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## Examining some aspects of balanced sampling in surveys

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### Supplementary Material

#### S1 A numerical example

To illustrate the effect of both sources of bias in (25), we conducted a limited simulation study. We generated a finite population  $P$  of size  $N = 500$ , consisting of two variables: an auxiliary variable  $x$ , whose values were available prior to sampling for all the population units and a characteristic of interest  $y$ . First, the  $x$ -values were generated from a normal distribution with mean 0 and variance equal to 1. Then, five of the  $x$ -values were replaced manually by the values 7.9, 8, 8.2, 8.3 and 8.5. Thus, the distribution of the  $x$ -values may be viewed as a mixture of two populations with 1% contamination. Then, given the  $x$ -values, the  $y$ -values were generated according to

$$y_i = 1 + 2(x_i - \bar{X})^2 + \epsilon_i,$$

where the errors  $\epsilon_i$  were generated from a normal distribution with mean and variance equal to 0 and 1, respectively and  $\bar{X} = N^{-1} \sum_{i \in P} x_i$ . From the population, we generated  $K = 10,000$  samples of size 25 according to the rejective procedure of Fuller (2009) using simple random sampling without replacement as the basic procedure; see Example 3.1 for more details. In each sample, we computed the Monte Carlo contribution (in %) to the bias of  $\hat{t}_{\text{reg}}^p$  for each of the two terms on the right hand-side of (25), referred below as term 1 and term 2, respectively, for several values of the balancing tolerance  $\gamma$ . In addition, for each value of  $\gamma$ , we computed the percent Monte Carlo relative bias of  $\hat{t}_{\text{reg}}^p$ . The results are shown in Table 1.

From Table 1, we note that, for small values of  $\gamma$ , the contribution of term 2 was small. For example for  $\gamma \leq 0.32$ , the contribution of Term 2 was ranging from 0% to 5%. This can be explained by the fact that a small value of  $\gamma$  corresponds to a high the rejection rate, which in turns

Table 1: Monte Carlo contribution (in %) of each term on the right hand-side of (25) and percent Monte Carlo relative bias of  $\hat{t}_{\text{reg}}^p$

$\gamma$	0.01	0.06	0.12	0.32	0.67	1.15	1.65	2.24	2.83
Rejection rate (in %)	95.0	84.0	74.0	54.0	30.0	13.0	5.0	1.0	0.3
Contribution (in %) Term 1	100	100	99	95	77	51	29	9	2
Contribution (in %) Term 2	0	0	1	5	23	49	71	91	98
Monte Carlo relative bias of $\hat{t}_{\text{reg}}^p$ (in %)	-8.8	-8.1	-9.1	-9.1	-8.8	-9.9	-11.2	-12.4	-12.3

implies that the term  $\mathbf{t}_z - \hat{\mathbf{t}}_z^p$  is close to zero; that is, the resulting estimator does not suffer from a small sample bias. On the other hand, a small value of  $\gamma$  tends to lead to inclusion probabilities  $\pi_i$  significantly different from the basic inclusion probabilities  $p_i$ , especially for units exhibiting a large  $x$ -value. As a result, the contribution of term 1 is expected to be large. In contrast, large values of  $\gamma$  correspond to a low rejection rate, in which case there is not much difference between a rejective sampling procedure and a non-rejective one. In this case, the  $\pi$ 's are expected to be close to the  $p_i$ 's and the resulting estimator would suffer from small sample bias only; that is, the contribution of term 2 is expected to be large. This is confirmed by the results shown in Table 1.

## S2 Other simulation results

Tables 3-11 below show the Monte Carlo results corresponding to a 50% rejection rate. The results are relatively similar to those obtained for a 90% rejection rate, except for Basic and Fuller. The latter showed some differences, especially in terms of RE. For example, when the  $x$ -values were generated according to a mixture distribution, Basic showed a value of RE equal to 57% for  $n = 25$  when the  $y$ -values were generated according to the linear model; see Table 4. The same was true, regardless of the distribution of the  $x$ -value. The differences between Basic and Fuller can be explained by the fact that, with a 50% rejection rate, the term  $\mathbf{t}_z - \hat{\mathbf{t}}_z^p$  on the right

