

# Supplemental file for “Efficient learning of nonparametric directed acyclic graph with statistical guarantee”

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## Technical proofs

**Proof of Theorem 1.** For any  $t = 0, \dots, T - 1$ , given  $\mathcal{S}_t$ , it is always true that  $E \text{Var}(x_j | \mathbf{x}_{\mathcal{S}_t}) = E \text{Var}(x_j | \mathbf{x}_{\text{pa}_j}) = \sigma_j^2$  for any  $j \in \mathcal{A}_t$ , due to the fact that  $\text{pa}_j \subset \mathcal{S}_t$  if  $j \in \mathcal{A}_t$ . Moreover, for any  $j \in \mathcal{V} \setminus \{\mathcal{S}_t \cup \mathcal{A}_t\}$ , by total variance, we have

$$\begin{aligned} E[\text{Var}(x_j | \mathbf{x}_{\mathcal{S}_t})] &= E[E[\text{Var}(x_j | \mathbf{x}_{\text{pa}_j}) | \mathbf{x}_{\mathcal{S}_t}]] + E[\text{Var}(E[x_j | \mathbf{x}_{\text{pa}_j}] | \mathbf{x}_{\mathcal{S}_t})] \\ &= \sigma_j^2 + E[\text{Var}(E[x_j | \mathbf{x}_{\text{pa}_j}] | \mathbf{x}_{\mathcal{S}_t})]. \end{aligned}$$

This completes the first part of Theorem 1. Additionally, by Assumption 1 in the main text, for any  $j, j' \in \mathcal{A}_t$ , we have

$$E[\text{Var}(x_j | \mathbf{x}_{\mathcal{S}_t})] = E[\text{Var}(x_{j'} | \mathbf{x}_{\mathcal{S}_t})] := \sigma_{t, \min}^2,$$

and for any  $k \in \mathcal{V} \setminus \{\mathcal{S}_t \cup \mathcal{A}_t\}$ , we have

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$$E[\text{Var}(x_k|\mathbf{x}_{\mathcal{S}_t})] = \sigma_k^2 + E[\text{Var}(E[x_k|\mathbf{x}_{\text{pa}_k}]|\mathbf{x}_{\mathcal{S}_t})] > \sigma_{t,\min}^2 + M_{\max}. \quad (1)$$

Clearly, all the nodes in  $\mathcal{A}_t$  can be exactly identified by evaluating the expected conditional variance. This completes the proof.  $\blacksquare$

**Proof of Theorem 3.** Note that the sample variance estimator

$$\widehat{\text{Var}}(x_k) = \frac{1}{n-1} \sum_{i=1}^n \left( x_{ik} - \frac{1}{n} \sum_{j=1}^n x_{jk} \right)^2 = \frac{1}{\binom{n}{2}} \sum_{i < j} \frac{1}{2} (x_{ik} - x_{jk})^2$$

is a U-statistics with kernel  $\frac{1}{2}(x_{ik} - x_{jk})^2$ . By the definition of  $C_{\mathcal{X}}$  that denotes the diameter of the support  $\mathcal{X}$ , then we have  $\frac{1}{2}(x_{ik} - x_{jk})^2 \leq \frac{1}{2}C_{\mathcal{X}}^2$ . Then, by McDiarmid's inequality, for any  $\zeta > 0$  and  $k \in \mathcal{V}$ , there holds

$$P\left(|\widehat{\text{Var}}(x_k) - \text{Var}(x_k)| > \zeta\right) \leq 2 \exp\left(-\frac{n\zeta^2}{2C_{\mathcal{X}}^4}\right). \quad (2)$$

Moreover, we define the following event

$$\mathcal{E}_0 = \left\{ \max_{k \in \mathcal{V}} |\widehat{\text{Var}}(x_k) - \text{Var}(x_k)| \leq \frac{M_{\max}}{4} \right\},$$

and use the notation  $\mathcal{E}_0^c$  to denote its complementary. By (2), we have

$$P(\mathcal{E}_0^c) \leq 2p \exp\left(-\frac{nM_{\max}^2}{32C_{\mathcal{X}}^4}\right). \quad (3)$$

Note that

$$\begin{aligned}
P(\mathcal{A}_0 \neq \widehat{\mathcal{A}}_0) &\leq P(\mathcal{A}_0 \neq \widehat{\mathcal{A}}_0, \mathcal{E}_0) + P(\mathcal{E}_0^c) \\
&\leq P\left(\exists k \in \mathcal{A}_0 \text{ such that } |\widehat{\text{Var}}(x_k) - \widehat{\sigma}_{\min}^{(0)}| \geq \epsilon_0, \mathcal{E}_0\right) \\
&\quad + P\left(\exists k \in \mathcal{V} \setminus \{\mathcal{A}_0\} \text{ such that } |\widehat{\text{Var}}(x_k) - \widehat{\sigma}_{\min}^{(0)}| < \epsilon_0, \mathcal{E}_0\right) + P(\mathcal{E}_0^c) \\
&= P_1 + P_2 + P(\mathcal{E}_0^c), \tag{4}
\end{aligned}$$

where  $\widehat{\sigma}_{\min}^{(0)} = \min_{j \in \mathcal{V}} \widehat{\text{Var}}(x_j)$ . For ease notation, we denote  $k_0 = \operatorname{argmin}_{k \in \mathcal{V}} \widehat{\text{Var}}(x_k)$ , and it always holds true that  $k_0 \in \mathcal{A}_0$ . If not, suppose that  $k_0 \in \mathcal{V} \setminus \{\mathcal{A}_0\}$  and for any  $j \in \mathcal{A}_0$ , under the event  $\mathcal{E}_0$  and by Theorem 1 in the main text, we have

$$\widehat{\text{Var}}(x_{k_0}) > \text{Var}(x_{k_0}) - \frac{M_{\max}}{2} > \text{Var}(x_j) + \frac{M_{\max}}{2} > \widehat{\text{Var}}(x_j),$$

which contradicts the definition that  $k_0 = \operatorname{argmin}_{k \in \mathcal{V}} \widehat{\text{Var}}(x_k)$ .

