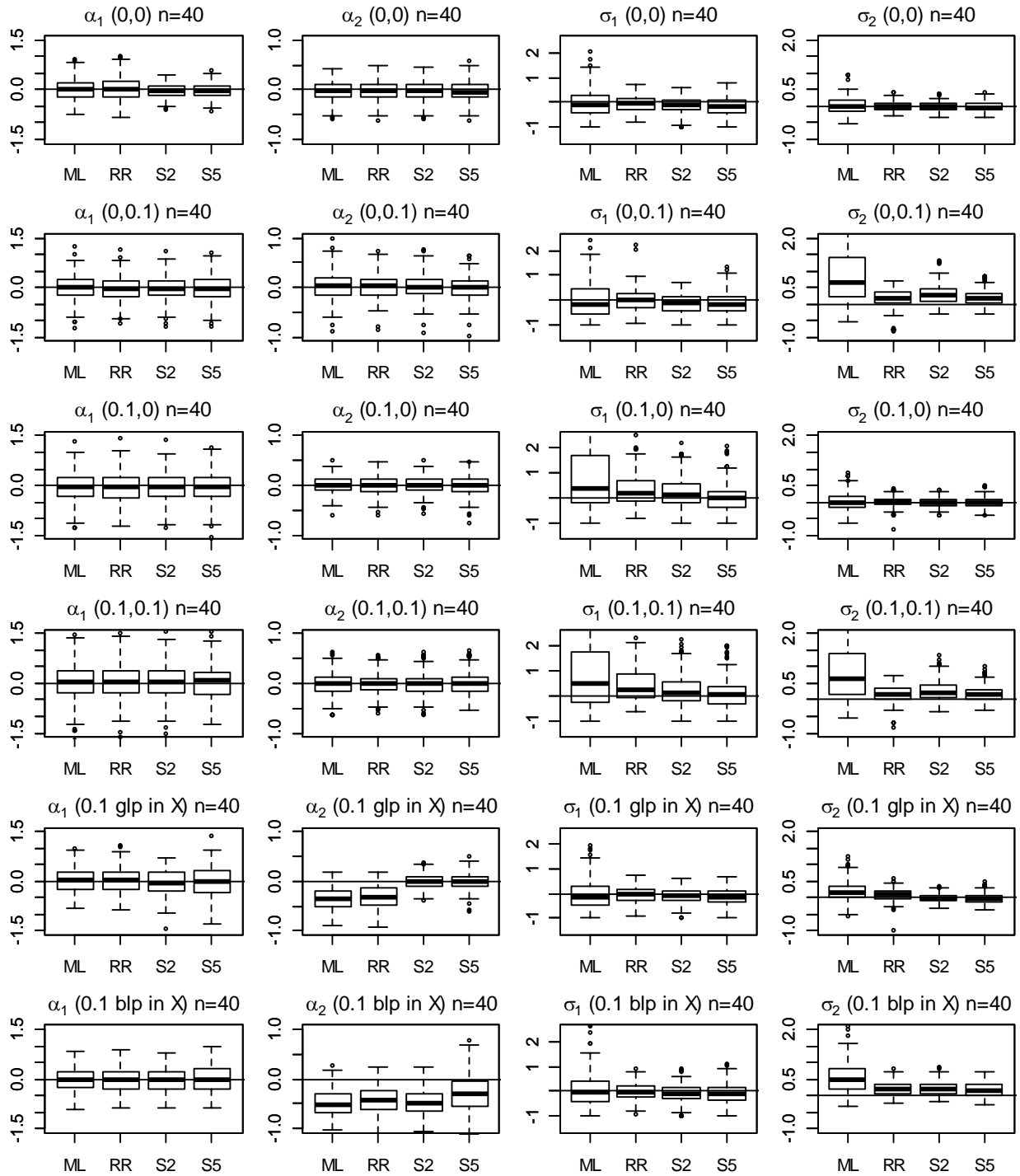
$$P_M = M^{-1} \sum_{i=1}^M \left[u(d_i) \mathbf{r}'_i \mathbf{V}_i^{-1/2} \left(\sum_{s=1}^l \mathbf{L}_{is} \right) \mathbf{V}_i^{-1/2} \mathbf{r}_i \right]$$

then the first term is $l \times 1$ vector $P_M \mathbf{Q}(\boldsymbol{\theta}) \boldsymbol{\theta}$ and system (3.2) is re-written as

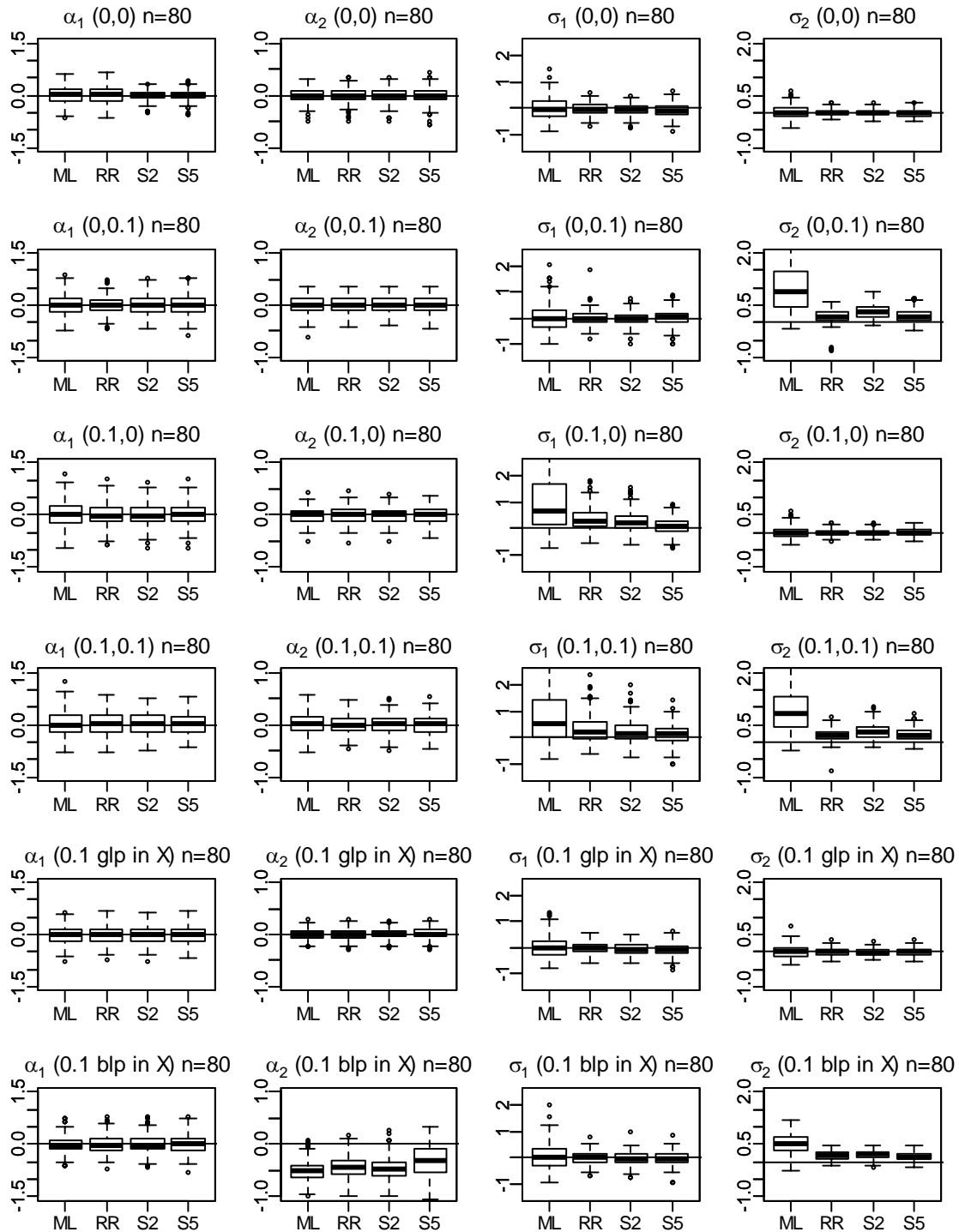
$$P_M \mathbf{Q}(\boldsymbol{\theta}) \boldsymbol{\theta} + \mathbf{U}_M + M^{-1} g(\mathbf{y}, \boldsymbol{\xi}) \mathbf{c} = \mathbf{0}.$$