S2. OTHER SIMULATION RESULTS

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			Linear			Quadratic			Exponential		
	Mean	Var.	Mean	Var.	Corr.	Mean	Var.	Corr.	Mean	Var.	Corr.
Normal	1.96	1.03	1.04	5.09	0.90	3.16	8.10	-0.01	4.54	28.0	0.82
Mixture	2.08	1.41	0.96	6.65	0.92	3.89	62.0	-0.39	16.5	15 900	0.53
Lognormal	1.52	2.80	1.05	12.5	0.96	6.62	6.27	0.77	5 570	8.7 10 <sup>9</sup>	0.47
			Bump			Anova			Logistic		
	Mean	Var.	Mean	Var.	Corr.	Mean	Var.	Corr.	Mean	Var.	Corr.
Normal			1.65	21.5	-0.01	16.0	64.6	0.70	0.48	0.25	0.62
Mixture			2.59	77.0	0.36	16.2	61.8	0.47	0.53	0.25	0.56
Lognormal			5.47	647	0.76	9.18	800	0.46	0.30	0.21	0.57

Table 2: Mean and variance of the auxiliary variable for three distributions, and mean, variance and coefficient of correlation with the auxiliary variable of the variable of interest for six models

hand-side of (24) may have been large for some samples since a low rejection rate corresponds to a large value of the balancing tolerance  $\gamma$ . In this case, using a calibration strategy such as Fuller generally "pays off" in terms of efficiency.

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-0	-0	-0	-0	-0	-0	-0	-1
	CV	23	19	24	19	20	19	42	19
	RE	63	100	62	100	85	98	20	95
Quadratic	RB	-2	-2	-0	-1	-0	-0	-0	-4
	CV	17	18	18	18	18	18	18	19
	RE	103	100	93	95	95	95	99	84
Exponential	RB	-1	-1	0	-0	-0	-0	0	-3
	CV	15	13	16	14	14	13	23	14
	RE	79	100	69	93	87	95	32	88
Bump	RB	-5	-7	-1	-3	-0	-1	-0	-12
	CV	55	56	56	56	56	56	56	59
	RE	102	100	100	99	101	101	101	85
Stratification	RB	0	1	-0	0	0	0	0	1
	CV	4	8	-2	2	0	1	1	16
	RE	90	100	95	100	95	98	52	88
Logistic	RB	-0	-0	-0	-0	0	0	-0	-0
	CV	17	16	17	16	16	16	20	16
	RE	92	100	93	101	101	103	63	97

Table 3: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 25$ , with a rejection rate equal to 50% and for a normal distribution of  $x$ .

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-1	-0	0	0	0	-0	0	-0
	CV	28	21	31	21	26	21	53	21
	RE	57	100	45	100	67	100	16	101
Quadratic	RB	-8	-10	-0	-6	0	-2	-0	-14
	CV	34	33	49	39	39	37	39	32
	RE	97	100	50	75	77	88	76	95
Exponential	RB	-25	-29	0	-20	2	-8	0	-38
	CV	115	108	177	130	139	124	148	101
	RE	90	100	40	72	65	81	57	107
Bump	RB	-13	-16	-0	-10	1	-4	-0	-23
	CV	59	58	80	66	66	63	66	58
	RE	99	100	57	80	81	90	82	91
Stratification	RB	1	2	-0	1	0	0	0	2
	CV	14	20	-0	11	1	6	1	29
	RE	97	100	101	89	95	94	76	91
Logistic	RB	1	1	0	1	-0	0	0	1
	CV	16	15	15	15	15	15	18	16
	RE	94	100	99	99	99	100	70	97

Table 4: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 25$ , with a rejection rate equal to 50% and for a mixture distribution of  $x$ .

