

Homogeneity Tests for High-dimensional Mean

Vectors and Covariance Matrices

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Supplementary Material

All Technical proofs are listed in this Supplementary Material.

S1 Proof of Lemma 2.1

Proof. First, we can obtain the following relations of equivalence.

$$\begin{aligned} & \boldsymbol{\mu}_1 = \cdots = \boldsymbol{\mu}_R = \boldsymbol{\mu} \\ \Leftrightarrow & \mathbf{E}(\mathbf{X}|Y = r) = \mathbf{E}(\mathbf{X}), \forall r = 1, 2, \dots, R \\ \Leftrightarrow & \text{var}_Y\{\mathbf{E}(\mathbf{X}|Y)\} = \sum_{r=1}^R p_r \|\mathbf{E}(\mathbf{X}|Y = r) - \mathbf{E}(\mathbf{X})\|^2 = 0 \\ \Leftrightarrow & \sum_{r=1}^R p_r \{\mathbf{E}(\mathbf{X}_1|Y = r) - \mathbf{E}(\mathbf{X}_1)\}^T \{\mathbf{E}(\mathbf{X}_2|Y_2 = r) - \mathbf{E}(\mathbf{X}_2)\} = 0 \\ \Leftrightarrow & \sum_{r=1}^R p_r \mathbf{E}(\mathbf{X}_1^T \mathbf{X}_2) \left\{ \frac{I(Y_1 = r)}{p_r} - 1 \right\} \left\{ \frac{I(Y_2 = r)}{p_r} - 1 \right\} = 0 \end{aligned}$$

$$\begin{aligned} &\Leftrightarrow \mathbb{E}(\mathbf{X}_1^T \mathbf{X}_2) \sum_{r=1}^R p_r \left\{ \frac{I(Y_1 = r)}{p_r} - 1 \right\} \left\{ \frac{I(Y_2 = r)}{p_r} - 1 \right\} = 0 \\ &\Leftrightarrow \mathbb{E}(\mathbf{X}_1^T \mathbf{X}_2) \left\{ \sum_{r=1}^R \frac{I(Y_1 = r)I(Y_2 = r)}{p_r} - 1 \right\} = 0. \end{aligned}$$

Hence, by Definition 1, $\mathcal{U}(\mathbf{X}|Y) = \text{var}_Y\{\mathbb{E}(\mathbf{X}|Y)\} \geq 0$ is true, and the equality holds if and only if $\boldsymbol{\mu}_1 = \cdots = \boldsymbol{\mu}_R$, which completes the proof. \square

S2 Proof of equation (2.2)

Proof. By the definition of $M_{n,p}$, we have

$$\begin{aligned} M_{n,p} &= \sum_{i \neq j} \mathbf{X}_i^T \mathbf{X}_j \left\{ \sum_{r=1}^R \frac{I(Y_i = r)I(Y_j = r)}{(N_r - 1)/(n - 1)} - \sum_{r=1}^R \sum_{s=1}^R I(Y_i = r)I(Y_j = s) \right\} \\ &= \sum_{r=1}^R \frac{(n - 1)}{(N_r - 1)} \sum_{i \neq j} \mathbf{X}_{ri}^T \mathbf{X}_{rj} - \sum_{r=1}^R \sum_{i \neq j} \mathbf{X}_{ri}^T \mathbf{X}_{rj} - \sum_{r \neq s} \sum_{i=1}^{N_r} \sum_{j=1}^{N_s} \mathbf{X}_{ri}^T \mathbf{X}_{sj} \\ &= \sum_{r \neq s} \frac{N_s}{(N_r - 1)} \sum_{i \neq j} \mathbf{X}_{ri}^T \mathbf{X}_{rj} - \sum_{r \neq s} \sum_{i=1}^{N_r} \sum_{j=1}^{N_s} \mathbf{X}_{ri}^T \mathbf{X}_{sj} \\ &= \sum_{r \neq s} N_r N_s \frac{\sum_{i \neq j} \mathbf{X}_{ri}^T \mathbf{X}_{rj}}{N_r(N_r - 1)} - \sum_{r \neq s} N_r N_s \frac{\sum_{i=1}^{N_r} \sum_{j=1}^{N_s} \mathbf{X}_{ri}^T \mathbf{X}_{sj}}{N_r N_s} \\ &= \sum_{r > s} N_r N_s \left\{ \frac{\sum_{i \neq j} \mathbf{X}_{ri}^T \mathbf{X}_{rj}}{N_r(N_r - 1)} + \frac{\sum_{i \neq j} \mathbf{X}_{si}^T \mathbf{X}_{sj}}{N_s(N_s - 1)} - 2 \frac{\sum_{i=1}^{N_r} \sum_{j=1}^{N_s} \mathbf{X}_{ri}^T \mathbf{X}_{sj}}{N_r N_s} \right\}, \end{aligned}$$

which completes the proof. \square

S3 Proof of Theorem 2.1

Proof. Write $g_{ij} = \sum_{r=1}^R I(Y_i = r)I(Y_j = r)/p_r - 1$ and $\hat{g}_{ij} = \sum_{r=1}^R I(Y_i = r)I(Y_j = r)/\hat{p}_r - 1$. Then reconstruct the statistic $M_{n,p}$ as follows:

$$\begin{aligned}
 M_{n,p} &= \sum_{i \neq j} \mathbf{X}_i^T \mathbf{X}_j \hat{g}_{ij} = \sum_{i \neq j} (\boldsymbol{\mu}_i + \boldsymbol{\Gamma}_i \mathbf{Z}_i)^T (\boldsymbol{\mu}_j + \boldsymbol{\Gamma}_j \mathbf{Z}_j) \hat{g}_{ij} \\
 &= \sum_{i \neq j} \boldsymbol{\mu}_i^T \boldsymbol{\mu}_j \hat{g}_{ij} + \sum_{i \neq j} \boldsymbol{\mu}_i^T \boldsymbol{\Gamma}_j \mathbf{Z}_j \hat{g}_{ij} + \sum_{i \neq j} \boldsymbol{\mu}_j^T \boldsymbol{\Gamma}_i \mathbf{Z}_i \hat{g}_{ij} \\
 &\quad + \sum_{i \neq j} \mathbf{Z}_i^T \boldsymbol{\Gamma}_i^T \boldsymbol{\Gamma}_j \mathbf{Z}_j (\hat{g}_{ij} - g_{ij}) + \sum_{i \neq j} \mathbf{Z}_i^T \boldsymbol{\Gamma}_i^T \boldsymbol{\Gamma}_j \mathbf{Z}_j g_{ij} \\
 &=: W_1 + W_2 + W_3 + W_4 + W_5.
 \end{aligned}$$