To bound  $P_1$ , we notice that under the event  $\mathcal{E}_0$ , for any  $j \in \mathcal{A}_0$ , there holds

$$\begin{aligned}
|\widehat{\text{Var}}(x_j) - \widehat{\text{Var}}(x_{k_0})| &= |\widehat{\text{Var}}(x_j) - \text{Var}(x_j) + \text{Var}(x_j) - \text{Var}(x_{k_0}) + \text{Var}(x_{k_0}) - \widehat{\text{Var}}(x_{k_0})| \\
&\leq |\widehat{\text{Var}}(x_j) - \text{Var}(x_j)| + |\text{Var}(x_j) - \text{Var}(x_{k_0})| + |\text{Var}(x_{k_0}) - \widehat{\text{Var}}(x_{k_0})| \\
&\leq \frac{M_{\max}}{4} + 0 + \frac{M_{\max}}{4} = \frac{M_{\max}}{2},
\end{aligned}$$

where the last inequity follows from Assumption 1 in the main text and the definition of  $\mathcal{E}_0$ . Thus, by taking  $\epsilon_0 = \frac{M_{\max}}{2}$ , we have  $P_1 = 0$ .

Next, we turn to bound  $P_2$ . Note that for any  $k \in \mathcal{V} \setminus \{\mathcal{A}_0\}$ , by Theorem 1 in the main text, there holds

$$|\text{Var}(x_k) - \text{Var}(x_{k_0})| \geq M_{\max},$$

and triangle inequality yields that

$$|\text{Var}(x_k) - \text{Var}(x_{k_0})| \leq |\text{Var}(x_k) - \widehat{\text{Var}}(x_k)| + |\widehat{\text{Var}}(x_k) - \widehat{\text{Var}}(x_{k_0})| + |\widehat{\text{Var}}(x_{k_0}) - \text{Var}(x_{k_0})|.$$

Then, under the event  $\mathcal{E}_0$ , we have

$$\begin{aligned} |\widehat{\text{Var}}(x_k) - \widehat{\text{Var}}(x_{k_0})| &\geq M_{\max} - |\text{Var}(x_k) - \widehat{\text{Var}}(x_k)| - |\widehat{\text{Var}}(x_{k_0}) - \text{Var}(x_{k_0})| \\ &\geq M_{\max} - \frac{M_{\max}}{4} - \frac{M_{\max}}{4} = \frac{M_{\max}}{2}. \end{aligned}$$

Thus, by taking  $\epsilon_0 = \frac{M_{\max}}{2}$ , there holds  $P_2 = 0$ .

Clearly, we have

$$P(\mathcal{A}_0 \neq \widehat{\mathcal{A}}_0) \leq P(\mathcal{E}_0^c) \leq 2p \exp\left(-\frac{nM_{\max}^2}{32C_{\mathcal{X}}^4}\right), \quad (5)$$

by taking  $\epsilon_0 = \frac{M_{\max}}{2}$ . This completes the proof.  $\blacksquare$

**Lemma S1.** *Suppose that Assumptions 1–3 in the main text are satisfied. Given the events  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}\}$  and  $\mathcal{J}$  and assume  $n\lambda \rightarrow \infty$ . Then, with probability at least  $1 - \delta_n$ , for any  $j \in \mathcal{V} \setminus \{\mathcal{S}_t\}$ , there holds*

$$\|\widehat{f}_j - f_{j, \mathcal{S}_t}^*\|_{\infty} \leq \frac{\kappa_1 C_{j0}}{\lambda \sqrt{n}} \log \frac{2}{\delta_n} + \kappa_1 \lambda^{r-1/2} \|L_{K,t}^{-r} f_{j, \mathcal{S}_t}^*\|_2,$$

where  $C_{j0} = 2\kappa_1 \max\left\{C_{\mathcal{X}} + 2\kappa_1 R, \sqrt{2(2\kappa_1^2 R^2 + \sigma_j^2)}\right\}$ .

**Proof of Lemma S1.** To begin with, we define the sampling operator  $S_{\mathbf{x}_{\mathcal{S}_t}} : \mathcal{H}_K \rightarrow \mathcal{R}^n$  associated with some copies of  $\mathbf{x}_{\mathcal{S}_t} \in \mathcal{X}_t$  as

$$S_{\mathbf{x}_{\mathcal{S}_t}}(f) = (f(\mathbf{x}_{1\mathcal{S}_t}), \dots, f(\mathbf{x}_{n\mathcal{S}_t}))^T,$$

and the adjoint of the sample operator as  $S_{\mathbf{x}_{\mathcal{S}_t}}^T : \mathcal{R}^n \rightarrow \mathcal{H}_K$  as

$$S_{\mathbf{x}_{\mathcal{S}_t}}^T \mathbf{c} = \sum_{i=1}^n c_i K_{\mathbf{x}_i \mathcal{S}_t},$$

where  $\mathbf{c} = (c_1, \dots, c_n)^T \in \mathcal{R}^n$ . Note that given the events  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}\}$  and  $\mathcal{J}$ , we have  $\mathcal{S}_t = \widehat{\mathcal{S}}_t$ , and  $\|\widehat{f}_j\|_K \leq R$ , for any  $j \in \mathcal{V} \setminus \{\widehat{\mathcal{S}}_t\}$ . Clearly, the solution of (3.2) in Section 3.1 of the main text can be written as

$$\widehat{f}_j = \left( \frac{1}{n} S_{\mathbf{x}_{\mathcal{S}_t}}^T S_{\mathbf{x}_{\mathcal{S}_t}} + \lambda \mathbf{I}_n \right)^{-1} \frac{1}{n} S_{\mathbf{x}_{\mathcal{S}_t}}^T \mathbf{x}_j,$$

where  $\mathbf{x}_j = (x_{1j}, \dots, x_{nj})^T$  and  $\mathbf{I}_n \in \mathcal{R}^{n \times n}$  denotes the identity matrix.

Moreover, we define an immediate function  $f_{\lambda,j}$  as

$$f_{\lambda,j} = \operatorname{argmin}_{f_j \in \mathcal{H}_K} E[x_j - f_j(\mathbf{x}_{\mathcal{S}_t})]^2 + \lambda \|f_j\|_K^2. \quad (6)$$

Note that solving (6) equals solving the following problem that

$$f_{\lambda,j} = \operatorname{argmin}_{f_j \in \mathcal{H}_K} \|f_{j,\mathcal{S}_t}^* - f_j\|_{\mathcal{L}^2(\mathcal{X}_{\mathcal{S}_t}, \rho_{\mathbf{x}_{\mathcal{S}_t}})}^2 + \lambda \|f_j\|_K^2, \quad (7)$$

by the fact that each node  $x_j$  is centered with mean zero and  $Ef(\mathbf{x}) = 0$  for all  $f \in \mathcal{H}_K$ . Thus, the solution of (6) can be derived as

$$f_{\lambda,j} = (L_{K,t} + \lambda I)^{-1} L_{K,t} f_{j,\mathcal{S}_t}^*,$$

where the integral operator  $L_{K,t}$  is defined in Section 4 of the main text.

