

# Sufficient variable screening for ultrahigh-dimensional right censored data via independence measures

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## Supplementary Material

In this supplementary file, section A.1 provides the comprehensive computation of equations (1), (2), and (3); Section A.2 gives an estimation of  $G(t|\mathbf{X})$  across various scenarios; Section A.3 presents rigorous proofs for Theorems 1, 2, and 3. Section A. 4 contains additional simulation instances, real data analysis, and corresponding results.

### A.1 Detailed Calculation of Equations (3.1), (3.2), and (3.3)

**Calculation of Equation (3.1)** To obtain that  $dcov_\alpha^2(X_\alpha, T) = dcov_\alpha^2(X_\alpha, Y)$ , we need to prove that,

$$\begin{aligned} S_1 &= E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|T - T'\|_{d_T} \right] = E (\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y'\|_{d_Y}), \\ S_2 &= E \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|T - T'\|_{d_T} \right] = E \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \|Y - Y'\|_{d_Y}, \\ S_3 &= E \left[ \frac{\delta\delta'\delta''}{G(T|\mathbf{X})G(T'|\mathbf{X}')G(T''|\mathbf{X}'')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|T - T''\|_{d_T} \right] = E (\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y''\|_{d_Y}), \end{aligned}$$

In detail,

$$\begin{aligned}
S_1 &= E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|T - T'\|_{d_T} \right) \\
&= E \left[ E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|T - T'\|_{d_T} \mid (\mathbf{X}, \mathbf{X}', Y, Y') \right) \right] \\
&= E \left[ \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \left( \frac{\delta\delta' \|T - T'\|_{d_T}}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \mid (\mathbf{X}, \mathbf{X}', Y, Y') \right) \right] \\
&= E \left[ \frac{\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y'\|_{d_T}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(\delta = 1, \delta' = 1 \mid (\mathbf{X}, \mathbf{X}', Y, Y')) \right] \\
&= E \left[ \frac{\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y'\|_{d_T}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(\delta = 1 \mid (\mathbf{X}, Y)) Pr(\delta' = 1 \mid (\mathbf{X}', Y')) \right] \\
&= E \left[ \frac{\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y'\|_{d_T}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(C > Y \mid (\mathbf{X}, Y)) Pr(C' > Y' \mid (\mathbf{X}', Y')) \right] \\
&= E(\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y'\|_{d_T})
\end{aligned}$$

$$\begin{aligned}
S_2 &= E \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|T - T'\|_{d_T} \right] \\
&= E \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \left[ E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|T - T'\|_{d_T} \mid (\mathbf{X}, \mathbf{X}', Y, Y') \right) \right] \\
&= E \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \left[ \frac{\|Y - Y'\|_{d_T}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(\delta = 1, \delta' = 1 \mid (\mathbf{X}, \mathbf{X}', Y, Y')) \right] \\
&= E \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} E \|Y - Y'\|_{d_T}
\end{aligned}$$

$$\begin{aligned}
S_3 &= E \left( \frac{\delta\delta'\delta''}{G(T|\mathbf{X})G(T'|\mathbf{X}')G(T''|\mathbf{X}'')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|T - T''\|_{d_T} \right) \\
&= E \left[ E \left( \frac{\delta\delta''}{G(T|\mathbf{X})G(T'|\mathbf{X}')G(T''|\mathbf{X}'')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|T - T''\|_{d_T} \mid (\mathbf{X}, \mathbf{X}', \mathbf{X}'', Y, Y', Y'') \right) \right] \\
&= E \left[ \frac{\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y''\|_{d_T}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')G(Y''|\mathbf{X}'')} Pr(\delta = 1, \delta' = 1, \delta'' = 1 \mid (\mathbf{X}, \mathbf{X}', \mathbf{X}'', Y, Y', Y'')) \right] \\
&= E(\|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \|Y - Y''\|_{d_T}),
\end{aligned}$$

where  $G(t|\mathbf{X}) \stackrel{\text{def}}{=} P(C > t|\mathbf{X})$  is the survival function of censoring time  $C$ .

**Calculation of Equation (3.2)** To obtain that  $dcov_\alpha^{*2}(\mathbf{U}, \mathbf{W}) = dcov_\alpha^2(\mathbf{U}, \mathbf{V})$ , we

need to prove that,

$$S_1 = E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \right) = E (\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}}),$$

$$S_2 = E\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \right] = E\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}E\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}},$$

$$S_3 = E \left( \frac{\delta\delta'\delta''}{G(T|\mathbf{X})G(T'|\mathbf{X}')G(T''|\mathbf{X}'')} \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{W} - \mathbf{W}''\|_{d_{\mathbf{V}}} \right) = E (\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}''\|_{d_{\mathbf{V}}}),$$

In detail,

$$\begin{aligned} S_1 &= E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \right) \\ &= E \left[ E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \mid (\mathbf{X}, \mathbf{X}', Y, Y') \right) \right] \\ &= E \left[ \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}} E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \mid (\mathbf{X}, \mathbf{X}', Y, Y') \right) \right] \\ &= E \left[ \frac{\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(\delta = 1, \delta' = 1 \mid (\mathbf{X}, \mathbf{X}', Y, Y')) \right] \\ &= E \left[ \frac{\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(\delta = 1 \mid (\mathbf{X}, Y)) Pr(\delta' = 1 \mid (\mathbf{X}', Y')) \right] \\ &= E (\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}}) \end{aligned}$$

$$\begin{aligned} S_2 &= E\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \right] \\ &= E\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}E \left[ E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{W} - \mathbf{W}'\|_{d_{\mathbf{V}}} \mid (\mathbf{X}, \mathbf{X}', Y, Y') \right) \right] \\ &= E\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}E \left[ \frac{\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} Pr(\delta = 1, \delta' = 1 \mid (\mathbf{X}, \mathbf{X}', Y, Y')) \right] \\ &= E\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}E\|\mathbf{V} - \mathbf{V}'\|_{d_{\mathbf{V}}} \end{aligned}$$

$$\begin{aligned} S_3 &= E \left( \frac{\delta\delta'\delta''}{G(T|\mathbf{X})G(T'|\mathbf{X}')G(T''|\mathbf{X}'')} \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{W} - \mathbf{W}''\|_{d_{\mathbf{V}}} \right) \\ &= E \left[ E \left( \frac{\delta\delta'\delta''}{G(T|\mathbf{X})G(T'|\mathbf{X}')G(T''|\mathbf{X}'')} \|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{W} - \mathbf{W}''\|_{d_{\mathbf{V}}} \mid (\mathbf{X}, \mathbf{X}', \mathbf{X}'', Y, Y', Y'') \right) \right] \\ &= E \left[ \frac{\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')G(Y''|\mathbf{X}'')} E(\delta\delta'\delta''\|\mathbf{W} - \mathbf{W}''\|_{d_{\mathbf{V}}} \mid (\mathbf{X}, \mathbf{X}', \mathbf{X}'', Y, Y', Y'')) \right] \\ &= E \left[ \frac{\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}''\|_{d_{\mathbf{V}}}}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')G(Y''|\mathbf{X}'')} Pr(\delta = 1, \delta' = 1, \delta'' = 1 \mid (\mathbf{X}, \mathbf{X}', \mathbf{X}'', Y, Y', Y'')) \right] \\ &= E (\|\mathbf{U} - \mathbf{U}'\|_{d_{\mathbf{U}}}\|\mathbf{V} - \mathbf{V}''\|_{d_{\mathbf{V}}}) \end{aligned}$$

### Calculation of Equation (3.3)

The numerator of  $S_{1,s}$  is:

$$\begin{aligned} & E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \|_{d_{\mathbf{X}_{-\alpha}}} \mathbf{1}\{T \in [l_{s-1}, l_s]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} \right] \\ = & E \left[ \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \|_{d_{\mathbf{X}_{-\alpha}}} E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \mathbf{1}\{T \in [l_{s-1}, l_s]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}') \right) \right] \end{aligned}$$

The last term inside the expectation above can be derived as follows:

$$\begin{aligned} & E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \mathbf{1}\{T \in [l_{s-1}, l_s]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}') \right) \\ = & E \left[ E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \mathbf{1}\{T \in [l_{s-1}, l_s]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} | (Y, Y', \mathbf{X}, \mathbf{X}') \right) | ((\mathbf{X}, \mathbf{X}')) \right] \\ = & E \left[ \frac{1}{G(Y|\mathbf{X})G(Y'|\mathbf{X}')} \mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} Pr(Y < C, Y' < C' | (Y, Y', \mathbf{X}, \mathbf{X}')) | ((\mathbf{X}, \mathbf{X}')) \right] \\ = & E (\mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}')), \end{aligned}$$

Thus, the numerator of  $S_{1,s}$  has the following equivalence, which shows its connection with the noncensored case:

$$\begin{aligned} & E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \|_{d_{\mathbf{X}_{-\alpha}}} \mathbf{1}\{T \in [l_{s-1}, l_s]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} \right] \\ = & E \left[ \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \|_{d_{\mathbf{X}_{-\alpha}}} E (\mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}')) \right] \\ = & E \left[ E \left( \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \|_{d_{\mathbf{X}_{-\alpha}}} \mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}') \right) \right] \\ = & E \left[ \|X_\alpha - X'_\alpha\|_{d_{X_\alpha}} \| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \|_{d_{\mathbf{X}_{-\alpha}}} \mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} \right] \end{aligned}$$

Following similar proofs as above, the denominator of  $S_{1,s}$  is:

$$\begin{aligned} & E \left[ \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \mathbf{1}\{T \in [l_{s-1}, l_s]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} \right] \\ = & E \left[ E \left( \frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \mathbf{1}\{T \in [l_s, l_{s+1}]\} \mathbf{1}\{T' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}') \right) \right] \\ = & E [E (\mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} | (\mathbf{X}, \mathbf{X}'))] \\ = & E (\mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\}) \\ = & Pr\{Y \in [l_{s-1}, l_s]\} Pr\{Y' \in [l_{s-1}, l_s]\}. \end{aligned}$$

Thus,

$$\begin{aligned}
S_{1,s} &= \frac{E \left[ \left\| X_\alpha - X'_\alpha \right\|_{d_{X_\alpha}} \left\| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \right\|_{d_{\mathbf{X}_{-\alpha}}} \mathbf{1}\{Y \in [l_{s-1}, l_s]\} \mathbf{1}\{Y' \in [l_{s-1}, l_s]\} \right]}{\Pr\{Y \in [l_{s-1}, l_s]\} \Pr\{Y' \in [l_{s-1}, l_s]\}} \\
&= E \left[ \left\| X_\alpha - X'_\alpha \right\|_{d_{X_\alpha}} \left\| \mathbf{X}_{-\alpha} - \mathbf{X}'_{-\alpha} \right\|_{d_{\mathbf{X}_{-\alpha}}} \mid Y, Y' \in [l_{s-1}, l_s] \right]
\end{aligned}$$

Following a similar approach, we can prove that  $S_{2,s}$  and  $S_{3,s}$  have similar results as  $S_{1,s}$ .

Therefore, equation (3.3) holds.

## A.2 Estimation for $G(t|\mathbf{X})$

To estimate  $G(t|\mathbf{X})$ , the simplest approach is to assume that  $G(t|\mathbf{X}) = G(t)$ . Following Zhou and Zhu (2017), we will explore other cases for estimating  $G(t|\mathbf{X})$  and identify the corresponding conditions required for the proof of Theorems.

1. Assume that  $G(t|\mathbf{X}) = G(t)$ .

Similar assumptions were also employed by He et al. (2013), Song et al. (2014), Zhou and Zhu (2017), and Zhang et al. (2018). With this assumption, estimating  $G(t|\mathbf{X})$  amounts to estimating  $G(t)$ .

2. Instead of using  $G(t|\mathbf{X})$ , we use  $G_\alpha(t|X_\alpha) = \Pr(C > t|X_\alpha)$ .

This assumption was suggested by He et al. (2013) and Zhou and Zhu (2017). The direct utilization of  $G(t|\mathbf{X})$  may encounter computational challenges when dealing with high dimensional data. Additionally, this assumption leads to rank consistency and sure screening properties. In the subsequent discussions, we present various estimators for  $G_\alpha(t|X_\alpha)$ , depending on the characteristics of  $X_\alpha$ . For details, a) Regarding a categorical or discrete random variable  $X_\alpha$  with a finite number of potential values, we can estimate  $G_\alpha(t|X_\alpha)$  within each level or category of  $X_\alpha$  through the classical Kaplan-Meier method. For example, if  $X_\alpha$  has  $H$  categories, the estimator for each category can be denoted as  $\widehat{G}_\alpha(t|X_\alpha = h)$ , where  $h = 1, 2, \dots, H$ .

- b) A continuous random variable  $X_\alpha$  can be sliced into  $H$  nonoverlapping and non-

random intervals. Employing similar ideas as in the categorical/discrete case, we can estimate  $G_\alpha(t|X_\alpha)$  using the Kaplan-Meier method. We denote the slices as  $I_1, \dots, I_H$ , and assume that  $G_\alpha(t|X_\alpha = x) = G_\alpha(t|X_\alpha \in I_h)$  for  $x \in I_h$ . The estimator for each slice can be denoted as  $\widehat{G}_\alpha(t|X_\alpha \in I_h)$ .

3. It is possible to posit a semiparametric model, such as a proportional hazards model, and then estimate  $G(T|\mathbf{X})$  based on the fitted model. This approach is straightforward to employ, and as long as the model is correctly specified, and the resulting estimator of the survival function is consistent. This assumption was suggested by Lu and Li (2011).

4. Assume that  $G(T|\mathbf{X}) = G(T|B\mathbf{X})$ , where  $B$  is a  $d \times p$  matrix with  $d < n$ .

To directly estimate  $G(T|\mathbf{X})$ , the kernel smoothing method is a useful nonparametric method. However, the kernel smoothing method performs optimally only in low dimensional scenarios. Hence, it is necessary to reduce the dimension first. The simplest approach is double slicing proposed by Li (1991) to reduce the dimension of  $\mathbf{X}$ . It is equivalent to the usual sliced inverse regression (SIR) method, except that  $Y$  is partitioned within each subsample. Although the double slicing SIR is simple to use, it imposes some restrictive conditions and relies on a nonparametric estimation of the conditional survival function of  $Y$  given  $\mathbf{X}$ , which itself is computationally complicated. Xia et al. (2010) developed a dimension reduction method for censored survival data based on kernel estimation of the conditional hazard function, but this method may suffer from the curse of dimensionality when the number of predictors is large.

Once the dimension is reduced from  $p$  to  $r$ , following Beran (1981) and Dabrowska (1989), we apply  $d$ -dimension kernel smoothing on  $\widehat{B}_d\mathbf{X}$  to obtain the estimate of  $G(t|\widehat{B}_d\mathbf{X})$ .

$$G(t|\widehat{B}_d\mathbf{X}) = \prod_{i=1}^n \left\{ 1 - \frac{w_{ni}(x; h_n)}{\sum_{k=1}^n I(T_k \geq T_i) w_{nk}(x; h_n)} \right\}^{I(T_i \leq t, \delta_i=0)}$$

with

$$w_{ni}(x; h_n) = \frac{K(h_n^{-1}\widehat{B}_d(\mathbf{X}_i - x))}{\sum_{j=1}^n K(h_n^{-1}\widehat{B}_d(\mathbf{X}_j - x))} \quad i = 1, \dots, n,$$

where  $K$  is a density function, and  $\{h_n\}$  is the bandwidth.

Furthermore, the conditions required for proving the Theorems are as follows, case by case, in detail:

(C3.1) If  $G(t | \mathbf{X}) = G(t)$ , then we assume  $Pr(t \leq Y \leq C) \geq \tau_0 > 0$  for  $t \in [0, T_{\max}]$ , where  $T_{\max}$  represents the maximum follow-up time. Additionally,  $\sup\{t : Pr(Y > t) > 0\} \geq \sup\{t : pr(C > t) > 0\}$  and  $G'_\alpha(t)$ , the first derivative of  $G_\alpha(t)$ , is uniformly bounded.

(C3.2) If we use  $G(t|X_\alpha)$  instead of  $G(t | \mathbf{X})$ , and assume that  $X_\alpha$  is categorical or discrete, then we assume  $Pr(t \leq Y \leq C | X_\alpha = h) \geq \tau_0 > 0$  for  $t \in [0, T_{\max}]$ , where  $T_{\max}$  is the maximum follow-up time. Additionally,  $\sup\{t : Pr(Y > t | X_\alpha = h) > 0\} \geq \sup\{t : pr(C > t | X_\alpha = h) > 0\}$  and  $G'_\alpha(t | X_\alpha = h)$ , the first derivative of  $G_\alpha(t | X_\alpha = h)$ , is uniformly bounded.

(C3.3) If we use  $G(t|X_\alpha)$  instead of  $G(t | \mathbf{X})$ , and if  $X_\alpha$  is continuous, we slice the range of  $X_\alpha$  into  $H$  nonoverlapping and nonrandom intervals, denoted as  $I_1, \dots, I_H$ , and assume that  $G(t|X_\alpha = x) = G(t|X_\alpha \in I_h)$ , for  $x \in I_h$ . Then we assume  $Pr(t \leq Y \leq C | X_\alpha \in I_h) \geq \tau_0 > 0$  for  $t \in [0, T_{\max}]$ , where  $T_{\max}$  is the maximum follow-up time. Additionally,  $\sup\{t : Pr(Y > t | X_\alpha \in I_h) > 0\} \geq \sup\{t : pr(C > t | X_\alpha \in I_h) > 0\}$  and  $G'_\alpha(t | X_\alpha \in I_h)$ , the first derivative of  $G_\alpha(t | X_\alpha \in I_h)$ , is uniformly bounded.

(C3.4) We assume  $\inf_x Pr(t \leq Y \leq C | \mathbf{X}) \geq \tau_0 > 0$  for  $t \in [0, T_{\max}]$ , where  $T_{\max}$  is the maximum follow-up time.  $G(t | \mathbf{X})$  has first derivatives with respect to  $t$ , which are uniformly bounded away from infinity, and  $G(t | \mathbf{X})$  has bounded (uniformly in  $t$ ) second-order partial derivatives with respect to  $\mathbf{X}$ . Additionally,  $t_0 \leq \sup\{t : G(t | \mathbf{X}) > 0\} \leq t_1$  uniformly in  $\mathbf{X}$  for some positive constants  $t_0$  and  $t_1$ , and  $\sup\{t : Pr(Y > t | \mathbf{X}) > 0\} \geq$

$\sup\{t : G(t | \mathbf{X}) > 0\}$  almost surely for  $\mathbf{X}$ .

Conditions (C3.1)-(C3.4) are commonly found in the survival analysis literature to ensure that the Kaplan-Meier estimator and its inverse function are well behaved. Conditions (C3.1)-(C3.3) coincide with the conditions in Zhou and Zhu (2017). Condition (C3.4) is an extension of (C3.3). It is similar to (C6') in He et al. (2013), but the dimension of  $x$  considered in (C6') in He et al. (2013) is one dimension.

### A.3 Proofs of Theorems

Firstly, we establish Theorem 2 and demonstrate that the results hold dimension-free with regard to variables. Subsequently, we establish the marginal conclusion of Theorem 1 as a specific instance of Theorem 2, which can be proven without difficulty. Finally, we prove the corresponding outcome of Theorem 3.

**Lemma 1 (Hoeffding (1963), Deviation bound for U-statistics)** *Let  $g(\mathbf{U}_1, \dots, \mathbf{U}_r)$  be a kernel of the U-statistic  $U_n$ , where  $U_n = \frac{1}{\binom{n}{r}} \sum_{i_r^n} g(\mathbf{U}_{i_1}, \dots, \mathbf{U}_{i_r})$ ,  $n \geq r$ ,  $\binom{n}{r} = \frac{n!}{(n-r)!}$  and  $\sum_{i_r^n}$  is taken over all  $r$ -tuples  $i_1, \dots, i_r$  drawn without replacement from  $1, \dots, n$ . If  $a \leq g(\mathbf{U}_1, \dots, \mathbf{U}_r) \leq b$ , then for any  $t > 0$  and  $w = [n/r]$  the largest integer contained in  $n/r$ , the following bound holds:*

$$P\{U_n - EU_n \geq t\} \leq \exp(-2wt^2/(b-a)^2). \quad (\text{S0.1})$$

If we treat  $G(t|\mathbf{X})$  as  $G(t)$ , then we can follow the lemma:

**Lemma 2 (Lo and Singh (1986))** *Under Condition (C3), the Kaplan-Meier estimator  $\widehat{G}(\cdot)$  satisfies: (1)  $\sup_{0 \leq t \leq T} |\widehat{G}(t) - G(t)| = O\{(\frac{\log n}{n})^{\frac{1}{2}}\}$  almost surely;*

*(2)  $\frac{1}{\widehat{G}(t)} - \frac{1}{G(t)} = \frac{1}{nG(t)^2} \sum_{g=1}^n \xi(T_g, \delta_g, t) + R_n(t)$ , where  $\xi(T_g, \delta_g, t)$ ,  $g = 1, \dots, n$  are i.i.d. random variables with mean zero and  $\sup_{0 \leq t \leq T} |R_n(t)| = O\{(\frac{\log n}{n})^{\frac{3}{4}}\}$  almost surely;*

*(3)  $\sup_{0 \leq t \leq T} |\frac{1}{\widehat{G}(t)} - \frac{1}{G(t)}| = O\{(\frac{\log n}{n})^{\frac{1}{2}}\}$  almost surely.*



If we reduce the dimension of  $\mathbf{X}$  to  $\widehat{B}_d\mathbf{X}$ , following Beran (1981) we apply d-dimension kernel smoothing on  $\widehat{B}_d\mathbf{X}$  to obtain the estimate  $G_n(t|\widehat{B}_d\mathbf{X})$ .