With elemental calculation, it is obtained that

$$\begin{aligned}
 W_1 &= \sum_{r=1}^R \boldsymbol{\mu}_r^T \boldsymbol{\mu}_r N_r (n - N_r) - \sum_{r \neq s} \boldsymbol{\mu}_r^T \boldsymbol{\mu}_s N_r N_s = \sum_{r > s} N_r N_s \|\boldsymbol{\mu}_r - \boldsymbol{\mu}_s\|^2, \\
 \mathbb{E}(W_2 | Y_1, \dots, Y_n) &= \mathbb{E}(W_3 | Y_1, \dots, Y_n) = \mathbb{E}(W_4 | Y_1, \dots, Y_n) = 0, \\
 \text{var}(W_2 | Y_1, \dots, Y_n) &= \sum_{j=1}^n \left(\sum_{i \neq j} \boldsymbol{\mu}_i^T \hat{g}_{ij} \right) \boldsymbol{\Sigma}_j \left(\sum_{k \neq j} \boldsymbol{\mu}_k \hat{g}_{kj} \right) \\
 &= \sum_{j=1}^n \left\{ \sum_{r \neq s} I(Y_j = r) N_s (\boldsymbol{\mu}_r - \boldsymbol{\mu}_s)^T \right\} \boldsymbol{\Sigma}_j \left\{ \sum_{r \neq s} I(Y_j = r) N_s (\boldsymbol{\mu}_r - \boldsymbol{\mu}_s) \right\} \\
 &= \sum_{r=1}^R N_r \left\{ \sum_{s \neq r} N_s (\boldsymbol{\mu}_r - \boldsymbol{\mu}_s)^T \right\} \boldsymbol{\Sigma}_r \left\{ \sum_{s \neq r} N_s (\boldsymbol{\mu}_r - \boldsymbol{\mu}_s) \right\} \\
 &\leq O(n^3) \sum_{r \neq s} (\boldsymbol{\mu}_r - \boldsymbol{\mu}_s)^T \boldsymbol{\Sigma}_r (\boldsymbol{\mu}_r - \boldsymbol{\mu}_s) = o(n^2 p), \\
 \text{var}(W_3 | Y_1, \dots, Y_n) &= \text{var}(W_2 | Y_1, \dots, Y_n), \\
 \text{var}(W_4 | Y_1, \dots, Y_n) &= 2 \sum_{r=1}^R \text{tr}(\boldsymbol{\Sigma}_r^2) N_r (N_r - 1) \frac{(\hat{p}_r - p_r)^2}{\hat{p}_r^2 p_r^2}
 \end{aligned}$$

$$\begin{aligned}
 &\leq 2O(n^2) \sum_{r=1}^R \text{tr}(\boldsymbol{\Sigma}_r^2) \frac{(\hat{p}_r - p_r)^2}{\hat{p}_r^2 p_r^2}, \\
 \text{var}(W_5) &= 2n(n-1) \left\{ \sum_{r=1}^R \text{tr}(\boldsymbol{\Sigma}_r^2) - 2 \sum_{r=1}^R p_r \text{tr}(\boldsymbol{\Sigma}_r^2) + \sum_{r=1}^R \sum_{s=1}^R p_r p_s \text{tr}(\boldsymbol{\Sigma}_r \boldsymbol{\Sigma}_s) \right\} \\
 &= 2n(n-1) \left\{ \sum_{r=1}^R (1-p_r)^2 \text{tr}(\boldsymbol{\Sigma}_r^2) + \sum_{r \neq s} p_r p_s \text{tr}(\boldsymbol{\Sigma}_r \boldsymbol{\Sigma}_s) \right\}.
 \end{aligned}$$

Write $Z_{ni} = 2 \sum_{j=1}^{i-1} g_{ij} \mathbf{Z}_i^T \boldsymbol{\Gamma}_i^T \boldsymbol{\Gamma}_j \mathbf{Z}_j$, and $\mathcal{F}_i = \sigma\left\{\left(\frac{\mathbf{Z}_1}{Y_1}\right), \dots, \left(\frac{\mathbf{Z}_i}{Y_i}\right)\right\}$ be the σ -field generated by $\{(\mathbf{Z}_j^T, Y_j), j \leq i\}$, $d_{n,p} = \text{var}(W_5)$, $v_{ni} = E(Z_{ni}^2 | \mathcal{F}_{i-1})$, $2 \leq i \leq n$, and $V_n = \sum_{i=2}^n v_{ni}$. Then $W_5 / \sqrt{d_{n,p}} = \sum_{i=2}^n Z_{ni} / \sqrt{d_{n,p}}$. It is easy to see that $E(Z_{ni} | \mathcal{F}_{i-1}) = 0$ and that is to say $\{\sum_{i=2}^k Z_{ni}, \mathcal{F}_k : 2 \leq j \leq n\}$ is zero mean martingale. The central limit theorem in Hall and Heyde (1980) will hold if

$$d_{n,p}^{-1} V_n \xrightarrow{p} 1, \quad (\text{S3.1})$$

and for any $\varepsilon > 0$

$$\sum_{i=2}^n d_{n,p}^{-1} E\{Z_{ni}^2 I(|Z_{ni}| > \varepsilon \sqrt{d_{n,p}}) | \mathcal{F}_{i-1}\} \xrightarrow{p} 0. \quad (\text{S3.2})$$