Simple algebra yields that

$$\begin{aligned}\widehat{f}_j - f_{\lambda,j} &= \left( \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} + \lambda \mathbf{I}_n \right)^{-1} \left( \frac{1}{n} S_{\mathbf{x}_{S_t}}^T \mathbf{x}_j - \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} f_{\lambda,j} - \lambda f_{\lambda,j} \right) \\ &= \left( \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} + \lambda \mathbf{I}_n \right)^{-1} \left( \frac{1}{n} \sum_{i=1}^n (x_{ij} - f_{\lambda,j}(\mathbf{x}_{iS_t})) K_{\mathbf{x}_{iS_t}} - L_{K,t}(f_{j,S_t}^* - f_{\lambda,j}) \right).\end{aligned}$$

Thus, we have

$$\|\widehat{f}_j - f_{\lambda,j}\|_K \leq \frac{1}{\lambda} \left\| \frac{1}{n} \sum_{i=1}^n (x_{ij} - f_{\lambda,j}(\mathbf{x}_{iS_t})) K_{\mathbf{x}_{iS_t}} - L_{K,t}(f_{j,S_t}^* - f_{\lambda,j}) \right\|_K.$$

For notation simplicity, we denote  $\xi_i = (x_{ij} - f_{\lambda,j}(\mathbf{x}_{iS_t})) K_{\mathbf{x}_{iS_t}}$ , which satisfies

$$\begin{aligned}E[\xi_i] &= L_{K,t}(f_{j,S_t}^* - f_{\lambda,j}), \quad \|\xi_i\|_K \leq \frac{\kappa_1}{2} (C_{\mathcal{X}} + 2\|f_{\lambda,j}\|_{\infty}) \\ \text{and } E[\|\xi_i\|_K^2] &\leq \kappa_1^2 \int (x_j - f_{\lambda,j}(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t} \cup \{j\}},\end{aligned}$$

and then by Lemma 2 of Smale and Zhou (2007), for any  $\delta_n \in (0, 1)$ , with probability at least  $1 - \delta_n$ , we have

$$\|\widehat{f}_j - f_{\lambda,j}\|_K \leq \frac{\kappa_1 (C_{\mathcal{X}} + 2\|f_{\lambda,j}\|_{\infty}) \log(2/\delta_n)}{\lambda n} + \frac{\kappa_1}{\lambda} \sqrt{\frac{2 \int (x_j - f_{\lambda,j}(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t} \cup \{j\}} \log(2/\delta_n)}{n}}.$$

It is clear that by plugging  $f_j = 0$  into (7), we have  $\|f_{\lambda,j}\|_K \leq \frac{\kappa_1 \|f_{j,S_t}^*\|_K}{\lambda^{1/2}}$ , and thus we have  $\|f_{\lambda,j}\|_{\infty} \leq \kappa_1 \|f_{\lambda,j}\|_K < \frac{\kappa_1^2 \|f_{j,S_t}^*\|_K}{\lambda^{1/2}}$ . To bound  $\int (x_j - f_{\lambda,j}(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t} \cup \{j\}}$ , simple calculation yields that for any  $f \in \mathcal{H}_K$ ,

$$\int (x_j - f(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t} \cup \{j\}} - \int (x_j - f_{j,S_t}^*(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t} \cup \{j\}} = \|f - f_{j,S_t}^*\|_{\mathcal{L}^2(\mathcal{X}_{S_t}, \rho_{\mathbf{x}_{S_t}})}^2. \quad (8)$$

By taking  $f = 0$ , there holds

$$\int (x_j - f_{j,S_t}^*(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t \cup \{j\}}} + \|0 - f_{j,S_t}^*\|_{\mathcal{L}^2(\mathcal{X}_{S_t}, \rho_{\mathbf{x}_{S_t}})}^2 = E[f_j^*(\mathbf{x}_{\text{pa}_j}) + n_j]^2 \leq \kappa_1^2 \|f_{j,S_t}^*\|_K^2 + \sigma_j^2,$$

where the last equality follows from the generating scheme of Model 1 in the main text and the last inequality follows from Assumption 3 in the main text. Moreover, we notice that from (7) and by the definition of  $f_{\lambda,j}$ , there holds

$$\|f_{j,S_t}^* - f_{\lambda,j}\|_{\mathcal{L}^2(\mathcal{X}_{S_t}, \rho_{\mathbf{x}_{S_t}})}^2 + \lambda \|f_{\lambda,j}\|_K^2 \leq \|f_{j,S_t}^* - 0\|_{\mathcal{L}^2(\mathcal{X}_{S_t}, \rho_{\mathbf{x}_{S_t}})}^2 + \lambda \|0\|_K^2 \leq \kappa_1^2 \|f_{j,S_t}^*\|_K^2,$$

and then, by plugging  $f = f_{\lambda,j}$  into (8), we have

$$\int (x_j - f_{\lambda,j}(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t \cup \{j\}}} = \int (x_j - f_{j,S_t}^*(\mathbf{x}_{S_t}))^2 d\rho_{\mathbf{x}_{S_t \cup \{j\}}} + \|f_{\lambda,j} - f_{j,S_t}^*\|_2^2 \leq 2\kappa_1^2 \|f_{j,S_t}^*\|_K^2 + \sigma_j^2,$$

Therefore, with probability  $1 - \delta_n$ , we have

$$\begin{aligned} \|\hat{f}_j - f_{\lambda,j}\|_K &\leq \frac{\kappa_1(C_{\mathcal{X}} + 2\kappa_1^2 \|f_{j,S_t}^*\|_K / \lambda^{1/2}) \log(2/\delta_n)}{\lambda n} + \frac{\kappa_1}{\lambda} \sqrt{\frac{2(2\kappa_1^2 \|f_{j,S_t}^*\|_K^2 + \sigma_j^2) \log(2/\delta_n)}{n}} \\ &\leq \frac{C_{j0} \log(2/\delta_n)}{\lambda \sqrt{n}}, \end{aligned} \quad (9)$$

where  $C_{j0} = 2\kappa_1 \max \left\{ C_{\mathcal{X}} + 2\kappa_1 \|f_{j,S_t}^*\|_K, \sqrt{2(2\kappa_1^2 \|f_{j,S_t}^*\|_K^2 + \sigma_j^2)} \right\}$  and the last inequality follows from the fact that  $n\lambda \rightarrow \infty$ .