Under suitable assumptions, Beran (1981) and Dabrowska (1989) established the consistency of their estimates at rates of convergence (which are slower than the square root of n rate) commonly observed in nonparametric regression. The estimation of the conditional distribution of the response variable, given the covariates in the presence of right censoring, typically use the estimator proposed by Beran (1981). This estimator is often referred to as the conditional Kaplan-Meier estimator, can be viewed as a smooth function of the kernel regression estimator that captures the conditional behavior of the observations (Dabrowska (1989)).

**Lemma 3 (Beran (1981), uniform consistency)** *Let  $H_{1n}(t|\mathbf{X} = x) = \sum_{i=1}^n I(T_i > t, \delta_i = 1)B_{ni}(x)$  and  $H_{2n}(t|\mathbf{X} = x) = \sum_{i=1}^n I(T_i > t)B_{ni}(x)$  be the strongly consistent estimators of  $H_1(t|x) = P(T > t, \delta = 1|\mathbf{X} = x)$  and  $H_2(t|x) = P(T > t|\mathbf{X} = x)$ , respectively, i.e., for  $\mu$ -almost all  $x$ ,  $|H_{in}(t^-|x) - H_i(t^-|x)| \rightarrow 0$  a.s. and  $|H_{in}(t^+|x) - H_i(t^+|x)| \rightarrow 0$  as  $n \rightarrow \infty$  with  $B_{ni}(x)$  is a random set of nonnegative weights depending on covariates only. Suppose  $L(z) < \sup\{t : H_2(t|\mathbf{X}) > 0\}$ . Then, for  $\mu$ -almost all  $x$ , as  $n \rightarrow \infty$ ,*

$$\sup_{0 \leq t \leq L(z)} |G_n(t|x) - G(t|x)| \rightarrow 0, \quad a.s.$$

We first prove Theorem 2, and demonstrate that the results are dimension-free of the variables step by step. Subsequently, from this result, we readily deduce the marginal conclusion presented in Theorem 1. To prove these theorems, it is imperative to initially establish Lemma 4. This lemma 4 establishes a uniform bound for the discrepancies between  $\widehat{u}_\alpha^*$  and the estimators  $u_\alpha^*$  acquired through the screening step, for all  $\alpha = 1, \dots, p$ . For random vectors  $\mathbf{U}$  and  $\mathbf{W}$  with dimensions  $d_u$  and  $d_w$  respectively, as expounded in Section 2.4 of the main paper, the following lemma ascertains the bound

for  $u_\alpha^*$  within the context of survival data. Specifically, by setting  $\mathbf{U} = X_\alpha$  and  $\mathbf{W} = (T, \mathbf{X}_{-\alpha})$ , we can establish Theorem 2; whereas assigning  $\mathbf{U} = X_\alpha$  and  $\mathbf{W} = T$  enables us to substantiate Theorem 1.

**Lemma 4 (Bound for  $u_\alpha^*$  in survival data)** *Under conditions (C1) and (C3), for any  $0 < (\nu + \gamma) < \frac{1}{2}$ , there exist positive constants  $c_1 > 0$  and  $c_2 > 0$  such that*

$$\begin{aligned} & Pr \left( |\hat{u}_\alpha^* - u_\alpha^*| \geq cn^{-\nu} \right) \\ & \leq O \left( [\exp \{-c_1 n^{1-2(\nu+\gamma)}\} + n \exp \{-c_2 n^\gamma \tau_0^2\}] \right). \end{aligned}$$

**Proof 1 (of Lemma 4)** *By the Cauchy-Schwartz inequality and condition (C3), we have*

$$\begin{aligned} S_1 &= E \left( \frac{\delta \delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')} \|\mathbf{U} - \mathbf{U}'\|_{d_u} \|\mathbf{W} - \mathbf{W}'\|_{d_w} \right) \\ &\leq \frac{1}{\tau_0^2} \{E[\|\mathbf{U} - \mathbf{U}'\|_{d_u}^2] E[\|\mathbf{W} - \mathbf{W}'\|_{d_w}^2]\}^{1/2} \quad (\text{S0.2}) \\ &\leq \frac{2}{\tau_0^2} \{E[\|\mathbf{U}\|_{d_u}^2] E[\|\mathbf{W}\|_{d_w}^2]\}^{1/2} = A < \infty \end{aligned}$$

*By condition (C1), we have that  $S_1$  is bounded by a constant  $A$ , and for any given  $\epsilon > 0$ , take  $n$  large enough such that  $\frac{S_1}{n} < 2\epsilon$ .*

*Step 1:*

$$\begin{aligned} \hat{S}_1 &= \frac{1}{n^2} \sum_{i,j=1}^n \frac{\delta_i \delta'_j}{\hat{G}_n(T_i|\mathbf{X}_i) \hat{G}_n(T'_j|\mathbf{X}'_j)} \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w}, \\ &= \frac{1}{n^2} \sum_{i,j=1}^n \frac{\delta_i \delta'_j}{G(T_i|\mathbf{X}_i) G(T'_j|\mathbf{X}'_j)} \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w} \\ &\quad + \frac{1}{n^2} \sum_{i,j=1}^n \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w} \delta_i \delta'_j \left( \frac{1}{\hat{G}_n(T_i|\mathbf{X}_i) \hat{G}_n(T'_j|\mathbf{X}'_j)} - \frac{1}{G(T_i|\mathbf{X}_i) G(T'_j|\mathbf{X}'_j)} \right), \\ &= \hat{S}_{11} + \hat{S}_{12} \end{aligned}$$

*For  $\hat{S}_{12}$ , we have*

$$\begin{aligned} \hat{S}_{12} &= \frac{1}{n^2} \sum_{i,j=1}^n \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w} \frac{\delta_i \delta'_j}{G(T_i|\mathbf{X}_i) G(T'_j|\mathbf{X}'_j)} \left( \frac{G(T_i|\mathbf{X}_i) G(T'_j|\mathbf{X}'_j)}{\hat{G}_n(T_i|\mathbf{X}_i) \hat{G}_n(T'_j|\mathbf{X}'_j)} - 1 \right), \\ &\leq \max_{i,j} \left| \frac{G(T_i|\mathbf{X}_i) G(T'_j|\mathbf{X}'_j)}{\hat{G}_n(T_i|\mathbf{X}_i) \hat{G}_n(T'_j|\mathbf{X}'_j)} - 1 \right| \frac{1}{n^2} \sum_{i,j=1}^n \frac{\delta_i \delta'_j}{G(T_j|\mathbf{X}_j) G(T'_j|\mathbf{X}'_j)} \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w} \\ &= \max_{i,j} \left| \frac{G(T_i|\mathbf{X}_i) G(T'_j|\mathbf{X}'_j)}{\hat{G}_n(T_i|\mathbf{X}_i) \hat{G}_n(T'_j|\mathbf{X}'_j)} - 1 \right| \hat{S}_{11} \end{aligned}$$

By Condition (C3) and Lemma 2 (or Lemma 3 combined with Taylor expansion), we

have that

$$\frac{G(T_i|\mathbf{X}_i)}{\widehat{G}_n(T_i|\mathbf{X}_i)} - 1 = \frac{1}{\widehat{G}(T_i|\mathbf{X}_i)}(G_n(T_i|\mathbf{X}_i) - \widehat{G}(T_i|\mathbf{X}_i)) \rightarrow 0,$$

when  $n$  goes to infinity. Based on the fact that

$$\frac{AB}{\widehat{A}\widehat{B}} - 1 = \left(\frac{A}{\widehat{A}} - 1\right)\left(\frac{B}{\widehat{B}} - 1\right) + \left(\frac{A}{\widehat{A}} - 1\right) + \left(\frac{B}{\widehat{B}} - 1\right)$$

and  $S_1$  is bounded by a constant, we have

$$\max_{i,j} \left| \frac{G(T_i|\mathbf{X}_i)G(T'_j|\mathbf{X}'_j)}{\widehat{G}_n(T_i|\mathbf{X}_i)\widehat{G}_n(T'_j|\mathbf{X}'_j)} - 1 \right| S_1 = o_p(1) < \epsilon, \text{ as } n \rightarrow \infty.$$

Our goal is to obtain the consistency of  $\widehat{S}_1$ , considering that

$$\begin{aligned} Pr(|\widehat{S}_1 - S_1| \geq 6\epsilon) &= Pr\left(\left|\widehat{S}_{11} + \max_{i,j} \left| \frac{G(T_i|\mathbf{X}_i)G(T'_j|\mathbf{X}'_j)}{\widehat{G}_n(T_i|\mathbf{X}_i)\widehat{G}_n(T'_j|\mathbf{X}'_j)} - 1 \right| \widehat{S}_{11} - S_1 \right| \geq 6\epsilon\right) \\ &\leq Pr\left(\left|(o_p(1) + 1)(\widehat{S}_{11} - S_1) + cS_1\right| \geq 6\epsilon\right) \\ &\leq Pr\left(|\widehat{S}_{11} - S_1| \geq \frac{5\epsilon}{o_p(1) + 1}\right) \\ &\leq Pr(|\widehat{S}_{11} - S_1| \geq 4\epsilon) \end{aligned}$$

To establish the uniform consistency of  $\widehat{S}_1$ , we need the uniform consistency of  $\widehat{S}_{11}$ .

Define

$$\widehat{S}_{11}^* = \frac{1}{n(n-1)} \sum_{i \neq j} \frac{\delta_i \delta'_j}{G(T_i|\mathbf{X}_i)G(T'_j|\mathbf{X}'_j)} \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w},$$

which is a usual  $U$ -statistic. For  $\widehat{S}_{11} = \widehat{S}_{11}^* \frac{n-1}{n}$ , it can be easily shown that

$$\begin{aligned} P(|\widehat{S}_{11} - S_1| \geq 4\epsilon) &= P\left(\left|\widehat{S}_{11}^* \frac{n-1}{n} - S_1 \frac{n-1}{n} - S_1 \frac{1}{n}\right| \geq 4\epsilon\right) \\ &\leq P\left(\left|\widehat{S}_{11}^* - S_1\right| \frac{n-1}{n} \geq 4\epsilon - S_1 \frac{1}{n}\right) \\ &\leq P(|\widehat{S}_{11}^* - S_1| \geq 2\epsilon). \end{aligned}$$

Therefore, we show the uniform consistency of  $\widehat{S}_{11}^*$ . Let  $g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) = \frac{\delta_i \delta'_j}{G(T_i|\mathbf{X}_i)G(T'_j|\mathbf{X}'_j)} \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w}$  be the kernel of  $\widehat{S}_{11}^*$ . We decompose the kernel

function  $g_1$  into two parts  $g_1 = g_1 \mathbf{1}(g_1 \leq M) + g_1 \mathbf{1}(g_1 > M)$ , where  $M$  will be specified later. Note that  $\widehat{S}_{11}^*$  can be written as

$$\begin{aligned} \widehat{S}_{11}^* &= \frac{1}{n(n-1)} \sum_{i \neq j} g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \mathbf{1}(g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \leq M) \\ &\quad + \frac{1}{n(n-1)} \sum_{i \neq j} g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \mathbf{1}(g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) > M) = \widehat{S}_{1,1}^* + \widehat{S}_{1,2}^*. \end{aligned}$$

Accordingly, we decompose  $S_1$  into two parts

$$\begin{aligned} S_1 &= E[g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \mathbf{1}(g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \leq M)] \\ &\quad + E[g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \mathbf{1}(g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) > M)] = S_{1,1} + S_{1,2}. \end{aligned}$$

Clearly,  $\widehat{S}_{1,1}^*$  and  $\widehat{S}_{1,2}^*$  are unbiased estimators of  $S_{1,1}$  and  $S_{1,2}$ , respectively.

By using the conclusion of Lemma 1 in this supplement, we can obtain the uniform consistency of the bounded two-order  $U$  statistic  $\widehat{S}_{1,1}^*$  directly. That is,

$$P\{|\widehat{S}_{1,1}^* - S_{1,1}| \geq \epsilon\} \leq 2 \exp(-2[n/2]\epsilon^2/M^2) \quad (\text{S0.3})$$

For  $S_{1,2}$ , with the Cauchy-Schwartz inequality

$$\begin{aligned} S_{1,2}^2 &\leq E\{g_1^2(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j)\} Pr\{g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) > M\} \\ &\leq E\{g_1^2(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j)\} * I\{g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) > M\} \end{aligned}$$

Using the fact that  $(a^2 + b^2)/2 \geq (a + b)^2/4 \geq |ab|$  and condition (C3), we have

$$\begin{aligned} &g_1^2(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \\ &= \left( \frac{\delta_i \delta'_j}{G(T_i | \mathbf{X}_i) G(T'_j | \mathbf{X}'_j)} \|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u} \|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w} \right)^2 \\ &\leq \frac{1}{\tau_0^4} [(\|\mathbf{U}_i\|_{d_u} + \|\mathbf{U}'_j\|_{d_u})(\|\mathbf{W}_i\|_{d_w} + \|\mathbf{W}'_j\|_{d_w})]^2 \\ &\leq \frac{1}{4\tau_0^4} [(\|\mathbf{U}_i\|_{d_u} + \|\mathbf{U}'_j\|_{d_u})^2 + (\|\mathbf{W}_i\|_{d_w} + \|\mathbf{W}'_j\|_{d_w})^2]^2 \\ &\leq \frac{1}{\tau_0^4} (\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{U}'_j\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 + \|\mathbf{W}'_j\|_{d_w}^2)^2 \end{aligned}$$

By condition (C1), we have  $E(g_1^2(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j)) < \infty$  then  $S_{1,2} \leq \epsilon/2$  when  $n$  is sufficiently large. Consequently,

$$Pr(|\widehat{S}_{1,2}^* - S_{1,2}| > \epsilon) \leq Pr(|\widehat{S}_{1,2}^*| > \epsilon/2). \quad (\text{S0.4})$$

It remains to bound the probability  $Pr(|\widehat{S}_{1,2}^*| > \epsilon/2)$ . We observe that the events satisfy

$$\{|\widehat{S}_{1,2}^*| > \epsilon/2\} \subseteq \{\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 > \frac{M\tau_0^2}{2}, \text{ for some } 1 \leq i \leq n\}. \quad (\text{S0.5})$$

Based on the fact that,

$$\begin{aligned} M &< \widehat{S}_{1,2}^* = \frac{1}{n(n-1)} \sum_{i \neq j} g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) \mathbf{1}(g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) > M) \\ &\leq E(\max_{i,j} g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j)) \mathbf{1}(g_1(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}'_j, \mathbf{W}'_j, \delta'_j) > M) \\ &\leq \frac{1}{\tau_0^2} (\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{U}'_j\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 + \|\mathbf{W}'_j\|_{d_w}^2). \end{aligned}$$

This means that if the event  $|\widehat{S}_{1,2}^*| > \epsilon/2$  occurs, the event  $\{\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 > \frac{M\tau_0^2}{2}, \text{ for some } 1 \leq i \leq n\}$  will occur. This verifies that (S0.5) is true.

By invoking the above condition (C1) that  $E\{\exp(s\|\mathbf{U}\|_{d_u}^2)\} < \infty$ ,  $E\{\exp(s\|\mathbf{W}\|_{d_w}^2)\} < \infty$ , and applying Markov's inequality for  $s > 0$ , there must exist a constant  $C$  such that

$$\begin{aligned} Pr(|\widehat{S}_{1,2}^*| > \epsilon/2) &\leq nPr(\|\mathbf{U}\|_{d_u}^2 + \|\mathbf{W}\|_{d_w}^2 \geq \frac{M\tau_0^2}{2}) \\ &\leq nPr(\|\mathbf{U}\|_{d_u}^2 \geq \frac{M\tau_0^2}{4}) + nPr(\|\mathbf{W}\|_{d_w}^2 \geq \frac{M\tau_0^2}{4}) \\ &\leq nE\left[\frac{\exp(s\|\mathbf{U}\|_{d_u}^2)}{\exp(sM\tau_0^2/4)}\right] + nE\left[\frac{\exp(s\|\mathbf{W}\|_{d_w}^2)}{\exp(sM\tau_0^2/4)}\right] \end{aligned}$$

Then we have

$$Pr(|\widehat{S}_{1,2}^*| > \epsilon/2) \leq 2nC \exp(-sM\tau_0^2/4).$$

If we choose that  $M = c_1 n^\gamma$ , with  $0 < \gamma < \frac{1}{2} - \nu$ , then

$$\begin{aligned} &Pr(|\widehat{S}_{11} - S_1| \geq 4\epsilon) \\ &\leq 2nC \exp(-sM\tau_0^2/4) + 2 \exp(-2[n/2]\epsilon^2/M^2) \\ &= 2nC \exp(-sc_1 n^\gamma \tau_0^2/4) + 2 \exp(-\epsilon^2 n^{1-2\gamma}). \end{aligned} \quad (\text{S0.6})$$

Combining (S0.6) and (S0.4), we have

$$\begin{aligned} &Pr(|\widehat{S}_1 - S_1| \geq 6\epsilon) \\ &\leq 2nC \exp(-cn^\gamma \tau_0^2/4) + 2 \exp(-\epsilon^2 n^{1-2\gamma}). \end{aligned} \quad (\text{S0.7})$$

Step 2. Next, we turn to  $\widehat{S}_2$ . We write  $S_2 = S_{2,1}S_{2,2}$ , where  $S_{2,1} = E\{\|\mathbf{U} - \mathbf{U}'\|_{d_u}\}$  and  $S_{2,2} = E\{\frac{\delta\delta'}{G(T|\mathbf{X})G(T'|\mathbf{X}')}\|\mathbf{W} - \mathbf{W}'\|_{d_w}\}$ . Correspondingly,  $\widehat{S}_2 = \widehat{S}_{2,1}\widehat{S}_{2,2}$ , where  $\widehat{S}_{2,1} = n^{-2}\sum_{i\neq j}\sum\|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u}$ , and  $\widehat{S}_{2,2} = n^{-2}\sum_{i\neq j}\sum\frac{\delta_i\delta'_j}{\widehat{G}(T_i|\mathbf{x}_i)\widehat{G}(T'_j|\mathbf{x}'_j)}\|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w}$ .

Following arguments for proving (S0.7) in the supplement, we can show that

$$\begin{aligned} Pr(|\widehat{S}_{2,1} - S_{2,1}| \geq 4\epsilon) &\leq 2nC \exp(-cn^\gamma\tau_0^2/4) + 2\exp(-\epsilon^2n^{1-2\gamma}), \text{ and} \\ Pr(|\widehat{S}_{2,2} - S_{2,2}| \geq 4\epsilon) &\leq 2nC \exp(-cn^\gamma\tau_0^2/4) + 2\exp(-\epsilon^2n^{1-2\gamma}). \end{aligned} \quad (\text{S0.8})$$

Condition (C1) ensures that  $S_{2,1} \leq \{E(\|\mathbf{U}_i - \mathbf{U}'_j\|_{d_u}^2)\}^{1/2} \leq \{4E(\|\mathbf{U}\|_{d_u}^2)\}^{1/2}$  and  $S_{2,2} \leq \{E(\|\mathbf{W}_i - \mathbf{W}'_j\|_{d_w}^2)\}^{1/2} \leq \{4E(\|\mathbf{W}\|_{d_w}^2)\}^{1/2}$  are uniformly bounded. That is,  $\max\{S_{2,1}, S_{2,2}\} \leq C$ , for some constant  $C$ . Using (S0.8) repeatedly, we can easily prove that

$$\begin{aligned} Pr\{(|\widehat{S}_{2,1} - S_{2,1})S_{2,2}| \geq \epsilon\} &\leq Pr(|\widehat{S}_{2,1} - S_{2,1}| \geq \epsilon/C) \\ &\leq 2\exp\{-\epsilon^2n^{1-2\gamma}/(16C^2)\} + 2nC \exp(-cn^\gamma\tau_0^2/4), \\ Pr(|S_{2,1}(\widehat{S}_{2,2} - S_{2,2})| \geq \epsilon) &\leq Pr(|\widehat{S}_{2,2} - S_{2,2}| \geq \epsilon/C) \\ &\leq 2\exp\{-\epsilon^2n^{1-2\gamma}/(16C^2)\} + 2nC \exp(-cn^\gamma\tau_0^2/4), \end{aligned} \quad (\text{S0.9})$$

and

$$\begin{aligned} &Pr\{(|\widehat{S}_{2,1} - S_{2,1})(\widehat{S}_{2,2} - S_{2,2})| \geq \epsilon\} \\ &\leq Pr(|\widehat{S}_{2,1} - S_{2,1}| \geq \sqrt{\epsilon}) + Pr(|\widehat{S}_{2,2} - S_{2,2}| \geq \sqrt{\epsilon}) \\ &\leq 4\exp(-\epsilon n^{1-2\gamma}/16) + 4nC \exp(-cn^\gamma\tau_0^2/4). \end{aligned} \quad (\text{S0.10})$$

It follows from Bonferroni's inequality that (S0.9) and (S0.10) imply that,

$$\begin{aligned} &Pr\left(|\widehat{S}_2 - S_2| \geq 3\epsilon\right) = Pr\left(|\widehat{S}_{2,1}\widehat{S}_{2,2} - S_{2,1}S_{2,2}| \geq 3\epsilon\right) \\ &\leq Pr\left\{(|\widehat{S}_{2,1} - S_{2,1})S_{2,2}| \geq \epsilon\right\} + Pr\left\{|S_{2,1}(\widehat{S}_{2,2} - S_{2,2})| \geq \epsilon\right\} \\ &\quad + Pr\left\{(|\widehat{S}_{2,1} - S_{2,1})(\widehat{S}_{2,2} - S_{2,2})| \geq \epsilon\right\} \\ &\leq 8\exp\{-\epsilon^2n^{1-2\gamma}/(16C^2)\} + 8nC \exp(-cn^\gamma\tau_0^2/4), \end{aligned} \quad (\text{S0.11})$$

where the last inequality holds when  $\epsilon$  is sufficiently small and  $C$  is sufficiently large.