References

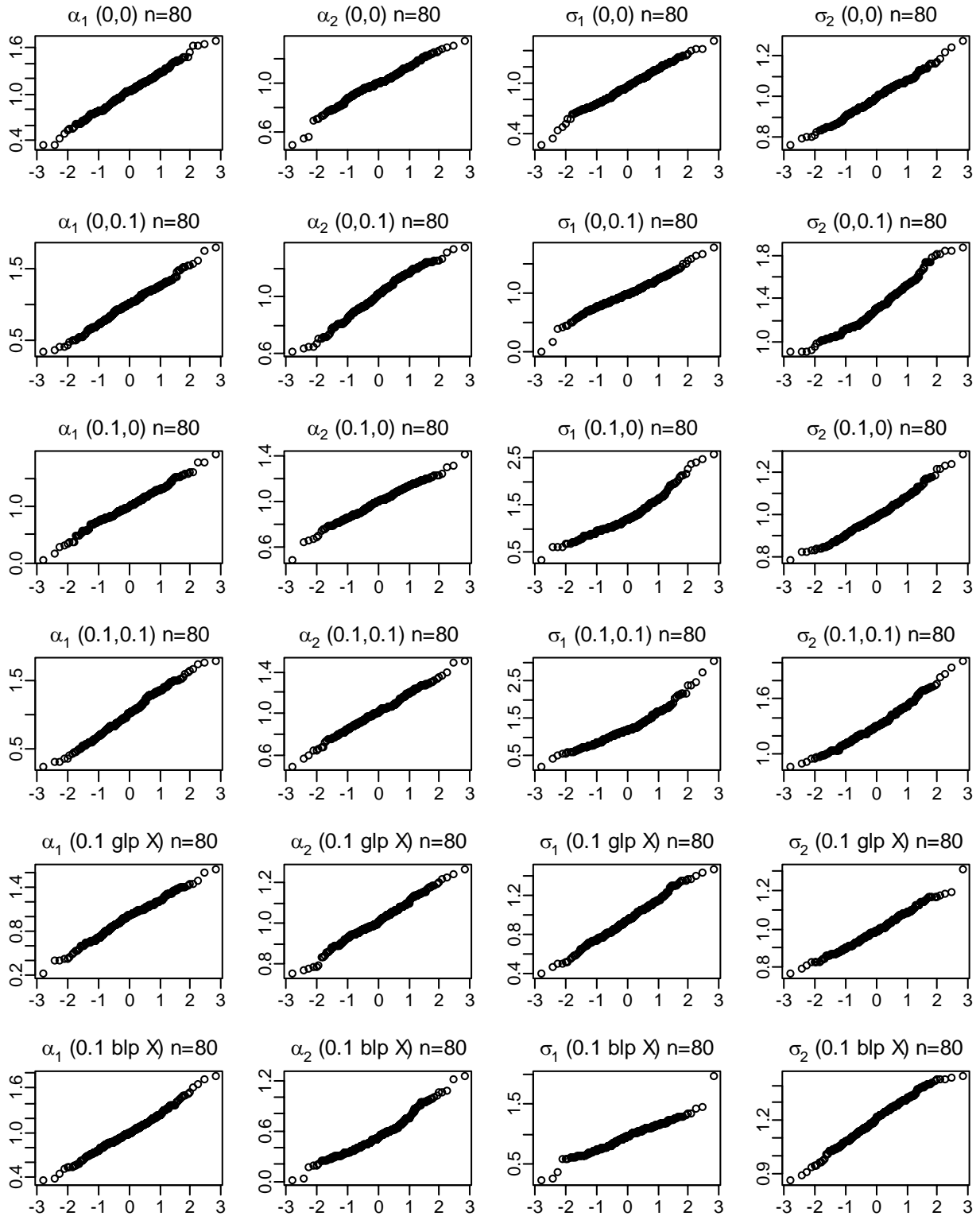
- Kackar, R. N. and Harville, D. A. (1981). Unbiasedness of two-stage estimation and prediction procedures for mixed linear models. *Comm. Statist. Theory Methods* **10**, 1249–1261.