Then, we turn to bound  $\|f_{\lambda,j} - f_{j,S_t}^*\|_K$  following similar treatments as in Smale and Zhou (2005). Specifically, for the integral operator  $L_{K,t}$  defined in Section 4 of the main text with normalized eigenpairs  $\{(\mu_k, \phi_k)\}_{k=1}^{\infty}$ , we have

$$L_{K,t}^{1/2} \phi_i = \sum_{j \geq 1} \mu_j^{1/2} \langle \phi_i, \phi_j \rangle_2 \phi_j = \mu_i^{1/2} \phi_i \in \mathcal{H}_K,$$

and

$$\|\mu_i^{1/2} \phi_i\|_K = \left( \sum_{j \geq 1} \frac{\langle \mu_i^{1/2} \phi_i, \phi_j \rangle_2^2}{\mu_j} \right)^{1/2} = \langle \phi_i, \phi_i \rangle_2 = 1.$$

Thus by Assumption 2 of the main text, there exists some function  $h_{j,t} = \sum_{i \geq 1} \langle h_{j,t}, \phi_i \rangle_2 \phi_i \in \mathcal{L}^2(\mathcal{X}_{\mathcal{S}_t}, \rho_{\mathbf{x}_{\mathcal{S}_t}})$  such that  $f_{j,\mathcal{S}_t}^* = L_{K,t}^r h_{j,t} = \sum_{i \geq 1} \mu_i^r \langle h_{j,t}, \phi_i \rangle_2 \phi_i \in \mathcal{H}_K$ .

Therefore, we have

$$f_{\lambda,j} - f_{j,\mathcal{S}_t}^* = (L_{K,t} + \lambda I)^{-1} (-\lambda f_{j,\mathcal{S}_t}^*) = - \sum_{i \geq 1} \frac{\lambda}{\lambda + \mu_i} \mu_i^r \langle h_{j,t}, \phi_i \rangle_2 \phi_i,$$

and then, the  $K$ -norm of  $f_{\lambda,j} - f_{j,\mathcal{S}_t}^*$  can be bounded as

$$\begin{aligned} \|f_{\lambda,j} - f_{j,\mathcal{S}_t}^*\|_K^2 &= \sum_{i \geq 1} \left( \frac{\lambda}{\lambda + \mu_i} \mu_i^{r-1/2} \langle h_{j,t}, \phi_i \rangle_2 \right)^2 \\ &= \lambda^{2r-1} \sum_{i \geq 1} \left( \frac{\lambda}{\lambda + \mu_i} \right)^{3-2r} \left( \frac{\mu_i}{\lambda + \mu_i} \right)^{2r-1} \langle h_{j,t}, \phi_i \rangle_2^2 \\ &\leq \lambda^{2r-1} \sum_{i \geq 1} \langle h_{j,t}, \phi_i \rangle_2^2 = \lambda^{2r-1} \|h_{j,t}\|_2^2 = \lambda^{2r-1} \|L_{K,t}^{-r} f_{j,\mathcal{S}_t}^*\|_2^2. \end{aligned} \quad (10)$$

Combining (9) and (10), under the events  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}\}$  and  $\mathcal{J}$ , with probability at least  $1 - \delta_n$ , we have

$$\begin{aligned} \|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_K &\leq \|\widehat{f}_j - f_{\lambda,j}\|_K + \|f_{\lambda,j} - f_{j,\mathcal{S}_t}^*\|_K \\ &\leq \frac{C_{j0}}{\lambda \sqrt{n}} \log \frac{2}{\delta_n} + \lambda^{r-1/2} \|L_{K,t}^{-r} f_{j,\mathcal{S}_t}^*\|_2. \end{aligned}$$

Moreover, we notice that  $\|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_\infty \leq \kappa_1 \|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_K$  by the reproducing property and the requirement that  $\|f_{j,\mathcal{S}_t}^*\|_K^2 \leq R/2$  in Section 4 of the main text. This completes the proof.  $\blacksquare$

**Proof of Theorem 4.** Given the event  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}\}$ , we have  $\mathcal{S}_t = \widehat{\mathcal{S}}_t$ . Then, for



any  $j \in \mathcal{V} \setminus \{\mathcal{S}_t\}$ , there holds

$$\begin{aligned}
& \left| \widehat{E} \widehat{\text{Var}}(x_j | \mathbf{x}_{\mathcal{S}_t}) - E \text{Var}(x_j | \mathbf{x}_{\mathcal{S}_t}) \right| \\
&= \left| \frac{1}{n} \sum_{i=1}^n (x_{ij})^2 - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{i\mathcal{S}_t}))^2 - E[x_j^2] + E[E[x_j | \mathbf{x}_{\mathcal{S}_t}]^2] \right| \\
&\leq \left| E[x_j^2] - \frac{1}{n} \sum_{i=1}^n (x_{ij})^2 \right| + \left| E[E[x_j | \mathbf{x}_{\mathcal{S}_t}]^2] - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{i\mathcal{S}_t}))^2 \right|. \tag{11}
\end{aligned}$$

To bound the first term of (11), we notice that each  $x_j$  is required to be centered with mean zero in Section 2 of the main text, which implies that zero belong to the support  $\mathcal{X}$ , and then  $x_{ij}^2$  are bounded by  $\frac{C_{\mathcal{X}}^2}{4}$  from the definition of  $C_{\mathcal{X}}$ , which denotes the diameter of the support  $\mathcal{X}$ . Then by the Hoeffding's inequality, for any  $\frac{\zeta}{2} > 0$ , there holds

$$P\left(\left|E[x_j^2] - \frac{1}{n} \sum_{i=1}^n (x_{ij})^2\right| > \frac{\zeta}{2}\right) \leq 2 \exp\left(-\frac{8n\zeta^2}{C_{\mathcal{X}}^4}\right). \tag{12}$$

Next, the second term of (11) can be decomposed as

$$\begin{aligned}
& \left| E[E[x_j | \mathbf{x}_{\mathcal{S}_t}]^2] - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{i\mathcal{S}_t}))^2 \right| \\
&\leq \left| E[E[x_j | \mathbf{x}_{\mathcal{S}_t}]^2 - \widehat{f}_j^2(\mathbf{x}_{\mathcal{S}_t})] \right| + \left| E[\widehat{f}_j^2(\mathbf{x}_{\mathcal{S}_t})] - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{i\mathcal{S}_t}))^2 \right| \\
&= \Delta_1 + \Delta_2,
\end{aligned}$$

and thus it suffices to bound  $\Delta_1$  and  $\Delta_2$  sequentially under the events  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}\}$  and  $\mathcal{J}$ .