Step 3. The uniform consistency of  $\widehat{S}_3$  remains to be shown. We first study the

following  $U$ -statistic:

$$\begin{aligned}
\widehat{S}_3^* &= \frac{1}{n(n-1)(n-2)} \sum_{i < j < l} \left\{ \frac{\delta_j' \delta_l''}{\widehat{G}_n(T_j' | \mathbf{X}_j') \widehat{G}_n(T_l'' | \mathbf{X}_l'')} \|\mathbf{U}_i - \mathbf{U}_j'\|_{d_u} \|\mathbf{W}_j' - \mathbf{W}_l''\|_{d_w} + \right. \\
&\quad \frac{\delta_j \delta_l''}{\widehat{G}_n(T_j' | \mathbf{X}_j') \widehat{G}_n(T_l'' | \mathbf{X}_l'')} \|\mathbf{U}_i - \mathbf{U}_l''\|_{d_u} \|\mathbf{W}_j' - \mathbf{W}_l''\|_{d_w} + \\
&\quad \frac{\delta_i \delta_l''}{\widehat{G}_n(T_i | \mathbf{X}_i) \widehat{G}_n(T_l'' | \mathbf{X}_l'')} \|\mathbf{U}_i - \mathbf{U}_j'\|_{d_u} \|\mathbf{W}_i - \mathbf{W}_l''\|_{d_w} + \\
&\quad \frac{\delta_i \delta_l''}{\widehat{G}_n(T_i | \mathbf{X}_i) \widehat{G}_n(T_l'' | \mathbf{X}_l'')} \|\mathbf{U}_l'' - \mathbf{U}_j'\|_{d_u} \|\mathbf{W}_i - \mathbf{W}_l''\|_{d_w} + \\
&\quad \frac{\delta_i \delta_j'}{\widehat{G}_n(T_i | \mathbf{X}_i) \widehat{G}_n(T_j' | \mathbf{X}_j')} \|\mathbf{U}_l'' - \mathbf{U}_j'\|_{d_u} \|\mathbf{W}_i - \mathbf{W}_j'\|_{d_w} + \\
&\quad \left. \frac{\delta_i \delta_j'}{\widehat{G}_n(T_i | \mathbf{X}_i) \widehat{G}_n(T_j' | \mathbf{X}_j')} \|\mathbf{U}_l'' - \mathbf{U}_i\|_{d_u} \|\mathbf{W}_i - \mathbf{W}_j'\|_{d_w} \right\} \\
&= \frac{6}{n(n-1)(n-2)} \sum_{i < j < l} g_3(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}_j', \mathbf{W}_j', \delta_j'; \mathbf{U}_l'', \mathbf{W}_l'', \delta_l''). \tag{S0.12}
\end{aligned}$$

Here,  $g_3(\mathbf{U}_i, \mathbf{W}_i, \delta_i; \mathbf{U}_j', \mathbf{W}_j', \delta_j'; \mathbf{U}_l'', \mathbf{W}_l'', \delta_l'')$  is the kernel of  $U$ -statistic  $\widehat{S}_3^*$ . Following the arguments to deal with  $\widehat{S}_1^*$ , we decompose  $g_3$  into two parts:  $g_3 = g_3 \mathbf{1}(g_3 > M) + g_3 \mathbf{1}(g_3 \leq M)$ . Accordingly,

$$\begin{aligned}
\widehat{S}_3^* &= \frac{6}{n(n-1)(n-2)} \sum_{i < j < l} g_3 \mathbf{1}(g_3 \leq M) + \frac{6}{n(n-1)(n-2)} \sum_{i < j < l} g_3 \mathbf{1}(g_3 > M) \\
&= \widehat{S}_{3,1}^* + \widehat{S}_{3,2}^*,
\end{aligned}$$

$$S_3 = E \{g_3 \mathbf{1}(g_3 \leq M)\} + E \{g_3 \mathbf{1}(g_3 > M)\} = S_{3,1} + S_{3,2}.$$

Following similar arguments for proving ( ), we can show that

$$\Pr(|\widehat{S}_{3,1}^* - S_{3,1}| \geq \epsilon) \leq 2 \exp(-2\epsilon^2 m' / M^2), \tag{S0.13}$$

where  $m' = \lfloor n/3 \rfloor$  because  $\widehat{S}_{3,1}^*$  is a third-order  $U$ -statistic.

Then we deal with  $\widehat{S}_{3,2}^*$ . We observe that  $g_3(\mathbf{U}_i, \mathbf{W}_i; \mathbf{U}_j', \mathbf{W}_j'; \mathbf{U}_l'', \mathbf{W}_l'') \leq \frac{2}{3\tau_0^2} (\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{U}_j'\|_{d_u}^2 + \|\mathbf{U}_l''\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 + \|\mathbf{W}_j'\|_{d_w}^2 + \|\mathbf{W}_l''\|_{d_w}^2)$ , which will be smaller than  $M$  if  $\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 \leq \frac{M\tau_0^2}{2}$ , for all  $1 \leq i \leq n$ . Thus, for any  $\epsilon > 0$ , the events satisfy

$$\{|\widehat{S}_{3,2}^*| > \epsilon/2\} \subseteq \{\|\mathbf{U}_i\|_{d_u}^2 + \|\mathbf{W}_i\|_{d_w}^2 > M\tau_0^2/2, \text{ for some } 1 \leq i \leq n\}.$$

By using similar arguments to prove (S0.6), it follows that

$$Pr(|\widehat{S}_{3,2}^* - S_{3,2}| > \epsilon) \leq Pr(|\widehat{S}_{3,2}^*| > \epsilon/2) \leq 2nC \exp(-cM\tau_0^2/4). \quad (\text{S0.14})$$

Then, we combine the results (S0.13) and (S0.14) with  $M = c_1 n^\gamma$  for some  $0 < \gamma < 1/2 - \nu$  to obtain that

$$Pr\left(\left|\widehat{S}_3^* - S_3\right| \geq 2\epsilon\right) \leq 2 \exp(-2\epsilon^2 n^{1-2\gamma}/3) + 2nC \exp(-cn^\gamma \tau_0^2/4). \quad (\text{S0.15})$$

By the definition of  $\widehat{S}_3$ ,

$$\widehat{S}_3 = \frac{(n-1)(n-2)}{n^2} \left\{ \widehat{S}_3^* + \frac{1}{(n-2)} \widehat{S}_1^* \right\}.$$

Thus, using similar techniques to deal with  $\widehat{S}_1$ , we can obtain that

$$Pr\left(\left|\widehat{S}_3 - S_3\right| \geq 4\epsilon\right) = Pr\left\{\left|\frac{(n-1)(n-2)}{n^2} (\widehat{S}_3^* - S_3) - \frac{3n-2}{n^2} S_3 + \frac{n-1}{n^2} (\widehat{S}_{11}^* - S_1) + \frac{n-1}{n^2} S_1\right| \geq 4\epsilon\right\}.$$

Using similar arguments for dealing with  $S_1$ , we can show that  $S_3$  is also uniformly bounded. Taking  $n$  large enough such that  $\{(3n-2)/n^2\}S_3 \leq \epsilon$  and  $\{(n-1)/n^2\}S_1 \leq \epsilon$ , then

$$\begin{aligned} Pr(|\widehat{S}_3 - S_3| \geq 4\epsilon) &\leq Pr(|\widehat{S}_3^* - S_3| \geq \epsilon) + Pr\{|\widehat{S}_{11}^* - S_1| \geq \epsilon\} \\ &\leq 4 \exp(-\epsilon^2 n^{1-2\gamma}/6) + 4nC \exp(-cn^\gamma \tau_0^2/4). \end{aligned} \quad (\text{S0.16})$$

The last inequality follows from (S0.7) and (S0.15). This, together with (S0.7), (S0.11) and Bonferroni's inequality, implies

$$\begin{aligned} &Pr\{|\widehat{S}_1 + \widehat{S}_2 - 2\widehat{S}_3 - (S_1 + S_2 - 2S_3)| \geq \epsilon\} \\ &\leq Pr(|\widehat{S}_1 - S_1| \geq \epsilon/4) + Pr(|\widehat{S}_2 - S_2| \geq \epsilon/4) + Pr(|\widehat{S}_3 - S_3| \geq \epsilon/4) \\ &= O\left\{\exp(-c_1 \epsilon^2 n^{1-2\gamma}) + n \exp(-c_2 n^\gamma \tau_0^2)\right\}, \end{aligned} \quad (\text{S0.17})$$

for some positive constants  $c_1$  and  $c_2$ . The convergence rate of the numerator of  $\widehat{u}_\alpha^*$ ,  $\alpha = 1, \dots, p$  is now achieved. Following similar arguments, we can obtain the convergence rate of the denominator. In effect the convergence rate of  $\widehat{u}_\alpha^*$  has the same form as (S0.17).



**Proof 2 (of Theorem 2)** Based on the bound of  $\widehat{u}^*$ , where  $\mathbf{U} = X_\alpha$  and  $\mathbf{W} = (T, \mathbf{X}_{-\alpha})$ , we have

$$Pr(|\widehat{u}_\alpha^* - u_\alpha^*| \geq \epsilon) \leq O \left\{ \exp(-c_1 \epsilon^2 n^{1-2\gamma}) + n \exp(-c_2 n^\gamma \tau_0^2) \right\},$$

for  $\alpha = 1, 2, \dots, p$ , with  $c$  being the positive constants. If we take  $\epsilon = cn^{-\nu}$  with  $0 < (\nu + \gamma) < \frac{1}{2}$ , then we have

$$\begin{aligned} Pr(\max_{1 \leq \alpha \leq p} |\widehat{u}_\alpha^* - u_\alpha^*| \geq cn^{-\nu}) &\leq p \max_{1 \leq \alpha \leq p} Pr(|\widehat{u}_\alpha^* - u_\alpha^*| \geq cn^{-\nu}) \\ &\leq O \left\{ p \left[ \exp(-c_1 n^{1-2(\nu+\gamma)}) + n \exp(-c_2 n^\gamma \tau_0^2) \right] \right\}. \end{aligned}$$

Denote  $\zeta^* = \min_{\alpha \in \mathcal{D}^*} u_\alpha^* - \max_{\alpha \in \widehat{\mathcal{D}}^*} u_\alpha^*$ , under condition (C2\*),  $\zeta^* \geq 2cn^{-\nu} \geq 0$ , Then

$$\begin{aligned} &Pr(\max_{\alpha \in \widehat{\mathcal{D}}^*} \widehat{u}_\alpha^* < \min_{\alpha \in \mathcal{D}^*} \widehat{u}_\alpha^*) \\ &= 1 - Pr(\max_{\alpha \in \widehat{\mathcal{D}}^*} \widehat{u}_\alpha^* \geq \min_{\alpha \in \mathcal{D}^*} \widehat{u}_\alpha^*) \\ &= 1 - Pr(\max_{\alpha \in \widehat{\mathcal{D}}^*} \widehat{u}_\alpha^* - \max_{\alpha \in \widehat{\mathcal{D}}^*} u_\alpha^* \geq \min_{\alpha \in \mathcal{D}^*} \widehat{u}_\alpha^* - \min_{\alpha \in \mathcal{D}^*} u_\alpha^* + \zeta^*) \\ &\geq 1 - \left[ Pr(\max_{\alpha \in \widehat{\mathcal{D}}^*} |\widehat{u}_\alpha^* - u_\alpha^*| \geq \zeta^*/2) + Pr(\max_{\alpha \in \mathcal{D}^*} |\widehat{u}_\alpha^* - u_\alpha^*| \geq \zeta^*/2) \right] \\ &\geq 1 - O \left\{ 2p \left[ \exp(-c_1 n^{1-2(\nu+\gamma)}) + n \exp(-c_2 n^\gamma \tau_0^2) \right] \right\}. \end{aligned}$$

Next we show that

$$Pr(\mathcal{D}^* \subseteq \widehat{\mathcal{D}}^*) \geq 1 - O(s_n \left[ \exp(-c_1 n^{1-2(\nu+\gamma)}) + n \exp(-c_2 n^\gamma \tau_0^2) \right]),$$

where  $s_n$  is the cardinality of  $\mathcal{D}^*$ .

If  $\mathcal{D}^* \not\subseteq \widehat{\mathcal{D}}^*$ , there must exist an  $\alpha \in \mathcal{D}^*$  such that  $u_\alpha^* \geq 2cn^{-\nu}$ , but  $\widehat{u}_\alpha^* \leq cn^{-\nu}$  to ensure that  $|u_\alpha^* - \widehat{u}_\alpha^*| \geq cn^{-\nu}$  for some  $\alpha \in \mathcal{D}^*$ . This means that the event satisfies  $\mathcal{D}^* \not\subseteq \widehat{\mathcal{D}}^* \subseteq \{|u_\alpha^* - \widehat{u}_\alpha^*| \geq cn^{-\nu}, \text{ for some } \alpha \in \mathcal{D}^*\}$ . Then we have  $\mathcal{E}_n = \{\max_{\alpha \in \mathcal{D}^*} |u_\alpha^* - \widehat{u}_\alpha^*| \leq cn^{-\nu}\} \subseteq \{\mathcal{D}^* \subseteq \widehat{\mathcal{D}}^*\}$ .

$$\begin{aligned} Pr(\mathcal{D}^* \subseteq \widehat{\mathcal{D}}^*) &\geq Pr(\mathcal{E}_n) = 1 - Pr(\mathcal{E}_n^c) \\ &= 1 - Pr(\min_{\alpha \in \mathcal{D}^*} |u_\alpha^* - \widehat{u}_\alpha^*| \geq cn^{-\nu}) \\ &= 1 - s_n Pr(|u_\alpha^* - \widehat{u}_\alpha^*| \geq cn^{-\nu}) \\ &\geq 1 - O(s_n \left[ \exp(-c_1 n^{1-2(\nu+\gamma)}) + n \exp(-c_2 n^\gamma \tau_0^2) \right]). \end{aligned}$$

To prove Theorem 1, we can use a special case of Theorem 2, with  $\mathbf{U} = X_\alpha$ , and  $\mathbf{W} = T$ . It includes the rank consistence of marginal measure  $\hat{u}$ , and the sure screening property. The details of Theorem 1 are included in the main paper.

To prove Theorem 3, apart from the additional conditions included in the main manuscript, the following Lemma 5 is also needed:

**Lemma 5 (Bound for  $u_{\alpha,s}^{\mathcal{J}}$  in survival data)** *For slice  $s$ , under conditions (C1), (C3), (C4\*\*), and (C5\*\*), for any  $0 < \gamma < \frac{1}{2} - \nu$ , there exist positive constants  $c_1 > 0$  and  $c_2 > 0$  such that*

$$\begin{aligned} & Pr \left( |\hat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{\mathcal{J}}| \geq cn^{-\nu} \right) \\ & \leq O \left( \left[ \exp \left\{ -c_1 \tau_0^4 n^{1-2(\nu+\gamma)} \right\} + n \exp \left( -c_2 n^\gamma \tau_0^2 \right) \right] \right). \end{aligned}$$

**Proof 3 (of Lemma 5)** *Similar to the bound proof for  $\hat{S}_1$  in Lemma 4, we can obtain the bound for the numerator of the first term  $\hat{S}_{1,s}$  in  $\hat{u}_{\alpha,s}^{\mathcal{J}}$ , and denote it as  $\hat{S}_{1,s,n}$ , except that each random variable is in the  $s$ -th slice, i.e.  $i, j, l \in \mathcal{V}_s$ . That is*

$$\begin{aligned} & Pr \left( |\hat{S}_{1,s,n} - S_{1,s,n}| \geq 6\epsilon \right) \\ & \leq 2nC \exp(-cn^\gamma \tau_0^2 / 4) + 2 \exp(-\epsilon^2 n^{1-2\gamma}). \end{aligned} \tag{S0.18}$$

*Consider the denominator of  $\hat{S}_{1,s}$  in  $\hat{u}_{\alpha,s}^{\mathcal{J}}$ , denoted as  $\hat{S}_{1,s,d}$ ,*

$$\begin{aligned} \hat{S}_{1,s,d} &= \frac{1}{n^2} \sum_{i,j \in \mathcal{V}_s} \frac{\delta_i \delta'_j}{\hat{G}_n(T_i | \mathbf{X}_i) \hat{G}_n(T'_j | \mathbf{X}'_j)} \\ &= \frac{1}{n^2} \sum_{i,j \in \mathcal{V}_s} \frac{\delta_i \delta'_j}{G(T_i | \mathbf{X}_i) G(T'_j | \mathbf{X}'_j)} \\ &\quad + \frac{1}{n^2} \sum_{i,j \in \mathcal{V}_s} \delta_i \delta'_j \left( \frac{1}{\hat{G}_n(T_i | \mathbf{X}_i) \hat{G}_n(T'_j | \mathbf{X}'_j)} - \frac{1}{G(T_i | \mathbf{X}_i) G(T'_j | \mathbf{X}'_j)} \right), \\ &= \hat{S}_{11,d} + \hat{S}_{12,d} \end{aligned}$$

*Under condition (C3),  $\hat{S}_{11,d} \geq \tau_0^2$  has a lower bound, and  $\hat{S}_{12,d} \rightarrow 0$  as  $n$  is sufficiently*

large, so

$$\begin{aligned}
& Pr(|\widehat{S}_{1,s} - S_{1,s}| \geq 6\epsilon) \\
& \leq Pr(|\widehat{S}_{1,s,n} - S_{1,s,n}| \geq 6\epsilon\tau_0^2) \\
& \leq 2nC \exp(-cn^\gamma\tau_0^2/4) + 2\exp(-\epsilon^2\tau_0^4n^{1-2\gamma}) \\
& \leq 2nC \exp(-c_2n^\gamma\tau_0^2/4) + 2\exp(-c_1\epsilon^2\tau_0^4n^{1-2\gamma}).
\end{aligned} \tag{S0.19}$$

The rest of the proof for Lemma 5 is almost the same as that of Lemma 4 for  $u_\alpha^*$ .

**Proof 4 (of Theorem 3)** Define  $u_{\alpha,s}^{**} = \mathcal{I}_{t \in [l_{s-1}, l_s]} \{(X_\alpha, \mathbf{X}_{-\alpha}) \mid T = t\}$ , for a specific  $t \in [l_{s-1}, l_s]$ ,

$$\begin{aligned}
Pr(|\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{**}| \geq \epsilon) & \leq Pr(|\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{\mathcal{J}}| \geq \frac{\epsilon}{2}) + Pr(|u_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{**}| \geq \frac{\epsilon}{2}) \\
& = Pr(|\widehat{\mathcal{I}}\{(X_\alpha, \mathbf{X}_{-\alpha}) \mid T \in [l_{s-1}, l_s]\} - \mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha}) \mid T \in [l_{s-1}, l_s]\}| \geq \frac{\epsilon}{2}) \\
& \quad + Pr(|\mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha}) \mid T \in [l_{s-1}, l_s]\} - \mathcal{I}_{t \in [l_{s-1}, l_s]} \{(X_\alpha, \mathbf{X}_{-\alpha}) \mid T = y\}| \geq \frac{\epsilon}{2}) \\
& = I_1 + I_2.
\end{aligned}$$

In term  $I_1$ , for each  $s$ ,  $s = 1, \dots, S$ , the rank consistency of  $\widehat{u}_{\alpha,s}^{\mathcal{J}}$  can be obtained directly, by replacing  $(T, \mathbf{X}_{-\alpha})$  in Theorem 2 with  $\mathbf{X}_{-\alpha}$ . That is, under conditions (C1) and (C3),

$$I_1 = Pr(|\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{\mathcal{J}}| \geq \frac{\epsilon}{2}) \leq O \left\{ \exp \left( -\frac{c_1}{4} \tau_0^4 \epsilon^2 n^{1-2\gamma} \right) + n \exp \left( -c_2 n^\gamma \tau_0^2 \right) \right\}$$

For the  $s$ -th sample quantile slice, if the empirical cumulative distribution of  $F_n(t)$  and the cumulative distribution  $F(t)$  are sufficiently close, i.e. for any  $\epsilon > 0$ , the event  $A = \sup_y |F_n(t) - F(t)| < \frac{\epsilon}{2}$  occurs, then, we have event  $B = \{\frac{1}{S} - \epsilon \leq P(T \in [l_{s-1}, l_s]) \leq \frac{1}{S} + \epsilon\}$ , and event  $B$  meets the condition of (C5\*\*). In fact, for any  $\epsilon > 0$ , if the event  $A = \sup_y |F_n(t) - F(t)| < \frac{\epsilon}{2}$ , following Lemma 4 in Mai and Zou (2015), we have

$$\begin{aligned}
Pr(l_{s-1} \leq T < l_s) & = Pr\left(\frac{s-1}{S} \leq F_n(T) < \frac{s}{S}\right) \\
& \leq Pr\left(\frac{s-1}{S} - \frac{\epsilon}{2} \leq F(T) < \frac{s}{S} + \frac{\epsilon}{2}\right) \\
& = \frac{1}{S} + \epsilon.
\end{aligned}$$

Similarly, we have

$$\begin{aligned}
Pr(l_{s-1} \leq T < l_s) &= Pr(F(l_{s-1}) \leq F(T) < F(l_s)) \\
&= F(l_s) - F(l_{s-1}) \\
&\geq F_n(l_s) - \frac{\epsilon}{2} - (F_n(l_{s-1}) + \frac{\epsilon}{2}) \\
&= \frac{1}{S} - \epsilon.
\end{aligned}$$

Furthermore, we have  $Pr(B) \geq Pr(A) \geq 1 - 2 \exp(-\frac{1}{2}n\epsilon^2)$  by the Dvoretzky-Kiefer-Wolfowitz inequality.