- Mukhopadhyay, P. (2009) *Multivariate statistical analysis*. New Jersey: World Scientific.
- Neudecker, H. (1969) Some theorems on matrix differentiation with special reference to Kronecker matrix products. *Journal of the American Statistical Association* **64**, 953-963.
- Shao, J (2003) *Mathematical Statistics*. New York: Springer.
- Varadhan, R. and Gilbert, P. (2009). BB: An R Package for Solving a Large System of Non-linear Equations and for Optimizing a High-Dimensional Nonlinear Objective Function. *Journal of Statistical Software* **32(4)**, 1-26.
- Yuan, K.-H. (1997) A theorem of uniform convergence of stochastic functions with applications, *Journal of Multivariate Analysis* **62**, 100-109.
- Yuan, K.-H. and Jennrich, R.I. (1998) Asymptotics of estimating equations under natural conditions. *Journal of Multivariate Analysis* **65**, 245-260.



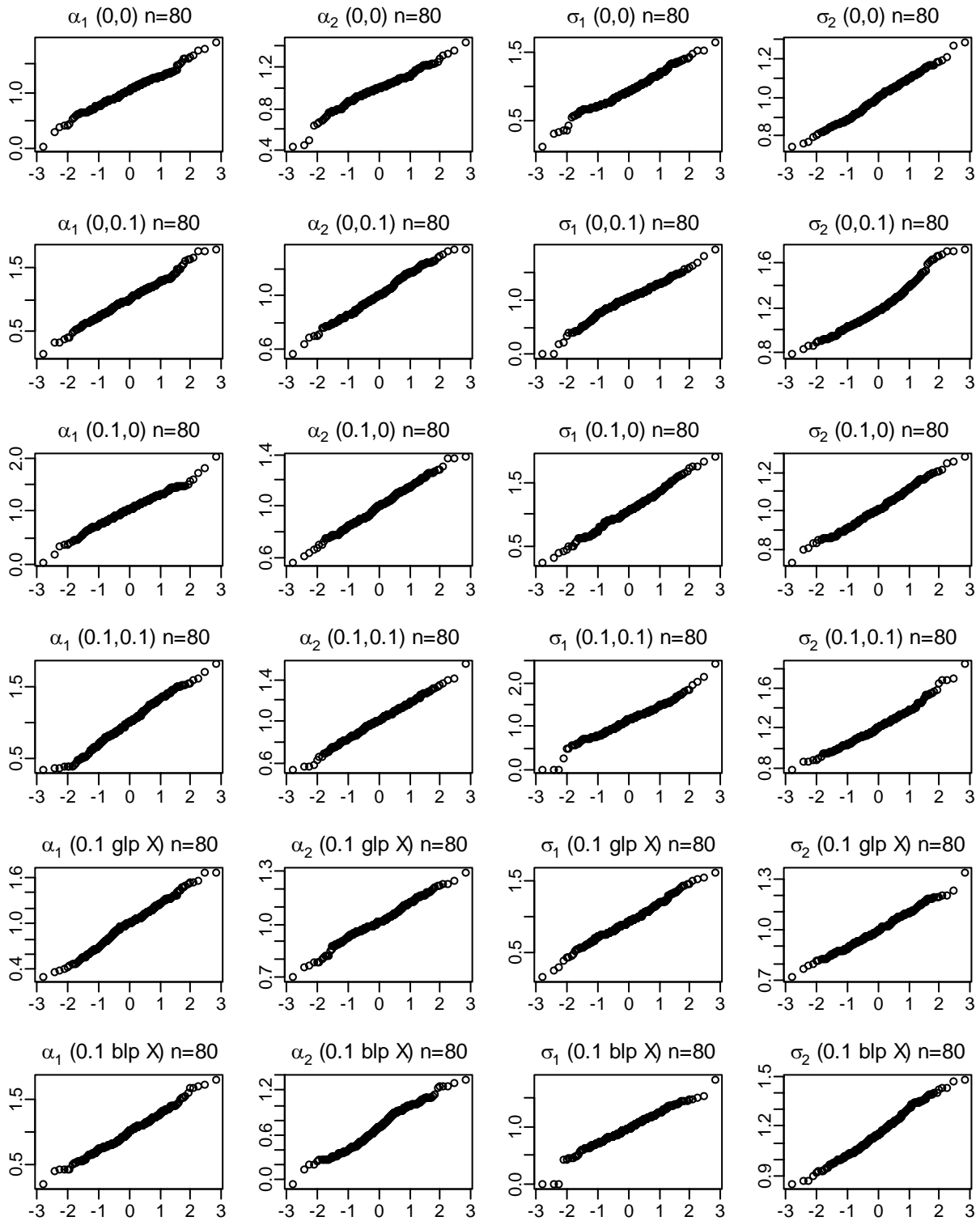