To bound  $\Delta_1$ , we notice that

$$\begin{aligned}
\Delta_1 &= \left| E[f_{j,S_t}^*(\mathbf{x}_{S_t})(f_{j,S_t}^*(\mathbf{x}_{S_t}) - \widehat{f}_j(\mathbf{x}_{S_t}))] + E[\widehat{f}_j(\mathbf{x}_{S_t})(f_{j,S_t}^*(\mathbf{x}_{S_t}) - \widehat{f}_j(\mathbf{x}_{S_t}))] \right| \\
&\leq \|f_{j,S_t}^* - \widehat{f}_j\|_\infty \left| \int |f_{j,S_t}^*(\mathbf{x}_{S_t})| d\rho_{\mathbf{x}_{S_t}} + \int |\widehat{f}_j(\mathbf{x}_{S_t})| d\rho_{\mathbf{x}_{S_t}} \right| \\
&\leq 2\kappa_1 \max\{\|\widehat{f}_j\|_K, \|f_{j,S_t}^*\|_K\} \|f_{j,S_t}^* - \widehat{f}_j\|_\infty \leq 2\kappa_1 R \|f_{j,S_t}^* - \widehat{f}_j\|_\infty,
\end{aligned}$$

where the last inequality follows from the reproducing property of RKHS, the requirement that  $\|f_{j,S_t}^*\|_K \leq R/2$  in Section 4 of the main text and under the event  $\mathcal{J}$  that  $\|\widehat{f}_j\|_K \leq R$ . Then, by Lemma S1, with probability at least  $1 - \delta_n/2$ , we have

$$\Delta_1 \leq 2\kappa_1^2 R \left( \frac{C_{j0}}{\lambda\sqrt{n}} \log \frac{4}{\delta_n} + \lambda^{r-1/2} \|L_{K,t}^{-r} f_{j,S_t}^*\|_2 \right). \quad (13)$$

To bound  $\Delta_2$ , we notice that

$$\begin{aligned}
\Delta_2 &= \left| \int \widehat{f}_j(\mathbf{x}_{S_t}) \langle \widehat{f}_j, K_{\mathbf{x}_{S_t}} \rangle_K d\rho_{\mathbf{x}_{S_t}} - \frac{1}{n} \sum_{i=1}^n \widehat{f}_j(\mathbf{x}_{iS_t}) \langle \widehat{f}_j, K_{\mathbf{x}_{iS_t}} \rangle_K \right| \\
&= \left| \left\langle \widehat{f}_j, \int \widehat{f}_j(\mathbf{x}_{S_t}) K_{\mathbf{x}_{S_t}} d\rho_{\mathbf{x}_{S_t}} \right\rangle_K - \frac{1}{n} \left\langle \widehat{f}_j, S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} \widehat{f}_j \right\rangle_K \right| \\
&= \left| \left\langle \widehat{f}_j, \int \widehat{f}_j(\mathbf{x}_{S_t}) K_{\mathbf{x}_{S_t}} d\rho_{\mathbf{x}_{S_t}} - \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} \widehat{f}_j \right\rangle_K \right| \\
&= \left| \left\langle \widehat{f}_j, \left( L_{K,t} - \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} \right) \widehat{f}_j \right\rangle_K \right| \leq \|\widehat{f}_j\|_K \left\| L_{K,t} - \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} \right\|_{HS},
\end{aligned}$$

where  $S_{\mathbf{x}_{S_t}}^T$  and  $S_{\mathbf{x}_{S_t}}$  denote the sampling operators defined in Lemma S1, and  $\|\cdot\|_{HS}$  denotes the norm endowed with a Hilbert space  $HS(K)$  containing all the Hilbert-Schmidt operators on  $\mathcal{H}_K$  and satisfying  $\|T\|_K \leq \|T\|_{HS}$  for any  $T \in HS(K)$ . Note that under the event  $\mathcal{J}$ , we have  $\|\widehat{f}_j\|_K \leq R$ . Moreover, by Lemma 18 of Rosasco et al. (2013), with probability at least  $1 - \delta_n/2$ ,

we have

$$\left\| L_{K,t} - \frac{1}{n} S_{\mathbf{x}_{S_t}}^T S_{\mathbf{x}_{S_t}} \right\|_{HS} \leq \frac{2\sqrt{2}\kappa_1^2}{\sqrt{n}} \log \frac{4}{\delta_n}.$$

Clearly, with probability at least  $1 - \delta_n/2$ , we have  $\Delta_2 \leq \frac{2R\sqrt{2}\kappa_1^2}{\sqrt{n}} \log \frac{4}{\delta_n}$ .

Combining the upper bounds of  $\Delta_1$  and  $\Delta_2$ , with probability at least  $1 - \delta_n$ , there holds

$$\begin{aligned} & \left| E[E[x_j|\mathbf{x}_{S_t}]^2] - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{iS_t}))^2 \right| \\ & \leq 2\kappa_1^2 R \left( \frac{C_{j0}}{\lambda\sqrt{n}} \log \frac{4}{\delta_n} + \lambda^{r-1/2} \|L_{K,t}^{-r} f_{j,S_t}^*\|_2 \right) + \frac{2R\sqrt{2}\kappa_1^2}{\sqrt{n}} \log \frac{4}{\delta_n} \\ & \leq 2\kappa_1^2 R \left( \frac{C_{j0} + \sqrt{2}}{\lambda\sqrt{n}} \log \frac{4}{\delta_n} + \lambda^{r-1/2} \|L_{K,t}^{-r} f_{j,S_t}^*\|_2 \right). \end{aligned}$$

Then, by taking  $\lambda = n^{-\frac{1}{2r+1}}$ , for any  $\delta_n \in (0, 1)$ , with probability at least  $1 - \delta_n$  there holds

$$\left| E[E[x_j|\mathbf{x}_{S_t}]^2] - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{iS_t}))^2 \right| \leq C_{jt} n^{-\frac{2r-1}{2(2r+1)}} \log \frac{4}{\delta_n},$$

where  $C_{jt} = 6\kappa_1^2 R \max \{C_{j0}, \sqrt{2}, \|L_{K,t}^{-r} f_{j,S_t}^*\|_2\}$ . Correspondingly, for any  $\zeta > 0$ , we have

$$\begin{aligned} & P\left( \left| E[E[x_j|\mathbf{x}_{S_t}]^2] - \frac{1}{n} \sum_{i=1}^n (\widehat{f}_j(\mathbf{x}_{iS_t}))^2 \right| > \frac{\zeta}{2} \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J} \right) \\ & \leq 4 \exp\left( -\frac{\zeta}{2C_{jt}} n^{\frac{2r-1}{2(2r+1)}} \right). \end{aligned} \tag{14}$$

Combining (12) and (14), for any  $\zeta > 0$ , there holds

$$\begin{aligned} & P\left( \left| E\text{Var}(x_j|\mathbf{x}_{S_t}) - \widehat{E}\widehat{\text{Var}}(x_j|\mathbf{x}_{S_t}) \right| > \zeta \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J} \right) \\ & \leq 2 \exp\left( -\frac{8n\zeta^2}{C_{jt}^4} \right) + 4 \exp\left( -\frac{\zeta}{2C_{jt}} n^{\frac{2r-1}{2(2r+1)}} \right). \end{aligned} \tag{15}$$

This completes the proof of the first part of Theorem 4.