Based on event  $B$ , i.e.  $\frac{1}{S} - \epsilon \leq P(T \in [l_{s-1}, l_s]) \leq \frac{1}{S} + \epsilon$ , the event  $D = \{|\mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\} - \mathcal{I}_{t \in [l_{s-1}, l_s]}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\}| \leq \frac{\epsilon}{2}\}$  occurs, i.e.  $B \subset D$ . Actually, for any  $t \in [l_{s-1}, l_s)$ ,

$$\begin{aligned}
&|\mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\} - \mathcal{I}_{t \in [l_{s-1}, l_s]}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\}| \\
&\leq \sup_{t \in [l_{s-1}, l_s)} \mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\} - \inf_{t \in [l_{s-1}, l_s)} \mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\}| \\
&\leq \frac{\epsilon}{2},
\end{aligned}$$

the last inequality follows from the condition (C5\*\*).

$$\begin{aligned}
Pr(D) &= Pr(|\mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\} - \mathcal{I}_{t \in [l_{s-1}, l_s]}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\}| \leq \frac{\epsilon}{2}) \\
&\geq Pr(B) \geq Pr(A) \geq 1 - 2 \exp(-\frac{1}{2}n\epsilon^2)
\end{aligned}$$

Therefore, we have that

$$\begin{aligned}
I_2 &= Pr(|\mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\} - \mathcal{I}_{t \in [l_{s-1}, l_s]}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\}| \geq \frac{\epsilon}{2}) \\
&= 1 - Pr(D) \leq 2 \exp(-\frac{1}{2}n\epsilon^2)
\end{aligned}$$

Combining the bounds for  $I_1$  and  $I_2$ , we have

$$Pr(|\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{**}| \geq \frac{\epsilon}{2}) \leq O \left\{ \exp(-c_1 \tau_0^4 \epsilon^2 n^{1-2\gamma}) + n \exp(-c_2 n^\gamma \tau_0^2) + \exp(-c_3 n \epsilon^2) \right\}$$

Similar to the proof of Theorem 2, if we take  $\epsilon = cn^{-\nu}$  for any  $0 < (\nu + \gamma) < \frac{1}{2}$ , there exists a positive constant  $c > 0$  such that

$$Pr\left(\max_{1 \leq \alpha \leq p} |\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{**}| \geq cn^{-\nu}\right) \leq O \left\{ p \left[ \exp(-c_1 \tau_0^4 n^{1-2(\gamma+\nu)}) + n \exp(-c_2 n^\gamma \tau_0^2) + \exp(-c_3 n^{1-2\nu}) \right] \right\}$$

Furthermore, we have

$$\begin{aligned}
& Pr(\max_{1 \leq \alpha \leq p} |\widehat{u}_\alpha^{**} - u_\alpha^{**}| \geq cn^{-\nu}) \\
& \leq Pr(\max_{1 \leq \alpha \leq p} |\frac{1}{S} \sum_{s=1}^S \widehat{u}_{\alpha,s}^{\mathcal{J}} - \frac{1}{S} \sum_{s=1}^S u_{\alpha,s}^{\mathcal{J}}| \geq cn^{-\nu}) \\
& \leq pPr(|\frac{1}{S} \sum_{s=1}^S \widehat{u}_{\alpha,s}^{\mathcal{J}} - \frac{1}{S} \sum_{s=1}^S u_{\alpha,s}^{\mathcal{J}}| \geq cn^{-\nu}) \\
& \leq pPr(\max_{1 \leq s \leq S} |\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{\mathcal{J}}| \geq cn^{-\nu}) \\
& = pPr(\max_{1 \leq s \leq S} |\widehat{\mathcal{I}}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\} - \mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\}| \geq cn^{-\nu}) \\
& \leq pPr(\sup_{t \in [l_{s-1}, l_s]} |\widehat{\mathcal{I}}\{(X_\alpha, \mathbf{X}_{-\alpha})|T \in [l_{s-1}, l_s]\} - \mathcal{I}\{(X_\alpha, \mathbf{X}_{-\alpha})|T = t\}| \geq cn^{-\nu}) \\
& \leq pPr(|\widehat{u}_{\alpha,s}^{\mathcal{J}} - u_{\alpha,s}^{\mathcal{J}}| \geq cn^{-\nu}) \\
& \leq O(p [\exp(-c_1 \tau_0^4 n^{1-2(\gamma+\nu)}) + n \exp(-c_2 n^\gamma \tau_0^2) + \exp(-c_3 n^{1-2\nu})])
\end{aligned}$$

Under condition (C2\*\*), denote  $\delta = \min_{\alpha \in \mathcal{D}^{**}} u_\alpha^{**} - \max_{\alpha \in \widehat{\mathcal{D}}^{**}} u_\alpha^{**}$ , then

$$\begin{aligned}
& Pr(\max_{\alpha \in \widehat{\mathcal{D}}^{**}} \widehat{u}_\alpha^{**} \geq \min_{\alpha \in \mathcal{D}^{**}} \widehat{u}_\alpha^{**}) \\
& = Pr(\max_{\alpha \in \widehat{\mathcal{D}}^{**}} \widehat{u}_\alpha^{**} - \max_{\alpha \in \mathcal{D}^{**}} u_\alpha^{**} \geq \min_{\alpha \in \mathcal{D}^{**}} \widehat{u}_\alpha^{**} - \min_{\alpha \in \mathcal{D}^{**}} u_\alpha^{**} + \delta) \\
& \leq Pr(\max_{\alpha \in \widehat{\mathcal{D}}^{**}} |\widehat{u}_\alpha^{**} - u_\alpha^{**}| \geq \delta/2) + Pr(\max_{\alpha \in \mathcal{D}^{**}} |\widehat{u}_\alpha^{**} - u_\alpha^{**}| \geq \delta/2) \\
& \leq 2p [\exp(-c_1 \tau_0^4 n^{1-2(\gamma+\nu)}) + n \exp(-c_2 n^\gamma \tau_0^2) + \exp(-c_3 n^{1-2\nu})].
\end{aligned}$$

Under conditions (C1), (C3), (C2\*\*), (C4\*\*), and (C5\*\*), similar to the proof of

Theorem 2, we have that

$$Pr(\mathcal{D}^{**} \subseteq \widehat{\mathcal{D}}^{**}) \geq 1 - O(s_n [\exp(-c_1 \tau_0^4 n^{1-2(\gamma+\nu)}) + n \exp(-c_2 n^\gamma \tau_0^2) + \exp(-c_3 n^{1-2\nu})]),$$

where  $s_n$  is the cardinality of  $\mathcal{D}^{**}$ .

#### A.4 Additional Simulation Examples

In this section, additional simulation examples and real data analysis results are presented.

Table 1 shows the results of Example 1 in the main text with  $p = 2000$ .

**Example 1** Assuming a general transformation model for survival time  $T$  described in Song et al. (2014) and Liu et al. (2018), we have  $H(T) = -\beta^\top \mathbf{X} + \epsilon$ , where  $H(t) = \log\{0.5(\exp(2t) - 1)\}$ , and  $\epsilon \sim N(0, 1)$ . Here,  $\beta^\top = (-1, -0.9, \mathbf{0}_6, 0.8, 1, \mathbf{0}_{p-10})$ , with  $\mathbf{0}_p$  referring to a zero vector of length  $p$ , such that only four predictors are active.  $\mathbf{X} = (X_1, \dots, X_p)^\top \sim N(0, \Sigma)$ , and  $\Sigma = (\sigma_{ij})_{p \times p}$  with  $\sigma_{ij} = 0.8^{|i-j|}$ . The censoring time  $C \sim U(0, 100)$ . We set  $n = 100, 200, p = 1000, 2000$ . The results are exhibited in Table 2.

**Example 2** This example follows Example 2 in main text. Tables (3)-(13) contain the additional simulation results for Example 2 (a) and the simulation results for Example 2 (b)-(d).

Table 14 represents the identified gene IDs based on different variable screening methods utilized on the training set for the real data analysis presented in the main text.

**Example 3** This is an additional illustration of real data analysis. We applied our proposed procedures to diffuse large-B-cell lymphoma(DLBCL) dataset, which was studied by Rosenwald et al. (2002). To identify genes that have an influence on patient survival risk, DLBCL is one of the most common types of lymphoma in adults in the United

Table 1: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 1 in main text with  $p = 2000$ .

	$\mathcal{P}_s$					$\mathcal{P}_a$	$\mathcal{P}_s$					$\mathcal{P}_a$						
	$X_1$	$X_2$	$X_3$	$X_4$	$X_5$	ALL	$X_1$	$X_2$	$X_3$	$X_4$	$X_5$	ALL	$X_1$	$X_2$	$X_3$	$X_4$	$X_5$	ALL
	$n = 100, p = 2000, CR \approx 0.019$						$n = 200, p = 2000, CR \approx 0.019$											
<i>CDC</i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>CDC<sub>1</sub></i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>CDC<sub>2</sub></i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>QaSIS</i>	0.935	1.000	0.995	0.990	0.945	0.870	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>CSIR</i>	0.525	0.755	0.775	0.705	0.600	0.325	0.990	0.995	1.000	0.995	0.985	0.970	1.000	1.000	1.000	1.000	1.000	1.000
<i>RCDCS</i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>CRSIS</i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>SVSIR</i>	0.960	1.000	0.995	0.995	0.980	0.945	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>CRIS</i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>KM</i>	0.990	0.990	1.000	1.000	0.990	0.985	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>IPOD</i>	0.985	1.000	1.000	1.000	0.995	0.980	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>LQ</i>	0.965	0.980	0.960	0.975	0.975	0.870	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
<i>ISIS<sub>C</sub></i>	0.840	0.888	0.934	0.918	0.824	0.530	0.984	0.976	0.986	0.986	0.978	0.910	1.000	1.000	1.000	1.000	1.000	1.000

States (Rosenwald et al., 2002). The dataset is available from: <http://www-stat.stanford.edu/~tibbs/superpc/staudt.html>. DLBCL data has also been utilized by Zhu et al. (2011), He et al. (2013) and Zhou and Zhu (2017). It consists of survival times for  $n = 240$  DLBCL patients following chemotherapy, as well as gene expression measurements for  $p = 7399$  genes obtained from complementary DNA microarrays for each individual patient. The survival rate after standard chemotherapy is approximately  $35 \sim 40\%$ . During the exploratory data analysis, all gene expression levels have been standardized to have a mean of zero and a standard deviation of one. Following Rosenwald et al. (2002), Zhu et al. (2011), He et al. (2013) and Zhou and Zhu (2017), we divided this dataset into a training set with  $n_1 = 160$  patients and a testing set with the remaining  $n_2 = 80$  patients. Specifically, various screening procedures were applied to the training data, and select out  $\lceil n_1 / \log(n_1) \rceil = 31$  genes. These selected covariates are denoted by  $\mathbf{X}_{\hat{\mathcal{D}}}$ . Table 15 in Appendix A.4 summarizes the first 31 screened genes. Gene IDs 2409 and 1969 were chosen by the one-stage screening procedures, while genes with IDs 2308 and 2306 were selected by all two-stage screening procedures. This suggests a strong association between these genes and patient survival times, although they were not chosen by the marginal methods.

The  $p$ -values of the log-rank test can be found in the final column of Table 16. Based on the  $p$ -values, all screening methods successfully identify subgroups that exhibit differences in survival times. Table 16 summarizes the  $C$ -statistic, along with its standard deviations, and the lower and upper bounds for different methods. The standard deviation (SD) of the  $C$ -statistic is obtained using a perturbation resampling method with 200 replicates. Figure 1 illustrates the  $C$ -statistic for various marginal methods, one-stage methods and two-stage methods. These methods demonstrate moderate predictive power, as the  $C$ -statistic is approximately 0.80 throughout. Notably, the one-stage and two-stage procedures yield

higher  $C$ -statistic values compared to the marginal methods, with the one-stage procedure exhibiting particularly promising results. The green dashed lines in the figure are generally similar or above the marginal line. Among all the methods, the CSIR and CRSIS methods exhibit slightly better overall performance. This confirms the importance of focusing on gene IDs 2308, 2306, 2409 and 1969, which were selected by the one-stage and two-stage procedures, in future research. Furthermore, the proposed method, which combines the CSIR measure in the marginal stage and the DC-based measure in the one-stage or two-stage, enhances the performance of CSIR in Zhou and Zhu (2017). Specifically,  $CRSIS_2^M$  stands out as the best among all competitors.

## Bibliography

- Bair, E. and Tibshirani, R. (2004). Semi-supervised methods to predict patient survival from gene expression data. *PLoS Biol* **2**, e108.
- Beran, R. (1981). *Nonparametric regression with randomly censored survival data*.
- Cook, R. D. and Ni, L. (2005). Sufficient dimension reduction via inverse regression: A minimum discrepancy approach. *Journal of the American Statistical Association* **100**, 410–428.
- Dabrowska, D. M. (1989). Uniform consistency of the kernel conditional kaplan-meier estimate. *The Annals of Statistics* **17**, 1157–1167.
- Gannoun, A. and Saracco, J. (2003). An asymptotic theory for sir  $\alpha$  method. *Statistica Sinica* **13**, 297–310.
- He, X., Wang, L. and Hong, H. G. (2013). Quantile-adaptive model-free variable screening for high-dimensional heterogeneous data. *Annals of Statistics* **41**, 342–369.



- Hoeffding, W. (1963). Probability inequalities for sums of bounded random variables. *Journal of the American Statistical Association*, **58**, 13–30.
- Li, K. C. (1991). Sliced inverse regression for dimension reduction. *Journal of the American Statistical Association* **86**, 316–327.
- Li, R., Zhong, W. and Zhu, L. (2012). Feature screening via distance correlation learning. *Journal of the American Statistical Association* **107**, 1129–1139.
- Liu, J., Zhang, J. and Zhao, X. (2018). A new nonparametric screening method for ultrahigh-dimensional survival data. *Computational Statistics and Data Analysis* **119**, 74–85.
- Lo, S. H. and Singh. K. (1986). The product-limit estimator and the bootstrap: some asymptotic representations. *Probability Theory and Related Fields* **71**, 455–465.
- Lu, W. and Li. L. (2011). Sufficient dimension reduction for censored regressions. *Biometrics* **67**, 513–523.
- Mai, Q. and Zou. H. (2015). The fused kolmogorov filter: A nonparametric model-free screening method. *The Annals of Statistics* **43**, 1471–1497.
- Rosenwald, A., Wright, G. and Chan, W. C. Connors, J. M. Campo, E. Fisher, R. I. Gascoyne, R. D. Muller-Hermelink, H. K. Smeland, E. B. Giltnane, J. M. . . . Staudt, L. M. (2002). The use of molecular profiling to predict survival after chemotherapy for diffuse large-b-cell lymphoma. *New England Journal of Medicine* **346**, 1937–1947.
- Song, R., Lu, W., Ma. S. and Jeng, X. J. (2014). Censored rank independence screening for high-dimensional survival data. *Biometrika* **101**, 799–814.

- Uno, H., Cai, T., Pencina, M. J., D'Agostino, R. B. and Wei, L. J. (2011). On the c-statistics for evaluating overall adequacy of risk prediction procedures with censored survival data. *Statistics in medicine* **30**, 1105–1117.
- Xia, Y., Zhang, D., and Xu, J. (2010). Dimension reduction and semiparametric estimation of survival models. *Journal of the American Statistical Association* **105**, 278–290.
- Zhang, J., Yin, G., Liu, Y. and Wu, Y. (2018). Censored cumulative residual independent screening for ultrahigh-dimensional survival data. *Lifetime data analysis* **24**, 273–292.
- Zhou, T. and Zhu, L. (2017). Model-free feature screening for ultrahigh dimensional censored regression. *Statistics and Computing* **27**, 947–961.
- Zhu, L. P., Li, L., Li, R. and Zhu, L. X. (2011). Model-free feature screening for ultrahigh-dimensional data. *Journal of the American Statistical Association* **106**, 1464–1475.

Table 2: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 1 in Appendix.

Model	$\mathcal{P}_s$					$\mathcal{P}_a$	Model	$\mathcal{P}_s$					$\mathcal{P}_a$
	$X_1$	$X_2$	$X_9$	$X_{10}$	ALL	$X_1$		$X_2$	$X_9$	$X_{10}$	ALL		
$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.008$							$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.008$						
<i>CDC</i>	1.000	1.000	0.994	1.000	0.994	0.994	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CDC</i> <sub>1</sub>	1.000	1.000	0.992	1.000	0.992	0.992	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CDC</i> <sub>2</sub>	1.000	1.000	0.992	1.000	0.992	0.992	1.000	1.000	1.000	1.000	1.000	1.000	
<i>QaSIS</i>	0.946	0.894	0.810	0.910	0.652	0.652	1.000	1.000	0.998	1.000	0.998	0.998	
<i>CSIR</i>	0.734	0.628	0.508	0.630	0.218	0.218	0.988	0.968	0.938	0.980	0.900	0.900	
<i>RCDCS</i>	1.000	1.000	0.992	0.998	0.992	0.992	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CRSIS</i>	1.000	1.000	0.998	1.000	0.998	0.998	1.000	1.000	1.000	1.000	1.000	1.000	
<i>SVSIR</i>	1.000	1.000	0.956	0.992	0.954	0.954	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CRIS</i>	1.000	1.000	0.998	1.000	0.998	0.998	1.000	1.000	1.000	1.000	1.000	1.000	
<i>KM</i>	0.990	0.976	0.946	0.988	0.916	0.916	1.000	1.000	1.000	1.000	1.000	1.000	
<i>IPOD</i>	0.996	0.982	0.934	0.986	0.904	0.904	1.000	1.000	1.000	1.000	1.000	1.000	
<i>LQ</i>	0.988	0.986	0.964	0.986	0.932	0.932	1.000	1.000	1.000	1.000	1.000	1.000	
<i>ISIS<sub>C</sub></i>	0.994	0.974	0.904	0.996	0.872	0.872	1.000	1.000	1.000	1.000	1.000	1.000	
$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.008$							$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.008$						
<i>CDC</i>	1.000	1.000	0.995	1.000	0.995	0.995	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CDC</i> <sub>1</sub>	1.000	1.000	0.995	1.000	0.995	0.995	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CDC</i> <sub>2</sub>	1.000	1.000	0.995	1.000	0.995	0.995	1.000	1.000	1.000	1.000	1.000	1.000	
<i>QaSIS</i>	0.920	0.825	0.615	0.825	0.410	0.410	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CSIR</i>	0.645	0.555	0.405	0.500	0.110	0.110	0.985	0.950	0.865	0.950	0.795	0.795	
<i>RCDCS</i>	1.000	1.000	0.991	1.000	0.991	0.991	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CRSIS</i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	
<i>SVSIR</i>	1.000	1.000	0.920	0.980	0.910	0.910	1.000	1.000	1.000	1.000	1.000	1.000	
<i>CRIS</i>	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	
<i>KM</i>	0.990	0.980	0.945	0.965	0.895	0.895	0.995	0.995	0.995	0.995	0.995	0.995	
<i>IPOD</i>	0.990	0.975	0.915	0.950	0.860	0.860	1.000	1.000	1.000	1.000	1.000	1.000	
<i>LQ</i>	0.990	0.990	0.955	0.970	0.905	0.905	1.000	1.000	1.000	1.000	1.000	1.000	
<i>ISIS<sub>C</sub></i>	0.990	0.964	0.898	0.970	0.834	0.834	1.000	1.000	1.000	1.000	1.000	1.000	

Table 3: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (a) Continued.