S2. OTHER SIMULATION RESULTS

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-3	-0	-0	-0	0	0	-0	-0
	CV	30	19	51	19	33	19	65	19
	RE	38	100	13	100	32	98	8	94
Quadratic	RB	-15	-17	0	-14	-0	-8	-0	-21
	CV	42	36	108	39	60	41	73	37
	RE	79	100	13	91	43	89	29	85
Exponential	RB	-66	-72	-0	-69	-1	-39	-3	-67
	CV	188	156	574	171	319	207	320	152
	RE	74	100	9	87	29	67	29	106
Bump	RB	-18	-21	0	-18	-0	-10	-0	-27
	CV	53	46	132	49	75	52	90	49
	RE	81	100	15	92	45	90	31	81
Stratification	RB	3	5	0	4	0	3	0	9
	CV	17	28	1	20	0	13	0	40
	RE	108	100	116	96	117	103	90	73
Logistic	RB	2	4	-0	3	-0	2	-0	5
	CV	25	25	24	25	24	25	30	26
	RE	99	100	107	100	104	101	70	90

Table 5: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 25$ , with a rejection rate equal to 50% and for a log-normal distribution of  $x$ .

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	0	-0	0	0	-0	-0	0	-0
	CV	16	13	16	13	14	13	29	13
	RE	64	100	64	100	91	100	20	100
Quadratic	RB	-1	-1	0	-0	-0	-0	0	-2
	CV	12	12	12	12	12	12	12	13
	RE	101	100	97	98	96	96	99	91
Exponential	RB	-0	-1	0	-0	-0	-0	0	-1
	CV	10	9	11	9	10	9	16	9
	RE	78	100	73	97	89	94	33	96
Bump	RB	-2	-3	0	-1	-1	-1	0	-5
	CV	38	38	38	38	38	38	38	39
	RE	101	100	100	99	100	100	101	93
Stratification	RB	0	0	-0	0	0	0	-0	1
	CV	3	5	-1	1	0	1	-0	10
	RE	89	100	91	100	97	98	51	93
Logistic	RB	0	0	0	0	-0	-0	-0	-0
	CV	11	11	11	11	11	11	14	11
	RE	92	100	93	101	97	98	61	98

Table 6: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 50$ , with a rejection rate equal to 50% and for a normal distribution of  $x$ .

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-0	-0	0	0	0	0	0	-0
	CV	19	14	19	14	16	14	36	14
	RE	58	100	54	100	80	103	16	99
Quadratic	RB	-3	-4	0	-2	0	-1	1	-7
	CV	25	25	29	27	26	26	27	25
	RE	99	100	76	85	91	96	85	94
Exponential	RB	-10	-13	1	-6	0	-3	1	-22
	CV	88	85	104	95	93	89	103	82
	RE	95	100	69	83	86	94	70	103
Bump	RB	-5	-7	1	-3	0	-1	1	-12
	CV	43	42	48	46	45	44	46	43
	RE	100	100	80	87	92	96	89	93
Stratification	RB	0	1	-0	0	-0	0	-0	1
	CV	8	12	-1	5	-0	2	-1	22
	RE	97	100	99	93	97	96	77	94
Logistic	RB	0	0	0	0	0	0	0	1
	CV	11	10	11	10	11	11	13	11
	RE	93	100	95	99	98	98	67	97

Table 7: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 50$ , with a rejection rate equal to 50% and for a mixture distribution of  $x$ .

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-1	0	0	0	0	0	-0	0
	CV	21	13	24	13	19	13	45	13
	RE	38	100	30	100	46	100	8	97
Quadratic	RB	-7	-8	0	-5	-0	-3	-0	-12
	CV	34	31	49	34	39	32	51	30
	RE	83	100	42	87	69	100	39	97
Exponential	RB	-31	-37	1	-29	-0	-17	-1	-48
	CV	184	172	274	193	218	187	225	149
	RE	89	100	41	82	65	88	61	127
Bump	RB	-8	-10	0	-7	-0	-4	-0	-15
	CV	43	39	60	42	48	40	62	38
	RE	84	100	44	87	71	100	41	96
Stratification	RB	1	2	0	1	-0	1	-0	4
	CV	11	17	1	11	-0	6	-1	28
	RE	103	100	107	96	113	106	85	83
Logistic	RB	1	2	0	1	0	1	-0	3
	CV	17	17	17	17	17	17	21	17
	RE	96	100	100	99	104	103	70	95

Table 8: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 50$ , with a rejection rate equal to 50% and for a log-normal distribution of  $x$ .