Supplementary Figure 1. Error distribution of parameter estimates for simulated data sets with $n=40$ without contamination (0,0), with 10% contamination in error terms (0,0.1), with 10% contamination in random effects (0.1,0), with 10% contamination in both error terms and random effects (0.1,0.1), with 10% of "good" (0.1 glp in X) and "bad" (0.1 blp in X) leverage points. Four estimation methods used are REML (ML), Robust REML II (RR), GS-estimator with breakdown parameter 0.2 (S2), GS-estimator with breakdown parameter 0.5 (S5).



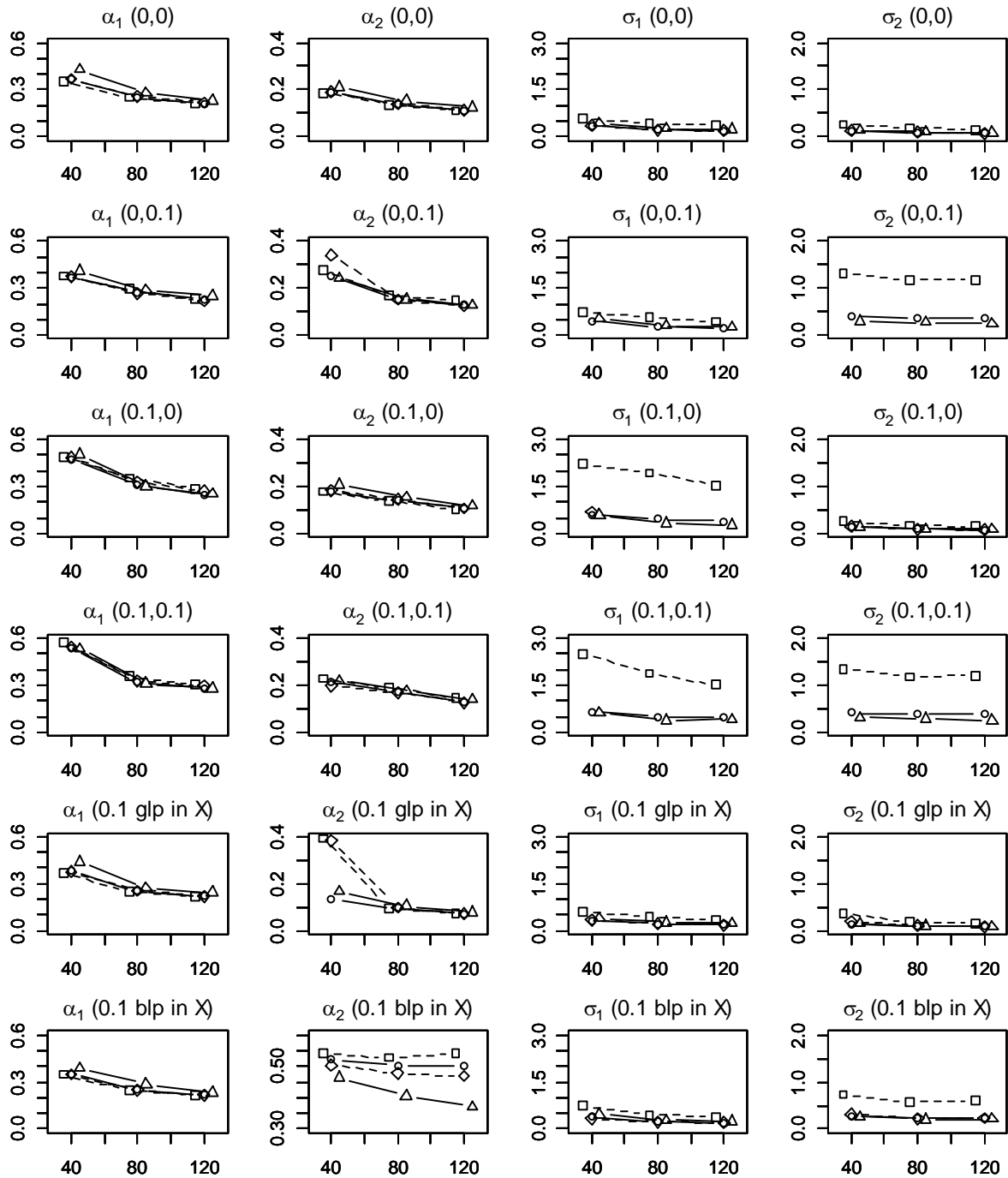
Supplementary Figure 2. Error distribution of parameter estimates for simulated data sets with $n=80$ without contamination (0,0), with 10% contamination in error terms (0,0.1), with 10% contamination in random effects (0.1,0), with 10% contamination in both error terms and random effects (0.1,0.1), and with 10% of "good" (0.1 glp