	$r = 0.5$					$r = 0.8$				
	$\mathcal{P}_s$		$\mathcal{P}_a$			$\mathcal{P}_s$		$\mathcal{P}_a$		
	$X_1$	$X_2$	$X_3$	$X_4$	ALL	$X_1$	$X_2$	$X_3$	$X_4$	ALL
	$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	0.916	0.93	0.914	0.026	0.02	0.678	0.714	0.670	0.028	0.000
<i>CDC</i> <sub>1</sub>	0.912	0.926	0.908	0.958	0.734	0.664	0.696	0.654	0.990	0.288
<i>CDC</i> <sub>2</sub>	0.912	0.926	0.908	0.990	0.762	0.664	0.698	0.654	0.988	0.290
<i>QaSIS</i>	0.562	0.566	0.530	0.314	0.038	0.102	0.106	0.110	0.356	0.000
<i>CSIR</i>	0.288	0.334	0.310	0.008	0.000	0.152	0.184	0.178	0.002	0.000
<i>RCDCS</i>	0.914	0.926	0.920	0.002	0.000	0.682	0.698	0.692	0.004	0.000
<i>CRSIS</i>	0.928	0.932	0.924	0.000	0.000	0.654	0.650	0.636	0.000	0.000
<i>SVSIR</i>	0.934	0.946	0.938	0.000	0.000	0.738	0.744	0.742	0.000	0.000
<i>CRIS</i>	0.930	0.946	0.944	0.000	0.000	0.770	0.762	0.766	0.000	0.000
<i>KM</i>	0.716	0.722	0.716	0.030	0.002	0.274	0.244	0.264	0.036	0.000
<i>IPOD</i>	0.770	0.782	0.774	0.024	0.008	0.338	0.306	0.314	0.044	0.000
<i>LQ</i>	0.616	0.616	0.578	0.064	0.006	0.114	0.096	0.106	0.048	0.000
<i>ISIS<sub>C</sub></i>	0.643	0.643	0.596	0.506	0.276	0.216	0.193	0.203	0.330	0.006
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	0.905	0.870	0.830	0.010	0.000	0.660	0.600	0.560	0.015	0.000
<i>CDC</i> <sub>1</sub>	0.905	0.870	0.820	0.940	0.615	0.635	0.585	0.535	0.950	0.170
<i>CDC</i> <sub>2</sub>	0.905	0.870	0.820	0.985	0.630	0.635	0.585	0.535	0.965	0.180
<i>QaSIS</i>	0.460	0.405	0.390	0.225	0.005	0.095	0.060	0.040	0.205	0.000
<i>CSIR</i>	0.210	0.265	0.280	0.000	0.000	0.070	0.120	0.150	0.000	0.000
<i>RCDCS</i>	0.870	0.875	0.830	0.000	0.000	0.580	0.655	0.575	0.000	0.000
<i>CRSIS</i>	0.905	0.900	0.860	0.000	0.000	0.590	0.575	0.515	0.000	0.000
<i>SVSIR</i>	0.900	0.875	0.855	0.000	0.000	0.650	0.695	0.635	0.000	0.000
<i>CRIS</i>	0.910	0.890	0.850	0.000	0.000	0.685	0.705	0.655	0.000	0.000
<i>KM</i>	0.660	0.655	0.575	0.030	0.005	0.175	0.180	0.180	0.005	0.000
<i>IPOD</i>	0.705	0.680	0.595	0.030	0.005	0.235	0.250	0.180	0.025	0.000
<i>LQ</i>	0.515	0.560	0.470	0.040	0.000	0.090	0.065	0.060	0.050	0.000
<i>ISIS<sub>C</sub></i>	0.486	0.486	0.450	0.300	0.103	0.150	0.173	0.140	0.196	0.000
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	1.000	1.000	1.000	0.035	0.035	0.955	0.945	0.900	0.055	0.025
<i>CDC</i> <sub>1</sub>	1.000	1.000	1.000	1.000	1.000	0.955	0.945	0.895	1.000	0.805
<i>CDC</i> <sub>2</sub>	1.000	1.000	1.000	1.000	1.000	0.955	0.945	0.895	1.000	0.805
<i>QaSIS</i>	0.905	0.915	0.890	0.575	0.410	0.195	0.245	0.205	0.655	0.015
<i>CSIR</i>	0.595	0.585	0.640	0.005	0.000	0.360	0.330	0.380	0.005	0.000
<i>RCDCS</i>	1.000	1.000	1.000	0.000	0.000	0.950	0.940	0.920	0.005	0.005
<i>CRSIS</i>	1.000	1.000	0.990	0.000	0.000	0.940	0.895	0.875	0.000	0.000
<i>SVSIR</i>	1.000	1.000	0.995	0.000	0.000	0.965	0.960	0.945	0.000	0.000
<i>CRIS</i>	1.000	1.000	1.000	0.000	0.000	0.975	0.975	0.950	0.000	0.000
<i>KM</i>	0.965	0.970	0.990	0.100	0.095	0.465	0.510	0.510	0.100	0.010
<i>IPOD</i>	0.975	0.975	0.995	0.090	0.085	0.575	0.555	0.560	0.135	0.015
<i>LQ</i>	0.975	0.980	0.975	0.140	0.135	0.220	0.285	0.275	0.170	0.010
<i>ISIS<sub>C</sub></i>	0.986	0.983	0.980	0.980	0.956	0.586	0.516	0.583	0.780	0.226
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
<i>CDC</i>	0.480	0.530	0.380	0.005	0.005	0.415	0.415	0.320	0.015	0.000
<i>CDC</i> <sub>1</sub>	0.465	0.505	0.355	0.670	0.060	0.400	0.405	0.315	0.765	0.030
<i>CDC</i> <sub>2</sub>	0.465	0.505	0.355	0.975	0.085	0.400	0.410	0.315	0.960	0.040
<i>QaSIS</i>	0.095	0.100	0.100	0.030	0.000	0.030	0.090	0.065	0.090	0.000
<i>CSIR</i>	0.210	0.305	0.330	0.000	0.000	0.075	0.130	0.155	0.000	0.000
<i>RCDCS</i>	0.835	0.820	0.740	0.000	0.000	0.535	0.570	0.495	0.005	0.000
<i>CRSIS</i>	0.87	0.85	0.76	0.00	0.00	0.525	0.520	0.475	0.000	0.000
<i>SVSIR</i>	0.890	0.865	0.840	0.000	0.000	0.635	0.660	0.615	0.000	0.000
<i>CRIS</i>	0.880	0.875	0.815	0.000	0.000	0.680	0.715	0.635	0.000	0.000
<i>KM</i>	0.645	0.670	0.575	0.025	0.000	0.180	0.180	0.175	0.005	0.000
<i>IPOD</i>	0.665	0.670	0.570	0.025	0.000	0.230	0.250	0.180	0.025	0.000
<i>LQ</i>	0.475	0.520	0.460	0.045	0.005	0.105	0.050	0.050	0.050	0.000
<i>ISIS<sub>C</sub></i>	0.393	0.410	0.343	0.253	0.046	0.126	0.140	0.123	0.183	0.000
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
<i>CDC</i>	0.750	0.675	0.820	0.010	0.005	0.770	0.780	0.760	0.040	0.010
<i>CDC</i> <sub>1</sub>	0.750	0.670	0.820	0.920	0.320	0.765	0.775	0.760	0.995	0.430
<i>CDC</i> <sub>2</sub>	0.750	0.670	0.825	1.000	0.360	0.765	0.775	0.760	1.000	0.435
<i>QaSIS</i>	0.515	0.425	0.510	0.380	0.045	0.150	0.210	0.210	0.555	0.005
<i>CSIR</i>	0.650	0.670	0.695	0.005	0.000	0.365	0.365	0.395	0.005	0.000
<i>RCDCS</i>	1.000	0.995	0.985	0.000	0.000	0.955	0.925	0.895	0.005	0.005
<i>CRSIS</i>	1.000	0.995	0.990	0.000	0.000	0.925	0.860	0.860	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.000	0.000	0.970	0.955	0.945	0.000	0.000
<i>CRIS</i>	1.000	1.000	1.000	0.000	0.000	0.965	0.975	0.950	0.000	0.000
<i>KM</i>	0.965	0.965	0.975	0.085	0.080	0.465	0.505	0.515	0.105	0.015
<i>IPOD</i>	0.975	0.970	0.980	0.095	0.090	0.550	0.545	0.550	0.130	0.015
<i>LQ</i>	0.975	0.960	0.955	0.130	0.115	0.260	0.285	0.255	0.180	0.000
<i>ISIS<sub>C</sub></i>	0.966	0.980	0.963	0.953	0.923	0.526	0.493	0.520	0.756	0.166

Table 4: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (b) with  $p = 1000$ .

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$
	$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	0.930	0.940	0.942	0.034	0.024	0.728	0.750	0.744	0.054	0.014
<i>CDC</i> <sub>1</sub>	0.916	0.940	0.936	0.952	0.772	0.720	0.734	0.724	0.980	0.372
<i>CDC</i> <sub>2</sub>	0.916	0.940	0.936	0.974	0.784	0.720	0.734	0.724	0.984	0.372
<i>QaSIS</i>	0.674	0.664	0.632	0.142	0.034	0.280	0.284	0.256	0.150	0.004
<i>CSIR</i>	0.316	0.330	0.312	0.002	0.000	0.142	0.166	0.134	0.002	0.000
<i>RCDCS</i>	0.916	0.926	0.922	0.046	0.034	0.688	0.714	0.704	0.076	0.018
<i>CRSIS</i>	0.932	0.956	0.936	0.000	0.000	0.680	0.646	0.664	0.000	0.000
<i>SVSIR</i>	0.958	0.958	0.944	0.020	0.016	0.782	0.788	0.776	0.028	0.002
<i>CRIS</i>	0.958	0.968	0.954	0.024	0.020	0.794	0.798	0.790	0.026	0.006
<i>KM</i>	0.780	0.774	0.788	0.030	0.010	0.400	0.432	0.386	0.058	0.002
<i>IPOD</i>	0.810	0.814	0.816	0.026	0.010	0.448	0.470	0.428	0.056	0.006
<i>LQ</i>	0.658	0.640	0.636	0.028	0.004	0.180	0.176	0.198	0.012	0.002
<i>ISIS<sub>C</sub></i>	0.796	0.733	0.773	0.676	0.323	0.370	0.320	0.350	0.440	0.020
	$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	1.000	1.000	1.000	0.124	0.124	0.986	0.983	0.968	0.199	0.175
<i>CDC</i> <sub>1</sub>	1.000	1.000	1.000	1.000	1.000	0.986	0.982	0.968	1.000	0.937
<i>CDC</i> <sub>2</sub>	1.000	1.000	1.000	1.000	1.000	0.986	0.982	0.968	1.000	0.937
<i>QaSIS</i>	0.978	0.982	0.978	0.394	0.372	0.624	0.578	0.588	0.498	0.120
<i>CSIR</i>	0.696	0.714	0.736	0.004	0.000	0.328	0.374	0.396	0.002	0.000
<i>RCDCS</i>	1.000	1.000	0.998	0.212	0.212	0.972	0.962	0.964	0.376	0.338
<i>CRSIS</i>	1.000	1.000	1.000	0.006	0.006	0.968	0.958	0.944	0.006	0.006
<i>SVSIR</i>	1.000	1.000	1.000	0.076	0.076	0.994	0.992	0.990	0.102	0.100
<i>CRIS</i>	1.000	1.000	1.000	0.066	0.066	0.996	0.992	0.988	0.090	0.088
<i>KM</i>	0.996	0.986	0.982	0.080	0.078	0.798	0.774	0.776	0.136	0.062
<i>IPOD</i>	0.998	0.990	0.990	0.086	0.084	0.832	0.824	0.848	0.150	0.084
<i>LQ</i>	0.974	0.974	0.952	0.062	0.052	0.568	0.560	0.538	0.046	0.012
<i>ISIS<sub>C</sub></i>	0.976	0.966	0.980	0.993	0.920	0.670	0.670	0.596	0.870	0.273
	$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
<i>CDC</i>	0.830	0.830	0.824	0.048	0.034	0.630	0.636	0.614	0.084	0.026
<i>CDC</i> <sub>1</sub>	0.820	0.824	0.822	0.866	0.494	0.622	0.636	0.610	0.844	0.218
<i>CDC</i> <sub>2</sub>	0.820	0.824	0.822	1	0.552	0.622	0.636	0.610	1.000	0.242
<i>QaSIS</i>	0.738	0.754	0.736	0.228	0.122	0.352	0.398	0.366	0.240	0.010
<i>CSIR</i>	0.708	0.738	0.750	0.004	0.000	0.334	0.382	0.394	0.002	0.000
<i>RCDCS</i>	1.000	0.998	0.998	0.186	0.186	0.950	0.934	0.940	0.332	0.266
<i>CRSIS</i>	1.000	0.998	1.000	0.006	0.006	0.944	0.932	0.908	0.004	0.004
<i>SVSIR</i>	1.000	1.000	0.998	0.078	0.078	0.990	0.984	0.980	0.112	0.106
<i>CRIS</i>	1.000	1.000	1.000	0.066	0.066	0.984	0.978	0.980	0.084	0.080
<i>KM</i>	0.992	0.988	0.978	0.078	0.076	0.780	0.770	0.782	0.142	0.068
<i>IPOD</i>	0.992	0.990	0.994	0.084	0.082	0.830	0.816	0.844	0.144	0.080
<i>LQ</i>	0.974	0.976	0.958	0.074	0.058	0.532	0.566	0.538	0.058	0.014
<i>ISIS<sub>C</sub></i>	0.986	0.986	0.973	0.976	0.926	0.660	0.676	0.593	0.886	0.223

Figure 1: Estimated  $C$  statistic for DLBCL data with different methods

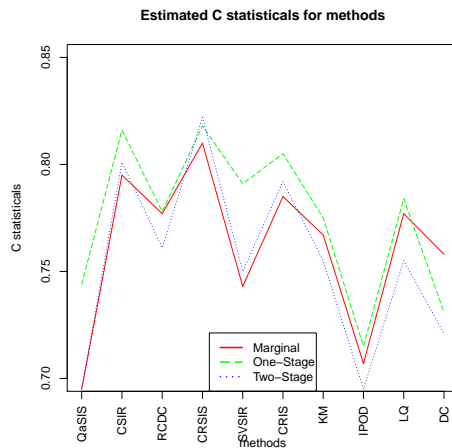


Table 5: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (b) with  $p = 2000$

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$\mathcal{P}_s$ $X_3$	$X_4$	$\mathcal{P}_a$ ALL	$X_1$	$X_2$	$\mathcal{P}_s$ $X_3$	$X_4$	$\mathcal{P}_a$ ALL
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	0.930	0.920	0.880	0.010	0.005	0.625	0.730	0.600	0.010	0.000
<i>CDC<sub>1</sub></i>	0.930	0.905	0.855	0.950	0.695	0.620	0.725	0.575	0.955	0.245
<i>CDC<sub>2</sub></i>	0.930	0.905	0.855	0.945	0.685	0.620	0.725	0.575	0.965	0.235
<i>QaSIS</i>	0.570	0.580	0.535	0.065	0.000	0.190	0.200	0.215	0.100	0.000
<i>CSIR</i>	0.190	0.255	0.290	0.000	0.000	0.060	0.120	0.105	0.000	0.000
<i>RCDCS</i>	0.895	0.885	0.830	0.015	0.005	0.560	0.630	0.560	0.020	0.005
<i>CRSIS</i>	0.930	0.940	0.875	0.000	0.000	0.590	0.600	0.525	0.000	0.000
<i>SVSIR</i>	0.930	0.930	0.870	0.000	0.000	0.670	0.765	0.645	0.000	0.000
<i>CRIS</i>	0.930	0.940	0.885	0.005	0.000	0.705	0.780	0.695	0.000	0.000
<i>KM</i>	0.690	0.725	0.590	0.015	0.000	0.275	0.325	0.240	0.020	0.000
<i>IPOD</i>	0.740	0.760	0.605	0.015	0.000	0.295	0.360	0.295	0.025	0.000
<i>LQ</i>	0.580	0.505	0.480	0.020	0.000	0.115	0.115	0.125	0.010	0.000
<i>ISIS<sub>C</sub></i>	0.760	0.736	0.700	0.5766667	0.273	0.300	0.296	0.280	0.313	0.013
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	1.000	1.000	1.000	0.070	0.070	0.980	0.950	0.935	0.145	0.125
<i>CDC<sub>1</sub></i>	1.000	1.000	1.000	1.000	1.000	0.980	0.950	0.935	1.000	0.870
<i>CDC<sub>2</sub></i>	1.000	1.000	1.000	1.000	1.000	0.980	0.950	0.935	1.000	0.870
<i>QaSIS</i>	0.955	0.945	0.955	0.265	0.230	0.500	0.485	0.485	0.325	0.040
<i>CSIR</i>	0.585	0.635	0.650	0.010	0.000	0.290	0.270	0.340	0.000	0.000
<i>RCDCS</i>	1.000	1.000	1.000	0.125	0.125	0.945	0.955	0.920	0.240	0.205
<i>CRSIS</i>	1.000	1.000	1.000	0.005	0.005	0.955	0.915	0.915	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.050	0.050	0.980	0.975	0.955	0.080	0.065
<i>CRIS</i>	1.000	1.000	1.000	0.050	0.050	0.990	0.985	0.970	0.040	0.035
<i>KM</i>	0.990	0.990	0.970	0.050	0.050	0.670	0.710	0.690	0.095	0.040
<i>IPOD</i>	1.000	0.985	0.985	0.055	0.055	0.730	0.785	0.775	0.100	0.045
<i>LQ</i>	0.965	0.970	0.950	0.055	0.050	0.410	0.435	0.455	0.020	0.000
<i>ISIS<sub>C</sub></i>	0.966	0.983	0.980	1.000	0.933	0.600	0.626	0.623	0.816	0.233
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
<i>CDC</i>	0.505	0.540	0.455	0.005	0.000	0.230	0.325	0.205	0.010	0.000
<i>CDC<sub>1</sub></i>	0.495	0.520	0.430	0.475	0.060	0.230	0.325	0.200	0.420	0.010
<i>CDC<sub>2</sub></i>	0.495	0.520	0.430	0.925	0.085	0.230	0.325	0.200	0.950	0.010
<i>QaSIS</i>	0.180	0.185	0.115	0.025	0.000	0.110	0.060	0.080	0.050	0.000
<i>CSIR</i>	0.195	0.270	0.310	0.000	0.000	0.080	0.130	0.105	0.000	0.000
<i>RCDCS</i>	0.865	0.875	0.780	0.015	0.005	0.485	0.580	0.445	0.025	0.000
<i>CRSIS</i>	0.895	0.915	0.820	0.000	0.000	0.515	0.555	0.405	0.000	0.000
<i>SVSIR</i>	0.915	0.915	0.855	0.005	0.000	0.640	0.755	0.605	0.005	0.000
<i>CRIS</i>	0.920	0.910	0.885	0.005	0.000	0.690	0.730	0.645	0.005	0.000
<i>KM</i>	0.700	0.745	0.565	0.015	0.000	0.250	0.345	0.225	0.020	0.000
<i>IPOD</i>	0.740	0.765	0.600	0.015	0.000	0.285	0.355	0.300	0.025	0.000
<i>LQ</i>	0.605	0.495	0.445	0.015	0.000	0.135	0.135	0.150	0.015	0.000
<i>ISIS<sub>C</sub></i>	0.720	0.696	0.630	0.483	0.186	0.250	0.267	0.230	0.273	0.006
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
<i>CDC</i>	0.805	0.745	0.820	0.035	0.020	0.585	0.540	0.505	0.055	0.015
<i>CDC<sub>1</sub></i>	0.805	0.740	0.820	0.845	0.400	0.580	0.530	0.500	0.795	0.140
<i>CDC<sub>2</sub></i>	0.805	0.740	0.820	1.000	0.455	0.580	0.530	0.505	1.000	0.145
<i>QaSIS</i>	0.650	0.625	0.675	0.095	0.025	0.305	0.265	0.230	0.145	0.000
<i>CSIR</i>	0.585	0.670	0.685	0.010	0.000	0.290	0.280	0.340	0.000	0.000
<i>RCDCS</i>	1.000	1.000	0.995	0.100	0.100	0.930	0.920	0.880	0.205	0.150
<i>CRSIS</i>	1.000	1.000	1.000	0.005	0.005	0.905	0.850	0.865	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.050	0.050	0.980	0.980	0.940	0.060	0.050
<i>CRIS</i>	1.000	1.000	1.000	0.045	0.045	0.980	0.975	0.950	0.040	0.030
<i>KM</i>	0.985	0.975	0.975	0.060	0.060	0.675	0.690	0.685	0.095	0.040
<i>IPOD</i>	1.000	0.980	0.985	0.055	0.055	0.735	0.760	0.760	0.095	0.045
<i>LQ</i>	0.965	0.965	0.945	0.040	0.030	0.470	0.435	0.435	0.045	0.000
<i>ISIS<sub>C</sub></i>	0.986	0.970	0.966	1.000	0.926	0.620	0.580	0.583	0.796	0.173

Table 6: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (c) with  $p = 1000$