Next, we define the following event

$$\mathcal{E}_{1t} = \left\{ \max_{j \in \mathcal{V} \setminus \{\mathcal{S}_t\}} |E\text{Var}(x_j | \mathbf{x}_{\mathcal{S}_t}) - \widehat{E}\widehat{\text{Var}}(x_j | \mathbf{x}_{\mathcal{S}_t})| \leq \frac{M_{\max}}{4} \right\},$$

and use the notation  $\mathcal{E}_{1t}^c$  to denote its complementary. Then, by (15), we have

$$\begin{aligned} & P(\mathcal{E}_{1t}^c \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \leq 2(p - |\mathcal{S}_t|) \exp\left(-\frac{nM_{\max}^2}{2C_{\mathcal{X}}^4}\right) + 4(p - |\mathcal{S}_t|) \exp\left(-\frac{M_{\max} n^{\frac{2r-1}{2(2r+1)}}}{8C_{jt}}\right). \end{aligned} \quad (16)$$

Note that

$$\begin{aligned} & P(\widehat{\mathcal{A}}_t \neq \mathcal{A}_t \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \leq P(\widehat{\mathcal{A}}_t \neq \mathcal{A}_t, \mathcal{E}_{1t} \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) + P(\mathcal{E}_{1t}^c \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \leq P(\exists j \in \mathcal{A}_t \text{ such that } |\widehat{E}\widehat{\text{Var}}(x_j | \mathbf{x}_{\mathcal{S}_t}) - \widehat{\sigma}_{\min}^{(t)}| \geq \epsilon_t, \mathcal{E}_{1t} \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \quad + P(\exists j \in \mathcal{V} \setminus \{\mathcal{S}_t \cup \mathcal{A}_t\} \text{ such that } |\widehat{E}\widehat{\text{Var}}(x_j | \mathbf{x}_{\mathcal{S}_t}) - \widehat{\sigma}_{\min}^{(t)}| < \epsilon_t, \mathcal{E}_{1t} \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \quad + P(\mathcal{E}_{1t}^c \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & = P_3 + P_4 + P(\mathcal{E}_{1t}^c \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}). \end{aligned} \quad (17)$$

Note that following the similar treatments as that of  $P_1$  and  $P_2$  in the proof of Theorem 3 in the main text and by taking  $\epsilon_t = \frac{M_{\max}}{2}$ , we have  $P_3 = 0$  and  $P_4 = 0$ . Finally, the bound (17) reduces to

$$\begin{aligned} & P(\widehat{\mathcal{A}}_t \neq \mathcal{A}_t \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \leq P(\mathcal{E}_{1t}^c \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) \\ & \leq 2(p - |\mathcal{S}_t|) \exp\left(-\frac{nM_{\max}^2}{2C_{\mathcal{X}}^4}\right) + 4(p - |\mathcal{S}_t|) \exp\left(-\frac{M_{\max} n^{\frac{2r-1}{2(2r+1)}}}{8C_{jt}}\right). \end{aligned}$$

This completes the proof. ■

**Proof of Lemma 1.** Given the event  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_t = \mathcal{A}_t\}$ ,  $t \geq 1$ , we have  $\mathcal{S}_t = \widehat{\mathcal{S}}_t$ . At first, for some positive constant  $C_3$ , we define the following event

$$\mathcal{E}_{2t} = \left\{ \max_{j \in \mathcal{A}_t, k \in \mathcal{S}_t} \left| \|\widehat{g}_{jk}\|_n^2 - \|g_{jk}^*\|_2^2 \right| \leq C_3 n^{-\frac{2r-1}{2(2r+1)}} \log(4|\mathcal{S}_t| \max\{n, |\mathcal{S}_t|\}) \right\},$$

and use the notation  $\mathcal{E}_{2t}^c$  to denote its complementary.

We notice that

$$\begin{aligned} & P\left(\{\mathcal{E}_j \neq \widehat{\mathcal{E}}_j : j \in \widehat{\mathcal{A}}_t\} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right) \\ & \leq P\left(\{\mathcal{E}_j \neq \widehat{\mathcal{E}}_j : j \in \widehat{\mathcal{A}}_t\}, \mathcal{E}_{2t} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right) \\ & \quad + P(\mathcal{E}_{2t}^c \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}). \end{aligned} \tag{18}$$

Note that by the definition that  $\widehat{\mathcal{E}}_j = \{k \rightarrow j, \|\widehat{g}_{jk}\|_n^2 > \nu_n^{(t)}\}$ , for any  $k \in \widehat{\mathcal{S}}_t$  and by Assumption 4 of the main text, for the first term of (18), there holds

$$\begin{aligned} & P\left(\{\mathcal{E}_j \neq \widehat{\mathcal{E}}_j : j \in \widehat{\mathcal{A}}_t\}, \mathcal{E}_{2t} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right) \\ & \leq P\left(\exists k \in \text{pa}_j \text{ such that } \|\widehat{g}_{jk}\|_n^2 \leq \nu_n^{(t)}, \mathcal{E}_{2t} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right) \\ & \quad + P\left(\exists k \in \widehat{\text{pa}}_j \text{ such that } \|g_{jk}^*\|_2^2 = 0, \mathcal{E}_{2t} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right) \\ & = P_5 + P_6, \end{aligned} \tag{19}$$

where  $\widehat{\text{pa}}_j = \{k : k \rightarrow j \in \widehat{\mathcal{E}}_j\}$ .

For the bound of  $P_5$ , by taking  $\nu_n^{(t)} = \frac{C_2}{2} n^{-\frac{2r-1}{2(2r+1)}} \left( \log(4|\mathcal{S}_t| \max\{n, |\mathcal{S}_t|\}) \right)^\beta$  and Assumption 3 in the main text, we have

$$\left| \|\widehat{g}_{jk}\|_n^2 - \|g_{jk}^*\|_2^2 \right| \geq \|g_{jk}^*\|_2^2 - \|\widehat{g}_{jk}\|_n^2 > 2\nu_n^{(t)} - \nu_n^{(t)} = \nu_n^{(t)},$$

which contradicts with  $\mathcal{E}_{2t}$ . Precisely, under  $\mathcal{E}_{2t}$ , we have

$$\max_{j \in \mathcal{A}_t, k \in \mathcal{S}_t} \left| \|\widehat{g}_{jk}\|_n^2 - \|g_{jk}^*\|_2^2 \right| \leq C_3 n^{-\frac{2r-1}{2(2r+1)}} \log(4|\mathcal{S}_t| \max\{n, |\mathcal{S}_t|\}),$$

and then the different rates of convergence lead to the contradiction. Thus, when  $n$  is sufficiently large,  $P_5 = 0$ . To bound  $P_6$ , it is obvious that  $\|\widehat{g}_{jk}\|_n^2 - \|g_{jk}^*\|_2^2 > \nu_n^{(t)}$ , which contradicts with  $\mathcal{E}_{2t}$  again, which yields that  $P_6 = 0$ .