in X) and "bad" (0.1 blp in X) leverage points. Four estimation methods used are REML (ML), Robust REML II (RR), GS-estimator with breakdown parameter 0.2 (S2), GS-estimator with breakdown parameter 0.5 (S5).



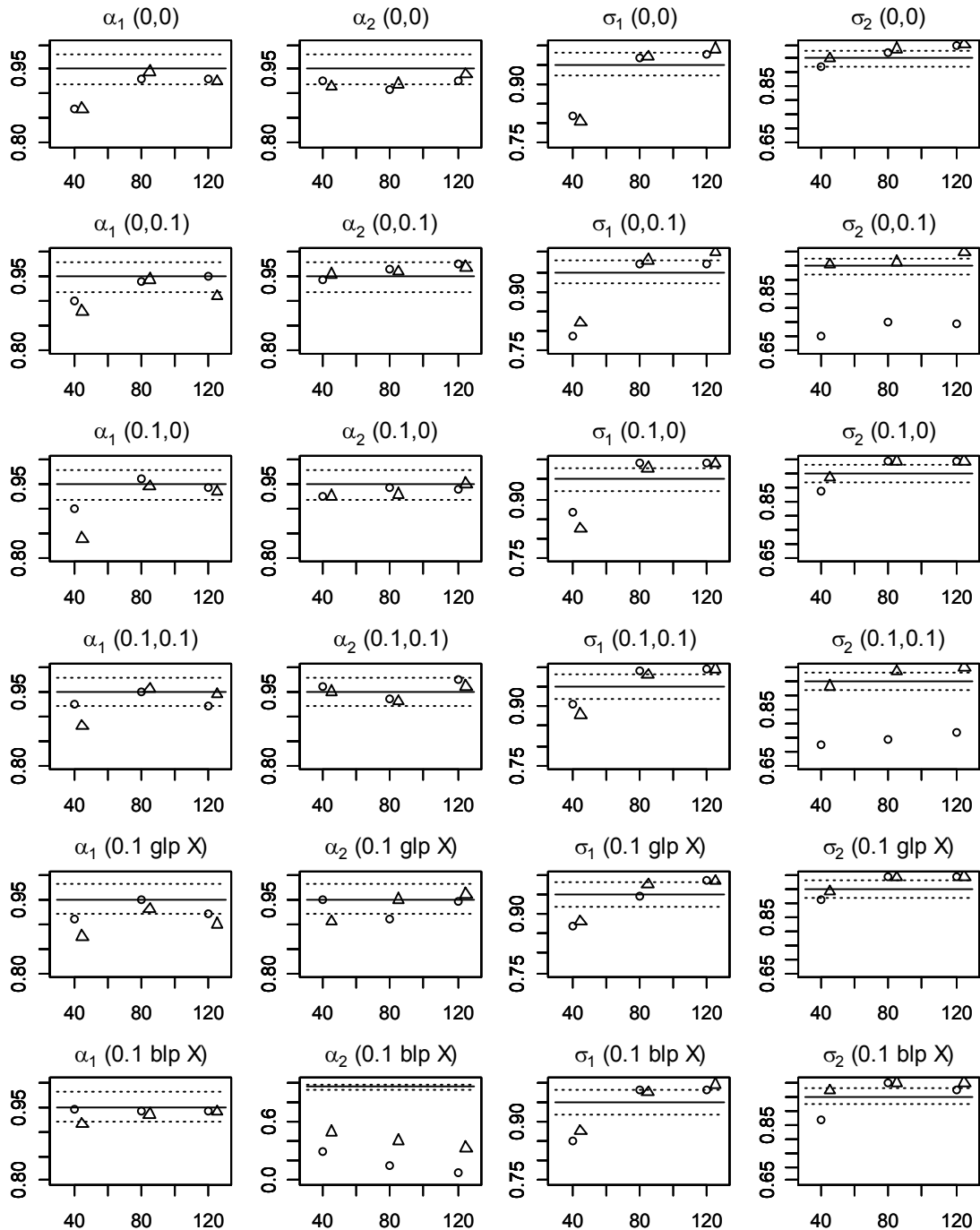
Supplementary Figure 3. QQplots of GS parameter estimates with breakdown parameter 0.2, $n=80$.



Supplementary Figure 4. QQplots of GS parameter estimates with breakdown parameter 0.5, $n=80$.



Supplementary Figure 5. Root mean squared errors (RMSE) in simulated data without contamination (0,0), with 10% contamination in error terms (0,0.1), 10% contamination in random effects (0.1,0), 10% contamination in both error terms and random effects (0.1,0.1), and 10% of "good" (0.1 glp in X) and "bad" (0.1 blp in X) leverage points. Estimation methods used are REML (squares, dashed lines), Robust REML II (diamonds, dashed lines), GS-estimator with breakdown parameter 0.2 (circles, solid lines), GS-estimator with breakdown parameter 0.5 (triangles, solid lines). Missing points for Robust REML II correspond to the mean squared errors outside the plotting range.



..

Uwrrigo gpvct{ 'Hli wtg'80 Coverage of asymptotic 95% confidence intervals for sample sizes 40, 80 and 120 in simulated data sets without contamination (0,0), with 10% contamination in error terms (0,0.1), 10% contamination in random effect of group (0.1,0), 10% contamination in both error terms and random effect of group (0.1,0.1), and 10% of "good" (0.1 glp in X) and "bad" (0.1 glp in X) leverage points. Results for GS-estimator with breakdown parameter 0.2 plotted with circles and for GS-estimator with breakdown parameter 0.5 plotted with triangles. Solid lines show nominal 95%, and the dotted lines limit the approximate acceptance region [0.92,0.98] for the null hypothesis $H_0: p = 0.95$.