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$
$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$										
<i>CDC</i>	0.910	0.908	0.916	0.016	0.012	0.688	0.682	0.646	0.060	0.006
<i>CDC</i> <sub>1</sub>	0.906	0.894	0.910	0.952	0.688	0.676	0.66	0.628	0.994	0.306
<i>CDC</i> <sub>2</sub>	0.906	0.894	0.910	1.000	0.728	0.676	0.66	0.628	1.000	0.306
<i>QaSIS</i>	0.754	0.678	0.696	0.006	0.002	0.256	0.230	0.190	0.006	0.000
<i>CSIR</i>	0.452	0.474	0.452	0.000	0.000	0.212	0.244	0.230	0.000	0.000
<i>RCDCS</i>	0.966	0.970	0.954	0.002	0.002	0.688	0.736	0.682	0.000	0.000
<i>CRSIS</i>	0.990	0.988	0.986	0.000	0.000	0.840	0.836	0.824	0.000	0.000
<i>SVSIR</i>	0.988	0.980	0.980	0.004	0.004	0.864	0.866	0.870	0.000	0.000
<i>CRIS</i>	0.988	0.982	0.972	0.004	0.004	0.788	0.796	0.786	0.000	0.000
<i>KM</i>	0.842	0.852	0.830	0.004	0.000	0.326	0.300	0.280	0.004	0.000
<i>IPOD</i>	0.854	0.866	0.844	0.002	0.002	0.318	0.306	0.282	0.004	0.000
<i>LQ</i>	0.904	0.908	0.904	0.006	0.004	0.388	0.394	0.348	0.002	0.000
<i>ISIS<sub>C</sub></i>	0.800	0.790	0.773	0.993	0.440	0.300	0.290	0.300	0.630	0.010
$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$										
<i>CDC</i>	1.000	1.000	0.998	0.006	0.006	0.966	0.974	0.964	0.064	0.046
<i>CDC</i> <sub>1</sub>	1.000	1.000	0.998	1.000	0.998	0.966	0.970	0.964	1.000	0.910
<i>CDC</i> <sub>2</sub>	1.000	1.000	0.998	1.000	0.998	0.966	0.970	0.964	1.000	0.910
<i>QaSIS</i>	0.992	0.990	0.990	0.036	0.036	0.650	0.662	0.632	0.004	0.000
<i>CSIR</i>	0.880	0.922	0.918	0.000	0.000	0.576	0.628	0.606	0.000	0.000
<i>RCDCS</i>	1.000	1.000	1.000	0.024	0.024	0.952	0.952	0.950	0.000	0.000
<i>CRSIS</i>	1.000	1.000	1.000	0.024	0.024	0.978	0.978	0.972	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.020	0.020	0.988	0.990	0.990	0.000	0.000
<i>CRIS</i>	1.000	1.000	1.000	0.022	0.022	0.972	0.966	0.968	0.000	0.000
<i>KM</i>	1.000	0.996	0.996	0.026	0.026	0.724	0.726	0.742	0.010	0.004
<i>IPOD</i>	0.998	0.998	0.998	0.032	0.032	0.732	0.748	0.746	0.010	0.008
<i>LQ</i>	1.000	1.000	1.000	0.046	0.046	0.844	0.820	0.832	0.010	0.006
<i>ISIS<sub>C</sub></i>	1.000	0.996	0.996	1.000	0.993	0.776	0.793	0.786	0.823	0.710
$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$										
<i>CDC</i>	0.630	0.594	0.654	0.054	0.004	0.544	0.548	0.572	0.114	0.000
<i>CDC</i> <sub>1</sub>	0.624	0.590	0.646	0.792	0.202	0.534	0.544	0.564	0.834	0.196
<i>CDC</i> <sub>2</sub>	0.624	0.590	0.646	1.000	0.276	0.536	0.544	0.564	1.000	0.256
<i>QaSIS</i>	0.516	0.530	0.512	0.034	0.004	0.240	0.246	0.214	0.008	0.000
<i>CSIR</i>	0.926	0.950	0.944	0.000	0.000	0.648	0.688	0.662	0.000	0.000
<i>RCDCS</i>	1.000	0.998	1.000	0.020	0.020	0.890	0.898	0.892	0.000	0.000
<i>CRSIS</i>	1.000	1.000	1.000	0.004	0.004	0.964	0.954	0.946	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.012	0.012	0.986	0.984	0.980	0.000	0.000
<i>CRIS</i>	1.000	1.000	1.000	0.008	0.008	0.956	0.956	0.954	0.000	0.000
<i>KM</i>	0.994	0.998	0.992	0.030	0.030	0.660	0.678	0.668	0.006	0.002
<i>IPOD</i>	0.994	1.000	0.996	0.028	0.028	0.676	0.692	0.694	0.006	0.006
<i>LQ</i>	1.000	1.000	1.000	0.044	0.044	0.800	0.782	0.788	0.006	0.006
<i>ISIS<sub>C</sub></i>	0.996	1.000	1.000	1.000	0.996	0.786	0.823	0.783	0.856	0.687

Table 7: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (c) with  $p = 2000$

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$ ALL	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$ ALL
$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$										
<i>CDC</i>	0.865	0.855	0.830	0.015	0.000	0.570	0.585	0.540	0.035	0.000
<i>CDC</i> <sub>1</sub>	0.865	0.835	0.825	0.935	0.550	0.560	0.565	0.535	0.965	0.155
<i>CDC</i> <sub>2</sub>	0.865	0.835	0.825	1.000	0.575	0.560	0.565	0.535	1.000	0.155
<i>QaSIS</i>	0.625	0.535	0.630	0.010	0.010	0.135	0.130	0.145	0.000	0.000
<i>CSIR</i>	0.350	0.365	0.420	0.000	0.000	0.165	0.195	0.220	0.000	0.000
<i>RCDCS</i>	0.940	0.940	0.950	0.000	0.000	0.590	0.645	0.595	0.000	0.000
<i>CRSIS</i>	0.985	0.990	0.980	0.000	0.000	0.810	0.800	0.770	0.000	0.000
<i>SVSIR</i>	0.980	0.980	0.975	0.000	0.000	0.795	0.825	0.780	0.000	0.000
<i>CRIS</i>	0.975	0.975	0.980	0.000	0.000	0.730	0.730	0.725	0.000	0.000
<i>KM</i>	0.800	0.770	0.775	0.010	0.010	0.265	0.225	0.200	0.000	0.000
<i>IPOD</i>	0.800	0.795	0.760	0.015	0.005	0.270	0.220	0.220	0.000	0.000
<i>LQ</i>	0.860	0.865	0.850	0.005	0.000	0.345	0.305	0.255	0.000	0.000
<i>ISIS<sub>C</sub></i>	0.776	0.743	0.713	0.993	0.360	0.236	0.276	0.246	0.677	0.010
$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$										
<i>CDC</i>	1.000	1.000	0.995	0.010	0.010	0.950	0.925	0.935	0.045	0.025
<i>CDC</i> <sub>1</sub>	1.000	1.000	0.995	1.000	0.995	0.950	0.925	0.925	1.000	0.835
<i>CDC</i> <sub>2</sub>	1.000	1.000	0.995	1.000	0.995	0.950	0.925	0.925	1.000	0.835
<i>QaSIS</i>	0.975	0.985	0.995	0.020	0.015	0.505	0.530	0.580	0.000	0.000
<i>CSIR</i>	0.835	0.825	0.845	0.000	0.000	0.535	0.520	0.560	0.000	0.000
<i>RCDCS</i>	1.000	1.000	1.000	0.005	0.005	0.945	0.945	0.955	0.000	0.000
<i>CRSIS</i>	1.000	1.000	1.000	0.005	0.005	0.985	0.980	0.980	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.010	0.010	0.990	0.990	0.980	0.000	0.000
<i>CRIS</i>	1.000	1.000	1.000	0.015	0.015	0.980	0.970	0.975	0.000	0.000
<i>KM</i>	0.980	0.990	0.985	0.015	0.015	0.610	0.655	0.635	0.005	0.000
<i>IPOD</i>	0.995	1.000	0.995	0.010	0.010	0.640	0.650	0.635	0.000	0.000
<i>LQ</i>	1.000	1.000	1.000	0.020	0.020	0.790	0.785	0.810	0.000	0.000
<i>ISIS<sub>C</sub></i>	1.000	0.996	0.993	1.000	0.990	0.773	0.783	0.770	0.837	0.673
$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$										
<i>CDC</i>	0.355	0.365	0.290	0.015	0.000	0.205	0.230	0.210	0.040	0.000
<i>CDC</i> <sub>1</sub>	0.350	0.345	0.290	0.415	0.020	0.200	0.220	0.205	0.425	0.005
<i>CDC</i> <sub>2</sub>	0.350	0.345	0.290	0.965	0.045	0.200	0.215	0.205	0.925	0.015
<i>QaSIS</i>	0.065	0.055	0.045	0.005	0.000	0.030	0.025	0.035	0.000	0.000
<i>CSIR</i>	0.445	0.440	0.515	0.000	0.000	0.235	0.225	0.310	0.000	0.000
<i>RCDCS</i>	0.860	0.880	0.870	0.000	0.000	0.500	0.450	0.440	0.000	0.000
<i>CRSIS</i>	0.950	0.945	0.930	0.000	0.000	0.660	0.680	0.615	0.000	0.000
<i>SVSIR</i>	0.975	0.970	0.960	0.000	0.000	0.760	0.775	0.720	0.000	0.000
<i>CRIS</i>	0.940	0.930	0.950	0.000	0.000	0.670	0.660	0.625	0.000	0.000
<i>KM</i>	0.755	0.700	0.715	0.010	0.010	0.215	0.200	0.190	0.000	0.000
<i>IPOD</i>	0.755	0.695	0.710	0.005	0.000	0.250	0.205	0.175	0.000	0.000
<i>LQ</i>	0.790	0.780	0.795	0.010	0.000	0.275	0.285	0.230	0.000	0.000
<i>ISIS<sub>C</sub></i>	0.593	0.623	0.550	0.980	0.143	0.220	0.230	0.213	0.653	0.006
$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$										
<i>CDC</i>	0.570	0.455	0.590	0.025	0.000	0.470	0.475	0.505	0.090	0.000
<i>CDC</i> <sub>1</sub>	0.570	0.455	0.585	0.760	0.115	0.465	0.480	0.505	0.800	0.115
<i>CDC</i> <sub>2</sub>	0.570	0.45	0.585	1.000	0.170	0.470	0.475	0.505	1.000	0.165
<i>QaSIS</i>	0.350	0.305	0.320	0.010	0.000	0.180	0.145	0.185	0.000	0.000
<i>CSIR</i>	0.900	0.870	0.915	0.000	0.000	0.575	0.600	0.650	0.000	0.000
<i>RCDCS</i>	1.000	1.000	1.000	0.000	0.000	0.875	0.820	0.870	0.000	0.000
<i>CRSIS</i>	1.000	1.000	1.000	0.005	0.005	0.925	0.925	0.940	0.000	0.000
<i>SVSIR</i>	1.000	1.000	1.000	0.000	0.000	0.985	0.990	0.980	0.000	0.000
<i>CRIS</i>	1.000	0.995	1.000	0.000	0.000	0.950	0.960	0.955	0.000	0.000
<i>KM</i>	0.990	0.985	0.990	0.015	0.015	0.580	0.550	0.565	0.005	0.000
<i>IPOD</i>	0.990	0.990	0.990	0.015	0.015	0.560	0.575	0.575	0.000	0.000
<i>LQ</i>	1.000	0.995	1.000	0.020	0.020	0.735	0.735	0.735	0.005	0.005
<i>ISIS<sub>C</sub></i>	0.993	0.980	0.983	1.000	0.956	0.756	0.756	0.720	0.833	0.583



Table 8: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (d) with  $p = 1000$ .

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$
	$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	0.804	0.786	0.792	0.016	0.002	0.558	0.558	0.548	0.052	0.000
<i>CDC</i> <sub>1</sub>	0.788	0.770	0.778	0.400	0.196	0.548	0.536	0.536	0.514	0.076
<i>CDC</i> <sub>2</sub>	0.788	0.770	0.778	0.888	0.422	0.548	0.534	0.536	0.922	0.132
<i>QaSIS</i>	0.300	0.256	0.224	0.008	0.000	0.154	0.126	0.130	0.018	0.000
<i>CSIR</i>	0.346	0.396	0.368	0.012	0.000	0.180	0.222	0.226	0.062	0.000
<i>RCDCS</i>	0.910	0.894	0.904	0.006	0.004	0.640	0.622	0.602	0.042	0.006
<i>CRSIS</i>	0.952	0.948	0.952	0.004	0.002	0.764	0.766	0.756	0.042	0.020
<i>SVSIR</i>	0.930	0.930	0.938	0.000	0.000	0.722	0.702	0.706	0.014	0.002
<i>CRIS</i>	0.924	0.930	0.938	0.004	0.002	0.724	0.690	0.696	0.044	0.008
<i>KM</i>	0.754	0.756	0.770	0.004	0.000	0.314	0.298	0.272	0.018	0.000
<i>IPOD</i>	0.610	0.578	0.576	0.008	0.004	0.228	0.232	0.204	0.014	0.000
<i>LQ</i>	0.778	0.774	0.776	0.004	0.002	0.420	0.394	0.372	0.016	0.000
<i>ISIS</i> <sub>C</sub>	0.480	0.490	0.470	0.933	0.100	0.256	0.240	0.286	0.656	0.020
	$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$									
<i>CDC</i>	0.990	0.992	0.992	0.010	0.010	0.926	0.900	0.904	0.048	0.028
<i>CDC</i> <sub>1</sub>	0.990	0.992	0.992	0.832	0.812	0.920	0.896	0.902	0.892	0.664
<i>CDC</i> <sub>2</sub>	0.990	0.992	0.992	1.000	0.974	0.920	0.896	0.902	1.000	0.754
<i>QaSIS</i>	0.624	0.612	0.636	0.002	0.002	0.456	0.378	0.428	0.026	0.006
<i>CSIR</i>	0.762	0.808	0.792	0.020	0.012	0.526	0.560	0.552	0.066	0.004
<i>RCDCS</i>	1.000	1.000	1.000	0.000	0.000	0.950	0.946	0.952	0.034	0.026
<i>CRSIS</i>	1.000	1.000	1.000	0.000	0.000	0.990	0.992	0.986	0.032	0.030
<i>SVSIR</i>	1.000	1.000	1.000	0.000	0.000	0.976	0.978	0.980	0.010	0.010
<i>CRIS</i>	1.000	1.000	1.000	0.000	0.000	0.970	0.980	0.978	0.040	0.036
<i>KM</i>	0.984	0.984	0.986	0.002	0.002	0.732	0.748	0.734	0.018	0.002
<i>IPOD</i>	0.894	0.902	0.906	0.014	0.008	0.512	0.458	0.484	0.036	0.000
<i>LQ</i>	0.990	0.998	1.000	0.000	0.000	0.874	0.872	0.856	0.020	0.008
<i>ISIS</i> <sub>C</sub>	0.970	0.956	0.976	1.000	0.923	0.803	0.806	0.823	0.853	0.733
	$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
<i>CDC</i>	0.762	0.744	0.762	0.016	0.000	0.700	0.688	0.686	0.092	0.008
<i>CDC</i> <sub>1</sub>	0.756	0.738	0.762	0.868	0.390	0.696	0.682	0.678	0.872	0.316
<i>CDC</i> <sub>2</sub>	0.756	0.738	0.762	1	0.430	0.698	0.682	0.678	1.000	0.372
<i>QaSIS</i>	0.400	0.372	0.366	0.018	0.002	0.274	0.224	0.276	0.034	0.002
<i>CSIR</i>	0.766	0.800	0.786	0.026	0.014	0.532	0.552	0.546	0.082	0.008
<i>RCDCS</i>	1.000	0.998	1.000	0.000	0.000	0.948	0.942	0.946	0.032	0.024
<i>CRSIS</i>	1.000	1.000	1.000	0.000	0.000	0.988	0.990	0.988	0.032	0.030
<i>SVSIR</i>	1.000	1.000	1.000	0.000	0.000	0.974	0.976	0.980	0.010	0.010
<i>CRIS</i>	1.000	1.000	1.000	0.002	0.002	0.972	0.976	0.976	0.054	0.046
<i>KM</i>	0.986	0.988	0.990	0.002	0.002	0.728	0.744	0.732	0.016	0.002
<i>IPOD</i>	0.940	0.930	0.958	0.008	0.008	0.588	0.574	0.600	0.038	0.000
<i>LQ</i>	0.994	1.000	0.998	0.000	0.000	0.870	0.872	0.854	0.016	0.004
<i>ISIS</i> <sub>C</sub>	0.990	0.983	0.997	1.000	0.973	0.783	0.816	0.786	0.836	0.713

Table 9: Simulation results of  $\mathcal{P}_s$  and  $\mathcal{P}_a$  for Example 2 (d) with  $p = 2000$

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$ ALL	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$ ALL
$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$										
<i>CDC</i>	0.685	0.740	0.745	0.005	0.000	0.395	0.475	0.425	0.035	0.000
<i>CDC</i> <sub>1</sub>	0.675	0.735	0.740	0.320	0.135	0.390	0.460	0.400	0.410	0.025
<i>CDC</i> <sub>2</sub>	0.675	0.735	0.740	0.850	0.300	0.390	0.460	0.400	0.890	0.060
<i>QaSIS</i>	0.150	0.135	0.160	0.010	0.000	0.100	0.075	0.105	0.010	0.000
<i>CSIR</i>	0.240	0.280	0.330	0.015	0.000	0.125	0.175	0.160	0.035	0.000
<i>RCDCS</i>	0.830	0.855	0.815	0.005	0.005	0.470	0.555	0.435	0.035	0.000
<i>CRSIS</i>	0.940	0.920	0.885	0.005	0.005	0.690	0.695	0.625	0.045	0.010
<i>SVSIR</i>	0.910	0.885	0.850	0.000	0.000	0.615	0.640	0.575	0.005	0.005
<i>CRIS</i>	0.905	0.885	0.855	0.005	0.005	0.590	0.635	0.555	0.030	0.000
<i>KM</i>	0.675	0.685	0.600	0.000	0.000	0.175	0.240	0.195	0.005	0.000
<i>IPOD</i>	0.525	0.480	0.460	0.010	0.000	0.120	0.175	0.125	0.010	0.000
<i>LQ</i>	0.665	0.630	0.670	0.000	0.000	0.290	0.320	0.235	0.010	0.000
<i>ISIS<sub>C</sub></i>	0.350	0.416	0.386	0.906	0.043	0.216	0.223	0.206	0.690	0.013
$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.01 \sim 0.03$										
<i>CDC</i>	0.985	1.000	0.965	0.000	0.000	0.835	0.900	0.830	0.025	0.015
<i>CDC</i> <sub>1</sub>	0.985	1.000	0.965	0.775	0.750	0.830	0.890	0.825	0.845	0.525
<i>CDC</i> <sub>2</sub>	0.985	1.000	0.965	1.000	0.950	0.830	0.890	0.825	1.000	0.620
<i>QaSIS</i>	0.505	0.490	0.510	0.000	0.000	0.360	0.285	0.305	0.010	0.000
<i>CSIR</i>	0.660	0.695	0.710	0.015	0.010	0.445	0.410	0.490	0.085	0.000
<i>RCDCS</i>	1.000	1.000	0.985	0.000	0.000	0.885	0.890	0.910	0.015	0.010
<i>CRSIS</i>	1.000	0.995	0.995	0.000	0.000	0.965	0.975	0.950	0.015	0.010
<i>SVSIR</i>	1.000	1.000	0.995	0.000	0.000	0.945	0.945	0.940	0.015	0.015
<i>CRIS</i>	1.000	1.000	0.995	0.000	0.000	0.935	0.945	0.950	0.015	0.015
<i>KM</i>	0.960	0.955	0.965	0.005	0.005	0.660	0.640	0.615	0.000	0.000
<i>IPOD</i>	0.850	0.835	0.820	0.010	0.005	0.405	0.380	0.390	0.005	0.000
<i>LQ</i>	1.000	0.990	0.990	0.000	0.000	0.825	0.805	0.780	0.005	0.000
<i>ISIS<sub>C</sub></i>	0.956	0.933	0.920	1.000	0.840	0.810	0.843	0.800	0.893	0.710
$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$										
<i>CDC</i>	0.370	0.425	0.340	0.015	0.000	0.280	0.260	0.280	0.045	0.000
<i>CDC</i> <sub>1</sub>	0.360	0.415	0.310	0.460	0.015	0.270	0.245	0.270	0.420	0.005
<i>CDC</i> <sub>2</sub>	0.360	0.415	0.310	0.855	0.030	0.270	0.245	0.270	0.885	0.010
<i>QaSIS</i>	0.090	0.070	0.050	0.000	0.000	0.045	0.055	0.030	0.000	0.000
<i>CSIR</i>	0.265	0.310	0.335	0.015	0.000	0.135	0.185	0.165	0.035	0.000
<i>RCDCS</i>	0.830	0.840	0.805	0.000	0.000	0.455	0.540	0.435	0.035	0.000
<i>CRSIS</i>	0.935	0.905	0.885	0.005	0.000	0.670	0.680	0.600	0.050	0.015
<i>SVSIR</i>	0.915	0.875	0.835	0.000	0.000	0.620	0.625	0.565	0.010	0.005
<i>CRIS</i>	0.910	0.875	0.840	0.015	0.010	0.595	0.615	0.540	0.065	0.005
<i>KM</i>	0.680	0.695	0.605	0.000	0.000	0.175	0.240	0.195	0.005	0.000
<i>IPOD</i>	0.495	0.540	0.450	0.005	0.000	0.115	0.180	0.130	0.000	0.000
<i>LQ</i>	0.730	0.665	0.685	0.000	0.000	0.305	0.320	0.255	0.010	0.000
<i>ISIS<sub>C</sub></i>	0.607	0.623	0.537	0.973	0.146	0.206	0.223	0.216	0.596	0.006
$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$										
<i>CDC</i>	0.695	0.730	0.690	0.005	0.000	0.650	0.585	0.595	0.065	0.000
<i>CDC</i> <sub>1</sub>	0.680	0.730	0.685	0.855	0.310	0.650	0.570	0.585	0.810	0.200
<i>CDC</i> <sub>2</sub>	0.680	0.730	0.685	1.000	0.325	0.650	0.570	0.585	1.000	0.240
<i>QaSIS</i>	0.250	0.305	0.270	0.000	0.000	0.185	0.205	0.140	0.005	0.000
<i>CSIR</i>	0.650	0.705	0.720	0.020	0.010	0.435	0.425	0.485	0.085	0.000
<i>RCDCS</i>	1.000	0.995	0.985	0.000	0.000	0.885	0.890	0.900	0.015	0.010
<i>CRSIS</i>	1.000	0.995	0.995	0.000	0.000	0.965	0.965	0.945	0.015	0.010
<i>SVSIR</i>	1.000	1.000	0.995	0.000	0.000	0.940	0.945	0.940	0.015	0.015
<i>CRIS</i>	1.000	0.995	0.995	0.015	0.015	0.920	0.940	0.935	0.040	0.035
<i>KM</i>	0.975	0.965	0.975	0.005	0.005	0.660	0.645	0.615	0.000	0.000
<i>IPOD</i>	0.930	0.895	0.915	0.005	0.000	0.475	0.445	0.475	0.005	0.000
<i>LQ</i>	1.000	0.995	0.995	0.000	0.000	0.815	0.800	0.780	0.005	0.000
<i>ISIS<sub>C</sub></i>	0.993	0.986	0.983	1.000	0.963	0.760	0.776	0.736	0.813	0.660

Table 10: Combination of existing methods and  $CSV S_1$ ,  $CSV S_2$  in Example 2 (a) continuous.