Now, we turn to bound  $P(\mathcal{E}_{2t}^c \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J})$ . At first, we define the sample operators for gradients  $\widehat{D}_{t,l} : \mathcal{H}_K \rightarrow \mathcal{R}^n$  as  $(\widehat{D}_{t,l}f)_i = \langle f, \partial_l K_{\mathbf{x}_{i\mathcal{S}_t}} \rangle_K$ , their adjoint operators  $\widehat{D}_{t,l}^* : \mathcal{R}^n \rightarrow \mathcal{H}_K$  as  $\widehat{D}_{t,l}^* \mathbf{c} = \frac{1}{n} \sum_{i=1}^n \partial_l K_{\mathbf{x}_{i\mathcal{S}_t}} c_i$ , and define the integral operators for gradients  $D_{t,l} : \mathcal{H}_K \rightarrow \mathcal{L}^2(\mathcal{X}_{\mathcal{S}_t}, \rho_{\mathbf{x}_{\mathcal{S}_t}})$  as  $D_{t,l}f = \langle f, \partial_l K_{\mathbf{x}_{\mathcal{S}_t}} \rangle_K$ ,  $D_{t,l}^* : \mathcal{L}^2(\mathcal{X}_{\mathcal{S}_t}, \rho_{\mathbf{x}_{\mathcal{S}_t}}) \rightarrow \mathcal{H}_K$  as  $D_{t,l}^* f = \int \partial_l K_{\mathbf{x}_{\mathcal{S}_t}} f(\mathbf{x}_{\mathcal{S}_t}) d\rho_{\mathbf{x}_{\mathcal{S}_t}}$ . Then, we have

$$D_{t,l}^* D_{t,l} f_{j,\mathcal{S}_t}^* = \int \partial_l K_{\mathbf{x}_{\mathcal{S}_t}} g_{jl}^*(\mathbf{x}_{\mathcal{S}_t}) d\rho_{\mathbf{x}_{\mathcal{S}_t}} \quad \text{and} \quad \widehat{D}_{t,l}^* \widehat{D}_{t,l} f_{j,\mathcal{S}_t}^* = \frac{1}{n} \sum_{i=1}^n \partial_l K_{\mathbf{x}_{i\mathcal{S}_t}} g_{jl}^*(\mathbf{x}_{i\mathcal{S}_t}).$$

Note that  $D_{t,l}^* D_{t,l}$  and  $\widehat{D}_{t,l}^* \widehat{D}_{t,l}$  are the Hilbert-Schmidt operators belonging to a Hilbert space endowed with norm  $\|\cdot\|_{HS}$ .

Moreover, we notice that for any  $j \in \mathcal{A}_t$  and  $k \in \mathcal{S}_t$

$$\begin{aligned} \left| \|\widehat{g}_{jk}\|_n^2 - \|g_{jk}^*\|_2^2 \right| &= \left| \frac{1}{n} \sum_{i=1}^n (\widehat{g}_{jk}(\mathbf{x}_{i\mathcal{S}_t}))^2 - \int (g_{jk}^*(\mathbf{x}_{\mathcal{S}_t}))^2 d\rho_{\mathbf{x}_{\mathcal{S}_t}} \right| \\ &= \left| \langle \widehat{f}_j, \frac{1}{n} \sum_{i=1}^n \widehat{g}_{jk}(\mathbf{x}_{i\mathcal{S}_t}) \partial_k K_{\mathbf{x}_{i\mathcal{S}_t}} \rangle_K - \langle f_{j,\mathcal{S}_t}^*, \int g_{jk}^*(\mathbf{x}_{\mathcal{S}_t}) \partial_k K_{\mathbf{x}_{\mathcal{S}_t}} d\rho_{\mathbf{x}_{\mathcal{S}_t}} \rangle_K \right| \\ &= \left| \langle \widehat{f}_j - f_{j,\mathcal{S}_t}^*, \widehat{D}_{t,k}^* \widehat{D}_{t,k} (\widehat{f}_j - f_{j,\mathcal{S}_t}^*) \rangle_K + \langle \widehat{D}_{t,k}^* \widehat{D}_{t,k} f_{j,\mathcal{S}_t}^*, \widehat{f}_j - f_{j,\mathcal{S}_t}^* \rangle_K \right. \\ &\quad \left. + \langle f_{j,\mathcal{S}_t}^*, \widehat{D}_{t,k}^* \widehat{D}_{t,k} (\widehat{f}_j - f_{j,\mathcal{S}_t}^*) \rangle_K + \langle f_{j,\mathcal{S}_t}^*, (\widehat{D}_{t,k}^* \widehat{D}_{t,k} - D_{t,k}^* D_{t,k}) f_{j,\mathcal{S}_t}^* \rangle_K \right| \\ &\leq \|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_K^2 \|\widehat{D}_{t,k}^* \widehat{D}_{t,k}\|_{HS} + 2\|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_K \|f_{j,\mathcal{S}_t}^*\|_K \|\widehat{D}_{t,k}^* \widehat{D}_{t,k}\|_{HS} \\ &\quad + \|\widehat{D}_{t,k}^* \widehat{D}_{t,k} - D_{t,k}^* D_{t,k}\|_{HS} \|f_{j,\mathcal{S}_t}^*\|_K^2. \end{aligned}$$

By Assumption 3 in the main text, direct calculation yields that

$$\max_{k \in \mathcal{S}_t} \|\widehat{D}_{t,k}^* \widehat{D}_{t,k}\|_{HS} = \max_{k \in \mathcal{S}_t} \|\partial_k K_{\mathbf{x}_{\mathcal{S}_t}}\|_K^2 \leq \kappa_2^2.$$