Denote $\tilde{g}_{ij}^r = p_r^{-1} \{I(Y_i = r) - p_r\} \{I(Y_j = r) - p_r\}$. Then it can be shown that

$$v_{ni} = 4 \sum_{j=1}^{i-1} \sum_{r=1}^R \tilde{g}_{jj}^r \mathbf{Z}_j^T \boldsymbol{\Gamma}_j^T \boldsymbol{\Sigma}_r \boldsymbol{\Gamma}_j \mathbf{Z}_j + 8 \sum_{1 \leq j < k < i} \sum_{r=1}^R \tilde{g}_{jk}^r \mathbf{Z}_j^T \boldsymbol{\Gamma}_j^T \boldsymbol{\Sigma}_r \boldsymbol{\Gamma}_k \mathbf{Z}_k,$$

and thus

$$d_{n,p}^{-1} V_n$$

$$\begin{aligned}
 &= d_{n,p}^{-1} \sum_{r=1}^R \left\{ 4 \sum_{j=1}^{n-1} (n-j) \tilde{g}_{jj}^r \mathbf{Z}_j^T \Gamma_j^T \Sigma_r \Gamma_j \mathbf{Z}_j + 8 \sum_{1 \leq j < k < n} (n-k) \tilde{g}_{jk}^r \mathbf{Z}_j^T \Gamma_j^T \Sigma_r \Gamma_k \mathbf{Z}_k \right\} \\
 &=: W_6 + W_7.
 \end{aligned}$$

Obviously, we have $E(W_6) = 1$ and

$$\begin{aligned}
 E(W_6^2) &= \frac{16}{d_{n,p}^2} E \left\{ \sum_{r=1}^R \sum_{s=1}^R \sum_{i=1}^{n-1} (n-i)^2 \tilde{g}_{ii}^r \tilde{g}_{ii}^s (\mathbf{Z}_i^T \Gamma_i^T \Sigma_r \Gamma_i \mathbf{Z}_i) (\mathbf{Z}_i^T \Gamma_i^T \Sigma_s \Gamma_i \mathbf{Z}_i) \right. \\
 &\quad \left. + \sum_{r=1}^R \sum_{s=1}^R \sum_{i \neq j}^{n-1} (n-i)(n-j) \tilde{g}_{ij}^r \tilde{g}_{ij}^s (\mathbf{Z}_j^T \Gamma_j^T \Sigma_r \Gamma_j \mathbf{Z}_j) (\mathbf{Z}_i^T \Gamma_i^T \Sigma_s \Gamma_i \mathbf{Z}_i) \right\} \\
 &= \frac{16}{d_{n,p}^2} E \left[\sum_{r=1}^R \sum_{i=1}^{n-1} (n-i)^2 p_r^{-2} \{I(Y_i = r) - p_r\}^4 (\mathbf{Z}_i^T \Gamma_i^T \Sigma_r \Gamma_i \mathbf{Z}_i)^2 \right. \\
 &\quad + \sum_{r \neq s} \sum_{i=1}^{n-1} (n-i)^2 \tilde{g}_{ii}^r \tilde{g}_{ii}^s (\mathbf{Z}_j^T \Gamma_j^T \Sigma_r \Gamma_j \mathbf{Z}_j) (\mathbf{Z}_i^T \Gamma_i^T \Sigma_s \Gamma_i \mathbf{Z}_i) \\
 &\quad + \sum_{r=1}^R \sum_{s=1}^R \sum_{i \neq j}^{n-1} (n-i)(n-j) \{ \text{tr}(\Sigma_r^2) - 2p_r \text{tr}(\Sigma_r) + p_r \text{tr}(\Sigma_r \Sigma) \} \\
 &\quad \left. \cdot \{ \text{tr}(\Sigma_s^2) - 2p_s \text{tr}(\Sigma_s) + p_s \text{tr}(\Sigma_s \Sigma) \} \right] \\
 &= \frac{O\{n^3 \text{tr}^2(\Sigma^2)\}}{d_{n,p}^2} + \frac{O\{n^3 \text{tr}^2(\Sigma^2)\}}{d_{n,p}^2} + \left\{ \frac{d_{n,p}^2}{d_{n,p}^2} + \frac{O\{n^3 \text{tr}^2(\Sigma^2)\}}{d_{n,p}^2} \right\} = 1 + o(1),
 \end{aligned}$$

which implies that $\text{var}(W_6) = o(1)$. Hence, we have $W_6 \xrightarrow{p} 1$. Similarly,

$E(W_7) = 0$ and

$$\begin{aligned}
 &\text{var}(W_7) \\
 &= \frac{64}{d_{n,p}^2} \sum_{r=1}^R \sum_{s=1}^R \sum_{i < j < n} \sum_{k < t < n} (n-j)(n-t) E \left(\tilde{g}_{ij}^r \tilde{g}_{kt}^s \mathbf{Z}_i^T \Gamma_i^T \Sigma_r \Gamma_j \mathbf{Z}_j \mathbf{Z}_k^T \Gamma_k^T \Sigma_s \Gamma_t \mathbf{Z}_t \right) \\
 &= \frac{64}{d_{n,p}^2} \sum_{r=1}^R \sum_{s=1}^R \sum_{i < j < n} (n-j)^2 E \left(\tilde{g}_{ij}^r \tilde{g}_{ij}^s \mathbf{Z}_i^T \Gamma_i^T \Sigma_r \Gamma_j \mathbf{Z}_j \mathbf{Z}_j^T \Gamma_j^T \Sigma_s \Gamma_i \mathbf{Z}_i \right)
 \end{aligned}$$

$$= \frac{64}{d_{n,p}^2} \sum_{r=1}^R \sum_{s=1}^R \sum_{i < j < n} (n-j)^2 \mathbb{E} \{ \tilde{g}_{ij}^r \tilde{g}_{ij}^s \text{tr}(\boldsymbol{\Sigma}_r \boldsymbol{\Sigma}_j \boldsymbol{\Sigma}_s \boldsymbol{\Sigma}_i) \} = \frac{o\{n^4 \text{tr}^2(\boldsymbol{\Sigma}^2)\}}{d_{n,p}^2} \rightarrow 0,$$

which implies $W_7 \xrightarrow{p} 0$. Thus, (S3.1) follows. Next, we shall show (S3.2).