Model	$r = 0.5$					$r = 0.8$				
	$X_1$	$\mathcal{P}_s$			$\mathcal{P}_a$	$X_1$	$\mathcal{P}_s$			$\mathcal{P}_a$
		$X_2$	$X_3$	$X_4$	ALL		$X_2$	$X_3$	$X_4$	ALL
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.075	0.090	0.095	0.660	0.000	0.030	0.085	0.065	0.790	0.000
$QaSIS_2^M$	0.075	0.090	0.090	0.975	0.000	0.030	0.085	0.065	0.960	0.000
$CSIR_1^M$	0.210	0.305	0.315	0.660	0.020	0.070	0.125	0.125	0.765	0.000
$CSIR_2^M$	0.210	0.305	0.315	0.975	0.035	0.070	0.125	0.125	0.960	0.000
$RCDCS_1^M$	0.835	0.815	0.735	0.660	0.310	0.520	0.570	0.485	0.765	0.090
$RCDCS_2^M$	0.835	0.815	0.735	0.975	0.450	0.520	0.575	0.485	0.965	0.110
$CRSIS_1^M$	0.865	0.850	0.760	0.660	0.385	0.510	0.510	0.455	0.765	0.080
$CRSIS_2^M$	0.865	0.850	0.760	0.975	0.550	0.510	0.515	0.455	0.960	0.095
$SVSIR_1^M$	0.870	0.860	0.835	0.660	0.425	0.620	0.650	0.595	0.765	0.190
$SVSIR_2^M$	0.870	0.860	0.835	0.975	0.605	0.620	0.655	0.595	0.965	0.230
$CRIS_1^M$	0.850	0.870	0.815	0.660	0.395	0.670	0.695	0.630	0.765	0.215
$CRIS_2^M$	0.850	0.870	0.815	0.975	0.580	0.670	0.700	0.630	0.960	0.270
$KM_1^M$	0.615	0.655	0.550	0.665	0.120	0.170	0.155	0.160	0.770	0.005
$KM_2^M$	0.615	0.655	0.550	0.975	0.185	0.170	0.155	0.160	0.960	0.005
$IPOD_1^M$	0.660	0.665	0.565	0.670	0.125	0.220	0.245	0.175	0.775	0.010
$IPOD_2^M$	0.660	0.665	0.565	0.975	0.190	0.220	0.245	0.175	0.960	0.010
$LQ_1^M$	0.445	0.485	0.455	0.660	0.040	0.105	0.050	0.050	0.775	0.000
$LQ_2^M$	0.445	0.485	0.455	0.975	0.060	0.105	0.050	0.050	0.965	0.000
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.510	0.410	0.505	0.930	0.140	0.145	0.205	0.180	0.995	0.010
$QaSIS_2^M$	0.505	0.410	0.510	1.000	0.150	0.145	0.205	0.185	1.000	0.010
$CSIR_1^M$	0.640	0.650	0.690	0.925	0.225	0.355	0.360	0.395	0.995	0.060
$CSIR_2^M$	0.640	0.650	0.695	1.000	0.250	0.355	0.360	0.400	1.000	0.060
$RCDCS_1^M$	1.000	0.995	0.985	0.915	0.900	0.950	0.925	0.890	0.995	0.770
$RCDCS_2^M$	1.000	0.995	0.985	1.000	0.980	0.950	0.925	0.890	1.000	0.775
$CRSIS_1^M$	1.000	0.995	0.99	0.915	0.900	0.920	0.850	0.855	0.995	0.645
$CRSIS_2^M$	1.000	0.995	0.99	1.000	0.985	0.920	0.850	0.855	1.000	0.650
$SVSIR_1^M$	1.000	1.000	0.995	0.915	0.910	0.965	0.945	0.945	0.995	0.860
$SVSIR_2^M$	1.000	1.000	0.995	1.000	0.995	0.965	0.945	0.945	1.000	0.865
$CRIS_1^M$	1.000	1.000	1.000	0.920	0.920	0.965	0.975	0.945	0.995	0.880
$CRIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.965	0.975	0.945	1.000	0.885
$KM_1^M$	0.965	0.965	0.965	0.930	0.845	0.455	0.475	0.500	0.995	0.115
$KM_2^M$	0.965	0.965	0.965	1.000	0.915	0.455	0.475	0.500	1.000	0.115
$IPOD_1^M$	0.975	0.970	0.970	0.930	0.845	0.550	0.530	0.530	0.995	0.140
$IPOD_2^M$	0.975	0.970	0.970	1.000	0.915	0.550	0.530	0.530	1.000	0.140
$LQ_1^M$	0.975	0.960	0.955	0.930	0.825	0.250	0.285	0.255	0.995	0.000
$LQ_2^M$	0.975	0.960	0.955	1.000	0.890	0.250	0.285	0.260	1.000	0.000

Table 11: Combination of existing methods and  $CSV S_1$ ,  $CSV S_2$  in Example 2 (b).

	$r = 0.5$					$r = 0.8$				
	$\mathcal{P}_s$	$\mathcal{P}_s$	$\mathcal{P}_s$	$\mathcal{P}_s$	$\mathcal{P}_a$	$\mathcal{P}_s$	$\mathcal{P}_s$	$\mathcal{P}_s$	$\mathcal{P}_s$	$\mathcal{P}_a$
	$X_1$	$X_2$	$X_3$	$X_4$	ALL	$X_1$	$X_2$	$X_3$	$X_4$	ALL
	$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.736	0.748	0.734	0.856	0.434	0.356	0.410	0.364	0.840	0.068
$QaSIS_2^M$	0.736	0.748	0.734	1.000	0.468	0.356	0.41	0.364	1.00	0.072
$CSIR_1^M$	0.704	0.730	0.738	0.848	0.332	0.324	0.372	0.388	0.814	0.054
$CSIR_2^M$	0.704	0.730	0.738	1.000	0.376	0.324	0.372	0.388	1.000	0.056
$RDCS_1^M$	1.000	0.998	0.998	0.872	0.868	0.948	0.934	0.938	0.856	0.714
$RDCS_2^M$	1.000	0.998	0.998	1.000	0.996	0.948	0.934	0.938	1.000	0.828
$CRSIS_1^M$	1.000	1.000	1.000	0.852	0.852	0.948	0.93	0.912	0.818	0.668
$CRSIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.948	0.93	0.912	1.000	0.810
$SVSIR_1^M$	1.000	1.000	0.998	0.860	0.858	0.988	0.980	0.978	0.818	0.778
$SVSIR_2^M$	1.000	1.000	0.998	1.000	0.998	0.988	0.980	0.978	1.000	0.946
$CRIS_1^M$	1.000	1.000	1.000	0.860	0.860	0.984	0.978	0.98	0.818	0.772
$CRIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.984	0.978	0.98	1.000	0.942
$KM_1^M$	0.992	0.986	0.974	0.86	0.824	0.768	0.766	0.778	0.828	0.366
$KM_2^M$	0.992	0.986	0.974	1.00	0.954	0.768	0.766	0.778	1.000	0.438
$IPOD_1^M$	0.992	0.988	0.986	0.852	0.824	0.824	0.794	0.832	0.834	0.448
$IPOD_2^M$	0.992	0.988	0.986	1.000	0.966	0.824	0.794	0.832	1.000	0.534
$LQ_1^M$	0.974	0.974	0.958	0.862	0.780	0.528	0.554	0.528	0.82	0.188
$LQ_2^M$	0.974	0.974	0.958	1.000	0.912	0.528	0.554	0.528	1.00	0.218
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.160	0.165	0.095	0.460	0.015	0.115	0.055	0.075	0.405	0.000
$QaSIS_2^M$	0.160	0.165	0.095	0.925	0.020	0.115	0.055	0.075	0.955	0.000
$CSIR_1^M$	0.195	0.260	0.300	0.455	0.005	0.080	0.120	0.105	0.395	0.000
$CSIR_2^M$	0.195	0.260	0.300	0.925	0.015	0.080	0.120	0.105	0.950	0.000
$RDCS_1^M$	0.855	0.865	0.775	0.470	0.295	0.460	0.555	0.435	0.405	0.050
$RDCS_2^M$	0.855	0.865	0.775	0.925	0.525	0.460	0.555	0.435	0.950	0.100
$CRSIS_1^M$	0.880	0.915	0.820	0.455	0.325	0.505	0.535	0.410	0.395	0.045
$CRSIS_2^M$	0.880	0.915	0.820	0.925	0.610	0.505	0.535	0.410	0.950	0.115
$SVSIR_1^M$	0.910	0.915	0.855	0.460	0.340	0.630	0.735	0.585	0.395	0.115
$SVSIR_2^M$	0.910	0.915	0.855	0.925	0.660	0.630	0.735	0.585	0.950	0.265
$CRIS_1^M$	0.910	0.910	0.865	0.465	0.335	0.665	0.705	0.635	0.400	0.125
$CRIS_2^M$	0.910	0.910	0.865	0.925	0.655	0.665	0.705	0.635	0.950	0.275
$KM_1^M$	0.700	0.745	0.560	0.465	0.155	0.225	0.330	0.225	0.410	0.005
$KM_2^M$	0.700	0.745	0.560	0.925	0.265	0.225	0.330	0.225	0.950	0.020
$IPOD_1^M$	0.720	0.750	0.595	0.465	0.165	0.270	0.350	0.295	0.405	0.015
$IPOD_2^M$	0.720	0.750	0.595	0.925	0.275	0.270	0.350	0.295	0.950	0.040
$LQ_1^M$	0.595	0.490	0.430	0.475	0.100	0.135	0.125	0.150	0.410	0.010
$LQ_2^M$	0.595	0.490	0.430	0.925	0.160	0.135	0.125	0.150	0.950	0.010
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.635	0.605	0.665	0.825	0.275	0.300	0.265	0.240	0.795	0.025
$QaSIS_2^M$	0.635	0.605	0.665	1.000	0.305	0.300	0.265	0.245	1.000	0.030
$CSIR_1^M$	0.585	0.655	0.680	0.825	0.200	0.280	0.270	0.340	0.780	0.015
$CSIR_2^M$	0.585	0.655	0.680	1.000	0.250	0.280	0.270	0.345	1.000	0.020
$RDCS_1^M$	0.995	1.000	0.995	0.845	0.845	0.930	0.910	0.875	0.800	0.605
$RDCS_2^M$	0.995	1.000	0.995	1.000	0.990	0.930	0.910	0.875	1.000	0.745
$CRSIS_1^M$	1.000	0.995	1.000	0.825	0.820	0.900	0.860	0.860	0.785	0.530
$CRSIS_2^M$	1.000	0.995	1.000	1.000	0.995	0.900	0.860	0.860	1.000	0.680
$SVSIR_1^M$	1.000	1.000	1.000	0.840	0.840	0.980	0.970	0.935	0.790	0.700
$SVSIR_2^M$	1.000	1.000	1.000	1.000	1.000	0.980	0.970	0.935	1.000	0.890
$CRIS_1^M$	1.000	1.000	1.000	0.840	0.840	0.980	0.970	0.950	0.785	0.705
$CRIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.980	0.970	0.950	1.000	0.900
$KM_1^M$	0.985	0.975	0.975	0.835	0.790	0.670	0.675	0.680	0.780	0.280
$KM_2^M$	0.985	0.975	0.975	1.000	0.940	0.670	0.675	0.680	1.000	0.350
$IPOD_1^M$	1.000	0.980	0.985	0.835	0.805	0.720	0.745	0.750	0.785	0.330
$IPOD_2^M$	1.000	0.980	0.985	1.000	0.965	0.720	0.745	0.750	1.000	0.430
$LQ_1^M$	0.960	0.960	0.940	0.830	0.725	0.465	0.425	0.420	0.785	0.095
$LQ_2^M$	0.960	0.960	0.940	1.000	0.870	0.465	0.425	0.420	1.000	0.110

Table 12: Combination of existing methods and  $CSV S_1$ ,  $CSV S_2$  in Example 2 (c).

Model	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$
	$n = 200, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.380	0.320	0.360	0.796	0.052	0.176	0.176	0.192	0.826	0.014
$QaSIS_2^M$	0.376	0.318	0.358	1.000	0.056	0.178	0.174	0.190	1.000	0.014
$CSIR_1^M$	0.852	0.888	0.874	0.79	0.518	0.594	0.634	0.640	0.846	0.232
$CSIR_2^M$	0.852	0.888	0.874	1.00	0.660	0.596	0.632	0.638	1.000	0.272
$RCDCS_1^M$	0.998	0.992	0.998	0.782	0.770	0.880	0.888	0.880	0.840	0.574
$RCDCS_2^M$	0.998	0.992	0.998	1.000	0.988	0.880	0.888	0.880	1.000	0.694
$CRSIS_1^M$	1.000	1.000	1.000	0.784	0.784	0.962	0.952	0.944	0.84	0.728
$CRSIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.962	0.952	0.944	1.00	0.872
$SVSIR_1^M$	1.000	1.000	1.000	0.786	0.786	0.976	0.978	0.970	0.828	0.764
$SVSIR_2^M$	1.000	1.000	1.000	1.000	1.000	0.976	0.978	0.970	1.000	0.924
$CRIS_1^M$	1.000	1.000	1.000	0.790	0.790	0.960	0.968	0.95	0.846	0.756
$CRIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.960	0.968	0.95	1.000	0.896
$KM_1^M$	0.980	0.986	0.984	0.786	0.746	0.636	0.642	0.622	0.832	0.238
$KM_2^M$	0.980	0.986	0.984	1.000	0.954	0.638	0.64	0.622	1.000	0.286
$IPOD_1^M$	0.980	0.988	0.992	0.788	0.752	0.630	0.676	0.632	0.834	0.246
$IPOD_2^M$	0.980	0.988	0.992	1.000	0.960	0.632	0.674	0.632	1.000	0.294
$LQ_1^M$	0.986	0.996	0.996	0.784	0.766	0.792	0.782	0.752	0.828	0.404
$LQ_2^M$	0.986	0.996	0.996	1.000	0.98	0.792	0.782	0.752	1.000	0.486
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.035	0.020	0.060	0.410	0.000	0.030	0.015	0.040	0.415	0.000
$QaSIS_2^M$	0.020	0.015	0.050	0.965	0.000	0.020	0.0150	0.040	0.915	0.000
$CSIR_1^M$	0.325	0.340	0.380	0.425	0.030	0.170	0.175	0.210	0.435	0.010
$CSIR_2^M$	0.320	0.335	0.375	0.965	0.045	0.165	0.175	0.210	0.915	0.015
$RCDCS_1^M$	0.734	0.748	0.732	0.434	0.168	0.395	0.410	0.350	0.425	0.035
$RCDCS_2^M$	0.732	0.748	0.730	0.960	0.384	0.395	0.405	0.350	0.920	0.065
$CRSIS_1^M$	0.875	0.860	0.760	0.420	0.265	0.535	0.560	0.475	0.430	0.055
$CRSIS_2^M$	0.875	0.860	0.755	0.970	0.575	0.540	0.555	0.475	0.915	0.145
$SVSIR_1^M$	0.586	0.602	0.590	0.418	0.094	0.395	0.410	0.350	0.425	0.035
$SVSIR_2^M$	0.586	0.600	0.590	0.920	0.184	0.395	0.405	0.350	0.920	0.065
$CRIS_1^M$	0.835	0.82	0.845	0.430	0.255	0.565	0.565	0.52	0.475	0.060
$CRIS_2^M$	0.835	0.82	0.845	0.965	0.570	0.565	0.560	0.52	0.925	0.135
$KM_1^M$	0.640	0.59	0.565	0.415	0.080	0.150	0.200	0.130	0.415	0.005
$KM_2^M$	0.635	0.59	0.560	0.965	0.175	0.150	0.195	0.130	0.915	0.010
$IPOD_1^M$	0.650	0.565	0.54	0.015	0.000	0.170	0.175	0.155	0.415	0.005
$IPOD_2^M$	0.650	0.565	0.54	0.965	0.165	0.170	0.170	0.155	0.915	0.010
$LQ_1^M$	0.685	0.64	0.645	0.005	0.000	0.245	0.260	0.175	0.420	0.010
$LQ_2^M$	0.685	0.64	0.645	0.965	0.235	0.245	0.255	0.175	0.915	0.025
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.220	0.180	0.185	0.760	0.015	0.090	0.130	0.145	0.800	0.005
$QaSIS_2^M$	0.215	0.170	0.185	1.000	0.015	0.095	0.120	0.145	1.000	0.005
$CSIR_1^M$	0.780	0.795	0.800	0.760	0.400	0.505	0.500	0.590	0.820	0.120
$CSIR_2^M$	0.780	0.795	0.800	1.000	0.500	0.505	0.495	0.590	1.000	0.170
$RCDCS_1^M$	0.990	0.990	0.980	0.765	0.740	0.845	0.790	0.830	0.815	0.440
$RCDCS_2^M$	0.990	0.990	0.980	1.000	0.960	0.845	0.780	0.830	1.000	0.555
$CRSIS_1^M$	0.995	0.990	0.980	0.760	0.730	0.950	0.915	0.940	0.810	0.660
$CRSIS_2^M$	0.995	0.990	0.980	1.000	0.965	0.950	0.910	0.940	1.000	0.820
$SVSIR_1^M$	1.000	1.000	0.995	0.760	0.760	0.945	0.925	0.940	0.800	0.665
$SVSIR_2^M$	1.000	1.000	0.995	1.000	0.995	0.945	0.925	0.940	1.000	0.835
$CRIS_1^M$	1.000	1.000	0.995	0.770	0.770	0.940	0.925	0.920	0.810	0.640
$CRIS_2^M$	1.000	1.000	0.995	1.000	0.995	0.940	0.925	0.920	1.000	0.815
$KM_1^M$	0.965	0.975	0.945	0.760	0.680	0.580	0.540	0.555	0.800	0.175
$KM_2^M$	0.965	0.975	0.945	1.000	0.890	0.585	0.535	0.555	1.000	0.225
$IPOD_1^M$	0.970	0.975	0.970	0.760	0.705	0.580	0.555	0.580	0.800	0.175
$IPOD_2^M$	0.970	0.975	0.970	1.000	0.915	0.585	0.550	0.580	1.000	0.240
$LQ_1^M$	0.995	0.975	0.990	0.760	0.735	0.755	0.720	0.710	0.805	0.330
$LQ_2^M$	0.995	0.975	0.990	1.000	0.960	0.755	0.715	0.710	1.000	0.415

Table 13: Combination of existing methods and  $CSV S_1$ ,  $CSV S_2$  in Example 2 (d)