Moreover, by Lemma 18 of Rosasco et al. (2013), for any  $\delta_n \in (0, 1)$ , with probability at least  $1 - \delta_n/2$ , we have

$$\max_{k \in \mathcal{S}_t} \|\widehat{D}_{t,k}^* \widehat{D}_{t,k} - D_{t,k}^* D_{t,k}\|_{HS} \leq \frac{2\sqrt{2}\kappa_2^2}{\sqrt{n}} \log \frac{4|\mathcal{S}_t|}{\delta_n},$$

given  $\{\widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_t = \mathcal{A}_t\}$ . Thus, by taking  $\delta_n = (\max\{n, |\mathcal{S}_t|\})^{-1}$ , with probability at least  $1 - \delta_n/2$ , there holds

$$\max_{k \in \mathcal{S}_t} \|\widehat{D}_k^* \widehat{D}_k - D_k^* D_k\|_{HS} \leq \frac{2\sqrt{2}\kappa_2^2}{\sqrt{n}} \log (4|\mathcal{S}_t| \max\{n, |\mathcal{S}_t|\}).$$

When  $\|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_K$  is sufficiently small and by taking  $\lambda = n^{-\frac{1}{2r+1}}$ , with probability at least  $1 - \delta_n$ , we have

$$\begin{aligned} & \max_{j \in \mathcal{A}_t, k \in \mathcal{S}_t} \left| \|\widehat{g}_{jk}\|_n^2 - \|g_{jk}\|_2^2 \right| \\ & \leq \max\{\kappa_2^2, \kappa_2^2 \|f_{j,\mathcal{S}_t}^*\|_K, \|f_{j,\mathcal{S}_t}^*\|_K^2\} \left( 3 \max_{j \in \mathcal{A}_t, k \in \mathcal{S}_t} \|\widehat{f}_j - f_{j,\mathcal{S}_t}^*\|_K + \max_{k \in \mathcal{S}_t} \|\widehat{D}_k^* \widehat{D}_k - D_k^* D_k\|_{HS} \right) \\ & \leq \max\{\kappa_2^2, \kappa_2^2 \|f_{j,\mathcal{S}_t}^*\|_K, \|f_{j,\mathcal{S}_t}^*\|_K^2\} \left( 3 \max_{j \in \mathcal{A}_t, k \in \mathcal{S}_t} \{C_{j0} + \|L_{K,t}^{-r} f_{j,\mathcal{S}_t}^*\|_2\} n^{-\frac{2r-1}{2(2r+1)}} \log \frac{4|\mathcal{S}_t|}{\delta_n} + \frac{2\sqrt{2}\kappa_2^2}{\sqrt{n}} \log \frac{4|\mathcal{S}_t|}{\delta_n} \right) \\ & \leq C_3 n^{-\frac{2r-1}{2(2r+1)}} \log (4|\mathcal{S}_t| \max\{n, |\mathcal{S}_t|\}), \end{aligned}$$

where  $C_3 = 3 \max\{\kappa_2^2, \kappa_2^2 \|f_{j,\mathcal{S}_t}^*\|_K, \|f_{j,\mathcal{S}_t}^*\|_K^2\} \max_{j \in \mathcal{A}_t, k \in \mathcal{S}_t} \{3C_{j0}, 3\|L_{K,t}^{-r} f_{j,\mathcal{S}_t}^*\|_2, 2\sqrt{2}\kappa_2^2\}$ . Thus,

we have

$$P(\mathcal{E}_{2t}^c | \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_t = \mathcal{A}_t, \mathcal{J}) \leq \frac{1}{\max\{n, |\mathcal{S}_t|\}},$$

Finally, by (18) and (19), we have

$$P\left(\{\mathcal{E}_j = \widehat{\mathcal{E}}_j : j \in \widehat{\mathcal{A}}_t\} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right) \geq 1 - \frac{1}{\max\{n, |\mathcal{S}_t|\}}.$$

This completes the proof. ■

**Proof for Theorem 5.** Note that

$$P(\widehat{\mathcal{G}} \neq \mathcal{G}) = P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}) + P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}^c).$$

For  $P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}^c)$ , we have

$$P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}^c) \leq P(\mathcal{J}^c) \leq Tp \max_{1 \leq t \leq T-1, j \in \mathcal{V} \setminus \{\mathcal{S}_t\}} P(\|\widehat{f}_j\|_K > R). \quad (20)$$

Note that by Theorem 1 and Lemma 3 of Smale and Zhou (2007), for any  $t$  and  $j \in \mathcal{V} \setminus \{\mathcal{S}_t\}$ , we have  $P(\|\widehat{f}_j\|_K > R) \leq \frac{1}{n}$  if the sample size satisfies  $n \geq \left(\frac{C_4}{R} \log 2n\right)^{\frac{2(2r+1)}{2r-1}}$  for some positive constant  $C_4$  and the  $K$ -norm of the target function is upper bounded by  $R/2$  as assumed in Section 4 of the main text, and thus  $P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}^c) \rightarrow 0$  as  $n \rightarrow \infty$ .



For  $P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J})$ , we have

$$\begin{aligned}
P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}) &\leq P\left(\bigcup_{t=0}^{T-1} \{\widehat{\mathcal{A}}_t \neq \mathcal{A}_t\} \cup \{\widehat{\mathcal{E}} \neq \mathcal{E}\}, \mathcal{J}\right) \\
&\leq P(\widehat{\mathcal{A}}_0 \neq \mathcal{A}_0) + \sum_{t=1}^{T-1} P(\widehat{\mathcal{A}}_t \neq \mathcal{A}_t \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) + \\
&\quad P(\widehat{\mathcal{E}} \neq \mathcal{E} \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{T-1} = \mathcal{A}_{T-1}, \mathcal{J}) \\
&\leq P(\widehat{\mathcal{A}}_0 \neq \mathcal{A}_0) + \sum_{t=1}^{T-1} P(\widehat{\mathcal{A}}_t \neq \mathcal{A}_t \mid \widehat{\mathcal{A}}_0 = \mathcal{A}_0, \dots, \widehat{\mathcal{A}}_{t-1} = \mathcal{A}_{t-1}, \mathcal{J}) + \\
&\quad \sum_{t=1}^{T-1} P\left(\{\mathcal{E}_j = \widehat{\mathcal{E}}_j : j \in \widehat{\mathcal{A}}_t\} \mid \mathcal{A}_0 = \widehat{\mathcal{A}}_0, \dots, \mathcal{A}_t = \widehat{\mathcal{A}}_t, \mathcal{J}\right).
\end{aligned}$$

Clearly, combining with Theorem 4 and Lemma 1 in the main text, we have  $P(\widehat{\mathcal{G}} \neq \mathcal{G}, \mathcal{J}) \rightarrow 0$  as  $n \rightarrow \infty$ . This completes the proof. ■

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