S2. OTHER SIMULATION RESULTS

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-0	-0	0	0	-0	-0	0	-0
	CV	11	9	11	9	9	9	19	9
	RE	66	100	65	100	100	105	20	102
Quadratic	RB	-0	-0	-0	-0	-0	-0	0	-1
	CV	8	8	8	8	8	8	8	8
	RE	101	100	99	99	102	102	103	98
Exponential	RB	-0	-0	-0	-0	-0	-0	0	-1
	CV	7	6	7	6	6	6	10	6
	RE	79	100	77	99	99	102	35	100
Bump	RB	-1	-1	-0	-0	0	0	0	-2
	CV	25	25	25	25	25	25	25	26
	RE	101	100	100	100	103	103	101	97
Stratification	RB	0	0	-0	0	0	0	-0	0
	CV	1	2	-1	0	0	1	-1	6
	RE	89	100	90	100	99	100	52	99
Logistic	RB	-0	-0	-0	-0	-0	-0	-0	-0
	CV	8	7	8	7	7	7	9	7
	RE	92	100	92	100	100	101	63	102

Table 9: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 100$ , with a rejection rate equal to 50% and for a normal distribution of  $x$ .

Model		Basic	Fuller	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-0	-0	0	-0	0	0	0	0
	CV	13	10	13	10	10	10	24	10
	RE	57	100	56	100	87	100	16	98
Quadratic	RB	-1	-2	-0	-1	0	-0	-0	-4
	CV	17	17	18	18	17	17	18	17
	RE	99	100	90	94	97	99	89	93
Exponential	RB	-5	-6	-0	-3	0	-1	-0	-11
	CV	60	59	64	62	61	60	68	60
	RE	96	100	85	92	95	98	75	95
Bump	RB	-2	-3	-0	-2	0	-0	-0	-6
	CV	29	29	30	30	29	29	30	30
	RE	100	100	92	95	98	100	92	93
Stratification	RB	0	0	0	0	-0	-0	0	1
	CV	7	9	1	4	-1	-0	0	16
	RE	97	100	97	97	100	98	78	95
Logistic	RB	0	0	-0	0	0	0	-0	0
	CV	7	7	7	7	7	7	8	7
	RE	94	100	95	100	103	103	69	99

Table 10: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 100$ , with a rejection rate equal to 50% and for a mixture distribution of  $x$ .

Model		Basic	<i>Fuller</i>	MC	MC-Reg	Cube	Cube-Reg	SRS	SRS-Reg
Linear	RB	-0	-0	0	-0	-0	-0	0	0
	CV	14	9	14	9	11	9	30	9
	RE	39	100	36	100	61	102	8	100
Quadratic	RB	-2	-3	0	-2	-0	-1	0	-6
	CV	24	22	27	23	24	21	34	22
	RE	84	100	67	91	88	107	43	93
Exponential	RB	-10	-14	2	-8	-2	-8	-1	-28
	CV	138	134	157	144	141	134	149	125
	RE	96	100	74	88	92	101	82	111
Bump	RB	-2	-4	1	-2	-0	-2	0	-8
	CV	30	28	34	29	29	27	42	28
	RE	85	100	68	92	89	107	45	92
Stratification	RB	1	1	0	0	0	0	0	2
	CV	6	9	1	5	0	2	2	20
	RE	99	100	100	98	110	107	81	88
Logistic	RB	0	1	0	0	0	0	-0	1
	CV	12	11	11	11	11	11	14	11
	RE	95	100	96	99	101	101	68	96

Table 11: Monte Carlo percent relative bias, percent coefficient of variation, percent bias ratio and percent relative efficiency of several estimators under three sampling designs of size  $n = 100$ , with a rejection rate equal to 50% and for a log-normal distribution of  $x$ .