Since

$$\mathbb{E}\{Z_{ni}^2 I(|Z_{ni}| > \varepsilon \sqrt{d_{n,p}}) | \mathcal{F}_{i-1}\} \leq \mathbb{E}(Z_{ni}^4 | \mathcal{F}_{i-1}) / d_{n,p},$$

it is sufficient to show

$$\sum_{i=2}^n \mathbb{E}(Z_{ni}^4) = o(d_{n,p}^2)$$

by the law of large numbers. Because $|\tilde{g}_{ij}|$ is bounded, then we have

$$\begin{aligned} \sum_{i=2}^n \mathbb{E}(Z_{ni}^4) &= O(n^2) \{ \mathbb{E}(\mathbf{Z}_1^\top \boldsymbol{\Gamma}_1^\top \boldsymbol{\Gamma}_2 \mathbf{Z}_2)^4 \} + O(n^3) \mathbb{E} \{ (\mathbf{Z}_1^\top \boldsymbol{\Gamma}_1^\top \boldsymbol{\Gamma}_2 \mathbf{Z}_2)^2 (\mathbf{Z}_3^\top \boldsymbol{\Gamma}_3^\top \boldsymbol{\Gamma}_4 \mathbf{Z}_4)^2 \} \\ &\leq O(n^3) \mathbb{E}(\mathbf{Z}_1^\top \boldsymbol{\Gamma}_1^\top \boldsymbol{\Gamma}_2 \mathbf{Z}_2)^4 = o(d_{n,p}^2), \end{aligned}$$

which completes the proof. \square

S4 Proof of Lemma 3.1

Proof. Similar to the proof of Lemma 2.1 in S1, we have

$$\begin{aligned} &\boldsymbol{\Sigma}_1 = \cdots = \boldsymbol{\Sigma}_R = \boldsymbol{\Sigma} \\ \iff &\frac{1}{2} \mathbb{E} \{ (\mathbf{X}_1 - \mathbf{X}_2)(\mathbf{X}_1 - \mathbf{X}_2)^\top | Y_1 = r, Y_2 = r \} \\ &= \frac{1}{2} \mathbb{E} \sum_{s=1}^R \frac{I(Y_1 = s) I(Y_2 = s)}{p_s} (\mathbf{X}_1 - \mathbf{X}_2)(\mathbf{X}_1 - \mathbf{X}_2)^\top, \quad \forall r = 1, 2, \dots, R \\ \iff &\frac{1}{2} \mathbb{E} (\mathbf{X}_1 - \mathbf{X}_2)(\mathbf{X}_1 - \mathbf{X}_2)^\top \frac{I(Y_1 = r) I(Y_2 = r)}{p_r^2} \end{aligned}$$

$$\begin{aligned}
&= \frac{1}{2} \mathbb{E} \sum_{s=1}^R \frac{I(Y_1 = s)I(Y_2 = s)}{p_s} (\mathbf{X}_1 - \mathbf{X}_2)(\mathbf{X}_1 - \mathbf{X}_2)^\top, \quad \forall r = 1, 2, \dots, R \\
\iff &\frac{1}{4} \sum_{r=1}^R p_r \text{tr} \left[\mathbb{E}(\mathbf{X}_1 - \mathbf{X}_2)(\mathbf{X}_1 - \mathbf{X}_2)^\top \right. \\
&\quad \left. \left\{ \frac{I(Y_1 = r)I(Y_2 = r)}{p_r^2} - \sum_{s=1}^R \frac{I(Y_1 = s)I(Y_2 = s)}{p_s} \right\}^2 \right] = 0 \\
\iff &\frac{1}{4} \mathbb{E} \{ (\mathbf{X}_1 - \mathbf{X}_2)^\top (\mathbf{X}_3 - \mathbf{X}_4) \}^2 \left\{ \sum_{r=1}^R I(Y_1 = r)I(Y_2 = r)I(Y_3 = r)I(Y_4 = r) \right. \\
&\quad \left. \frac{(1 - p_r)}{p_r^3} - \sum_{r \neq s} \frac{I(Y_1 = r)I(Y_2 = r)}{p_r} \frac{I(Y_3 = s)I(Y_4 = s)}{p_s} \right\} = 0,
\end{aligned}$$

which implies that $\mathcal{V}(\mathbf{X}|Y) \geq 0$ is true, and the equality holds if and only if $\Sigma_1 = \dots = \Sigma_R$. \square

S5 Proof of Theorem 3.1

Proof. Under condition 2, it is obtained that

$$\begin{aligned}
T_{n,p} &= \frac{1}{4} \sum_{(i_1, i_2, i_3, i_4)}^* \{ (\Gamma_{i_1} \mathbf{Z}_{i_1} - \Gamma_{i_2} \mathbf{Z}_{i_2})^\top (\Gamma_{i_3} \mathbf{Z}_{i_3} - \Gamma_{i_4} \mathbf{Z}_{i_4}) \}^2 \hat{f}_{i_1 i_2 i_3 i_4} \\
&= \sum_{(i_1, i_2, i_3, i_4)}^* \left\{ (\mathbf{Z}_{i_1}^\top \Gamma_{i_1}^\top \Gamma_{i_3} \mathbf{Z}_{i_3})^2 \hat{f}_{i_1 i_2 i_3 i_4} - 2(\mathbf{Z}_{i_1}^\top \Gamma_{i_1}^\top \Gamma_{i_3} \mathbf{Z}_{i_3} \mathbf{Z}_{i_1}^\top \Gamma_{i_1}^\top \Gamma_{i_4} \mathbf{Z}_{i_4}) \hat{f}_{i_1 i_2 i_3 i_4} \right. \\
&\quad \left. + (\mathbf{Z}_{i_1}^\top \Gamma_{i_1}^\top \Gamma_{i_3} \mathbf{Z}_{i_3} \mathbf{Z}_{i_2}^\top \Gamma_{i_2}^\top \Gamma_{i_4} \mathbf{Z}_{i_4}) \hat{f}_{i_1 i_2 i_3 i_4} \right\} =: Q_1 + Q_2 + Q_3.
\end{aligned}$$

First, we consider Q_2 and Q_3 . It is easy to find that $\mathbb{E}(Q_2|Y_1, \dots, Y_n) =$