	$r = 0.5$					$r = 0.8$				
	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$ ALL	$X_1$	$X_2$	$X_3$	$X_4$	$\mathcal{P}_a$ ALL
	$n = 100, p = 1000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.400	0.366	0.362	0.866	0.048	0.268	0.218	0.274	0.872	0.026
$QaSIS_2^M$	0.400	0.366	0.362	1.000	0.052	0.270	0.218	0.274	1.000	0.026
$CSIR_1^M$	0.756	0.792	0.780	0.872	0.412	0.526	0.540	0.530	0.886	0.152
$CSIR_2^M$	0.756	0.792	0.780	1.000	0.470	0.526	0.540	0.530	1.000	0.176
$RDCS_1^M$	0.998	0.998	1.000	0.866	0.864	0.942	0.94	0.94	0.876	0.728
$RDCS_2^M$	0.998	0.998	1.000	1.000	0.996	0.942	0.94	0.94	1.000	0.832
$CRSIS_1^M$	1.000	1.000	1.000	0.866	0.866	0.986	0.990	0.984	0.874	0.836
$CRSIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.986	0.990	0.984	1.000	0.960
$SVSIR_1^M$	1.000	1.000	1.000	0.866	0.866	0.972	0.974	0.976	0.870	0.800
$SVSIR_2^M$	1.000	1.000	1.000	1.000	1.000	0.972	0.974	0.976	1.000	0.922
$CRIS_1^M$	1.000	1.000	1.000	0.866	0.866	0.968	0.974	0.976	0.878	0.804
$CRIS_2^M$	1.000	1.000	1.000	1.000	1.000	0.968	0.974	0.976	1.000	0.918
$KM_1^M$	0.986	0.988	0.988	0.866	0.834	0.718	0.736	0.726	0.870	0.344
$KM_2^M$	0.986	0.988	0.988	1.000	0.966	0.718	0.736	0.726	1.000	0.394
$IPOD_1^M$	0.938	0.926	0.956	0.866	0.720	0.574	0.56	0.586	0.874	0.182
$IPOD_2^M$	0.938	0.926	0.956	1.000	0.826	0.574	0.56	0.586	1.000	0.208
$LQ_1^M$	0.994	0.998	0.998	0.866	0.856	0.866	0.870	0.836	0.870	0.558
$LQ_2^M$	0.994	0.998	0.998	1.000	0.990	0.866	0.870	0.836	1.000	0.640
	$n = 100, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.070	0.075	0.090	0.895	0.000	0.035	0.035	0.050	0.905	0.000
$QaSIS_2^M$	0.070	0.075	0.090	1.000	0.000	0.035	0.035	0.050	1.000	0.000
$CSIR_1^M$	0.260	0.305	0.335	0.895	0.020	0.115	0.170	0.165	0.910	0.010
$CSIR_2^M$	0.260	0.305	0.335	1.000	0.030	0.115	0.170	0.165	1.000	0.010
$RDCS_1^M$	0.825	0.830	0.790	0.455	0.245	0.430	0.520	0.425	0.400	0.020
$RDCS_2^M$	0.825	0.830	0.790	0.850	0.450	0.430	0.520	0.425	0.880	0.075
$CRSIS_1^M$	0.935	0.905	0.855	0.460	0.340	0.655	0.655	0.595	0.420	0.110
$CRSIS_2^M$	0.935	0.905	0.855	0.850	0.630	0.655	0.655	0.595	0.880	0.240
$SVSIR_1^M$	0.905	0.860	0.830	0.455	0.315	0.600	0.610	0.540	0.400	0.070
$SVSIR_2^M$	0.905	0.860	0.830	0.850	0.565	0.600	0.610	0.540	0.875	0.170
$CRIS_1^M$	0.895	0.870	0.835	0.460	0.320	0.580	0.605	0.520	0.425	0.075
$CRIS_2^M$	0.895	0.870	0.835	0.850	0.565	0.580	0.605	0.520	0.885	0.165
$KM_1^M$	0.640	0.640	0.600	0.895	0.190	0.170	0.205	0.125	0.905	0.005
$KM_2^M$	0.640	0.640	0.600	1.000	0.215	0.170	0.205	0.125	1.000	0.005
$IPOD_1^M$	0.490	0.495	0.445	0.890	0.075	0.110	0.180	0.110	0.905	0.005
$IPOD_2^M$	0.490	0.495	0.445	1.000	0.090	0.110	0.180	0.110	1.000	0.005
$LQ_1^M$	0.770	0.710	0.700	0.890	0.310	0.245	0.305	0.215	0.905	0.045
$LQ_2^M$	0.770	0.710	0.700	1.000	0.350	0.245	0.305	0.215	1.000	0.045
	$n = 200, p = 2000, d = \lceil n/\log(n) \rceil, CR \approx 0.1 \sim 0.3$									
$QaSIS_1^M$	0.235	0.210	0.310	0.990	0.005	0.175	0.165	0.170	0.980	0.010
$QaSIS_2^M$	0.235	0.210	0.315	1.000	0.005	0.175	0.165	0.170	1.000	0.010
$CSIR_1^M$	0.690	0.720	0.745	0.990	0.345	0.460	0.420	0.465	0.980	0.100
$CSIR_2^M$	0.690	0.720	0.750	1.000	0.350	0.460	0.420	0.470	1.000	0.105
$RDCS_1^M$	1.000	0.995	0.985	0.855	0.840	0.880	0.890	0.900	0.815	0.600
$RDCS_2^M$	1.000	0.995	0.985	1.000	0.980	0.880	0.890	0.900	1.000	0.710
$CRSIS_1^M$	1.000	0.995	0.995	0.855	0.845	0.965	0.965	0.940	0.815	0.720
$CRSIS_2^M$	1.000	0.995	0.995	1.000	0.990	0.965	0.965	0.940	1.000	0.875
$SVSIR_1^M$	1.000	1.000	0.995	0.855	0.850	0.940	0.945	0.940	0.815	0.705
$SVSIR_2^M$	1.000	1.000	0.995	1.000	0.995	0.940	0.945	0.940	1.000	0.830
$CRIS_1^M$	1.000	0.995	0.995	0.860	0.850	0.920	0.935	0.935	0.815	0.680
$CRIS_2^M$	1.000	0.995	0.995	1.000	0.990	0.920	0.935	0.935	1.000	0.805
$KM_1^M$	0.975	0.980	0.970	0.990	0.935	0.600	0.610	0.585	0.980	0.255
$KM_2^M$	0.975	0.980	0.970	1.000	0.940	0.600	0.610	0.585	1.000	0.265
$IPOD_1^M$	0.905	0.905	0.895	0.990	0.720	0.470	0.450	0.430	0.980	0.115
$IPOD_2^M$	0.905	0.905	0.895	1.000	0.725	0.470	0.450	0.430	1.000	0.120
$LQ_1^M$	1.000	1.000	0.995	0.990	0.985	0.815	0.765	0.755	0.980	0.500
$LQ_2^M$	1.000	1.000	0.995	1.000	0.995	0.815	0.765	0.755	1.000	0.515

Table 14: Selected genes ID based on different screening methods.

QaSIS	4398	8136	3122	1190	2968	9287	8380	2344	530	8908	5336	6364	5828	575	7523	3455	1987	9068	7241	7634	483	4066
QaSIS <sub>1</sub> <sup>M</sup>	4398	8136	3122	1190	2968	9287	8380	2344	530	8908	5336	6364	5828	575	7523	3455	1987	9068	7241	7634	6231	491
QaSIS <sub>2</sub> <sup>M</sup>	4398	8136	3122	1190	2968	9287	8380	2344	530	8908	5336	6364	5828	575	7523	3455	1987	9068	7241	7634	2425	6594
CSIR	1804	870	7951	5746	4054	6666	7027	5894	7207	1000	9479	1746	5756	7366	4118	9074	2720	168	8460	6395	947	5975
CSIR <sub>1</sub> <sup>M</sup>	1804	870	7951	5746	4054	6666	7027	5894	7207	1000	9479	1746	5756	7366	4118	9074	2720	168	8460	6395	6231	491
CSIR <sub>2</sub> <sup>M</sup>	1804	870	7951	5746	4054	6666	7027	5894	7207	1000	9479	1746	5756	7366	4118	9074	2720	168	8460	6395	2425	6594
RCDCS	3633	5746	8312	2397	3846	5960	7295	1425	7207	3342	8150	7583	1951	7895	1977	2620	1000	1453	8727	6649	2429	870
RCDCS <sub>1</sub> <sup>M</sup>	3633	5746	8312	2397	3846	5960	7295	1425	7207	3342	8150	7583	1951	7895	1977	2620	1000	1453	8727	6649	5926	447
RCDCS <sub>2</sub> <sup>M</sup>	3633	5746	8312	2397	3846	5960	7295	1425	7207	3342	8150	7583	1951	7895	1977	2620	1000	1453	8727	6649	2292	6278
CRSIS	659	4873	2956	5746	6017	2353	1977	1000	5960	8150	1460	7895	2397	4805	9786	3633	1211	8814	6649	2720	10375	8312
CRSIS <sub>1</sub> <sup>M</sup>	659	4873	2956	5746	6017	2353	1977	1000	5960	8150	1460	7895	2397	4805	9786	3633	1211	8814	6649	2720	5926	447
CRSIS <sub>2</sub> <sup>M</sup>	659	4873	2956	5746	6017	2353	1977	1000	5960	8150	1460	7895	2397	4805	9786	3633	1211	8814	6649	2720	2292	6278
SVSIR	1000	870	8312	5149	3846	9485	10375	947	7559	1785	2966	6567	1512	2167	5894	1631	10567	3575	7528	8523	3041	2041
SVSIR <sub>1</sub> <sup>M</sup>	1000	870	8312	5149	3846	9485	10375	947	7559	1785	2966	6567	1512	2167	5894	1631	10567	3575	7528	8523	5926	447
SVSIR <sub>2</sub> <sup>M</sup>	1000	870	8312	5149	3846	9485	10375	947	7559	1785	2966	6567	1512	2167	5894	1631	10567	3575	7528	8523	2292	6278
CRIS	7674	7207	8326	1746	8460	7215	7366	10209	3342	3739	9065	1425	4081	3633	2367	8024	4833	7684	2335	9074	9912	5131
CRIS <sub>1</sub> <sup>M</sup>	7674	7207	8326	1746	8460	7215	7366	10209	3342	3739	9065	1425	4081	3633	2367	8024	4833	7684	2335	9074	5926	447
CRIS <sub>2</sub> <sup>M</sup>	7674	7207	8326	1746	8460	7215	7366	10209	3342	3739	9065	1425	4081	3633	2367	8024	4833	7684	2335	9074	2292	6278
KM	3012	9148	8814	2720	8727	9700	7469	5975	1746	3633	7951	56	1460	5746	2956	10096	5149	168	3033	2731	5303	4136
KM <sub>1</sub> <sup>M</sup>	3012	9148	8814	2720	8727	9700	7469	5975	1746	3633	7951	56	1460	5746	2956	10096	5149	168	3033	2731	6231	491
KM <sub>2</sub> <sup>M</sup>	3012	9148	8814	2720	8727	9700	7469	5975	1746	3633	7951	56	1460	5746	2956	10096	5149	168	3033	2731	2425	6594
IPOD	8727	3633	10375	7559	9700	9074	5450	9479	5241	8312	3033	8487	6923	9653	6440	3353	6055	3684	3805	947	1327	5821
IPOD <sub>1</sub> <sup>M</sup>	8727	3633	10375	7559	9700	9074	5450	9479	5241	8312	3033	8487	6923	9653	6440	3353	6055	3684	3805	947	6231	491
IPOD <sub>2</sub> <sup>M</sup>	8727	3633	10375	7559	9700	9074	5450	9479	5241	8312	3033	8487	6923	9653	6440	3353	6055	3684	3805	947	2425	6594
LQ	9399	5975	870	8814	4118	1460	2720	1327	1746	2531	5756	5969	10567	675	7469	2041	8326	9367	9279	9283	3633	8523
LQ <sub>1</sub> <sup>M</sup>	9399	5975	870	8814	4118	1460	2720	1327	1746	2531	5756	5969	10567	675	7469	2041	8326	9367	9279	9283	6231	491
LQ <sub>2</sub> <sup>M</sup>	9399	5975	870	8814	4118	1460	2720	1327	1746	2531	5756	5969	10567	675	7469	2041	8326	9367	9279	9283	2425	6594
CDC	5926	447	10375	1986	7515	6941	7592	4018	1357	659	9913	3787	3187	2892	2409	2213	9449	6095	8640	2353	7322	5226
CDC <sub>1</sub> <sup>M</sup>	5926	447	10375	1986	7515	6941	7592	4018	1357	659	9913	3787	3187	2892	2409	2213	9449	6095	8640	2353	7322	5226
CDC <sub>2</sub> <sup>M</sup>	5926	447	10375	1986	7515	6941	7592	4018	1357	659	9913	3787	3187	2892	2409	2213	9449	6095	8640	2353	2292	6278
ISIS <sub>c</sub>	581	635	1781	1977	2001	2595	2605	3327	4585	4704	4741	4982	5653	5719	5772	6980	7335	9257	9353	9953	10027	10090

Table 15: Selected gene IDs for DLBCL data based on different screening methods.

<i>QaSIS</i>	5849	871	62176145	10215957	126	129	161411402521	2517619623386976	616	1328	116325981361	7118598022522501	528911552154320225886904	799
<i>QaSIS</i> <sub>1</sub> <sup>M</sup>	5849	871	62176145	10215957	126	129	161411402521	2517619623386976	616	1328	116325981361	7118598022522501	528911552154320225882409	1969
<i>QaSIS</i> <sub>2</sub> <sup>M</sup>	5849	871	62176145	10215957	126	129	161411402521	2517619623386976	616	1328	116325981361	7118598022522501	528911552154320225882308	2306
<i>CSIR</i> <sub>1</sub> <sup>M</sup>	1831	25796365	182519943822	41317098	145638203821	1841	613422391660	4202181938012541	6133163916621671	2532379916641681	174038023813	80		
<i>CSIR</i> <sub>1</sub> <sup>M</sup>	1831	25796365	182519943822	41317098	145638203821	1841	613422391660	4202181938012541	6133163916621671	2532379916641681	174038022409	1969		
<i>CSIR</i> <sub>2</sub> <sup>M</sup>	1831	25796365	182519943822	41317098	145638203821	1841	613422391660	4202181938012541	6133163916621671	2532379916641681	174038022308	2306		
<i>RCDCS</i>	2579	50	4131182519943787	181924791097	958	394	1965343025762672	1023253210864232	10293822725118321831	7055184169031995	414847335775			
<i>RCDCS</i> <sub>1</sub> <sup>M</sup>	2579	50	4131182519943787	181924791097	958	394	1965343025762672	1023253210864232	10293822725118321831	7055184169031995	41482409	1969		
<i>RCDCS</i> <sub>2</sub> <sup>M</sup>	2579	50	4131182519943787	181924791097	958	394	1965343025762672	1023253210864232	10293822725118321831	7055184169031995	41482308	2306		
<i>CRSIS</i>	1819	182526721456	2901994394	257626742671	1841	50	3685775198624792673	63652958725134703367	4131253210293821	19652570562119003806				
<i>CRSIS</i> <sub>1</sub> <sup>M</sup>	1819	182526721456	2901994394	257626742671	1841	50	3685775198624792673	63652958725134703367	4131253210293821	1965257056212409	1969			
<i>CRSIS</i> <sub>2</sub> <sup>M</sup>	1819	182526721456	2901994394	257626742671	1841	50	3685775198624792673	63652958725134703367	4131253210293821	1965257056212308	2306			
<i>SVSIR</i>	4131	613438223787	290242326135	419361334149	4202383238203821	69034148	41927227119238243774	4416379941833810	41503775619342003825	5026				
<i>SVSIR</i> <sub>1</sub> <sup>M</sup>	4131	613438223787	290242326135	419361334149	4202383238203821	69034148	41927227119238243774	4416379941833810	41503775619342002409	1969				
<i>SVSIR</i> <sub>2</sub> <sup>M</sup>	4131	613438223787	290242326135	419361334149	4202383238203821	69034148	41927227119238243774	4416379941833810	41503775619342002308	2306				
<i>CRIS</i>	2579	145625322425	701926722958	256972511653	80	21821664	2901986166226946740	5970161221062959	2426662518747380	114816635621	50	3146		
<i>CRIS</i> <sub>1</sub> <sup>M</sup>	2579	145625322425	701926722958	256972511653	80	21821664	2901986166226946740	5970161221062959	2426662518747380	1148166356212409	1969			
<i>CRIS</i> <sub>2</sub> <sup>M</sup>	2579	145625322425	701926722958	256972511653	80	21821664	2901986166226946740	5970161221062959	2426662518747380	1148166356212308	2306			
<i>KM</i>	1653	22554149	419322391235	183142775655	651918304131	571515726498	133257925692442	4148215525321775	166437871645	50183239175599	1669			
<i>KM</i> <sub>1</sub> <sup>M</sup>	1653	22554149	419322391235	183142775655	651918304131	571515726498	133257925692442	4148215525321775	166437871645	50183239172409	1969			
<i>KM</i> <sub>2</sub> <sup>M</sup>	1653	22554149	419322391235	183142775655	651918304131	571515726498	133257925692442	4148215525321775	166437871645	50183239172308	2306			
<i>IPOD</i>	6519	165318301831	6733773	164518324149	536135402005	404825692239	473341334131	370721554148	215672815655	63382079	40364823	25115654	5334	
<i>IPOD</i> <sub>1</sub> <sup>M</sup>	6519	165318301831	6733773	164518324149	536135402005	404825692239	473341334131	370721554148	215672815655	63382079	40364823	25112409	1969	
<i>IPOD</i> <sub>2</sub> <sup>M</sup>	6519	165318301831	6733773	164518324149	536135402005	404825692239	473341334131	370721554148	215672815655	63382079	40364823	25112308	2306	
<i>LQ</i>	2579	613456551994	413118195298	695610865364	10234148	50	41499585775	427725324232	281183169031985	420226741830	29023832	37871671	1596	
<i>LQ</i> <sub>1</sub> <sup>M</sup>	2579	613456551994	413118195298	695610865364	10234148	50	41499585775	427725324232	281183169031985	420226741830	29023832	37872409	1969	
<i>LQ</i> <sub>2</sub> <sup>M</sup>	2579	613456551994	413118195298	695610865364	10234148	50	41499585775	427725324232	281183169031985	420226741830	29023832	37872308	2306	

Table 16: The C-statistic for DLBCL data with different screening methods.

	C-statistic	standard error	lower bound of $C$	upper bound of $C$	$p$ value
<i>QaSIS</i>	0.695	0.084	0.531	0.859	0.00132
<i>QaSIS</i> <sub>1</sub> <sup>M</sup>	0.744	0.090	0.567	0.921	4.281e-05
<i>QaSIS</i> <sub>2</sub> <sup>M</sup>	0.697	0.074	0.553	0.842	0.000342
<i>CSIR</i>	0.795	0.068	0.662	0.928	9.179e-09
<i>CSIR</i> <sub>1</sub> <sup>M</sup>	0.816	0.078	0.663	0.969	4.438e-08
<i>CSIR</i> <sub>2</sub> <sup>M</sup>	0.801	0.077	0.651	0.951	9.321e-10
<i>RCDCS</i>	0.777	0.078	0.625	0.929	3.069e-06
<i>RCDCS</i> <sub>1</sub> <sup>M</sup>	0.778	0.070	0.641	0.915	7.637e-06
<i>RCDCS</i> <sub>2</sub> <sup>M</sup>	0.761	0.067	0.630	0.892	2.146e-06
<i>CRSIS</i>	0.810	0.082	0.650	0.971	2.240e-10
<i>CRSIS</i> <sub>1</sub> <sup>M</sup>	0.818	0.070	0.681	0.954	3.463e-14
<i>CRSIS</i> <sub>2</sub> <sup>M</sup>	0.822	0.062	0.699	0.944	1.217e-08
<i>SVSIR</i>	0.743	0.071	0.604	0.883	2.458e-05
<i>SVSIR</i> <sub>1</sub> <sup>M</sup>	0.791	0.075	0.644	0.938	3.239e-05
<i>SVSIR</i> <sub>2</sub> <sup>M</sup>	0.750	0.077	0.599	0.901	1.886e-05
<i>CRIS</i>	0.785	0.079	0.630	0.939	3.451e-06
<i>CRIS</i> <sub>1</sub> <sup>M</sup>	0.805	0.076	0.657	0.953	1.095e-08
<i>CRIS</i> <sub>2</sub> <sup>M</sup>	0.792	0.099	0.598	0.986	8.451e-05
<i>KM</i>	0.767	0.075	0.621	0.913	0.000438
<i>KM</i> <sub>1</sub> <sup>M</sup>	0.775	0.074	0.630	0.920	2.211e-06
<i>KM</i> <sub>2</sub> <sup>M</sup>	0.755	0.073	0.612	0.898	1.188e-06
<i>IPOD</i>	0.707	0.071	0.567	0.846	0.00043
<i>IPOD</i> <sub>1</sub> <sup>M</sup>	0.715	0.086	0.546	0.883	7.912e-06
<i>IPOD</i> <sub>2</sub> <sup>M</sup>	0.695	0.074	0.550	0.839	0.00034
<i>LQ</i>	0.777	0.069	0.641	0.913	1.081e-05
<i>LQ</i> <sub>1</sub> <sup>M</sup>	0.784	0.084	0.619	0.950	1.243e-08
<i>LQ</i> <sub>2</sub> <sup>M</sup>	0.755	0.069	0.621	0.890	2.106e-06