### S3 Proof of equation (30)

We first compute an expansion of  $pr(-\gamma \leq X \leq \gamma) = F_X(\gamma) - F_X(-\gamma)$ . From the formal Edgeworth expansion (see Thompson, 1997, equation (3.41)), we get

$$\begin{aligned} F_X(\gamma) &= \psi(\gamma) - \phi(\gamma) \left\{ \frac{\kappa_3}{6}(\gamma^2 - 1) \right\} \\ &\quad - \phi(\gamma) \left\{ \frac{\kappa_4}{24}(\gamma^3 - 3\gamma) + \frac{\kappa_3^2}{72}(\gamma^5 - 10\gamma^3 + 15\gamma) \right\} + o(d^{-1}), \\ F_X(-\gamma) &= 1 - \psi(\gamma) - \phi(\gamma) \left\{ \frac{\kappa_3}{6}(\gamma^2 - 1) \right\} \\ &\quad + \phi(\gamma) \left\{ \frac{\kappa_4}{12}(\gamma^3 - 3\gamma) + \frac{\kappa_3^2}{36}(\gamma^5 - 10\gamma^3 + 15\gamma) \right\} + o(d^{-1}), \end{aligned} \tag{S3.1}$$

where  $\psi(\cdot)$  and  $\phi(\cdot)$  are the cumulative distribution function and the probability density function of a standard normal distribution with

$$\begin{aligned} \kappa_3 &\equiv \mu_3(X) = d^{-3/2} \sum_{i \in P} \tilde{x}_i^3 p_i(1 - p_i)(1 - 2p_i), \\ \kappa_4 &\equiv \mu_4(X) - 3\{\mu_2(X)\}^2 = d^{-2} \sum_{i \in P} \tilde{x}_i^4 p_i(1 - p_i)\{1 - 6p_i(1 - p_i)\}, \end{aligned}$$

and  $\mu_m(X)$  denoting the centered moment of order  $m$  of the random variable  $X$ . Equations (S3.1) lead to

$$\begin{aligned} pr(-\gamma \leq X \leq \gamma) &= \{2\psi(\gamma) - 1\} \\ &\quad - \phi(\gamma) \left\{ \frac{\kappa_4}{12}(\gamma^3 - 3\gamma) + \frac{\kappa_3^2}{36}(\gamma^5 - 10\gamma^3 + 15\gamma) \right\} + o(d^{-1}). \end{aligned} \tag{S3.2}$$

We now obtain an expansion of  $pr(-\gamma \leq X \leq \gamma | I_{bi} = 1)$ . We have

$$pr(X \leq \gamma | I_{bi} = 1) = pr(X_i \leq \gamma_i),$$

where

$$X_i = \frac{1}{\sqrt{d - \tilde{x}_i^2 p_i(1 - p_i)}} \sum_{j \neq i \in U} \tilde{x}_j (I_{bj} - p_j)$$

and

$$\gamma_i = \frac{1}{\sqrt{1 - d^{-1} \tilde{x}_i^2 p_i(1 - p_i)}} \left\{ \gamma - d^{-1/2} \tilde{x}_i (1 - p_i) \right\}.$$

Applying once again the formal Edgeworth expansion to  $X_i$  leads to

$$\begin{aligned}
 F_{X_i}(\gamma_i) &= \psi(\gamma_i) - \phi(\gamma_i) \left\{ \frac{\kappa_{3i}}{6}(\gamma_i^2 - 1) \right\} \\
 &\quad - \phi(\gamma_i) \left\{ \frac{\kappa_{4i}}{24}(\gamma_i^3 - 3\gamma_i) + \frac{\kappa_{3i}^2}{72}(\gamma_i^5 - 10\gamma_i^3 + 15\gamma_i) \right\} + o(d^{-1}),
 \end{aligned} \tag{S3.3}$$

where

$$\kappa_{3i} \equiv \mu_3(X_i) = \{d - \tilde{x}_i^2 p_i(1 - p_i)\}^{-3/2} \sum_{j \neq i \in P} \tilde{x}_j^3 p_j(1 - p_j)(1 - 2p_j)$$

and

$$\kappa_{4i} \equiv \mu_4(X_i) - 3\{\mu_2(X_i)\}^2 = \{d - \tilde{x}_i^2 p_i(1 - p_i)\}^{-2} \sum_{j \neq i \in P} \tilde{x}_j^4 p_j(1 - p_j)\{1 - 6p_j(1 - p_j)\}.$$