$\mathbb{E}(Q_3|Y_1, \dots, Y_n) = 0$. Write

$$\hat{h}_{i_1 i_2 i_3} = \sum_{r=1}^R \frac{(n - N_r)I(Y_{i_1} = Y_{i_2} = Y_{i_3} = r)}{(N_r - 2)(N_r - 1)/(n - 1)^2} - \sum_{r \neq s} \frac{I(Y_{i_1} = r)I(Y_{i_2} = Y_{i_3} = s)}{(N_s - 1)/(n - 1)^2}.$$

Then, we have $Q_2 = -2 \sum_{(i_1, i_2, i_3)}^* (\mathbf{Z}_{i_1}^T \Gamma_{i_1}^T \Gamma_{i_2} \mathbf{Z}_{i_2} \mathbf{Z}_{i_1}^T \Gamma_{i_1}^T \Gamma_{i_3} \mathbf{Z}_{i_3}) \hat{h}_{i_1 i_2 i_3}$, and

$$\begin{aligned}
 \mathbb{E}(Q_2^2 | Y_1, \dots, Y_n) &= 4\mathbb{E} \left\{ \sum_{(i_1, i_2, i_3)}^* \sum_{(j_1, j_2, j_3)}^* (\mathbf{Z}_{i_1}^T \Gamma_{i_1}^T \Gamma_{i_2} \mathbf{Z}_{i_2} \mathbf{Z}_{i_1}^T \Gamma_{i_1}^T \Gamma_{i_3} \mathbf{Z}_{i_3}) \right. \\
 &\quad \left. (\mathbf{Z}_{j_1}^T \Gamma_{j_1}^T \Gamma_{j_2} \mathbf{Z}_{j_2} \mathbf{Z}_{j_1}^T \Gamma_{j_1}^T \Gamma_{j_3} \mathbf{Z}_{j_3}) \hat{h}_{i_1 i_2 i_3} \hat{h}_{j_1 j_2 j_3} | Y_1, \dots, Y_n \right\} \\
 &\leq O(1) \sum_{(i_1, i_2, i_3, i_4)}^* \text{tr}(\Sigma_{i_1} \Sigma_{i_3} \Sigma_{i_4} \Sigma_{i_2}) \hat{h}_{i_1 i_2 i_3} \hat{h}_{i_4 i_2 i_3} + O(1) \sum_{(i_1, i_2, i_3)}^* \text{tr}(\Sigma_{i_1} \Sigma_{i_2}) \text{tr}(\Sigma_{i_1} \Sigma_{i_3}) \hat{h}_{i_1 i_2 i_3}^2 \\
 &\leq o\{\text{tr}^2(\Sigma^2)\} \sum_{(i_1, i_2, i_3, i_4)}^* \hat{h}_{i_1 i_2 i_3}^2 + O\{\text{tr}^2(\Sigma^2)\} \sum_{(i_1, i_2, i_3)}^* \hat{h}_{i_1 i_2 i_3}^2 = o(\delta_{n,p}). \\
 \mathbb{E}(Q_3^2 | Y_1, \dots, Y_n) &\leq O(1) \sum_{(i_1, i_2, i_3, i_4)}^* \text{tr}(\Sigma_{i_1} \Sigma_{i_3}) \text{tr}(\Sigma_{i_2} \Sigma_{i_4}) \hat{f}_{i_1 i_2 i_3 i_4}^2 \\
 &\quad + O(1) \sum_{(i_1, i_2, i_3, i_4)}^* \text{tr}(\Sigma_{i_1} \Sigma_{i_3} \Sigma_{i_4} \Sigma_{i_2}) \hat{f}_{i_1 i_2 i_3 i_4} \hat{f}_{i_1 i_3 i_2 i_4} \\
 &\leq O\{\text{tr}^2(\Sigma^2)\} \sum_{(i_1, i_2, i_3, i_4)}^* \hat{f}_{i_1 i_2 i_3 i_4}^2 = O\{n^4 \text{tr}^2(\Sigma^2)\}.
 \end{aligned}$$

Given Y_1, \dots, Y_n ,

$$\begin{aligned}
 \mathbb{E}(T_{n,p} | Y_1, \dots, Y_n) &= \mathbb{E}(Q_1 | Y_1, \dots, Y_n) \\
 &= \sum_{r=1}^R \text{tr}(\Sigma_r^2) (n-1)^2 N_r (n - N_r) - \sum_{r \neq s} \text{tr}(\Sigma_r \Sigma_s) (n-1)^2 N_r N_s \\
 &= (n-1)^2 \left\{ \sum_{r=1}^R N_r (n - N_r) \text{tr}(\Sigma_r^2) - \sum_{r \neq s} \text{tr}(\Sigma_r \Sigma_s) N_r N_s \right\} \\
 &= (n-1)^2 \sum_{r > s} N_r N_s \text{tr} \{ (\Sigma_r - \Sigma_s)^2 \}.
 \end{aligned}$$

By elemental calculation, it is obtained that

$$Q_1 - \mathbb{E}(Q_1 | Y_1, \dots, Y_n) = (n-1)^2 \sum_{i \neq j} \{ (\mathbf{Z}_i^T \Gamma_i^T \Gamma_j \mathbf{Z}_j)^2 - \text{tr}(\Sigma_i \Sigma_j) \} \hat{g}_{ij}.$$

Write $\xi_{ij} = (\mathbf{Z}_i^T \mathbf{\Gamma}_i^T \mathbf{\Gamma}_j \mathbf{Z}_j)^2$, $\xi_i^j = \mathbb{E}(\xi_{ij} | \mathbf{Z}_i, Y_j)$ and $\xi^{ij} = \mathbb{E}(\xi_{ij} | Y_i, Y_j)$. Then

$$\begin{aligned} Q_1 - \mathbb{E}(Q_1 | Y_1, \dots, Y_n) &= (n-1)^2 \sum_{i \neq j} \{ (\xi_{ij} - \xi^{ij})(\hat{g}_{ij} - g_{ij}) + (\xi_{ij} - \xi^{ij})g_{ij} \} \\ &= (n-1)^2 \sum_{i \neq j} \left\{ (\xi_{ij} - \xi_i^j - \xi_j^i + \xi^{ij})(\hat{g}_{ij} - g_{ij}) + (\xi_i^j + \xi_j^i - 2\xi^{ij})(\hat{g}_{ij} - g_{ij}) \right. \\ &\quad \left. + (\xi_{ij} - \xi^{ij})g_{ij} \right\} =: Q_{11} + Q_{12} + Q_{13}. \end{aligned}$$