By neglecting terms of smaller order than  $d^{-1}$ , we obtain after some algebra

$$\begin{aligned}
 F_{X_i}(\gamma_i) &= \psi(\gamma) - \phi(\gamma) \left\{ \frac{1}{\sqrt{d}} \tilde{x}_i(1 - p_i) + \frac{1}{6} \kappa_3(\gamma^2 - 1) \right\} \\
 &\quad + \left[ \frac{1}{2d} \tilde{x}_i^2(1 - p_i) (\gamma p_i \phi(\gamma) + (1 - p_i) \phi'(\gamma)) \right. \\
 &\quad + \frac{\kappa_3 \gamma \phi(\gamma)}{3\sqrt{d}} \tilde{x}_i(1 - p_i) - \frac{\kappa_4 \phi(\gamma)}{24} (\gamma^3 - 3\gamma) \\
 &\quad \left. - \phi(\gamma) \frac{\kappa_3^2}{72} (\gamma^5 - 10\gamma^3 + 15\gamma) + \frac{1}{6} \frac{\kappa_3(\gamma^2 - 1)}{\sqrt{d}} \tilde{x}_i(1 - p_i) \phi'(\gamma) \right] \\
 &\quad + o(d^{-1}).
 \end{aligned} \tag{S3.4}$$

Also,  $\text{pr}(X \leq -\gamma | I_{bi} = 1) = F_{X_i}(-\tilde{\gamma}_i)$ , where

$$\tilde{\gamma}_i = \frac{1}{\sqrt{1 - d^{-1} \tilde{x}_i^2 p_i(1 - p_i)}} \{ \gamma + d^{-1/2} \tilde{x}_i(1 - p_i) \}.$$

Similar arguments lead to

$$\begin{aligned}
F_{X_i}(-\tilde{\gamma}_i) &= \{1 - \psi(\gamma)\} - \phi(\gamma) \left\{ \frac{1}{\sqrt{d}} \tilde{x}_i(1 - p_i) + \frac{1}{6} \kappa_3(\gamma^2 - 1) \right\} \\
&+ \left[ \frac{1}{2d} \tilde{x}_i^2(1 - p_i) \{-\gamma p_i \phi(\gamma) - (1 - p_i) \phi'(\gamma)\} \right. \\
&\quad - \frac{\kappa_3 \gamma \phi(\gamma)}{3\sqrt{d}} \tilde{x}_i(1 - p_i) + \frac{\kappa_4 \phi(\gamma)}{24} (\gamma^3 - 3\gamma) \\
&\quad \left. + \phi(\gamma) \frac{\kappa_3^2}{72} (\gamma^5 - 10\gamma^3 + 15\gamma) - \frac{1}{6} \frac{\kappa_3(\gamma^2 - 1)}{\sqrt{d}} \tilde{x}_i(1 - p_i) \phi'(\gamma) \right] \\
&+ o(d^{-1}).
\end{aligned} \tag{S3.5}$$

We obtain

$$\begin{aligned}
pr(-\gamma \leq X \leq \gamma | I_i = 1) &= \{2\psi(\gamma) - 1\} + \frac{1}{d} \tilde{x}_i^2(1 - p_i) \{\gamma p_i \phi(\gamma) + (1 - p_i) \phi'(\gamma)\} \\
&+ 2 \left\{ \frac{\kappa_3 \gamma \phi(\gamma)}{3\sqrt{d}} \tilde{x}_i(1 - p_i) - \frac{\kappa_4 \phi(\gamma)}{24} (\gamma^3 - 3\gamma) \right. \\
&\quad \left. - \phi(\gamma) \frac{\kappa_3^2}{72} (\gamma^5 - 10\gamma^3 + 15\gamma) + \frac{1}{6} \frac{\kappa_3(\gamma^2 - 1)}{\sqrt{d}} \tilde{x}_i(1 - p_i) \phi'(\gamma) \right\} \\
&+ o(d^{-1}).
\end{aligned} \tag{S3.6}$$

Plugging (S3.2) and (S3.6) into (29) and making use of the identity  $\phi'(\gamma) = -\gamma\phi(\gamma)$ , we obtain (30).

## References

- Fuller, W.A. (2009). Some design properties of a rejective sampling procedure. *Biometrika* **96**, 933–944.
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