Then, $\mathbb{E}(Q_{1k} | Y_1, \dots, Y_n) = 0$ for $k = 1, 2, 3$. And

$$\begin{aligned} \text{var}(Q_{11} | Y_1, \dots, Y_n) &= 2(n-1)^4 \sum_{i \neq j} \hat{g}_{ij}^2 \mathbb{E} \{ (\xi_{ij} - \xi_i^j - \xi_j^i + \xi^{ij})^2 | Y_1, \dots, Y_n \} \\ &= 2(n-1)^4 \sum_{r=1}^R N_r(N_r-1) \frac{(\hat{p}_r - p_r)^2}{\hat{p}_r^2 p_r^2} \mathbb{E} \{ (\xi_{12} - \xi_1^r - \xi_2^r + \xi^{rr})^2 | Y_1 = Y_2 = r \} \\ &\leq O\{n^4 \text{tr}^2(\mathbf{\Sigma}^2)\} \sum_{r=1}^R N_r(N_r-1) \frac{(\hat{p}_r - p_r)^2}{\hat{p}_r^2 p_r^2}. \end{aligned}$$

It is not hard to obtain that

$$\begin{aligned} Q_{12} &= 2(n-1)^2 \sum_{i \neq j} (\xi_i^j - \xi^{ij})(\hat{g}_{ij} - g_{ij}) \\ &= 2(n-1)^2 \sum_{i \neq j} (\xi_i^j - \xi^{ij}) \left\{ \sum_{r=1}^R I(Y_i = Y_j = r)(\hat{p}_r^{-1} - p_r^{-1}) \right\} \\ &= 2(n-1)^2 \left\{ \sum_{r=1}^R \sum_{i=1}^n (\xi_i^r - \xi^{rr}) I(Y_i = r)(N_r - 1)(\hat{p}_r^{-1} - p_r^{-1}) \right\}. \end{aligned}$$

Then, we have

$$\begin{aligned} \text{var}(Q_{12} | Y_1, \dots, Y_n) &= 4(n-1)^4 \sum_{r=1}^R \sum_{i=1}^R I(Y_i = r)(N_r - 1)^2 (\hat{p}_r^{-1} - p_r^{-1})^2 \mathbb{E} \{ (\xi_i^r - \xi^{rr})^2 | Y_1, \dots, Y_n \} \end{aligned}$$

$$\begin{aligned}
 &= 4(n-1)^4 \sum_{r=1}^R N_r(N_r-1)^2 (\hat{p}_r^{-1} - p_r^{-1})^2 \mathbf{E} \{ (\xi_1^r - \xi^{rr})^2 | Y_1 = r \} \\
 &= 4(n-1)^4 \sum_{r=1}^R N_r(N_r-1)^2 \frac{(\hat{p}_r - p_r)^2}{\hat{p}_r^2 p_r^2} \{ 2\text{tr}(\boldsymbol{\Sigma}_r^4) + \Delta \text{tr}(\boldsymbol{\Gamma}_r^T \boldsymbol{\Gamma}_r \text{diag}(\boldsymbol{\Gamma}_r^T \boldsymbol{\Sigma}_r \boldsymbol{\Gamma}_r) \boldsymbol{\Gamma}_r^T \boldsymbol{\Gamma}_r) \}.
 \end{aligned}$$

Similar to the analysis in proof of Theorem 2.1, it is obtained that $Q_{1k} = o\{n^3 \text{tr}(\boldsymbol{\Sigma}^2)\}$ for $k = 1, 2$. Up to now, the proof completes if we have $Q_{13}/\sqrt{\delta_{n,p}} \xrightarrow{d} \mathcal{N}(0, 1)$ as $n, p \rightarrow \infty$.

Let $\mathbf{E}_i(\cdot)$ denote the conditional expectation given \mathcal{F}_i . Define $D_{ni} = (\mathbf{E}_i - \mathbf{E}_{i-1})Q_{13}$ and $\delta_{n,p} = \text{var}(Q_{13})$. Then $Q_{13}/\sqrt{\delta_{n,p}} = \sum_{i=2}^n D_{ni}/\sqrt{\delta_{n,p}}$, and $\{D_{nk}, 1 \leq k \leq n\}$ is a martingale difference sequence with respect to the σ -fields $\{\mathcal{F}_k, 1 \leq k \leq n\}$. By elemental calculation, it is obtained that

$$\begin{aligned}
 D_{ni} &= 2(n-1)^2 \sum_{j=1}^{i-1} \left\{ (\xi_{ij} - \xi^{ij})g_{ij} - \sum_{r=1}^R (I(Y_j = r) - p_r)(\xi_j^r - \xi^{jr}) \right\} \\
 &\quad + (n-i) \sum_{r=1}^R \{I(Y_i = r) - p_r\} (\xi_i^r - \xi^{ri}).
 \end{aligned}$$

With elemental calculation, we have

$$\begin{aligned}
 &\text{var} \{ (\mathbf{Z}_1^T \mathbf{M}^T \mathbf{Z}_2)^2 \} \\
 &= 3\mathbf{E} \{ (\mathbf{Z}_1^T \mathbf{M}^T \mathbf{M} \mathbf{Z}_1)^2 \} + \Delta \text{tr} [\mathbf{E} \{ (\mathbf{M} \mathbf{Z}_1 \mathbf{Z}_1^T \mathbf{M}^T) \circ (\mathbf{M} \mathbf{Z}_1 \mathbf{Z}_1^T \mathbf{M}^T) \}] - \text{tr}^2(\mathbf{M}^T \mathbf{M}) \\
 &= 6\text{tr}(\mathbf{M}^T \mathbf{M} \mathbf{M}^T \mathbf{M}) + 3 \Delta \text{tr}(\mathbf{M}^T \mathbf{M} \circ \mathbf{M}^T \mathbf{M}) + 2\text{tr}^2(\mathbf{M}^T \mathbf{M}) \\
 &\quad + \Delta^2 \sum_i \sum_j m_{ij}^4 + 3 \Delta \text{tr}(\mathbf{M} \mathbf{M}^T \circ \mathbf{M} \mathbf{M}^T),
 \end{aligned}$$

where $\mathbf{M} = (m_{ij})$ is a given matrix. Notice the fact that $\sum_i \sum_j m_{ij}^4 \leq$

$\text{tr}(\mathbf{M}\mathbf{M}^T \circ \mathbf{M}\mathbf{M}^T)$, then we have

$$\begin{aligned}
 \delta_{n,p} &= (n-1)^4 \left\{ 4 \sum_{(i,j,k)}^* (\xi_{ij} - \xi^{ij})(\xi_{ik} - \xi^{ik}) g_{ij} g_{ik} + 2 \sum_{i \neq j} (\xi_{ij} - \xi^{ij})^2 g_{ij}^2 \right\} \\
 &= (n-1)^4 \left[8n(n-1)(n-2) \sum_{r=1}^R p_r \text{tr} \{ (\boldsymbol{\Sigma}_r^2 - \boldsymbol{\Sigma}_r \boldsymbol{\Sigma})^2 \} \right. \\
 &\quad + 4\Delta n(n-1)(n-2) \sum_{r=1}^R p_r \text{tr} \{ \boldsymbol{\Gamma}_r^T (\boldsymbol{\Sigma}_r - \boldsymbol{\Sigma}) \boldsymbol{\Gamma}_r \circ \boldsymbol{\Gamma}_r^T (\boldsymbol{\Sigma}_r - \boldsymbol{\Sigma}) \boldsymbol{\Gamma}_r \} \\
 &\quad + 2n(n-1) \sum_{r=1}^R (1-p_r)^2 \{ 2\text{tr}^2(\boldsymbol{\Sigma}_r^2) + O(p) \} \\
 &\quad \left. + 2n(n-1) \sum_{r \neq s} p_r p_s \{ 2\text{tr}^2(\boldsymbol{\Sigma}_r \boldsymbol{\Sigma}_s) + O(p) \} \right] \\
 &= 4n^6 \left\{ \sum_{r=1}^R (1-p_r)^2 \text{tr}^2(\boldsymbol{\Sigma}_r^2) + \sum_{r \neq s} p_r p_s \text{tr}^2(\boldsymbol{\Sigma}_r \boldsymbol{\Sigma}_s) \right\} (1+o(1)) \\
 &\quad + 8n^7 \sum_{r=1}^R p_r \text{tr} \{ (\boldsymbol{\Sigma}_r^2 - \boldsymbol{\Sigma}_r \boldsymbol{\Sigma})^2 \} (1+o(1)) \\
 &\quad + 4\Delta n^7 \sum_{r=1}^R p_r \text{tr} \{ \boldsymbol{\Gamma}_r^T (\boldsymbol{\Sigma}_r - \boldsymbol{\Sigma}) \boldsymbol{\Gamma}_r \circ \boldsymbol{\Gamma}_r^T (\boldsymbol{\Sigma}_r - \boldsymbol{\Sigma}) \boldsymbol{\Gamma}_r \} (1+o(1)).
 \end{aligned}$$

Similar to the analysis in the proof of Theorem 2.1, $Q_{1k}/\sqrt{\delta_{n,p}} = o_P(1)$

for $k = 1, 2$. Under H_0 , $\delta_{n,p} = 4(R-1)n^6 \text{tr}^2(\boldsymbol{\Sigma}^2) \{1+o(1)\}$. Write $u_{ni} =$

$\text{E}(D_{ni}^2 | \mathcal{F}_{i-1})$, $2 \leq i \leq n$, and $U_n = \sum_{i=2}^n u_{ni}$. On the basis of the proof

of Theorem 2.1, it is sufficient to show $\delta_{n,p}^{-1} U_n \xrightarrow{p} 1$, and for any $\varepsilon > 0$

$\sum_{i=2}^n \delta_{n,p}^{-1} \text{E} \{ D_{ni}^2 I(|D_{ni}| > \varepsilon \sqrt{\delta_{n,p}}) | \mathcal{F}_{i-1} \} \xrightarrow{p} 0$. $u_{ni} = \text{E}_{i-1}(D_{ni}^2)$.

$$U_n = \sum_{i=2}^n \left\{ 4(n-1)^4 \sum_{j=1}^{i-1} \sum_{k=1}^{i-1} \text{E}_{i-1}(\xi_{ij} - \xi^{ij})(\xi_{ik} - \xi^{ik}) g_{ij} g_{ik} \right.$$

$$\begin{aligned}
 & -4(n-1)^4 \sum_{j=1}^{i-1} \sum_{k=1}^{i-1} \sum_{r=1}^R \sum_{s=1}^R (I(Y_j = r) - p_r)(I(Y_k = s) - p_s)(\xi_j^r - \xi^{rj})(\xi_k^s - \xi^{sk}) \\
 & +4(n-1)^4(n-i)^2 \sum_{r=1}^R \sum_{s=1}^R \mathbb{E}_{i-1}(I(Y_i = r) - p_r)(I(Y_i = s) - p_s)(\xi_i^r - \xi^{ri})(\xi_i^s - \xi^{si}) \\
 & +8(n-1)^4(n-i) \sum_{j=1}^{i-1} \mathbb{E}_{i-1}(\xi_{ij} - \xi^{ij})g_{ij} \sum_{r=1}^R (I(Y_i = r) - p_r)(\xi_i^r - \xi^{ri}) \Big\} \\
 = & 8(n-1)^4 \sum_{j=1}^{n-1} (n-j) \sum_{r=1}^R \tilde{g}_{jj}^r (\mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Sigma}_r \mathbf{\Gamma}_j \mathbf{Z}_j)^2 \\
 & +4\Delta(n-1)^4 \sum_{j=1}^{n-1} (n-j) \sum_{r=1}^R \tilde{g}_{jj}^r \text{tr}(\mathbf{\Gamma}_r^T \mathbf{\Gamma}_j \mathbf{Z}_j \mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Gamma}_r \circ \mathbf{\Gamma}_r^T \mathbf{\Gamma}_j \mathbf{Z}_j \mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Gamma}_r) \\
 & +4(n-1)^4 \sum_{j=1}^{n-1} (n-j) \sum_{r=1}^R \tilde{g}_{jj}^r (1-p_r)(\xi_j^r - \xi^{jr})^2 \\
 & +16(n-1)^4 \sum_{1 \leq j < k < n} (n-k) \sum_{r=1}^R \tilde{g}_{jk}^r (\mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Sigma}_r \mathbf{\Gamma}_k \mathbf{Z}_k)^2 \\
 & +8\Delta(n-1)^4 \sum_{1 \leq j < k < n} (n-k) \sum_{r=1}^R \tilde{g}_{jk}^r \text{tr} \{ \mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Gamma}_r \text{diag}(\mathbf{\Gamma}_r^T \mathbf{\Gamma}_k \mathbf{Z}_k \mathbf{Z}_k^T \mathbf{\Gamma}_k^T \mathbf{\Gamma}_r) \mathbf{\Gamma}_r^T \mathbf{\Gamma}_j \mathbf{Z}_j \} \\
 & +8(n-1)^4 \sum_{1 \leq j < k < n} (n-k) \sum_{r=1}^R \tilde{g}_{jk}^r (1-p_r)(\xi_j^r - \xi^{jr})(\xi_k^r - \xi^{kr}) \\
 & +8(n-1)^4 \sum_{i=2}^n (n-i)^2 \sum_{r=1}^R p_r \text{tr} \{ (\mathbf{\Sigma}_r^2 - \mathbf{\Sigma}_r \mathbf{\Sigma})^2 \} \\
 & +4\Delta(n-1)^4 \sum_{i=2}^n (n-i)^2 \sum_{r=1}^R p_r \text{tr} \{ \mathbf{\Gamma}_r^T (\mathbf{\Sigma}_r - \mathbf{\Sigma}) \mathbf{\Gamma}_r \circ \mathbf{\Gamma}_r^T (\mathbf{\Sigma}_r - \mathbf{\Sigma}) \mathbf{\Gamma}_r \} \\
 & +16(n-1)^4 \sum_{i=2}^n (n-i) \sum_{j=1}^{i-1} \sum_{r=1}^R \{ I(Y_j = r) - p_r \} \mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Sigma}_r (\mathbf{\Sigma}_r - \mathbf{\Sigma}) \mathbf{\Sigma}_r \mathbf{\Gamma}_j \mathbf{Z}_j \\
 & +8\Delta(n-1)^4 \sum_{i=2}^n (n-i) \sum_{j=1}^{i-1} \sum_{r=1}^R \{ I(Y_j = r) - p_r \} \mathbf{Z}_j^T \mathbf{\Gamma}_j^T \mathbf{\Gamma}_r \text{diag} \{ \mathbf{\Gamma}_r^T (\mathbf{\Sigma}_r - \mathbf{\Sigma}) \mathbf{\Gamma}_r \} \mathbf{\Gamma}_r^T \mathbf{\Gamma}_j \mathbf{Z}_j
 \end{aligned}$$

$$=: \sum_{k=1}^{10} Q_{13}^{(k)}$$

It is elemental to check that $E(U_n) = \delta_{n,p}(1 + o(1))$ and $E(|Q_{13}^{(k)}|) = O(n^6 p), k = 2, 3$. To prove $\delta_{n,p}^{-1} U_n \xrightarrow{P} 1$, it is sufficient to obtain that $\text{var}(Q_{13}^{(k)}) = o(\delta_{n,p}^2)$ for $k = 1, 4, 5, 6, 9, 10$. For $Q_{13}^{(1)}$, we have

$$\begin{aligned} \text{var}(Q_{13}^{(1)}) &\leq O(n^8)O(n^3) \sum_{r=1}^R \sum_{s=1}^R (\tilde{g}_{11}^r)^2 E(\mathbf{Z}_j^T \Gamma_j^T \Sigma_r \Gamma_j \mathbf{Z}_j)^2 (\mathbf{Z}_k^T \Gamma_k^T \Sigma_r \Gamma_k \mathbf{Z}_k)^2 \\ &= O\{n^{11} \text{tr}^4(\Sigma^2)\}, \end{aligned}$$

$$\text{var}(Q_{13}^{(4)}) \leq O(n^{13}) \sum_{r \neq s} \text{tr} \{(\Sigma_r^2 - \Sigma_s \Sigma)^2\} + O(n^{12}) o\{\text{tr}^2(\Sigma^2)\} = o(\delta_{n,p}^2).$$

Similar to the analysis of $Q_{13}^{(4)}$, $\text{var}(Q_{13}^{(k)}) = o(\delta_{n,p}^2)$ follows for $k = 5, 6, 9, 10$.

Next we want to prove that $\sum_{i=2}^n \delta_{n,p}^{-1} E\{D_{ni}^2 I(|D_{ni}| > \varepsilon \sqrt{\delta_{n,p}}) | \mathcal{F}_{i-1}\} \xrightarrow{P} 0$, and it is sufficient to show $\delta_{n,p}^{-2} \sum_{i=2}^n E(D_{ni}^4) \rightarrow 0$. By the definition of D_{ni} ,

it is obtained that

$$\begin{aligned} \sum_{i=2}^n E(D_{ni}^4) &\leq O(1) \left[\sum_{i=2}^n E \left\{ 2(n-1)^2 \sum_{j=1}^{i-1} (\xi_{ij} - \xi_j^i) g_{ij} \right\}^4 \right. \\ &\quad + \sum_{i=2}^n E \left\{ 2(n-1)^2 \sum_{j=1}^{i-1} (\xi_j^i - \xi^{ij}) g_{ij} \right\}^4 \\ &\quad + \sum_{i=2}^n E \left\{ 2(n-1)^2 \sum_{j=1}^{i-1} (\xi_j^r - \xi^{jr}) (I(Y_j = r) - p_r) \right\}^4 \\ &\quad \left. + \sum_{i=2}^n (n-i)^4 E \left\{ \sum_{r=1}^R (I(Y_i = r) - p_r) (\xi_i^r - \xi^{ir}) \right\}^4 \right] \\ &\leq O\{n^{11} \text{tr}^4(\Sigma^2)\} = o(\delta_{n,p}^2), \end{aligned}$$

which completes the proof. \square